Title: Gentrification drives patterns of alpha and beta diversity in cities

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**Abstract:**

While there is increasing recognition that social processes in cities like gentrification have ecological consequences, we lack nuanced understanding of the ways gentrification affects urban biodiversity. We analyzed a large camera trap dataset of mammals (> 500g) to evaluate how gentrification impacts species richness and community composition across 23 U.S. cities. After controlling for the negative effect of impervious cover, gentrified parts of a city had the highest mammal species richness. Change in community composition was associated with gentrification in a few cities, which were mostly located along the West Coast. Our results indicate that the impacts of gentrification extend to non-human animals, which provides further evidence that some aspects of nature in cities, such as wildlife, are chronically inaccessible to marginalized human populations.

**One-Sentence Summary:** Across the United States, gentrified parts of a city have more mammal species than non-gentrified parts of a city.

**Main Text:**

In urban areas, the processes that cause unequal access to environmental resources among people have ecological consequences (*1-4*). Gentrification, the process of neighborhood change that includes demographic and economic shifts in historically disinvested neighborhoods whereby white, more educated, and higher income residents displace under-resourced residents (*5*), exemplifies such a process. Ecologically, gentrification-induced changes to a neighborhood could alter which species are locally present in various ways (i.e., richness or alpha diversity). First, as gentrification brings an influx of wealthier residents in a neighborhood, alpha diversity could increase as residents invest more resources into landscape management (i.e., the luxury effect, *6*). Second, gentrification can accompany an influx of green infrastructure such as city parks, which can also increase alpha diversity (*7,8*) Yet, gentrification also displaces the people that previously lived in the area, further intensifying inequities with people’s experience with biodiversity (*9*). Thus, identifying whether gentrification leads to higher alpha diversity would 1) further demonstrate how changes in human populations contribute to changes in biodiversity and 2) provide accumulating evidence that some aspects of nature, such as wildlife, are chronically inaccessible to marginalized human populations.

Simply hypothesizing links between gentrification and alpha diversity, however, would fail to fully recognize how wildlife responds to changes in human populations and urban design. For example, species assemblages in gentrified versus non-gentrified areas may have similar alpha diversity (i.e., same number of species) but differ in the identity of the species present in those assemblages (i.e., high beta diversity). This could exacerbate environmental equity issues if species present in gentrified areas are considered desirable (e.g., songbirds), while those that persist in non-gentrified areas are considered nuisance species (e.g., rats) that are likely to generate human-wildlife conflict (*10*) or spread zoonotic disease (*11*). Furthermore, cities are not identical, and thus, the relative impact of gentrification on biodiversity likely differs among cities (*12*). Therefore, disentangling the effect gentrification may have in exacerbating environmental inequities requires us to quantify how gentrification affects multiple species diversity metrics across cities.

To understand how changes in biodiversity are linked to gentrification, we tested the hypothesis that gentrification is associated with changes in alpha and beta diversity of medium to large mammals (> 500 g). Mammals are an excellent focal group to explore changes in biodiversity that accompany gentrification because they respond strongly to urban landscape heterogeneity (*12*) and are often conservation and nuisance management targets (*13*). Given that gentrification-induced changes to a neighborhood could be associated with both alpha and beta diversity, we evaluated three predictions on the relative contribution gentrification could have on these diversity metrics (Fig. 1).



**Figure 1.** Four theoretical ways that mammal richness (i.e., alpha diversity) and shifts in community composition (i.e., beta diversity) could change in response to gentrification, where shapes in sub-figures A through D represent different species. A) Alpha diversity could remain constant but species assemblages could completely differ between gentrified and non-gentrified areas, resulting in no change in alpha diversity but the highest beta diversity. B) More species could be gained than lost in gentrified areas, resulting in a large increase in alpha diversity and a smaller increase in beta diversity relative to full replacement. C) Communities in non-gentrified areas could be a nested subset of those in gentrified areas, which could result in large increases in alpha diversity but low beta diversity. D) The null prediction: no change; gentrification is not associated with changes in alpha and beta diversity. E) The change in alpha diversity (x axis) as well as beta diversity, where a value of 0 indicates identical communities at a pair of sites and 1 indicates completely different communities at a pair of sites (y axis).

