

# ECON220C Discussion Section 4

## Staggered DID

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# Roadmap

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1. Parallel Trend & No Anticipation
2. Staggered DID Framework
3. Exercise TWFE vs DID

# Setup

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- **Goal:** measure the impact of a policy intervention using differences-in-differences method. Assume there are only two periods, program participation only occurs between periods one and two.
- **Framework:** potential outcome. Remember from 220B  $y_{it} = d_{it}y_{it}(1) + (1 - d_{it})y_{it}(0)$ . Our target:

$$ATT \equiv E[y_{it}(1) - y_{it}(0) | d_{it} = 1]$$

- Before going full-mode on potential outcome + staggered did framework, let's introduce two key assumptions: **parallel trend** and **no anticipation effect**.

# Let's Construct the Model

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**Model:**  $y_{it} = \alpha_i + \lambda_t + \beta d_{it} + u_{it} \implies$  Can we obtain this from Potential outcome notation?

1.  $y_{it}(1) = E[y_{it}(1)] + u_{it}(1)$  and same for  $y_{it}(0)$
2. Combine with  $y_{it} = d_{it}y_{it}(1) + (1 - d_{it})y_{it}(0)$
3. We get our **individual** and **time** fixed effect model.

$$y_{it}(0) = E[y_{it}(0)] + u_{it}(0)$$

$$y_{it}(1) = E[y_{it}(1)] + u_{it}(1)$$

$$y_{it} = x_{it} y_{it}(1) + (1 - x_{it}) y_{it}(0)$$

$$y_{it} = x_{it} (E[y_{it}(1)] + u_{it}(1)) + (1 - x_{it}) (E[y_{it}(0)] + u_{it}(0))$$

$$= \underbrace{E[y_{it}(0)]}_{\alpha_i + \lambda_t} + x_{it} \underbrace{(E[y_{it}(1)] - E[y_{it}(0)])}_{\beta} + u_{it}$$

$$= \alpha_i + \lambda_t + x_{it} \beta + u_{it}$$

## Identification $\beta_{DID}$

As before, the model is  $y_{it} = \alpha_i + \lambda_t + \beta d_{it} + u_{it}$ . The **identification strategy** for  $\beta$  is the following:

1. Take average  $\forall i$  of  $y_{i1}$  and  $y_{i2}$  or first difference (WITHIN TREATED (T) OR CONTROL GROUP (C) )  
 $\Delta y_{T2} = y_{T2} - y_{T1} = (\lambda_2 - \lambda_1) + x_{T2} \beta + \Delta u_{T2}$

$$\Delta y_{C1} = y_{C2} - y_{C1} = (\lambda_2 - \lambda_1) + \Delta u_{C2}$$

2. Remove time and individual fixed effects using double-differencing

$$\Delta y_{T2} - \Delta y_{C2} = (\lambda_2 - \lambda_1) + x_{T2} \beta + \Delta u_{T2} - (\lambda_2 - \lambda_1) - \Delta u_{C2}$$

$$\Delta y_{T2} - \Delta y_{C2} = \beta + \Delta u_{T2} - \Delta u_{C2} \quad \text{WE WANT TO GET RID OF THIS}$$

3. Parallel trends assumption

$$\text{IF: } E[\Delta u_{T2} - \Delta u_{C2}] = 0, \text{ THAT IS } E[u_{T2} - u_{T1}] = E[u_{C2} - u_{C1}]$$

$$\text{THEN } \beta_{DID} = E[\Delta y_{T2}] - E[\Delta y_{C2}] \Rightarrow \hat{\beta}_{DID} = \left( \frac{1}{N_T} \sum_{i \in T} (y_{i2} - y_{i1}) \right) - \left( \frac{1}{N_C} \sum_{i \in C} (y_{i2} - y_{i1}) \right)$$

$$\beta_{DID} = ATT?$$


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- First, let's rewrite  $\beta_{DID} = E[\Delta y_{T2} - \Delta y_{C2}]$  in terms of **potential outcomes**:

$$\begin{aligned}\beta_{DID} &= E[\Delta y_{T2}] - E[\Delta y_{C2}] = E[y_{T2} - y_{T1}] - E[y_{C2} - y_{C1}] \\ &= E[y_{i2}(1) - y_{i1}(1) | x_{i2} = 1] - E[y_{i2}(0) - y_{i1}(0) | x_{i2} = 0]\end{aligned}$$

