## FAS6337C - Lab 3

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## Growth Models in R.

Data for this laboratory are from two Florida Spotted Seatrout Cynoscion nebulosus populations (Indian River and Charlotte Harbor). The data set contains only females from each population. Fish were collected from 1986 to 1988 by Murphy and Taylor (1994, Transactions of the American Fisheries Society 123:482-497).

```
setwd("/workspaces/schooling/population_dynamics/lab_3/")
trout_data <- read.table("data/trout.txt", header=T, sep="")
head(trout_data)</pre>
```

```
##
                 bay tl
                            sex annuli age yearsold
## 1 CharlotteHarbor 387 FALSE
## 2 CharlotteHarbor 355 FALSE
                                     1
                                         1
                                               1.652
## 3 CharlotteHarbor 355 FALSE
                                         1
                                               1.652
## 4 CharlotteHarbor 320 FALSE
                                         1
                                     1
                                               1.652
## 5 CharlotteHarbor 335 FALSE
                                     1
                                         1
                                              1.652
## 6 CharlotteHarbor 410 FALSE
                                     1
                                          1
                                               1.652
```

The objectives of this laboratory are:

- 1. To determine the best growth model for each population and estimate  $L_{\infty}$ ,  $t_0$ , and k;
- 2. Estimate the mean length at age for each population;
- 3. Use an AIC model selection method to compare growth curves and parameters between populations.

Please conduct the following analyses in R and answer the questions. Each question or bullet point should be responded to either with text, a table, or a plot. Use trout.txt as your data file.

# 1. Use nonlinear least squares (nls) to estimate the von Bertalanffy growth parameters separately for Charlotte Harbor and for Indian River populations.

```
ch_data <- na.omit(
    trout_data[trout_data$bay == 'CharlotteHarbor',]
)
ch_data <- ch_data[order(ch_data$yearsold),]
ir_data <- na.omit(
    trout_data[trout_data$bay == 'IndianRiver',]
)
ir_data <- ir_data[order(ir_data$yearsold),]

do_least_squares_fit <- function(fish_data, Linf, vbk, tknot) {
    t1 <- as.numeric(fish_data$t1)
    yearsold <- as.numeric(fish_data$yearsold)
    result <- nls(
        t1 ~ Linf * (1 - exp(-vbk * (yearsold - tknot))),
        data=fish_data,</pre>
```

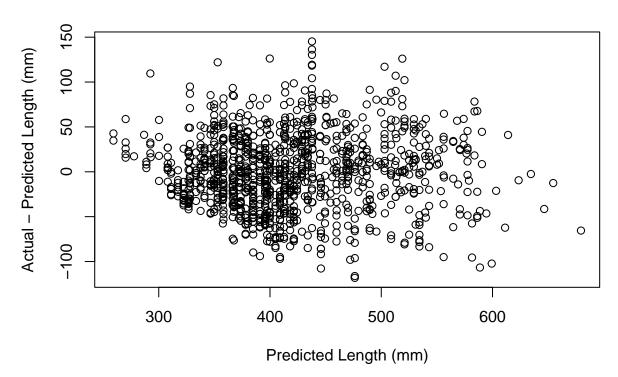
```
start=list(Linf=Linf, vbk=vbk, tknot=tknot)
 )
 return(result)
}
ch_result = do_least_squares_fit(ch_data, 800, 0.2, -1.4)
ir_result = do_least_squares_fit(ir_data, 950, 0.19, -1.2)
summary(ch_result)
##
## Formula: tl ~ Linf * (1 - exp(-vbk * (yearsold - tknot)))
## Parameters:
         Estimate Std. Error t value Pr(>|t|)
##
## Linf 788.88587 44.03635 17.914 < 2e-16 ***
                              7.651 4.34e-14 ***
## vbk
          0.20546
                   0.02685
## tknot -1.45233
                   0.17197 -8.445 < 2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 42.07 on 1099 degrees of freedom
##
## Number of iterations to convergence: 4
## Achieved convergence tolerance: 2.277e-06
summary(ir_result)
##
## Formula: tl ~ Linf * (1 - exp(-vbk * (yearsold - tknot)))
##
## Parameters:
        Estimate Std. Error t value Pr(>|t|)
## Linf 977.2375 63.8731 15.300 < 2e-16 ***
         0.1894
                  0.0261 7.259 7.03e-13 ***
## vbk
## tknot -1.2178
                     0.1550 -7.857 8.77e-15 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 58.34 on 1192 degrees of freedom
## Number of iterations to convergence: 2
## Achieved convergence tolerance: 5.089e-06
ch_data$nls_tl <- ch_result$m$predict(ch_data$yearsold)</pre>
ir_data$nls_tl <- ir_result$m$predict(ir_data$yearsold)</pre>
ch_data$nls_res <- ch_data$tl - ch_data$nls_tl</pre>
ir_data$nls_res <- ir_data$tl - ir_data$nls_tl</pre>
head(ch data)
##
                             sex annuli age yearsold nls_tl nls_res
                   bay tl
## 2150 CharlotteHarbor 302 FALSE
                                               0.487 259.2628 42.73723
                                      0
                                         0
## 2160 CharlotteHarbor 294 FALSE
                                               0.487 259.2628 34.73723
                                      0
                                          0
```

```
## 1165 CharlotteHarbor 329 FALSE 1 0 0.589 270.2466 58.75343
## 1188 CharlotteHarbor 303 FALSE 0 0 0.589 270.2466 32.75343
## 1189 CharlotteHarbor 286 FALSE 0 0 0.589 270.2466 15.75343
## 1190 CharlotteHarbor 296 FALSE 0 0 0.589 270.2466 25.75343
```

