

# Mathematics for Data Science

## Lecture 3: Point Estimation

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Mohamad GHASSANY

EFREI Paris

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# Introduction to Statistical Inference

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- ▶ **Statistics** is the science of collecting, processing and analyzing data derived from the observation of random phenomena.
- ▶ Data analysis is used to **describe** the phenomena studied, **make predictions** and **make decisions** about them. In this way, statistics is an essential tool for understanding and managing complex phenomena.
- ▶ The data studied can be of any nature, which makes statistics useful in all disciplinary fields.

The fundamental point is that the data present uncertainties and **variations**.

Statistical methods are divided into two classes:

- ▶ **Descriptive statistics, exploratory statistics or data analysis**, aims to summarize the information contained in the data in a synthetic and efficient way. Probabilities play only a minor role here.
- ▶ **Inferential statistics** goes beyond the simple description of data. Its purpose is to **make predictions** and **make decisions** based on observations. In general, it is necessary to propose **probabilistic models** of the studied random phenomenon and to know how to manage the risks of errors. Probabilities play a fundamental role here.
- ▶ **Probability** can be considered as a branch of pure mathematics, based on the theory of measurement, abstract and completely disconnected from reality.
- ▶ **Applied probability** proposes **probabilistic models** of the course of concrete random phenomena. One can then, **prior to any experiment**, make predictions about what will happen.

**Example:** it is usual to model the duration of the good functioning or life of a system, let's say a light bulb, by a random variable  $X$  of exponential law of parameter  $\lambda$ . Having adopted this probabilistic model, we can perform all the calculations we want. For example:

- ▶ The probability that the bulb has not yet failed at date  $t$  is  $P(X > t) = e^{-\lambda t}$ .
- ▶ The average lifetime is  $E(X) = 1/\lambda$ .
- ▶ If  $n$  identical light bulbs are turned on at the same time, and they work independently of each other, the number  $N_t$  of light bulbs that will fail before a time  $t$  is a random variable of binomial distribution  $\mathcal{B}(n, P(X \leq t)) = \mathcal{B}(n, 1 - e^{-\lambda t})$ . Thus we expect that, on average,  $E(N_t) = n(1 - e^{-\lambda t})$  bulbs will fail between 0 and  $t$ .

In practice, if we want to use the theoretical results stated above, we have to make sure that we have chosen a good model, i.e. that the life span of these bulbs is a random variable with an exponential law, and, on the other hand, we have to be able to calculate the value of the parameter  $\lambda$  in some way. It is statistics that will allow us to solve these problems. To do this, we need to do an **experiment**, **collect data** and **analyze** them.

We therefore set up what we call a **test** or an **experiment**. We run  $n = 10$  identical bulbs in parallel and independently of each other, under the same experimental conditions, and we record their lifetimes. Let's say that we obtain the following lifetimes, expressed in hours: 91.6, 35.7, 251.3, 24.3, 5.4, 67.3, 170.9, 9.5, 118.4, 57.1

Let us note  $x_1, \dots, x_n$  these observations. We will therefore consider that  $x_1, \dots, x_n$  are the **samples** of random variables  $X_1, \dots, X_n$ .

This means that after the experiment, the lifetime has been observed. We say that  $x_i$  is a sample (a realization) of  $X_i$  on the test performed.

Since the bulbs are identical, it is natural to suppose that  $X_i$  have the same law. This means that the same random phenomenon is observed several times.

We can also assume that the  $X_i$  are independent random variables. We can then ask the following questions:

- ❶ With respect to these observations, is it reasonable to assume that the lifetime of a light bulb is a random variable with an exponential distribution? If not, what other law would be more appropriate? This is a **fit test** (Chi-square test) problem.
- ❷ If the exponential distribution model has been chosen, how can we propose a good value (or set of values) for the parameter  $\lambda$ ? This is a parametric **estimation** problem.
- ❸ In this case, can we guarantee that  $\lambda$  is less than a fixed value  $\lambda_0$ ? This will guarantee that  $E(X) = 1/\lambda \geq 1/\lambda_0$ , in other words that the bulbs will be sufficiently reliable. This is a **parametric hypothesis testing** problem.
- ❹ If we have 100 light bulbs, how many failures can we expect in less than 50 hours? This is a **prediction** problem.