- $\beta_{DID} - \underline{ATT} =$

$$\begin{aligned}& (E[y_{i2}(1) - y_{i1}(1) | x_{i2} = 1] - E[y_{i2}(0) - y_{i1}(0) | x_{i2} = 0]) \\ & - E[y_{i2}(1) - y_{i2}(0) | x_{i2} = 1] \\ &= (\underbrace{E[y_{i2}(0) | d_{it} = 1]}_{\text{NOT IDENTIFIED}} - E[y_{i2}(1) | x_{i2} = 1]) - E[y_{i2}(0) - y_{i1}(0) | x_{i2} = 0]\end{aligned}$$

$$\beta_{DID} = ATT?$$

LET'S USE THIS

- ① • **No Anticipation effect:**  $E[y_{i1}(1)|d_{i2} = 1] = E[y_{i1}(0)|d_{i2} = 1]$ , in plain english knowing that I would receive the treatment is not changing my behavior in the pre-treatment period.
- $$= (E[y_{i2}(0)|d_{i2} = 1] - E[y_{i2}(1)|x_{i2} = 1]) - E[y_{i2}(0) - y_{i1}(0)|x_{i2} = 0]$$
- $$= (E[y_{i2}(0)|d_{i2} = 1] - E[y_{i2}(0)|x_{i2} = 1]) - E[y_{i2}(0) - y_{i1}(0)|x_{i2} = 0]$$
- ② • **New parallel trends** assumption:

$$E[y_{i2}(0) - y_{i1}(0)|d_{i2} = 1] = E[y_{i2}(0) - y_{i1}(0)|d_{i2} = 0]$$

What is telling us? Is the same as before?

BY ① + ② WE GET  $\beta_{DID} - ATT = 0$

$$\therefore E[y_{i2}(0) - y_{i1}(0)|x_{i2} = 1] - E[y_{i2}(0) - y_{i1}(0)|x_{i2} = 0]$$

# Empirical Studies: DID Estimator and Dummies

$$\hat{\beta}_{DID} = \left( \frac{1}{N_T} \sum_{i \in T} (y_{i2} - y_{i1}) \right) - \left( \frac{1}{N_C} \sum_{i \in C} (y_{i2} - y_{i1}) \right) \quad \text{CAN BE OBTAINED BY RUNNING}$$

$$y_{it} = \beta_0 + \mathbb{1}\{\text{TREATED}\}_{it} \cdot \beta_1 + \mathbb{1}\{t=2\}_{it} \cdot \beta_2 + \mathbb{1}\{\text{TREATED} \cap t=2\}_{it} \cdot \beta_3$$

$$\therefore \hat{\beta}_{DID} = \left[ \underbrace{(\hat{\beta}_3 + \hat{\beta}_2 + \hat{\beta}_1 + \hat{\beta}_0)}_{\bar{y}_{T2}} - \underbrace{(\hat{\beta}_0 + \hat{\beta}_1)}_{\bar{y}_{T1}} \right] - \left[ \underbrace{(\hat{\beta}_2 + \hat{\beta}_0)}_{\bar{y}_{C2}} - \underbrace{\hat{\beta}_0}_{\bar{y}_{C1}} \right]$$

→ EQUIVALENT TO TWFE ESTIMATOR

$$\Delta y_{i2} = \underbrace{\beta_0}_{\text{what is this}} + \underbrace{\Delta x_{i2} \beta}_{\Delta x_{i2} = x_{i2} \rightarrow \text{CLEAR?}} + u_{i2}$$

what is this  $\Delta x_{i2} = x_{i2} \rightarrow \text{CLEAR?}$



# Staggered ID - Setup

- **Setup:** multiple time periods, treatment could happen at any time but once treated, always treated. New notation:  $y_{it}(g)$  where  $g$  is the period when unit  $i$  gets treated (if  $t < g$  **not yet treated**),  $y_{it}(\infty)$  means **never treated**.
- Example: 4 treatment groups and three time periods.