Plot the residuals against the predicted values from the nls for each population.

```
with(
  ch_data, {
    plot(
       nls_res ~ nls_tl,
       main='Charlotte Harbor Residuals Plot',
       xlab='Predicted Length (mm)',
       ylab='Actual - Predicted Length (mm)'
    )
}
```

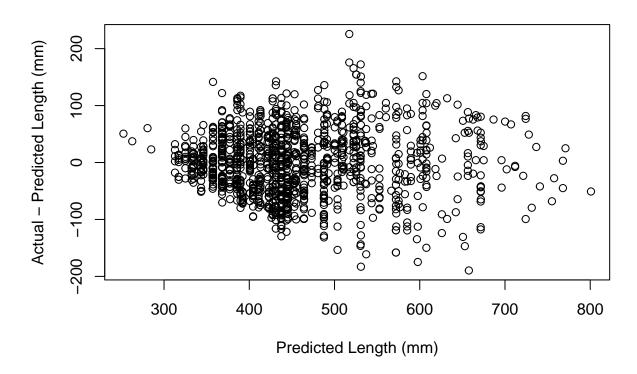
## **Charlotte Harbor Residuals Plot**



```
with(
  ir_data, {
   plot(
     nls_res ~ nls_tl,
     main='Indian River Residuals Plot',
     xlab='Predicted Length (mm)',
     ylab='Actual - Predicted Length (mm)'
)
```

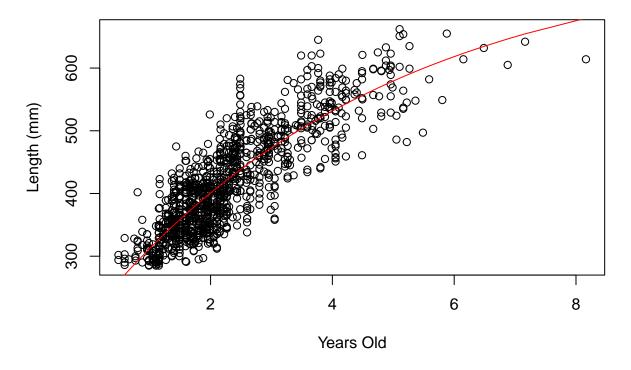
)

## **Indian River Residuals Plot**



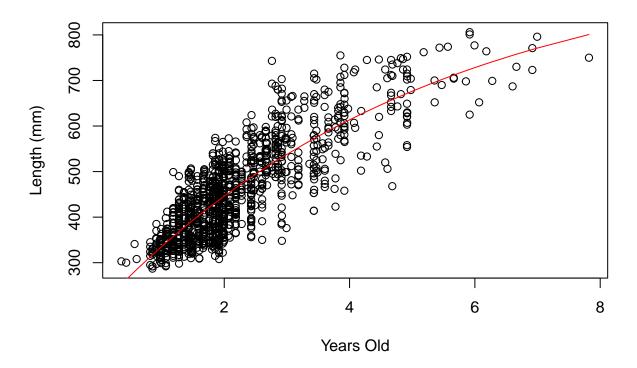
```
with(
  ch_data, {
    plot(
       tl ~ yearsold,
       main='Charlotte Harbor Fit',
       xlab='Years Old',
       ylab='Length (mm)'
    )
    lines(nls_tl ~ yearsold, col='red')
}
```

## **Charlotte Harbor Fit**



```
with(
  ir_data, {
  plot(
     tl ~ yearsold,
     main='Indian River Fit',
     xlab='Years Old',
     ylab='Length (mm)'
  )
  lines(nls_tl ~ yearsold, col='red')
}
```

## **Indian River Fit**



### Does it appear that a von Bertalanffy curve is reasonable for each population?

The conclusion here is the same as in the last lab. The fit is reasonable but far from ideal. There is a clear pattern from under predicting to overpredicting as length increases so there's definitely a pattern to the residuals that (in theory) could be captured in a better model.

There's also quite a lot of variance left in these estimates, and that variance itself is patterned with our gear selectivity creating variance toward smaller predicted lengths.