## Definition: Random sample

The random variables  $X_1, X_2, \dots, X_n$  are a random sample of size  $n$  if

- a) the  $X_i$ 's are independent random variables
- b) every  $X_i$  has the same probability distribution.

An observation (realization) of the sample is  $(x_1, \dots, x_n)$ .

## Definition: Statistic

A **statistic** is any function of the observations in a random sample.

$$T(X) = T(X_1, \dots, X_n)$$

For example, each of  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ ,  $X_1^2$  or  $(X_1, X_3 + X_4, 2 \ln X_6)$  is a statistic.

Since a statistic is a random variable, it has a probability distribution.

The probability distribution of a statistic is called a *sampling distribution*.

If  $X_1, \dots, X_n$  is a random sample of size  $n$  taken from a population (either finite or infinite) with mean  $\mu$  and finite variance  $\sigma^2$ , and if  $\bar{X}_n$  is the sample mean, the limiting form of the distribution of

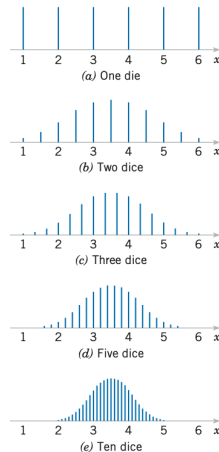
$$Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}}$$

as  $n \rightarrow \infty$ , is the **standard normal distribution**  $\mathcal{N}(0, 1)$ .

If we have two independent populations with means  $\mu_1$  and  $\mu_2$  and variances  $\sigma_1^2$  and  $\sigma_2^2$ , and if  $\bar{X}_1$  and  $\bar{X}_2$  are the sample means of two independent random samples of sizes  $n_1$  and  $n_2$  from these populations, then the sampling distribution of

$$Z = \frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\sigma_1^2/n_1 + \sigma_2^2/n_2}}$$

is **approximately standard normal** if the conditions of the central limit theorem apply. If the two populations are normal, the sampling distribution of  $Z$  is exactly standard normal.



*Figure 1: Distributions of average scores from throwing dice.*



**Definition: Point estimator**

A point estimate of some population parameter  $\theta$  is a single numerical value  $\hat{\theta}$  of a statistic  $T_n$ . The statistic  $T_n$  is called the point estimator.

A point estimator is a random variable. An estimation is a value.

Estimation problems occur frequently in engineering. We often need to estimate

- ▶ The mean  $\mu$  of a single population
- ▶ The variance  $\sigma^2$  (or standard deviation  $\sigma$ ) of a single population
- ▶ The proportion  $p$  of items in a population that belong to a class of interest
- ▶ The difference in means of two populations,  $\mu_1 - \mu_2$
- ▶ The difference in two population proportions,  $p_1 - p_2$

An estimator should be “close” in some sense to the true value of the unknown parameter.

Definition: Bias of an Estimator

The point estimator  $T_n$  is an unbiased estimator for the parameter  $\theta$  if

$$E(T_n) = \theta$$

If the estimator is not unbiased, then the difference  $E(T_n) - \theta$  is called the bias of the estimator  $T_n$ .

Formally, we say that  $T_n$  is an unbiased estimator of  $\theta$  if the expected value of  $T_n$  is equal to  $\theta$ . This is equivalent to saying that the mean of the probability distribution of  $T_n$  (or the mean of the sampling distribution of  $T_n$ ) is equal to  $\theta$ .

When an estimator is unbiased, the bias is zero; that is,  $E(T_n) - \theta = 0$ .

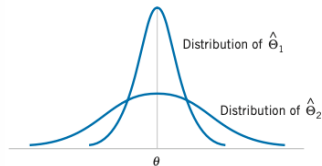
The bias measures a systematic error of estimation. If  $E(T_n) - \theta < 0$ ,  $T_n$  tends to under-estimate  $\theta$ .