To test our hypothesis and evaluate our predictions across a variety of urban landscapes, we used camera trap data from 23 US cities that are part of the Urban Wildlife Information Network (UWIN; *14*), a systematic multi-city biodiversity monitoring survey. Overall, we compiled three years of data between 2019 and 2021, which represented 188,909 camera trap days at 999 sites across 23 cities that span the contiguous United States (Fig. 2). Our analysis included 21 mammals across 11 families (See Table S1 for the species in our analysis and Table S2 for a summary of species detected across cities). This unique dataset allowed us to quantify variation in how gentrification influenced alpha and beta diversity across a wide range of North American mammals distributed over diverse urban landscapes and compare city-specific estimates to our own theoretical predictions.



**Figure 2. Locations of the 23 cities used to assess differences in mammal communities among gentrified and non-gentrified parts of a city.** Cities are represented by dots. Dot color illustrates the relative effect of gentrification on alpha and beta diversity at average sites in each city that vary in their gentrification status. Gentrification had a more pronounced effect on alpha diversity overall. However, gentrification in West Coast cities had a stronger effect on beta diversity, central U.S. cities had a mixture, and East Cost cities had a stronger alpha diversity effect. See fig. S1 for this map with city names included.

**Alpha diversity results**

We used a log-linear model that accounted for uncertainty in species richness estimates to quantify associations between alpha diversity and gentrification (see Tables S3-S10 for additional information on how we quantified gentrification). The model included two covariates: a binary covariate to indicate whether a site was within 500m of a gentrified Census tract and an urban intensity metric: the proportion of impervious cover within a 1 km buffer of each site. Additionally, we included the interaction between these covariates, and used city-level random effects for all intercept and slope terms. Among cities, the minimum impervious cover at gentrified sites was 22.5%, on average. The expected alpha diversity at non-gentrified sites with 22.5% impervious cover was 7.73 species (95% CI = 6.47, 9.22; 90% CI = 6.67, 8.95) while alpha diversity at a similar gentrified site was about 13% higher (8.74 species; 95% CI = 6.85, 11.12; 90% CI = 7.13, 10.69). This difference equates to roughly 1 to 2 more species at gentrified sites than non-gentrified sites within a city. The among-city model parameters agreed with this pattern: gentrification was likely associated with higher alpha diversity (*βgent* = 0.17; 95% CI = -0.01, 0.36; 90% CI = 0.02, 0.33; Fig. 3), and the interaction term suggested this effect was strongest at lower levels of impervious cover (*βgentXimp* = -0.22; 95% CI = -0.53, 0.09; 90% CI = -0.48, 0.04, Table S11, Fig. S2-S24). Conversely, higher levels of impervious cover across cities had the largest and most pronounced negative effect on alpha diversity (*βimp* = -0.49; 95% CI = -0.71, -0.27; 90% CI = -0.67, -0.31). Therefore, the effect of impervious cover on alpha diversity was roughly 2.87 times greater than that of gentrification. Alpha diversity decreased by about 3.5 species when impervious cover at a site increased from 0 to 80%.

A diagram of different types of gentrification

Description automatically generated

**Figure 3**. **Impervious cover had a greater effect than gentrification on alpha and beta diversity, although there was some indication that alpha diversity was greatest in gentrified areas with lower impervious cover.** Vertical black lines represent among-city estimates, the gray-filled rectangles are 95% credible intervals, and blue dots are city-specific estimates for each model parameter. Alpha diversity model parameters are on the log scale while beta diversity model parameters are on the clog scale (i.e,. -log(1 - x)). As a result, the beta diversity model parameters are constrained to be 0.

At the city level, 10 of 23 cities had > 0.9 probability that gentrification and alpha diversity positively covaried, whereas impervious cover was associated with decreased species richness in 18 of 23 cities (Table S13). Furthermore, cities in which we detected a gentrification effect on alpha diversity were predominantly located in the East Coast and Central United States (Fig. 4).