	TREATMENT DECISION				OBSERVED OUTCOME			
	$t=1$	$t=2$	$t=3$	:	$t=1$	$t=2$	$t=3$	
GROUP 1	<u>1</u>	1	1	:	<u><math>y_{i1}(1)</math></u>	$y_{i2}(1)$	$y_{i3}(1)$	
GROUP 2	0	<u>1</u>	1	:	$y_{i1}(2)$	<u><math>y_{i2}(2)</math></u>	$y_{i3}(2)$	
GROUP 3	0	0	<u>1</u>	:	$y_{i1}(3)$	$y_{i2}(3)$	<u><math>y_{i3}(3)</math></u>	
GROUP 4	<u>0</u>	<u>0</u>	<u>0</u>	:	<u><math>y_{i1}(\infty)</math></u>	<u><math>y_{i2}(\infty)</math></u>	<u><math>y_{i3}(\infty)</math></u>	$\Rightarrow$ NEVER TREATED, CONTROL GROUP

# Staggered DID - Assumptions

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PEOPLE TAKING  
TREATMENT AT TIME  $g$  AND  $g'$

## 1. **Parallel trends:**

$$E[y_{it}(\infty) - y_{it-1}(\infty) | G_i = g] = E[y_{jt}(\infty) - y_{jt-1}(\infty) | G_j = g']$$

here  $g, g' < \infty \forall t$ . If treatment hadn't happened, all adoption cohorts would have evolved in parallel at all periods. We are considering people in cohorts  $i$  and  $j$ , treated at  $g, g'$ .

## 2. **No anticipation effect** (same as before)

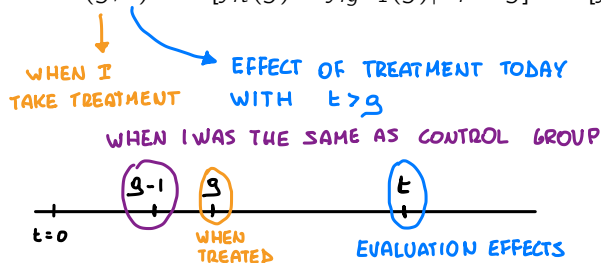
$$E[y_{it}(g) | G_i = g] = E[y_{it}(\infty) | G_i = g] \quad \underbrace{\forall t < g}$$

KNOWING THAT I GO  
INTO TREATMENT DOES NOT  
CHANGE MY BEHAVIOR BEFORE

# Staggered DID - Goal

- **ATT(g, t)** =  $E[y_{it}(g) - y_{it}(\infty) | G_i = g]$  this is the object we aim to identify and estimate.
- **Claim:** given conditions above,  $ATT(g, t)$  is identified as (we will prove it for a simple case later)

$$ATT(g, t) = E[y_{it}(g) - y_{ig-1}(g) | G_i = g] - E[y_{it}(\infty) - y_{ig-1}(\infty) | G_i = \infty]$$



## Exercise

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Consider the **staggered DID** methodology for panel data in the potential outcomes framework. There are two periods  $t = 1, 2$  and three potential treatment trajectories  $\{d_1 = (1, 1); d_2 = (0, 1); d_3 = (0, 0)\}$  as listed in the table below:

$d_i$	t = 1	t=2		t = 1	t=2
group 1	1	0	group 1	$\bar{Y}_{d_1,1} = 0.3$	$\bar{Y}_{d_1,2} = 1.2$
group 2	0	1	group 2	$\bar{Y}_{d_2,1} = 0$	$\bar{Y}_{d_2,2} = 0.6$
group 3	0	0	group 3	$\bar{Y}_{d_3,1} = 0.3$	$\bar{Y}_{d_3,2} = 0.3$

There are three potential outcomes  $\{Y_{i1}(d_i), Y_{i2}(d_i)\}$  given the three different treatment status  $d_i$ . The sample is iid and each group has size  $n_1 = n_2 = n_3 = 1/3$ . The table on the right reports the average value of the outcome of interest for group  $d$  and time  $t$ .

## Exercise - Q1

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Let  $ATT(d_2, t) = E[Y_{i2}(d_2) - Y_{i2}(d_3) | Di = d_2]$  be the average treatment effect on treated at time 2 for the group of individuals who chose  $d_2$  as the treatment trajectory. Explain how you would identify  $ATT(d_2, t)$ .