2. For the Charlotte Harbor population, create a function that estimates the negative log likelihood of the normal distribution. Minimize the negative log likelihood of the normal distribution to estimate the parameters of the von Bertalanffy growth equation.

```
predict_length <- function(yearsold, Linf, vbk, tknot) {
   pred_tl <- Linf * (1 - exp(-vbk * (yearsold - tknot)))
    return(pred_tl)
}

get_likelihood <- function(yearsold, tl, Linf, vbk, tknot, sig) {
   pred_tl <- predict_length(yearsold, Linf, vbk, tknot)
   NLL <- -1 * sum(dnorm(tl, pred_tl, sig, log=T), na.rm=T)
   return(NLL)
}

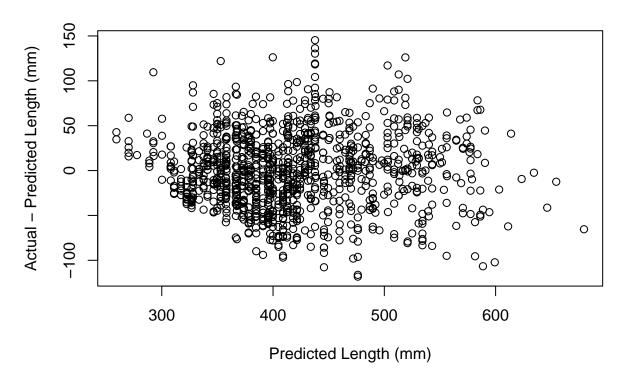
do_likelihood_fit <- function(fish_data, Linf, vbk, tknot, sig) {</pre>
```

```
tl <- as.numeric(fish_data$tl)</pre>
  yearsold <- as.numeric(fish_data$yearsold)</pre>
  lLinf <- log(Linf)</pre>
  lsig <- log(sig)</pre>
  objective <- function(v) {
    Linf \leftarrow \exp(v[1])
    vbk \leftarrow v[2]
    tknot \leftarrow v[3]
    sig \leftarrow exp(v[4])
    NLL <- get_likelihood(yearsold, tl, Linf, vbk, tknot, sig)</pre>
    return(NLL)
  }
  v <- c(lLinf, vbk, tknot, lsig)</pre>
  fit <- optim(v, objective, hessian=T)</pre>
  print("Fit Summary")
  print(fit)
  covm <- solve(fit$hessian)</pre>
  pred <- c(
    exp(fit$par[1]),
    fit$par[2],
    fit$par[3],
    exp(fit$par[4])
  )
  print(pred)
  return(list(pred, covm, fit$par))
ch_ml_result <- do_likelihood_fit(ch_data, 800, 0.2, -1.4, 40)</pre>
## [1] "Fit Summary"
## $par
## [1] 6.6703456 0.2055542 -1.4520493 3.7381685
##
## $value
## [1] 5683.029
## $counts
## function gradient
##
         157
                    NA
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
##
                                   [,2]
                                                   [,3]
                                                                [,4]
                   [,1]
## [1,] 112145.777336 348609.06203 -19240.314501
                                                           7.352657
```

Plot the residuals against the predicted values from the likelihood estimation

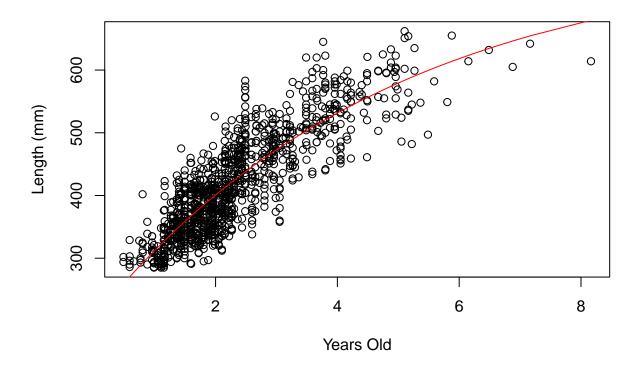
```
with(
  ch_data, {
    plot(
        ml_res ~ ml_tl,
        main='Charlotte Harbor Residuals Plot',
        xlab='Predicted Length (mm)',
        ylab='Actual - Predicted Length (mm)'
    )
}
```

## **Charlotte Harbor Residuals Plot**



```
with(
  ch_data, {
    plot(
       tl ~ yearsold,
       main='Charlotte Harbor Fit',
       xlab='Years Old',
       ylab='Length (mm)'
    )
    lines(ml_tl ~ yearsold, col='red')
}
```

## **Charlotte Harbor Fit**



Does the likelihood estimation appear to return VBGM parameter estimates that are a reasonable model for the data?

 $t_0$  is clearly off and the model seems to be overpredicting  $L_{\infty}$ . Furthermore there is clearly changing variance with both age and length so the  $\sigma$  parameter is also contentious. Overall the fit to the data given is reasonable.

Compare your estimates to the between the nls (question 1) and the likelihood estimation.

Besides having added  $\sigma$  this is the same fit. So the estimates are the same.

3. Generate the confidence intervals for each parameter from the model using the output from nls and the Hessian matrix from Optim.

Report the confidence intervals for each parameter from each method in a table

```
ml_stderr <- sqrt(diag(ch_covm))

ALPHA <- 0.05

ML_L95 <- (ch_ml_par - qnorm(1-(ALPHA/2)) * ml_stderr)

ML_U95 <- (ch_ml_par + qnorm(1-(ALPHA/2)) * ml_stderr)

ML_U95[1] <- exp(ML_U95[1])

ML_L95[1] <- exp(ML_L95[1])

ML_U95[4] <- exp(ML_U95[4])

ML_L95[4] <- exp(ML_U95[4])</pre>
```

```
ch_nls_par <- summary(ch_result)$parameters[,1]</pre>
ch_nls_stderr <- summary(ch_result)$parameters[,2]</pre>
NLS_L95 <- (ch_nls_par - qnorm(1-(ALPHA/2)) * ch_nls_stderr)</pre>
NLS_U95 <- (ch_nls_par + qnorm(1-(ALPHA/2)) * ch_nls_stderr)
NLS U95['sig'] <- NaN
NLS_L95['sig'] <- NaN
rbind(ML_U95, ML_L95, NLS_U95, NLS_L95)
##
               Linf
                           vbk
                                   tknot
## ML_U95 870.2427 0.2528076 -1.146851 43.81258
## ML L95 714.7402 0.1583008 -1.757248 40.30259
## NLS_U95 875.1955 0.2580933 -1.115282
                                               NaN
## NLS_L95 702.5762 0.1528279 -1.789379
                                               NaN
```

#### Compare the estimated confidence intervals between methods.

 $\sigma$  obviously offers no comparison as it is not in both methods.  $t_0$  and k are both very similar between the two methods.  $L_{\infty}$  is the only one that stands out but it was fit as the log of itself during maximum likelihood estimation and so it's expected that the confidence interval for  $L_{\infty}$  is asymmetric in the likelihood estimation.