**Figure 4.** **West Coast cities (more negative longitude) had smaller changes in alpha diversity but greater beta diversity, while the remaining cities had greater changes in alpha diversity and lower beta diversity.** Dots represent the expected beta diversity and change in alpha diversity between an average gentrified and non-gentrified site in each of the 23 cities. The x axis represents the change in alpha diversity as a function of gentrification, with positive values indicating greater species richness at gentrified sites. The y axis is Sørensen’s dissimilarity index, where 0 and 1 respectively represent completely identical and different communities between sites.

**Beta diversity results**

We used a generalized dissimilarity model that accounted for uncertainty in pairwise-dissimilarity between pairs of sites within each city. Our response variable was Sørensen’s dissimilarity index, where a value of 0 means the mammal communities at a pair of sites were identical and 1 means the two sites had completely dissimilar mammal communities. We included impervious cover and a binary covariate to indicate whether a site was near a gentrified Census tract as covariates, controlled for geographic distance between sites within a city, and used city-level random effects for all intercept and slope terms. On average, gentrification had a negligible association with beta diversity. For two adjacent sites with identical levels of impervious cover, the among-city difference in beta diversity at gentrified and non-gentrified sites was effectively zero ( 0.01; 95% CI = 0.00, 0.06; 90% CI = 0.00, 0.05).

On average, impervious cover had the greatest effect on beta diversity (*β­imp* = 0.07; 95% CI = 0.02, 0.19; 90% CI = 0.02, 0.18), followed by the geographic distance between sites within a city (*βdistance* = 0.05; 95% CI = 0.01, 0.13; 90% CI = 0.02, 0.13), and then gentrification (*βgent* = 0.01; 95% CI = 0.00, 0.07; 90% CI = 0.01, 0.05; Fig. 3). As such, the average effect of impervious cover on beta diversity was 5.83 times greater than the average effect of gentrification. For example, when comparing sites at opposite ends of the impervious cover gradient (e.g., hi vs. low impervious cover), beta diversity was almost two times greater than sites with the same level of impervious cover (e.g., low vs. low impervious cover). While holding other covariates at their mean, the beta diversity between sites with the highest (80%) and lowest (0%) impervious cover was 0.16 (95% CI = 0.07, 0.31; 90% CI = 0.08, 0.29) whereas sites with the exact same level of impervious cover was 0.09 (95% CI = 0.02, 0.23; 90% CI = 0.03, 0.21).

While the average among-city estimates indicated a minimal association of gentrification with beta diversity, beta diversity did strongly covary with gentrification in some cities (Fig. 1, Fig. 3, Fig. S2 – S24). Overall, it appears that gentrification in West Coast cities had a greater association with beta diversity than alpha diversity: 4 of the 5 largest associations between gentrification and beta diversity were observed on the West Coast (Fig. 4). Los Angeles, California, for example, had the strongest association between gentrification and beta diversity (median of summed spline coefficients = 0.13; 95% CI = 0.08, 0.19; 90% CI = 0.09, 0.18). In Los Angeles, the beta diversity between gentrified and non-gentrified sites was 1.28 times greater than sites with the same gentrification status (95% CI = 1.16, 1.40; 90% CI = 1.18, 1.38). In other words, sites in Los Angeles that only differed in their gentrification status had mammal communities that were about 60% similar while sites that did not differ in their gentrification status were about 69% similar.

