We want to show that if (joint density)

X, Y ~ BN (Mx, My, 02, 03, p)

then (marginal densities)

X~ N(Mx, ox) Y~N(My, ox)

We show this by computing

 $f_{\mathbf{x}}(\mathbf{x}) = \int_{-\infty}^{\infty} f_{\mathbf{x}_{\mathbf{y}}}(\mathbf{x}, \mathbf{y}) d\mathbf{y}$

Mathematical Interlude Factoid 1: Completing the Square ax2+6x+c=0 ()

 $X^2 + \frac{b}{a}X + \frac{c}{a} = 0$

Take $\left(\frac{1}{2}\frac{b}{a}\right)^2$ add it in and substract it out to obtain the completed square in \times $\chi^2 + \frac{b}{a} \times + \frac{c}{a} + \frac{1}{4}\frac{b^2}{a^2} - \frac{1}{4}\frac{b^2}{a^2} = 0$

quare in X

$$X^{2} + \frac{b}{a}X + \frac{c}{a} + \frac{1}{4}\frac{b^{2}}{a^{2}} - \frac{1}{4}\frac{b^{2}}{a^{2}} = 0$$

$$x^{2} + \frac{6}{a}x + \frac{1}{4}\frac{6^{2}}{a^{2}} + \frac{c}{a} - \frac{1}{4}\frac{6^{2}}{a^{2}} = 0$$

$$\left(x + \frac{b}{2a}\right)^{2} = \frac{1}{4} \frac{b^{2}}{a^{2}} - \frac{c}{a} \implies x = -\frac{b}{2a} + \left(\frac{b^{2} - 4ac}{2a}\right)^{2} + \frac{b}{2a}$$
Tormula

Factord 2:

$$\int_{-\infty}^{\infty} e^{-\frac{1}{2}} \frac{x^2}{\sigma^2} \int dx = (z_{11}\sigma)^{\frac{1}{2}}$$
 the normalizing constant for a Gaussian density $\times n N(0, \sigma^2)$

Factoid 3:

 $\int_{-\infty}^{\infty} \exp\{-\frac{1}{2}\frac{(x-\mu)}{o^2}\} dx = \int_{-\infty}^{\infty} \exp\{-\frac{1}{2}\frac{x'^2}{o^2}\} dx$ x ∈ (-∞, ∞) X = X-M $X' \in (-\infty, \infty)$

Shifting x by u does not change the value of the integral

$$f_{X}(x) = \int_{-\infty}^{\infty} f(x,y) dy$$

$$= \left[2\pi \sigma_{X} \sigma_{Y} \left(1 - \rho_{2} \right)^{\frac{1}{2}} \right]^{-1} x \int_{-\infty}^{\infty} \exp \left\{ -\frac{1}{2} \left(1 - \rho_{2} \right) \left[\frac{(x - M_{X})^{2}}{\sigma_{X}^{2}} + \frac{(y - M_{Y})^{2}}{\sigma_{X}^{2}} - \frac{2}{\rho_{X}^{2} M_{X}^{2} (y - M_{Y})^{2}} \right] - 2\rho(x - M_{X}^{2} (y - M_{Y}^{2})^{2}) + 2\rho(x - M_{Y}^{2} (y - \rho_{X}^{2})^{2}) +$$

 $\int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2} \frac{[v-\rho u]^{2}}{(l-\rho^{2})}\right\} dv = \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2} \frac{v'^{2}}{(l-\rho^{2})}\right\} dv'$ Apply Factoid 2 $\int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2} \frac{v'^{2}}{(l-\rho^{2})}\right\} dv' = \left(2\pi(l-\rho^{2})\right)^{\frac{1}{2}}$

C.F.Q.D.

Hence

$$f_{X}(x) = \left(\frac{1}{2\pi\sigma_{X}^{2}}\right)^{\frac{1}{2}} e^{x} \rho \left\{-\frac{1}{2}u^{2}\right\} \cdot \frac{1}{\left[2\pi\left(1-\rho^{2}\right)\right]^{\frac{1}{2}}} \cdot \left[2\pi\left(1-\rho^{2}\right)\right]^{\frac{1}{2}}$$

$$= \left(\frac{1}{2\pi\sigma_{X}^{2}}\right)^{\frac{1}{2}} e^{x} \rho \left\{-\frac{1}{2}\left(\frac{x-\mu_{X}}{\sigma_{X}^{2}}\right)^{2}\right\}$$

$$Because u = \left(\frac{x-\mu_{X}}{\sigma_{X}}\right)$$

$$Q. E. D.$$