Suppose that  $X$  is a random variable with mean  $\mu$  and variance  $\sigma^2$ . Let  $X_1, \dots, X_n$  be a random sample of size  $n$  from the population represented by  $X$ .

- ▶ Show that the sample mean  $\bar{X}_n$  is an unbiased estimator of  $\mu$ .
- ▶ Suggest an unbiased estimator of  $\sigma^2$ .

## Minimum Variance Unbiased Estimator

If we consider all unbiased estimators of  $\theta$ , the one with the smallest variance is called the **minimum variance unbiased estimator** (MVUE).



*Figure 2: The sampling distributions of two unbiased estimators.*

If  $X_1, \dots, X_n$  is a random sample of size  $n$  from a normal distribution with mean  $\mu$  and variance  $\sigma^2$ , the sample mean  $\bar{X}_n$  is the MVUE for  $\mu$ .

## Definition: Mean Squared Error of an Estimator

The mean squared error of an estimator  $T_n$  of the parameter  $\theta$  is defined as

$$MSE(T_n) = E[(T_n - \theta)^2]$$

The mean squared error can be rewritten as follows:

$$\begin{aligned} MSE(T_n) &= E[(T_n - \theta)^2] = E[(T_n - E(T_n) + E(T_n) - \theta)^2] \\ &= E[(T_n - E(T_n))^2] + 2E[T_n - E(T_n)]E[E(T_n) - \theta] + E[(E(T_n) - \theta)^2] \\ &= \text{Var}(T_n) + [E(T_n) - \theta]^2 \\ &= \text{Variance of the estimator} + \text{squared bias} \end{aligned}$$

## Methods of Point Estimation

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In this section, we discuss methods for obtaining point estimators: **the method of moments** and **the method of maximum likelihood**.

- ▶ Maximum likelihood estimates are generally preferable to moment estimators because they have better efficiency properties.
- ▶ However, moment estimators are sometimes easier to compute.
- ▶ Both methods can produce unbiased point estimators.

The general idea behind the method of moments is to **equate population moments**, which are defined in terms of expected values, to the corresponding **sample moments**.

The population moments will be functions of the unknown parameters. Then these equations are solved to yield estimators of the unknown parameters.

If  $E(X) = \phi(\theta)$ , where  $\phi$  is an invertible function, the moment estimator of  $\theta$  is  $\hat{\theta}_n = \phi^{-1}(\bar{X}_n)$ .

For example, if the parameter to estimate is the expected value of  $X_i$ , the **moment estimator** of  $E(X)$  is the sample mean  $\bar{X}_n$ .



## Examples:

- ▶ Exponential distribution
- ▶ Normal distribution
- ▶ Gamma distribution

Suppose that  $X$  is a random variable with probability distribution depending on a single unknown parameter  $\theta$ . Let  $x_1, \dots, x_n$  be the observed values in a random sample of size  $n$ . Then the **likelihood function** of the sample is

$$\mathcal{L}(\theta; x_1, \dots, x_n) = \begin{cases} P(X_1 = x_1, \dots, X_n = x_n; \theta) & \text{if } X_i \text{ are discrete} \\ f_{X_1, \dots, X_n}(x_1, \dots, x_n; \theta) & \text{if } X_i \text{ are continuous} \end{cases}$$

After supposing that all the  $X_i$  are independant:

$$\mathcal{L}(\theta; x_1, \dots, x_n) = \begin{cases} \prod_{i=1}^n P(X_i = x_i; \theta) = \prod_{i=1}^n P(X = x_i; \theta) & \text{if } X_i \text{ are discrete} \\ \prod_{i=1}^n f_{X_i}(x_i; \theta) = \prod_{i=1}^n f(x_i; \theta) & \text{if } X_i \text{ are continuous} \end{cases}$$

## Maximum likelihood estimator

Note that the likelihood function is now a function of only the unknown parameter  $\theta$ . The **maximum likelihood estimator (MLE)** of  $\theta$  is the value of  $\theta$  that maximizes the likelihood function  $\mathcal{L}(\theta)$  (or its  $\ln$ ).

## Examples:

- ▶ Bernoulli distribution
- ▶ Exponential distribution
- ▶ Normal distribution