# 4. Conduct the same analysis for the Indian River population (questions 1-3). Answer the same questions/bullet points.

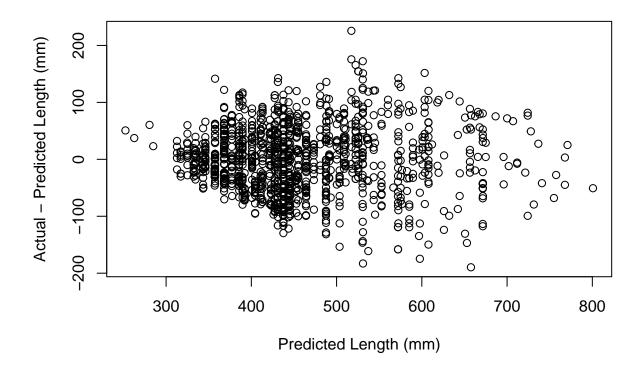
```
ir_ml_result <- do_likelihood_fit(ir_data, 950, 0.19, -1.2, 40)</pre>
## [1] "Fit Summary"
## $par
       6.8842492 0.1896156 -1.2168596 4.0649088
## [1]
##
## $value
## [1] 6553.262
##
## $counts
## function gradient
##
        177
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
##
                  [,1]
                                 [,2]
                                                            [,4]
## [1,] 76708.491347 276090.71644 -15324.743368
                                                        2.288774
## [2,] 276090.716445 1009606.09770 -57676.210383
                                                       11.120881
## [3,] -15324.743368 -57676.21038
                                        3507.467652
                                                       -0.400318
## [4,]
             2.288774
                            11.12088
                                          -0.400318 2390.262904
##
## [1] 976.7680293
                      0.1896156 -1.2168596
ir_pred_params <- ir_ml_result[[1]]</pre>
ir_covm <- ir_ml_result[[2]]</pre>
```

```
ir_ml_par <- ir_ml_result[[3]]
ir_data$ml_tl <- predict_length(
    ir_data$yearsold, ir_pred_params[1], ir_pred_params[2], ir_pred_params[3]
)
ir_data$ml_res <- ir_data$tl - ir_data$ml_tl</pre>
```

Plot the residuals against the predicted values from the likelihood estimation

```
with(
  ir_data, {
   plot(
     ml_res ~ ml_tl,
     main='Indian River Residuals Plot',
     xlab='Predicted Length (mm)',
     ylab='Actual - Predicted Length (mm)'
   )
}
```

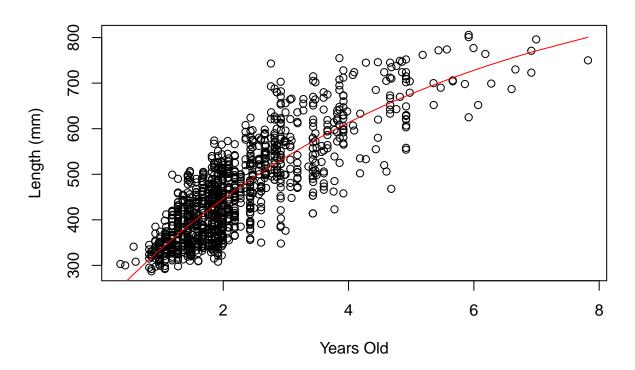
## **Indian River Residuals Plot**



```
with(
  ir_data, {
  plot(
    tl ~ yearsold,
    main='Indian River Fit',
    xlab='Years Old',
```

```
ylab='Length (mm)'
)
lines(ml_tl ~ yearsold, col='red')
}
)
```

## **Indian River Fit**



Does the likelihood estimation appear to return VBGM parameter estimates that are a reasonable model for the data?

 $t_0$  is clearly off again. There is still clearly changing variance with both age and length so the  $\sigma$  parameter is also contentious. Overall the fit to the data given is reasonable.

Compare your estimates to the between the nls (question 1) and the likelihood estimation.

Besides having added  $\sigma$  this is the same fit. So the estimates are the same.

Report the confidence intervals for each parameter from each method in a table

```
ml_stderr <- sqrt(diag(ir_covm))

ALPHA <- 0.05

ML_L95 <- (ir_ml_par - qnorm(1-(ALPHA/2)) * ml_stderr)

ML_U95 <- (ir_ml_par + qnorm(1-(ALPHA/2)) * ml_stderr)

ML_U95[1] <- exp(ML_U95[1])

ML_L95[1] <- exp(ML_L95[1])</pre>
```

```
ML_U95[4] \leftarrow exp(ML_U95[4])
ML_L95[4] \leftarrow exp(ML_L95[4])
ir_nls_par <- summary(ir_result)$parameters[,1]</pre>
ir_nls_stderr <- summary(ir_result)$parameters[,2]</pre>
NLS_L95 <- (ir_nls_par - qnorm(1-(ALPHA/2)) * ir_nls_stderr)
NLS_U95 <- (ir_nls_par + qnorm(1-(ALPHA/2)) * ir_nls_stderr)
NLS_U95['sig'] <- NaN
NLS_L95['sig'] <- NaN
rbind(ML_U95, ML_L95, NLS_U95, NLS_L95)
##
                 Linf
                            vbk
                                      tknot
                                                  sig
## ML_U95 1101.7064 0.2376684 -0.9299628 60.64261
## ML_L95
            865.9982 0.1415629 -1.5037564 55.97022
## NLS_U95 1102.4265 0.2405846 -0.9139983
                                                  NaN
## NLS_L95 852.0484 0.1382837 -1.5216089
                                                  NaN
```