## Exercise - Q1

NOT OBSERVED

Answer

$$\begin{aligned} \text{ATT}(d_2, t) &= E[Y_{i2}(d_2) - \overbrace{Y_{i2}(d_3)}^{\text{NOT OBSERVED}} \mid d_i = d_2] \\ &= E[Y_{i2}(d_2) \mid d_i = d_2] - E[Y_{i2}(d_3) \mid d_i = d_2] \\ &= E[Y_{i2}(d_2) \mid d_i = d_2] - E[Y_{i2}(d_3) \mid d_i = d_2] \\ &\quad + E[Y_{i2}(d_3) \mid d_i = d_3] - E[Y_{i2}(d_3) \mid d_i = d_3] \end{aligned}$$

ASSUME PARALLEL TRENDS

$$E[Y_{i2}(d_3) - Y_{i1}(d_3) \mid d_i = d_2] = E[Y_{i2}(d_3) - Y_{i1}(d_3) \mid d_i = d_3]$$

EQUAL TO

$$E[Y_{i2}(d_3) \mid d_i = d_2] - E[Y_{i2}(d_3) \mid d_i = d_3] = E[Y_{i1}(d_3) \mid d_i = d_2] - E[Y_{i1}(d_3) \mid d_i = d_3]$$

## Exercise - Q1

Answer

LHS OF SELECTION BIAS (WITH MINUS SIGN)

$$ATT(d_2, t) = E[Y_{i2}(d_2) | d_i = d_2] - E[Y_{i2}(d_3) | d_i = d_2]$$

$$+ E[Y_{i2}(d_3) | d_i = d_3] - E[Y_{i2}(d_3) | d_i = d_3]$$

$$= E[Y_{i2}(d_2) | d_i = d_2] - \underline{E[Y_{i2}(d_3) | d_i = d_2]} + E[Y_{i2}(d_3) | d_i = d_3] - E[Y_{i2}(d_3) | d_i = d_3]$$

WE  
SUBSTITUTE  
RHS

NOW WE IMPOSE NO ANTICIPATION EFFECT

$$\underline{E[Y_{i2}(d_3) | d_i = d_2]} = E[Y_{i1}(d_2) | d_i = d_2]$$

AND WE SUBSTITUTE IN ATT TO GET

$$ATT(d_2, t) = \{E[Y_{i2}(d_2) | d_i = d_2] - E[Y_{i1}(d_2) | d_i = d_2]\} - \{E[Y_{i2}(d_3) - Y_{i1}(d_3) | d_i = d_3]\}$$

## Exercise - Q2

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Explain how would you estimate  $ATT(d_2, t)$ . Please first provide a formula for your proposed estimator and then give a numerical answer based on the information provided.



## Exercise - Q2

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Answer

PUT AN HAT :

$$\begin{aligned} \text{ATT}(d_2, t) &= \{E[Y_{i2}(d_2) | d_i = d_2] - E[Y_{i1}(d_2) | d_i = d_1]\} - \{E[Y_{i2}(d_3) - Y_{i1}(d_3) | d_i = d_3]\} \\ &= \left\{ \frac{1}{N_T} \sum (y_{T2} - y_{T1}) \right\} - \left\{ \frac{1}{N_C} \sum (y_{C2} - y_{C1}) \right\} \\ &= \{ \bar{y}_{d_2,2} - \bar{y}_{d_2,1} \} - \{ \bar{y}_{d_3,2} - \bar{y}_{d_3,1} \} \\ &= (0.6 - 0.0) - (0.3 - 0.3) \end{aligned}$$

## Exercise - Q3

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Suppose we estimate the following model  $Y_{it}(D_i) = \alpha_i + \lambda_t + D_{it}\beta + e_i$  by the TWFE estimator which requires a two way demeaned version of the dummy variable of the treatment. Find  $\tilde{D}_{it}$  and the TWFE estimate  $\hat{\beta}_{TWFE}$ . Please provide numerical answer.

$$\tilde{D}_{it} = D_{it} - \bar{D}_i - \bar{D}_t + \bar{D} \Rightarrow \text{FIRST OF ALL : WHY ?}$$

BY FWL THEOREM: given  $y_{it} = x_{it}\beta + w_{it}\delta + u_{it}$

IF INTERESTED IN BETA, THE FOLLOWING ARE EQUIVALENT:

(1)  $y_{it} \sim x_{it} + w_{it}$

(2)  $\tilde{y}_{it} \sim \tilde{x}_{it}$  RESIDUALIZED WRT  $w_{it}$

(3)  $y_{it} \sim \tilde{x}_{it} \Rightarrow \text{FWL + ANGRIST (2009)}$

## Exercise - Q3 OUR MODEL $y_{it} = \alpha_i + \lambda_t + \Delta_{it} \beta + u_{it}$

Answer  $\tilde{y}_{it}$  IS  $\tilde{\Delta}_{it}$  IN OUR CASE

$\tilde{w}_{it}$  ARE  $\alpha_i$  AND  $\lambda_t$ . HOW DO WE REMOVE THEM? AUX. REG:  $\Delta_{it} = \alpha_i^D + \lambda_t^D + u_{it}^D$

$$1. \bar{\Delta}_i := \frac{1}{T} \sum_t (\alpha_i^D + \lambda_t^D + u_{it}^D) = \alpha_i^D + \bar{\lambda}_t^D + \frac{1}{T} \sum_t u_{it}^D$$

$$\begin{aligned} 2. \Delta_{it} - \bar{\Delta}_i &= (\alpha_i^D + \lambda_t^D + u_{it}^D) - \left( \alpha_i^D + \bar{\lambda}_t^D + \frac{1}{T} \sum_t u_{it}^D \right) \\ &= \lambda_t^D - \bar{\lambda}_t^D + u_{it}^D - \frac{1}{T} \sum_t u_{it}^D \end{aligned}$$

$$3. \bar{\Delta}_t := \frac{1}{N} \sum_i (\alpha_i^D + \lambda_t^D + u_{it}^D) = \bar{\alpha}_i^D + \lambda_t^D + \frac{1}{N} \sum_i u_{it}^D$$

$$\begin{aligned} 4. \Delta_{it} - \bar{\Delta}_i - \bar{\Delta}_t &= (\lambda_t^D - \bar{\lambda}_t^D + u_{it}^D - \frac{1}{T} \sum_t u_{it}^D) - (\bar{\alpha}_i^D + \lambda_t^D + \frac{1}{N} \sum_i u_{it}^D) \\ &= -\bar{\alpha}_i^D - \bar{\lambda}_t^D + u_{it}^D - \frac{1}{T} \sum_t u_{it}^D - \frac{1}{N} \sum_i u_{it}^D \end{aligned}$$

## Exercise - Q3

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Answer

$$\begin{aligned} 5. \quad \bar{D} &= \frac{1}{NT} \sum_i \sum_t (\alpha_i^D + \lambda_t^D + \mu_{it}^D) \\ &= \left( \frac{1}{NT} \sum_i T \alpha_i^D \right) + \left( \frac{1}{NT} \sum_t N \lambda_t^D \right) + \frac{1}{NT} \sum_i \sum_t \mu_{it}^D \\ &= \left( \frac{1}{N} \sum_i \alpha_i^D \right) + \left( \frac{1}{T} \sum_t \lambda_t^D \right) + \frac{1}{NT} \sum_i \sum_t \mu_{it}^D = \bar{\alpha}_i^D + \bar{\lambda}_t^D + \frac{1}{NT} \sum_i \sum_t \mu_{it}^D \end{aligned}$$

$$\begin{aligned} 6. \quad D_{it} - \bar{D}_i - \bar{D}_t + \bar{D} &= -\bar{\alpha}_i^D - \bar{\lambda}_t^D + \mu_{it}^D - \frac{1}{T} \sum_t \mu_{it}^D - \frac{1}{N} \sum_i \mu_{it}^D \\ &\quad + \bar{\alpha}_i^D + \bar{\lambda}_t^D + \frac{1}{NT} \sum_i \sum_t \mu_{it}^D \\ &= \mu_{it}^D - \frac{1}{N} \sum_i \mu_{it}^D - \frac{1}{T} \sum_t \mu_{it}^D + \frac{1}{NT} \sum_i \sum_t \mu_{it}^D := \tilde{D}_{it} \end{aligned}$$

# Exercise - Q3

		$\bar{D}_{t=1}$	$\bar{D}_{t=2}$	
Answer	$d_1$	1	1	$1 = \bar{D}_{i=1}$
	$d_2$	0	1	$1/5 = \bar{D}_{i=2}$
	$d_3$	0	0	$0 = \bar{D}_{i=3}$
		$1/3$	$2/3$	$\bar{D} = 1/2$

$$\tilde{D}_{it} = \begin{vmatrix} 1 & 1 \\ 0 & 1 \\ 0 & 0 \end{vmatrix} - \begin{vmatrix} 1 & 1 \\ 1/2 & 1/2 \\ 0 & 0 \end{vmatrix} - \begin{vmatrix} 1/3 & 2/3 \\ 1/3 & 2/3 \\ 1/3 & 2/3 \end{vmatrix} + \begin{vmatrix} 1/2 & 1/2 \\ 1/2 & 1/2 \\ 1/2 & 1/2 \end{vmatrix}$$

$$\tilde{D}_{it} = \begin{vmatrix} 1/6 & -1/2 \\ -1/3 & 1/3 \\ 1/6 & -1/6 \end{vmatrix}$$