Finally, cities varied in their relationship between impervious cover and beta diversity. Phoenix, Arizona had the largest effect (median of summed spline coefficients = 0.79; 95% CI = 0.72, 0.86; 90% CI = 0.73, 0.85). Beta diversity between a pair of sites at opposite ends of Phoenix’s impervious gradient was 0.72 (95% CI = 0.70, 0.73; 90% CI = 0.70,0.73) while the beta diversity between sites with identical levels of impervious cover was 1.91 times lower (median = 0.38; 95% CI = 0.36, 0.39; 90% CI = 0.36, 0.39). Conversely, Indianapolis, Indiana had the smallest effect (median of summed spline coefficients = 0.07; 95% CI = 0.02, 0.13; 90% CI = 0.03, 0.11). When making the same comparison in Indianapolis, the estimated beta diversity for sites at opposite ends of their impervious cover gradient was 0.27 (95% CI = 0.24, 0.30; 90% CI = 0.25, 0.29) which was similar to the estimated beta diversity between sites with identical levels of impervious land cover (median = 0.22; 95% CI = 0.24, 0.30; 90% CI = 0.25, 0.29). For more city-specific beta diversity results, see fig. S2-S24.

**Species that covaried with gentrification**

As the alpha and beta diversity estimates were compiled from the posterior of a multi-city multi-species occupancy model, we also quantified how each species responded to gentrification. Of the 21 species we analyzed, 11 of them positively covaried with gentrification at the 0.90 credible level (Fig. 5). As we centered and scaled impervious cover, negative parameter estimates for the interaction between gentrification and impervious cover indicate that many of these species were more likely to occupy gentrified areas when impervious cover was low (Fig. 5). See Table S11-S36 for parameter estimates from all models.



**Figure 5.** **The average effect of gentrification on the occupancy of the mammals in this study.** Overall, 10 species were more common in gentrified areas when impervious land cover was low, as evidenced by the strongly negative gentrification X impervious slope terms for those species. Three species were more common in gentrified parts of a city overall, and there were no species who negatively covaried with gentrification. Dots represent median estimates for each species and horizontal lines are 90% credible intervals.

**Discussion**

Our results indicate that gentrification, coupled with impervious cover, shape mammal distributions across U.S. cities. Compared to gentrification, however, impervious cover is the dominant form of environmental variation that impacts mammal distributions and greatly reduces alpha diversity. Further, sites at opposite ends of the impervious gradient in cities have the greatest beta diversity. While this finding confirms decades of urban ecological research (*15*), our analysis shows how generalizable this outcome is across a wide geographic range including multiple U.S. cities. Gentrification, on the other hand, represents a secondary axis that may lessen the negative effect of impervious cover on mammals. Such changes to the mammal community are likely due to wealth-associated increases in vegetative cover such as street trees, professional landscaping, and irrigation (*1,16*), which collectively improve habitat quality for mammals (*17*). Because gentrification also displaces less-affluent human residents, our results provide further evidence that aspects of nature are less accessible to some urban residents, which underlies the need for cities to develop equitable policies to avoid displacement and nature-enabled dispossession of marginalized communities (*1-4*). That we found gentrification-related shifts in mammal communities across a wide range of U.S. cities indicates that the impacts of gentrification extend to non-human animals, which highlights the broader implications and importance of this process.

Gentrification may have a smaller effect on mammal diversity than impervious cover for many reasons. First, areas with high impervious cover are likely unsuitable for most mammals (*6*). In fact, our analysis suggests that gentrification provides the greatest increase in alpha diversity when impervious cover was low. In such locations, higher availability of non-impervious cover coupled with increased resources (e.g., time, money) likely facilitates greening at the residential parcel level, which scales up to enhance neighborhood wildlife habitat. The lack of non-impervious cover in intensely urbanized neighborhoods may preclude this effect. Thus, social processes that occur within cities can non-linearly modulate the magnitude of effect the built environment has on urban biodiversity. Second, the non-gentrified locations we used for comparison represent a mix of sites that were either historically vulnerable to gentrification but did not gentrify or were never vulnerable to gentrification and so could not gentrify (e.g., an already affluent neighborhood). This mix could have made it more difficult to quantify gentrification-induced changes as the sites that were used for comparison varied from one another. Third, gentrification is a multidimensional and dynamic process that we distilled into a binary metric for sake of analysis. As many gentrification metrics exist, all of which quantify gentrification differently, further research is needed to understand gentrification in its different forms across cities (*18*). Nevertheless, we observed an association between our gentrification metric and patterns of mammal diversity across many U.S. cities. Such results indicate that simple gentrification metrics could be used to quantify how gentrification influences other taxa. To facilitate such investigations, we encourage others to tap into the wealth of data products that exist to track demographics over time in cities, beyond what is readily available from the decadal Census. As one example, the IPUMs dataset created by the National Historical Geographic Information System has spatially georectified census data across many years, making it easier for researchers to conduct large-scale comparative research (*19*).

