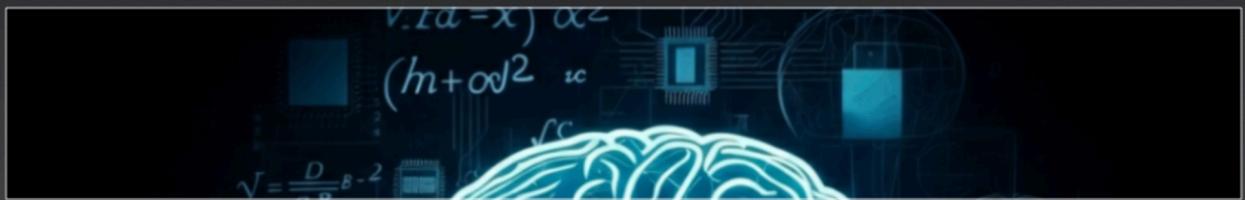


MODELLING NEURAL SYSTEMS



COMPUTATIONAL MODELLING OF NEURONS AND MICROCIRCUITS

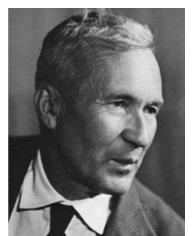
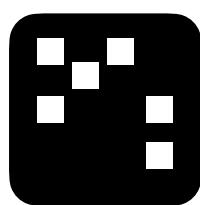
Prof. Ing. Michele GIUGLIANO, PhD

Stochastic processes

Probability Theory - *Experiment*

- The set Ω of **all possible outcomes of an experiment** is called the **sample space** for the experiment.
- e.g., *Tossing a coin*: $\Omega = \{ H, T \}$.
- e.g., *Tossing 2 coins*: $\Omega = \{ HH, HT, TH, TT \}$.
- An **event** A : **any collection of possible outcomes** of an exp.
An event can be a subset of Ω , it can be Ω itself, it can be the null set \emptyset .
- *Tossing 2 coins*: e.g., $A = \{ HH, HT \}$ (i.e. event “*first coin is heads*”).

Ω



A. Kolmogorov

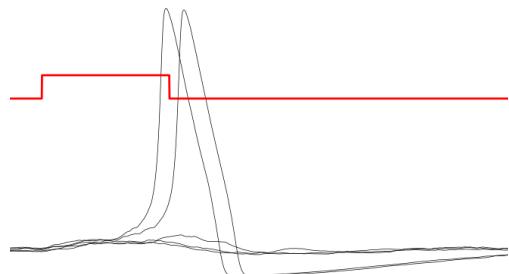
Probability Theory - Experiment

- **Experiment:** I show very briefly a picture of a dog to a monkey, as a stimulus. Then, I observe - during and immediately after such a presentation - a specific neuron (in the IT cortex) and whether it responds with APs or remains silent.



Ω

It fires It stays at rest



- **Probability:** something I associate to an event.

Probability: relative frequency definition

$$P(A) = \lim \frac{\text{num of times } A \text{ can occur}}{\text{total num of trials}}$$

- **Def: probability** is the long-run *relative frequency* of an event, **after a very large number of (independent, repeated) trials.**
- **Calculation:** by dividing num. of times an event occurs by total num of trials.
- **Nature:** Empirical or *a posteriori*, as it is derived from experience.
- **Usefulness:** Practical for real-world problems, where probabilities need to be estimated from data (e.g. insurance company analysing car accidents).
- **0** means the event is almost impossible, and a probability of **1** almost certain
- **Problems with this definition:** An actual limiting value may not exist. Moreover, it relies on having *enough* data for a *reliable* estimate.