## Exercise - Q3

Answer

PROBLEM : WE DON'T HAVE INDIVIDUAL  
OBSERVATION BUT  
GROUP VALUES

$$\begin{aligned}
 \hat{\beta}_{TWFE} &= \frac{\hat{\text{Cov}}(Y_{it}, \tilde{D}_{it})}{\hat{\text{Var}}(\tilde{D}_{it})} = \frac{\frac{1}{N} \sum_{i=1}^N \sum_{t=1}^2 \tilde{D}_{it} Y_{it}}{\frac{1}{N} \sum_{i=1}^N \sum_{t=1}^2 (\tilde{D}_{it}^2)} \\
 &= \frac{\frac{1}{N} \sum_{t=1}^2 \sum_{i=1}^N \tilde{D}_{it} Y_{it}}{\frac{1}{N} \sum_{t=1}^2 \sum_{i=1}^N (\tilde{D}_{it})^2} = \frac{\sum_{g=1}^3 \sum_{t=1}^2 \frac{n_g}{m_g} \frac{1}{N} \sum_{i \in g} \tilde{D}_{it} Y_{it}}{\sum_{g=1}^3 \sum_{t=1}^2 \frac{n_g}{m_g} \frac{1}{n} \sum_{i \in g} (\tilde{D}_{it})^2} \\
 &= \frac{\sum_{g=1}^3 \sum_{t=1}^2 \frac{n_g}{N} \tilde{D}_{gt} \left( \frac{1}{m_g} \sum_{i \in g} Y_{it} \right)}{\sum_{g=1}^3 \sum_{t=1}^2 \frac{n_g}{n} \frac{1}{m_g} \sum_{i \in g} (\tilde{D}_{it})^2} = \frac{\sum_{g=1}^3 \sum_{t=1}^2 \omega_g \tilde{D}_{gt} \bar{Y}_{gt}}{\sum_{g=1}^3 \sum_{t=1}^2 \omega_g \frac{1}{m_g} n_g \tilde{D}_{gt}^2}
 \end{aligned}$$

## Exercise - Q3

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Answer

$$\hat{\beta}_{TWFE} = \frac{\sum_{g=1}^3 \sum_{t=1}^2 \omega_g \tilde{D}_{gt} \bar{Y}_{gt}}{\sum_{g=1}^3 \sum_{t=1}^2 \omega_g \tilde{D}_{gt}^2} \quad := \kappa$$

$$= \frac{1}{\kappa} \sum_{g=1}^3 \sum_{t=1}^2 \omega_g \tilde{D}_{gt} \bar{Y}_{gt}$$

$$= \sum_{g=1}^3 \sum_{t=1}^2 \frac{\omega_g \tilde{D}_{gt}}{\kappa} \bar{Y}_{gt} = \sum_{g=1}^3 \sum_{t=1}^2 \alpha_{gt} \cdot \bar{Y}_{gt}$$

$$= 0.15$$

## Exercise - Q4

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Verify that  $\hat{\beta}_{TWFE} = \hat{\beta}_{2,DID} - \frac{1}{2}\hat{\beta}_{1,DID}$  where  $\hat{\beta}_{2,DID}$  is the estimate in Q2, and  $\hat{\beta}_{1,DID} = (\bar{Y}_{d_1,2} - \bar{Y}_{d_1,1}) - (\bar{Y}_{d_3,2} - \bar{Y}_{d_3,1})$

$$\hat{\beta}_{1,DID} = (1.2 - 0.3) - (0.3 - 0.3) = 0.9$$

$$\hat{\beta}_{2,DID} - \frac{1}{2} \hat{\beta}_{1,DID} = 0.6 - \frac{1}{2} 0.9 = 0.15 \quad \checkmark$$



## Exercise - Q4

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**Answer**