#### Compare the estimated confidence intervals between methods.

It's interesting that here all parameters are extremely close in terms of bounds - even  $L_{\infty}$ . However the confidence interval for  $L_{\infty}$  is still asymmetric for the maximum likelihood estimation.

5. Using the likelihood-based model parameter estimates for each population, calculate the mean length at each age for age 0 to age 8 for each population.

Report the mean length at age for each population.

```
age <- seq(0, 8, 1)
ch_tl <- predict_length(
   age, ch_pred_params[1], ch_pred_params[2], ch_pred_params[3]
)
ir_tl <- predict_length(
   age, ir_pred_params[1], ir_pred_params[2], ir_pred_params[3]
)
cbind(age, ch_tl, ir_tl)</pre>
```

```
##
         age
                ch_tl
                         ir_tl
##
   [1,]
           0 203.5166 201.2637
   [2,]
           1 312.2402 335.2111
##
   [3,]
           2 400.7624 446.0227
##
   [4,]
           3 472.8369 537.6946
##
   [5,]
           4 531.5196 613.5326
##
   [6,]
           5 579.2988 676.2717
##
    [7,]
           6 618.2005 728.1743
##
   [8,]
           7 649.8741 771.1122
   [9,]
           8 675.6626 806.6337
```

#### Compare the length at age between the two populations.

Clearly Indian River is growing faster and larger than the Charlotte Harbor. It's also pretty unbiologically sound that at age 0 they are 200mm already given they should be more like 6mm.

## 6. Set up an AIC table:

```
map_columns <- function(v, cols) {</pre>
    c \leftarrow rep(0, 8)
    i <- 0
    for (col in cols) {
         i <- i + 1
         if (endsWith(col, 'Linf')) {
              j <- 1
         } else if (endsWith(col, 'vbk')) {
             j <- 3
         } else if (endsWith(col, 'tknot')) {
              j <- 5
         } else {
              j <- 7
         }
         if (startsWith(col, 'ch_')) {
             c[j] \leftarrow v[i]
         } else if (startsWith(col, 'ir_')) {
             c[j+1] \leftarrow v[i]
         } else {
             c[j] <- v[i]
              c[j+1] \leftarrow v[i]
    }
    return(c)
}
map_columns(c(1, 3, 2, 3, 4), c('ir_Linf', 'ch_Linf', 'vbk', 'tknot', 'sig'))
## [1] 3 1 2 2 3 3 4 4
do_likelihood_fit <- function(v, cols, runs) {</pre>
  objective <- function(v) {</pre>
    c <- map_columns(v, cols)</pre>
    ch_Linf <- exp(c[1])</pre>
    ch_vbk \leftarrow c[3]
    ch_tknot <- c[5]
    ch_sig \leftarrow exp(c[7])
    ir_Linf <- exp(c[2])</pre>
    ir_vbk \leftarrow c[4]
    ir_tknot <- c[6]</pre>
    ir_sig \leftarrow exp(c[8])
    ch_NLL <- get_likelihood(ch_data$yearsold, ch_data$tl, ch_Linf, ch_vbk, ch_tknot, ch_sig)</pre>
    ir_NLL <- get_likelihood(ir_data$yearsold, ir_data$tl, ir_Linf, ir_vbk, ir_tknot, ir_sig)</pre>
    NLL <- ch_NLL + ir_NLL</pre>
    return(NLL)
  }
```

```
for (i in 1:runs) {
    fit <- optim(v, objective, hessian=T)
    v <- fit$par
}
return(fit)
}

summarize_aic <- function(fit, cols) {
    dof <- length(cols)
    nll <- fit$value
    aic <- 2*nll + 2*dof
    row <- c(paste(cols, collapse=','), dof, nll, aic)
    return(row)
}</pre>
```

Determine the log likelihood, the number of parameters, and the AIC for:

```
v <- c(
  6.7, 0.2, -1.5, 3.7,
  6.9, 0.2, -1.2, 4.1
)
cols <- c(
  'ch_Linf', 'ch_vbk', 'ch_tknot', 'ch_sig',
  'ir_Linf', 'ir_vbk', 'ir_tknot', 'ir_sig'
)
(fit <- do_likelihood_fit(v, cols, 25))</pre>
```