Probability: axiomatic definition

- **Def:** Probability is a **function** that assigns a number **p** to each possible event **A** in a sample space Ω . **p** satisfies three fundamental **axioms**:

- Axiom 1 (*non-negativity*): $p(A) \geq 0$

- Axiom 2 (*certainty*): $p(\Omega) = 1$

- Axiom 3 (*additivity*): $p(A \cup B \cup C) = p(A) + p(B) + p(C)$

(for any collection of mutually exclusive, i.e. disjoint, events)

- It follows that $P(c_A) = 1 - A$ and that $P(\emptyset) = 0$

- e.g., *Tossing a coin*: $\Omega = \{H, T\}$, $P(H) = 1/3$, $P(T) = 2/3$

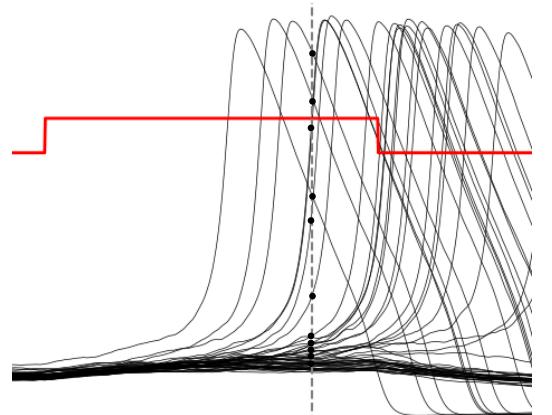
Random Variables (RV)

- a RV **X** : a **function** that for each element of Ω assigns a (real) number.
- It is a *numerical summary* of a random outcome of an experiment.
- **Notations:** we denote RVs by uppercase letters such as **X**, and use lowercase such as **x** for *potential* values and *realized* values.
- **Watch out!** a RV and its *realized* value are NOT the same thing.
The former is a **function** (from outcomes to values).
The latter is a **value**, associated with a specific outcome.
- For a RV **X**, we define its **cumulative distribution function** (CDF) as

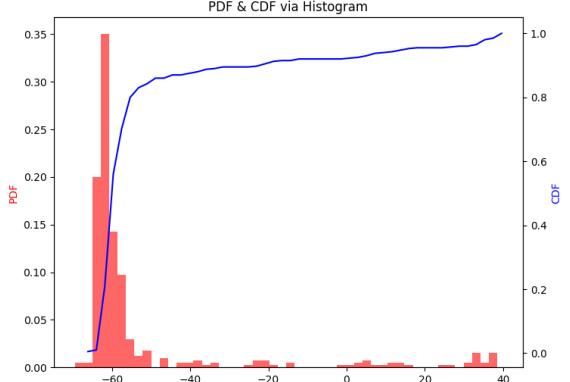
$$F_X(x) = p(X \leq x)$$

Random Variables (RV)

- **Experiment:** I show very briefly a picture of a dog to a monkey, as a stimulus. Then, I monitor during and immediately after such a presentation, a specific neuron (in the monkey's IT cortex) responds with APs or remains silent.



- **RV:** The numerical value of the membrane potential *4 ms* following the stimulus' onset.

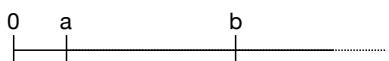


Random Variables (RV) - CDF

- $F_X(x) = p(X \leq x)$ **Cumulative distribution function**

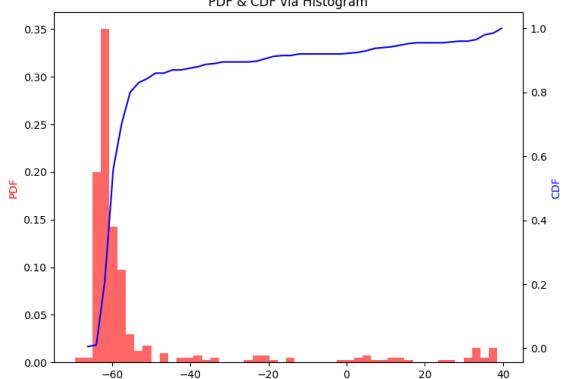
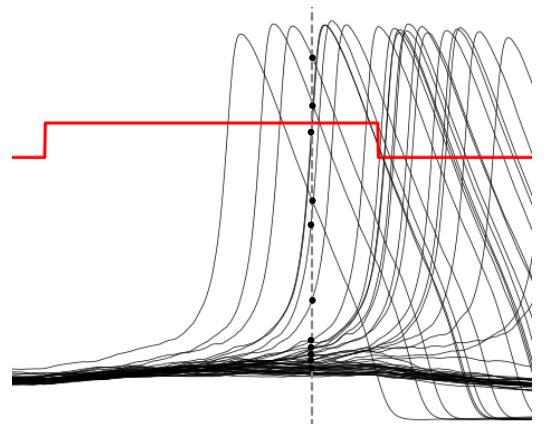
$$\bullet 0 \leq F_X(x) \leq 1$$