We found that our theoretical predictions (Fig. 1) were more extreme than the nuanced results we observed in our data (Fig. 4). While some cities experienced different species assemblages in gentrified versus non-gentrified areas, wildlife communities were never completely dissimilar (i.e., beta diversity never reached a value of 1; Fig. 1A). We also never witnessed large shifts in alpha diversity and high beta diversity simultaneously (Fig. 1B). Instead, gentrification in many cities was associated with increased alpha diversity but low beta diversity (Fig. 1C) and could indicate that wildlife communities in gentrified and not gentrified areas are nested subsets of one another. Such a pattern could arise due to the presence of common urban generalists throughout most North American cities, such as Northern raccoon (*Procyon lotor*). These results could also indicate that gentrification is a subtle and ongoing process that likely invokes change over long time periods. Our results may thus identify the start of a process that could become more pronounced over time.

We observed a spatial pattern in the relative association of gentrification with alpha and beta diversity, with West Coast cities showing distinct differences from East Coast cities and central U.S. cities representing a mixture of the two (Fig. 2). Cities closer to the East Coast are typically older and potentially more biotically homogenized (*20*); gentrification increased species richness but had minimal effects on community composition in these cities. On the other hand, in West Coast cities, which are younger and were affected by extreme weather events during our data collection period (e.g., drought and El Niño), gentrification impacted community composition but not species richness. Such extreme weather events likely led to more variable wildlife communities as species responded to changes in resource availability, such as water. It is possible that gentrified parts of a city benefited from additional water resources or cooling provided by trees during droughts (*17*). This could result in wildlife communities that differed more from non-gentrified areas if species responded to such resource allocation. In our analysis, we considered the average wildlife community across sampling periods and seasons, minimizing the potential impact of drought on our results. However, a more detailed and longer-term investigation into how drought influences urban species distributions is warranted. Arid cities, for example, exhibit a stronger correlation between local species richness and per capita income (*21*), while species in hotter cities are more negatively influenced by urban intensity (*22*). Exploring how the magnitude of such patterns change during extreme weather events in arid and non-arid cities alike could provide additional insight into how urban biodiversity may respond to a warming climate.

In closing, decades-old land-use decisions in cities can lead to a legacy effect that influences the current distribution of environmental resources available to both humans and wildlife. Similarly, our results show that socioeconomic processes, such as gentrification influence human demographics and wildlife communities. These factors shape where species occur in cities and who can observe or interact with them (negatively or positively). To address these issues, there is a critical need for updated land development and management policies as well as legal mechanisms to prioritize social equity (*23*). For instance, while urban greenspace is considered a public good in theory, in practice it often becomes a commodity primarily accessible to affluent white communities that displaces marginalized communities (*24*–26). Uncoupling urban greenspace development from Western capitalism is a challenging task but involves reframing greenspace development as an essential component of city maintenance rather than an economic development strategy (*23*). It is crucial to prioritize environmental equity in decision-making processes as the choices we make today will shape our cities for decades to come, particularly because cities will continue to house most of the global human population. Large-scale research networks like UWIN are well-positioned to assist decision-makers in understanding social-ecological disparities across different scales, providing valuable insights for building equitable and biodiverse cities.

**Materials and Methods**

*Biological sampling*

We used data from 23 UWIN cities in the United States (U.S.) for this study. Each city followed the same systematic study design, placing motion-triggered camera traps in urban greenspace along an urbanization gradient (*14*). Mammal data for this study came from 12 distinct sampling periods between 2019 and 2021. Camera deployments in each sampling period were about 35 days (sd = 13.01) and began on the first of January, April, July, or October of each year. Because UWIN cities joined the network at different times, the number of sampling periods among cities varied (median = 7, minimum = 2, maximum = 12). The median number of unique camera-trapping sites per city was 35 (minimum = 23, maximum = 104).