A model with different parameters  $(L_{\infty}, t_0, k, \text{ and } \sigma)$  for each population.

```
## $par
## [1] 6.6698015 0.2058859 -1.4495751 3.7381019 6.8934140 0.1860532 -1.2361554
## [8] 4.0651624
##
## $value
## [1] 12236.3
##
## $counts
## function gradient
       215
##
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
                               [,2]
                 [,1]
                                             [,3]
                                                                         [,5]
## [1,] 112181.383961 347993.25653 -1.925144e+04 -6.426805e+00
                                                                     0.000000
## [2,] 347993.256531 1098196.84390 -6.250397e+04 -1.593776e+01
                                                                     0.00000
## [3,] -19251.438278 -62503.97185 3.756849e+03 9.573509e-01
                                                                     0.000000
## [4,]
           -6.426805
                         -15.93776 9.573509e-01 2.203885e+03
                                                                     0.000000
## [5,]
            0.000000
                            0.00000 0.000000e+00 0.000000e+00 76678.357183
```

```
## [6,]
            0.000000
                           0.00000 2.273737e-07 -2.273737e-07 282775.217390
## [7,]
            0.000000
                           0.00000 0.000000e+00 0.000000e+00 -15298.653825
## [8,]
            0.000000
                           0.00000 0.000000e+00 0.000000e+00
                                                                   -3.567621
##
                 [,6]
                               [,7]
                                            [,8]
## [1,] 0.000000e+00 0.000000e+00
                                      0.0000000
        0.000000e+00 0.000000e+00
                                    0.0000000
## [2,]
## [3.] 2.273737e-07 0.000000e+00
                                    0.0000000
## [4,] -2.273737e-07 0.000000e+00
                                    0.000000
## [5,] 2.827752e+05 -1.529865e+04 -3.5676208
## [6,] 1.058861e+06 -5.893007e+04
                                    6.2779386
## [7,] -5.893007e+04 3.489742e+03 -0.1100952
## [8,] 6.277939e+00 -1.100952e-01 2389.0876109
(all_free <- summarize_aic(fit, cols))</pre>
## [1] "ch_Linf,ch_vbk,ch_tknot,ch_sig,ir_Linf,ir_vbk,ir_tknot,ir_sig"
## [2] "8"
## [3] "12236.300720374"
## [4] "24488.6014407481"
v <- c(
 6.7, 0.2, -1.5, 3.7
cols <- c(
 'Linf', 'vbk', 'tknot', 'sig'
(fit <- do_likelihood_fit(v, cols, 25))</pre>
A model with shared parameters for both populations.
## $par
## [1] 6.8171686 0.1797327 -1.4666141 4.0528655
##
## $value
## [1] 12569.06
##
## $counts
## function gradient
##
        99
                 NA
##
## $convergence
## [1] 0
##
```

**##** [1,] 137876.9646 520338.81924 -25168.801018

**##** [2,] 520338.8192 1990716.25051 -98792.191938

[,1]

**##** [3,] -25168.8010 -98792.19194

-7.3917

[,2]

-18.50501

## \$message ## NULL ##

## \$hessian

##

## [4,]

[,3]

5182.480004

[,4]

-7.391700

1.192387

-18.505014

1.192387 4595.299049

```
## [1] "Linf,vbk,tknot,sig" "4" "12569.0570197932" ## [4] "25146.1140395863"
```

### Which hypothesis had the most support based on AIC?

As in the last lab, all parameters shared has the most support.

## 7. Continue adding to your AIC table.

Determine the log likelihood function, the log likelihood, the number of parameters, and the AIC for

```
v <- c(
  6.7, 0.2, -1.5, 3.7,
  0.2, -1.2, 4.1
)
cols <- c(
  'Linf', 'ch_vbk', 'ch_tknot', 'ch_sig',
  'ir_vbk', 'ir_tknot', 'ir_sig'
)
(fit <- do_likelihood_fit(v, cols, 25))</pre>
```