$$\bullet \lim_{x \rightarrow -\infty} F_X(x) = 0 \quad \lim_{x \rightarrow +\infty} F_X(x) = 1$$



$$P(X \leq b) = P(X \leq a) + P(a < X \leq b)$$

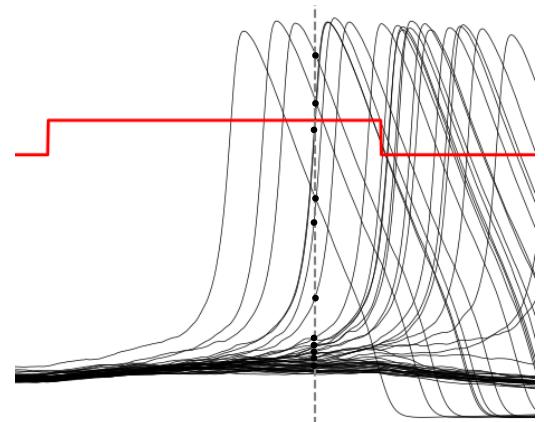
$$\bullet P(a < X \leq b) = F_X(b) - F_X(a)$$



Random Variables (RV) - PDF

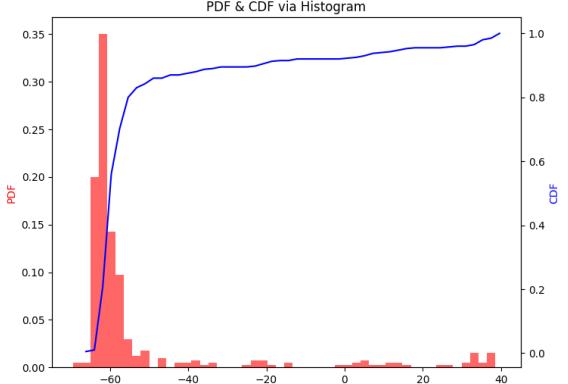
- $P(x < X \leq x + dx) = F_X(x + dx) - F_X(x)$

- $f_X(x) = \frac{dF_X(x)}{dx}$ **Probability (distribution) density function**



- $P(x < X \leq x + dx) = dx f_X(x)$

- $\int_{-\infty}^{+\infty} f_X(x)dx = 1 \quad f_X(x) > 0$



Discrete-valued Random Variables

- a discrete-valued RV X : a **function** that take only finite (or countably many) values, x_1, x_2, \dots, x_N (N could be infinite, but it is NOT a continuum).
- probabilities takes the form

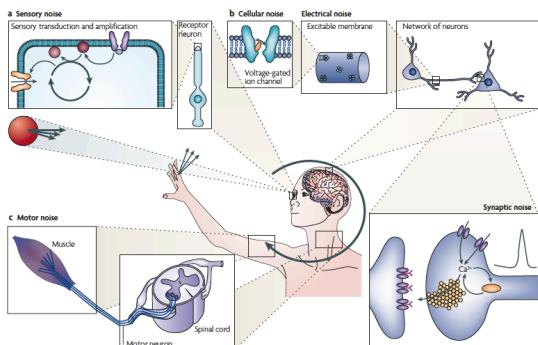
$$p(X = x_1) = p_1 \quad p(X = x_2) = p_2 \quad \dots, p(X = x_N) = p_N$$

$$p_1 + p_2 + \dots + p_N = 1$$

- For a discrete values RV X , its **CDF** $F_X(x)$ is a **step (with values 0 or p_i)**

“Noise” and “random variables” in the brain

J. Physiol. (1952), 117, 109–128



SPONTANEOUS SUBTHRESHOLD ACTIVITY AT MOTOR NERVE ENDINGS

BY P. FATT AND B. KATZ

From the Biophysics and Physiology Departments, University College, London

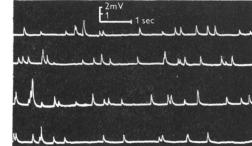
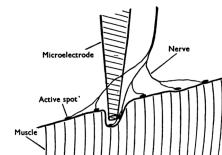


Fig. 2. Example of miniature e.p.p.'s in a muscle treated with 10^{-4} prostigmine bromide.