Mammals in camera trap images were identified to species by trained experts. However, flying squirrel, gray squirrel, and cottontail rabbit species were summarized to either the subgenus or genus level given challenges in identifying them to the species level from camera trap images (*27*). For each camera deployment we counted the number of days each species was detected and the number of operational camera days, which were then used to estimate species occupancy and detectability within our multi-city multi-species occupancy model (*6,28*).

Overall, 48 mammal species were photographed but our multi-city, multi-species occupancy model converged when limited to the 21 species that were most frequently detected (a minimum of 75 detection days across at least 3 cities). Gray squirrels (Sciurus carolinensis or Sciurus griseus) were detected most often (~41,300 detection days) while flying squirrels (Glaucomys sp.) were detected the least (79 detection days). See Table S1 for the names of the species included in our analysis and Table S2 for a summary of all species detected across cities.

*Social-environmental variables*

We calculated two independent variables and included both in all models. First, to represent a gradient of urban intensity we calculated the percent impervious cover within 1 km of each site from the 2019 National Land Cover Database imperviousness dataset (*29*). Second, we determined if each site was within 500 m of a gentrifying Census tract. To quantify gentrification across a wide range of cities we modified a two-step process described by Chapple et al. (*30*). For the first step, we identified Census tracts that were vulnerable to gentrification in 2010 as tracts with at least 500 residents and two of these three qualities: 1) a median income less than the city’s average income, 2) a proportion of college-educated residents less than the city average, and 3) a proportion of nonwhite residents greater than the city average. To calculate gentrification vulnerability we used the 2010 US decennial Census data via the tidycensus package in R v 4.2.0 (*31,32*). For the second step, we used the 2019 American Community Survey (*33*) data to determine if a vulnerable Census tract became gentrified. Here, vulnerable tracts from the first step were identified as gentrified if they experienced a greater increase in median income between 2010 and 2019 than the average change across a city–after correcting for inflation–as well as one of two qualities: a change in college-educated residents or a change in the proportion of non-Hispanic white residents between 2010 and 2019 that exceeded the average change across the city. For additional details and summaries regarding this gentrification metric see the “Additional gentrification metric details” of the supplemental information, where we provide additional summaries of the variables used to quantify gentrification (Tables *S3-S10*).

*Associations between gentrification and social-environmental variables*

Among cities, on average, 25% (sd = 11%) of camera sites were within 500m of a gentrified Census tract. At sites near gentrified Census tracts, 46% (sd = 20%) of land cover was impervious, on average, while sites not near gentrified Census tracts had an average of 25% (sd = 21%) impervious cover. Within cities, Urbana, Illinois had the lowest percent of sites within 500m of a gentrified Census tract (3%) and Phoenix, Arizona had the highest (50%).

With respect to the 2019 distribution of the variables we used to quantify gentrification across cities, the median per capita income of gentrified Census tracts (mean = $68,785, sd = $28,193) was roughly $30,000 less than non-gentrified Census tracts (mean = $98,678, sd = $50,777). The proportion of non-Hispanic white residents living in gentrified Census tracts (mean = 0.28, sd = 0.26) was lower than non-gentrified Census tracts (mean = 0.48, sd = 0.30), and the proportion of people with a college degree in gentrified Census tracts (mean = 0.34, sd = 0.18) was slightly lower than non-gentrified Census tracts (mean = 0.48, sd = 0.23). Thus, gentrified Census Tracts still have lower incomes, fewer non-Hispanic white residents, and fewer college educated residents than non-gentrified Census Tracts. However, gentrified Census tracts saw greater than average shifts in these variables over time, such that the residents living there have become whiter, richer, and more educated.