Models that share only one of the VBGM parameters  $(L_{\infty}, t_0, k, \text{ and } \sigma)$  for each population (hint: this should result in four additional models for each population).

```
## $par
       6.8068761 0.1532091 -1.8059854 3.7401227 0.2239987 -1.0478161 4.0655605
## [1]
##
## $value
## [1] 12239.52
##
## $counts
## function gradient
        192
##
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
##
                [,1]
                              [,2]
                                            [,3]
                                                           [,4]
                                                                         [,5]
## [1,] 188324.30850
                     5.088257e+05 -1.887554e+04 -5.441198e+01 2.218592e+05
## [2,] 508825.67475 2.337803e+06 -8.853964e+04 8.041794e+00 -9.094947e-07
## [3,] -18875.53594 -8.853964e+04
                                    3.518782e+03 3.565174e-01
                                                                0.000000e+00
## [4,]
           -54.41198 8.041794e+00
                                    3.565174e-01 2.204096e+03 -9.094947e-07
## [5,] 221859.22437 -9.094947e-07
                                    0.000000e+00 -9.094947e-07
                                                                6.574851e+05
## [6,] -15462.87936
                      0.000000e+00
                                    0.000000e+00 0.000000e+00 -4.738922e+04
## [7,]
            54.78169
                      0.000000e+00
                                    0.000000e+00 0.000000e+00 2.943767e+00
                 [,6]
                               [,7]
## [1,] -1.546288e+04
                      5.478169e+01
## [2,]
        0.000000e+00
                       0.000000e+00
## [3,]
        0.000000e+00 0.000000e+00
## [4,]
        0.000000e+00 0.000000e+00
```

```
## [5,] -4.738922e+04 2.943767e+00
## [6,] 3.650500e+03 -2.950856e-03
## [7,] -2.950856e-03 2.390953e+03
(Linf_shared <- summarize_aic(fit, cols))</pre>
## [1] "Linf,ch_vbk,ch_tknot,ch_sig,ir_vbk,ir_tknot,ir_sig"
## [2] "7"
## [3] "12239.5224904294"
## [4] "24493.0449808588"
v <- c(
 6.7, 0.2, -1.5, 3.7,
 6.9, -1.2, 4.1
)
cols <- c(
 'ch_Linf', 'vbk', 'ch_tknot', 'ch_sig',
 'ir_Linf', 'ir_tknot', 'ir_sig'
(fit <- do_likelihood_fit(v, cols, 25))</pre>
## $par
## [1] 6.6866966 0.1977578 -1.5017712 3.7381292 6.8648702 -1.1711513 4.0649778
##
## $value
## [1] 12236.39
##
## $counts
## function gradient
       180
##
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
                               [,2]
##
                 [,1]
                                             [,3]
                                                         [, 4]
                                                                       [,5]
## [1,] 1.121581e+05 366804.57392 -1.918373e+04
                                                  4.9503260 -4.547474e-07
## [2,] 3.668046e+05 2126473.25099 -6.550165e+04
                                                   39.7806680 2.615699e+05
## [3,] -1.918373e+04 -65501.65115 3.713248e+03
                                                    0.1034027 4.547474e-07
## [4,] 4.950326e+00
                          39.78067 1.034027e-01 2203.9542305 0.000000e+00
## [5,] -4.547474e-07 261569.89134 4.547474e-07 0.0000000 7.669719e+04
## [6,] -2.273737e-07 -54995.95528 2.273737e-07
                                                    0.0000000 -1.538221e+04
## [7,] 0.00000e+00
                         -10.83637 0.000000e+00
                                                    0.0000000 2.695255e+00
                 [,6]
                             [,7]
## [1,] -2.273737e-07
                        0.000000
## [2,] -5.499596e+04 -10.836374
## [3,] 2.273737e-07
                      0.000000
## [4,] 0.00000e+00
                      0.000000
## [5,] -1.538221e+04
                        2.695255
## [6,] 3.551922e+03 -1.031214
## [7,] -1.031214e+00 2390.151049
```

```
(vbk_shared <- summarize_aic(fit, cols))</pre>
## [1] "ch_Linf,vbk,ch_tknot,ch_sig,ir_Linf,ir_tknot,ir_sig"
## [2] "7"
## [3] "12236.3934261983"
## [4] "24486.7868523965"
v <- c(
 6.7, 0.2, -1.5, 3.7,
  6.9, 0.2, 4.1
cols <- c(
  'ch_Linf', 'ch_vbk', 'tknot', 'ch_sig',
  'ir_Linf', 'ir_vbk', 'ir_sig'
(fit <- do_likelihood_fit(v, cols, 25))</pre>
## $par
## [1] 6.6411004 0.2216115 -1.3462903 3.7384119 6.9343917 0.1699523 4.0654174
##
## $value
## [1] 12236.86
##
## $counts
## function gradient
##
       194
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
##
                               [,2]
                                           [,3]
                                                        [,4]
                 [,1]
## [1,] 1.121026e+05 315760.971000 -19409.8411
                                                  -0.5502118 0.000000e+00
## [2,] 3.157610e+05 907209.514896 -57515.5350
                                                 9.6947774 0.000000e+00
## [3,] -1.940984e+04 -57515.534999
                                    7230.1371 -10.1015044 -1.510730e+04
## [4,] -5.502118e-01
                           9.694777
                                       -10.1015 2203.4393414 0.000000e+00
## [5,] 0.00000e+00
                           0.000000 -15107.3000
                                                   0.0000000 7.663296e+04
## [6,]
        0.000000e+00
                           0.000000 -64885.4245
                                                   0.0000000 3.167108e+05
## [7,] 0.00000e+00
                           0.000000
                                        10.1164
                                                   0.0000000 1.951494e-01
##
                 [,6]
                              [,7]
## [1,] 0.00000e+00
                        0.0000000
## [2,]
        0.000000e+00
                       0.0000000
## [3,] -6.488542e+04
                       10.1163982
        0.000000e+00
## [4,]
                         0.0000000
## [5,]
        3.167108e+05
                         0.1951494
## [6,]
        1.325922e+06
                         5.3842418
## [7,] 5.384242e+00 2389.1819192
(tknot_shared <- summarize_aic(fit, cols))</pre>
## [1] "ch_Linf,ch_vbk,tknot,ch_sig,ir_Linf,ir_vbk,ir_sig"
```