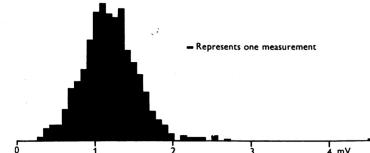


Fig. 13. Distribution of amplitudes. Muscle treated with prostigmine. Same 800 miniature potentials as used in Figs. 11 and 12.

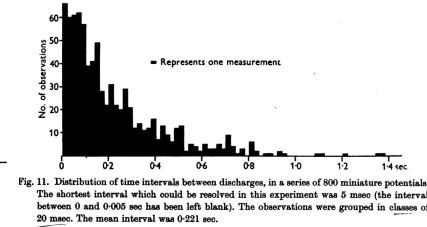


Fig. 11. Distribution of time intervals between discharges, in a series of 800 miniature potentials. The shortest interval which could be resolved in this experiment was 5 msec (the interval between 0 and 0.005 sec has been left blank). The observations were grouped in classes of 20 msec. The mean interval was 0.221 sec.

Faisal et al., 2008

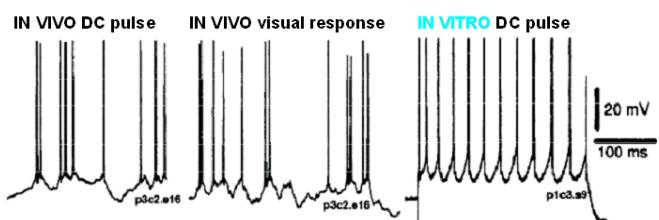
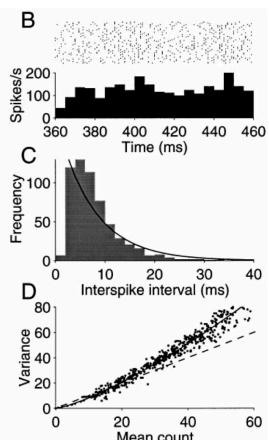
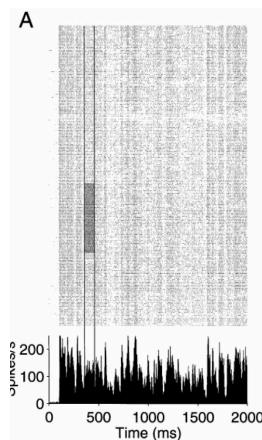
Fatt & Katz, 1952

Neuronal Firing appears to be irregular

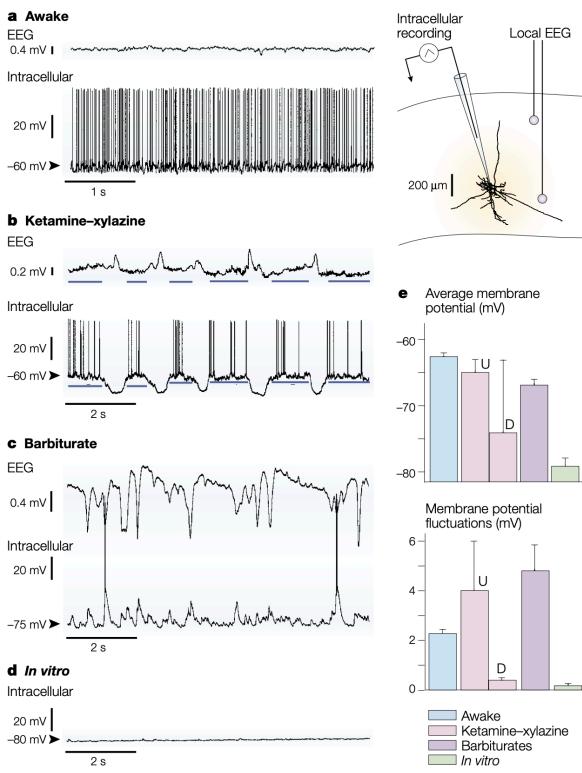
The Journal of Neuroscience, May 15, 1998, 18(10):3870-3896

The Variable Discharge of Cortical Neurons: Implications for Connectivity, Computation, and Information Coding