Gentrification may also be associated with an increase in urban greenspace. As such, we quantified whether gentrified Census tracts had a greater increase in the proportion of greenspace (i.e., developed, open space from NLCD data) over the same time frame we used to quantify gentrification (i.e., 2010 to 2019). We did not find this to be true. After averaging the proportional increase in urban greenspace across gentrified and not-gentrified Census tracts in each city, the among-city range in both types of Census tracts was effectively zero (min = -0.01, max = 0.00).

*Statistical analysis*

We used a meta-analytic approach to quantify associations between gentrification and impervious cover and patterns of alpha and beta diversity across U.S. cities, using a Bayesian approach for all models (see following sections for more thorough explanation of our three statistical models). However, unlike more common meta-analyses, which must contend with issues of publication bias that can distort results (*34*), our analysis used all available UWIN data to parameterize both alpha- and beta diversity models, resulting in a more unbiased and data-driven evaluation of our hypothesis. We explain the modeling procedure below, see supporting information for a complete description of each model.

To do so, we first fitted a Bayesian multi-city, multi-species occupancy model that included a first-order autoregressive term to account for repeat sampling across primary sampling periods within each city (*6,28*). This model had three separate logit-linear predictors: one to indicate a species presence within a city’s species pool, one for site-level occupancy, and one for site-level detection probability. Following Magle et al. (*6*), we included the distance of each city to the known margin of a species’ geographic range in the first linear predictor, with positive and negative numbers respectively indicating cities within and outside a species range. Range data came from IUCN red list data (*35*). For site-level occupancy and detection, we included impervious cover, gentrification, and the interaction between the two as slope terms in the model. All species-level parameters shared information among species and cities via their random effect structure. Following a 1,000 step adaptation phase and a 125,000 step burn-in, we sampled the posterior 120,000 times across 4 chains. We thinned chains by 3 for a total of 40,000 posterior samples. We assessed model convergence through a visual inspection of traceplots and ensured Gelman-Rubin diagnostics were < 1.10 (*36*). Following model convergence, we simulated species occupancy at each site across the entire study area from 5000 random samples of the occupancy model’s posterior distribution.

For the alpha diversity model we calculated 1) the expected species richness at each site and 2) the standard deviation in this estimate across the 5,000 posterior samples. To limit the effect of individual years on these estimates we calculated species richness at a site across all possible sampling periods. This resulted in one estimate per site across cities. We then fitted a varying intercept, varying slope log-linear model to these data, which treated species richness as the response variable but also incorporated the associated uncertainty in this estimate (*37*). Intercept and slope terms were treated as city-level random effects. We included impervious cover, gentrification, and the interaction between the two as covariates. Following a 1,000 step adaptation and 10,000 step burn-in phase, we sampled the posterior 160,000 times across 4 chains. We thinned chains by 2 for a total of 80,000 posterior samples.

For the beta diversity model we calculated 1) pairwise community dissimilarity between pairs of sites within each city (i.e., Sørensen’s dissimilarity index) and 2) the standard deviation in this estimate across the 5,000 posterior samples (*38*). Like the alpha diversity model, beta diversity estimates were made across all primary sampling periods. We then fitted a varying intercept, varying slope generalized dissimilarity model to these data, which treated pairwise dissimilarity between each pair of sites in a city as the response variable (*39,40*). This model used a clog link function and had an inverse link function of 1 − exp(−*µ*), where *µ* is the linear predictor for one data point. Similar to the alpha diversity model, the beta diversity model incorporated the associated uncertainty in the beta diversity estimate. Intercepts and slopes were treated as city-level random effects. Because community composition may be more similar in nearby sites, we included geographic distance between site pairs as a covariate. We also included differences in impervious cover and gentrification between sites as covariates. However, because this model uses I-spline basis functions to incorporate possible non-linear responses along environmental gradients we could not include an interaction between gentrification and impervious cover in this model (*39,40*). Following a 1,000 step adaptation and 2,000 step burn-in phase, we sampled the posterior 240,000 times across 4 chains. We thinned chains by 3 for a total of 80,000 posterior samples.

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All data and code are available on GitHub here: https://github.com/mfidino/uwin-gentrification. Pending acceptance of this manuscript we will archive all data and code on Zenodo and cite it with it’s DOI.