## [2] "7"

```
## [3] "12236.8592624778"
## [4] "24487.7185249557"
v <- c(
 6.7, 0.2, -1.5, 3.7,
 6.9, 0.2, -1.2
cols <- c(
 'ch_Linf', 'ch_vbk', 'ch_tknot', 'sig',
 'ir_Linf', 'ir_vbk', 'ir_tknot'
(fit <- do_likelihood_fit(v, cols, 25))</pre>
## $par
## [1] 6.6676369 0.2069044 -1.4436927 3.9341872 6.8839452 0.1897212 -1.2165016
##
## $value
## [1] 12295.98
##
## $counts
## function gradient
##
       182
                 NA
##
## $convergence
## [1] 0
##
## $message
## NULL
##
## $hessian
##
                 [,1]
                               [,2]
                                             [,3]
                                                           [,4]
                                                                         [,5]
## [1,] 7.578281e+04 2.335499e+05 -1.300820e+04 -0.55997748 0.000000e+00
## [2,] 2.335499e+05 7.323867e+05 -4.197229e+04 -2.61466539 -4.547474e-07
## [3,] -1.300820e+04 -4.197229e+04 2.540665e+03
                                                  0.22086579 4.547474e-07
## [4,] -5.599775e-01 -2.614665e+00 2.208658e-01 4593.40523503 2.110082e+00
## [5,] 0.000000e+00 -4.547474e-07 4.547474e-07 2.11008182 9.963053e+04
## [6,] -2.273737e-07 4.547474e-07 -2.273737e-07 9.33348656 3.583302e+05
## [7,] -4.547474e-07 0.000000e+00 4.547474e-07
                                                   -0.07147719 -1.990348e+04
##
                 [,6]
                               [,7]
## [1,] -2.273737e-07 -4.547474e-07
## [2,] 4.547474e-07 0.000000e+00
## [3,] -2.273737e-07 4.547474e-07
## [4,] 9.333487e+00 -7.147719e-02
## [5,] 3.583302e+05 -1.990348e+04
## [6,] 1.309389e+06 -7.485600e+04
## [7,] -7.485600e+04 4.555389e+03
(sig shared <- summarize aic(fit, cols))</pre>
## [1] "ch_Linf,ch_vbk,ch_tknot,sig,ir_Linf,ir_vbk,ir_tknot"
## [2] "7"
## [3] "12295.979579067"
## [4] "24605.959158134"
```

8. Calculate the  $\triangle$ AIC for all the models you trialed (report your AIC table to answer questions 6, 7, and 8).

```
aic_table <- data.frame(rbind(</pre>
    all_free,
    all_shared,
    Linf_shared,
    vbk_shared,
    tknot_shared,
    sig_shared
))
colnames(aic table) <- c('cols', 'dof', 'nll', 'aic')</pre>
aic table aic <- as.numeric (aic table aic)
aic_table$delta <- aic_table$aic - min(aic_table$aic)</pre>
(aic_table <- aic_table[order(aic_table$delta),])</pre>
##
                                                                             cols dof
## vbk shared
                           ch_Linf,vbk,ch_tknot,ch_sig,ir_Linf,ir_tknot,ir_sig
## tknot_shared
                             ch_Linf,ch_vbk,tknot,ch_sig,ir_Linf,ir_vbk,ir_sig
## all free
                ch_Linf,ch_vbk,ch_tknot,ch_sig,ir_Linf,ir_vbk,ir_tknot,ir_sig
                                                                                    8
## Linf_shared
                            Linf,ch_vbk,ch_tknot,ch_sig,ir_vbk,ir_tknot,ir_sig
                                                                                    7
## sig_shared
                           ch_Linf,ch_vbk,ch_tknot,sig,ir_Linf,ir_vbk,ir_tknot
                                                                                    7
## all_shared
                                                              Linf, vbk, tknot, sig
##
                              nll
                                        aic
                                                  delta
## vbk_shared
                12236.3934261983 24486.79
                                              0.0000000
## tknot_shared 12236.8592624778 24487.72
                                              0.9316726
## all_free
                  12236.300720374 24488.60
                                              1.8145884
## Linf_shared 12239.5224904294 24493.04
                                              6.2581285
## sig_shared
                 12295.979579067 24605.96 119.1723057
## all_shared
                12569.0570197932 25146.11 659.3271872
```

### Which hypothesis had the most support based on AIC?

Sharing k had the best support.

#### What are the top models based on the AIC?

Sharing k, sharing  $t_0$ , sharing  $L_{\infty}$ , and all free were the best.

#### Do these top models make sense given the biology of Spotted Seatrout?

Sharing  $t_0$  makes sense but the  $t_0$  here is certainly unbiological. Growth rates would definitely change from place to place, but I think we're still able to fit reasonably well because  $t_0$  and  $L_{\infty}$  are still free in the shared k case and we know these are all extremely correlated with one another. Same argument for  $L_{\infty}$ , I'd expect it to be different but the model is likely taking advantage of the other parameters to still fit. In general across these models the wild  $t_0$  makes these rather unbiological.

# 9. Adjust the slot limits for the following scenarios and answer the following questions:

MinLL	MaxLL	Bias	Over	Under	Parameters
300 (0.05)	$\infty$	< 2 years old undersampled	< 4	N/A	$L_{\infty} = 914, k = 0.178, t_0 = -1.09, \sigma = 110.0$
0	700 (0.05)	> 3 years old undersampled	0 < age < 3	> 3	$L_{\infty} = 667, k = 0.335, t_0 = -0.39, \sigma = 66.2$
380 (0)	480 (0)	Almost no sampling for age > 6 or < 2, extremely poor sampling across the board.	< 3	> 3	$L_{\infty} = 446, k = 0.939, t_0 = -1.15, \sigma = 26.4$
380 (0)	530 (0.03)	No sampling for age < 2 biased toward middling or smaller fish throughout.	< 3	> 3	$L_{\infty} = 532, k = 0.222, t_0 = -5.56, \sigma = 40.2$

## Across scenarios, which had the biggest impact on $L_{\infty}$ ? on k? on $t_0$ ?

- $L_{\infty}$  Definitely case 3, the total lack of sampling of large fish threw things off entirely.
- k Likewise case 3.
- $t_0$  case 4, the complete lack of samples on the low end with samples on the high end gave us a weird almost straight relationship between age and growth.