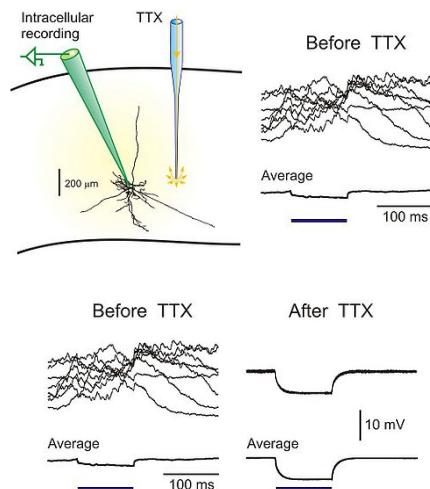
Michael N. Shadlen¹ and William T. Newsome^{2*}
¹Department of Physiology and Biophysics and Regional Primate Research Center, University of Washington, Seattle, Washington 98195-7290, and ²Howard Hughes Medical Institute and Department of Neurobiology, Stanford University School of Medicine, Stanford, California 94305



Holt, Softky, Koch, Douglas, J. Neurophysiol. 1996
Shadlen & Newsome, 1998



Synaptic inputs seems to be its culprit

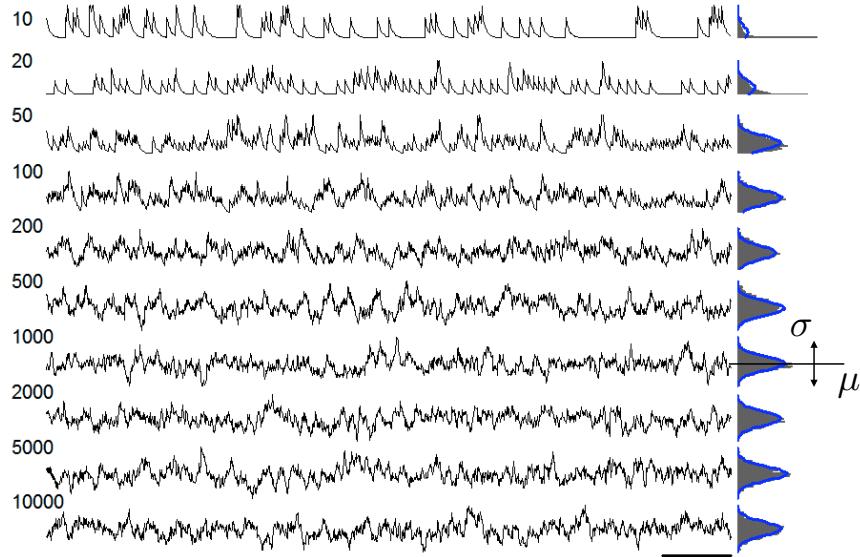


Destexhe et al., 2003

Plan for the day

- Refresher on stochastic processes
- Ornstein-Uhlenbeck's stochastic processes
- Stein's stochastic process
- Diffusion approximation

The *diffusion approximation*



Giugliano et al., 2008, Biological Cybernetics
Arsiero et al., 2007, J Neuroscience
Giugliano et al., 2004, J Neurophysiology

Probability Theory and Random Variables

- **Random Variable (RV):** A variable whose value is subject to variations due to chance (e.g., the mem. potential of a neuron at one moment in time).
- **Stochastic Process (of time) $X(t)$:** a collection of RVs, indexed **by time**.
 - At every (time) point t , $X(t)$ is a RV.
- **Trajectory (Realization):** One specific "run" of the experiment.

Gaussian White "noise" $\xi(t)$

- $\xi(t)$ is a continuous-time, continuous-valued stochastic process (of time)

- $\xi(t)$, at every t , is a RV whose prob. density function (PDF) is a Gaussian

$$f_\xi(Z) = \frac{1}{\sqrt{2\pi}} e^{-Z^2/2} \quad \text{mean is \textbf{zero}, variance is \textbf{unitary}}$$

autocorrelation func. is a Dirac's **delta**

- $\langle \xi(t) \rangle = 0, \quad \langle \xi(t)^2 \rangle = 1, \quad \langle \xi(t)\xi(t+T) \rangle = \delta(T)$

- **It does NOT exist in nature!** Real (biological) noise has a **finite** (auto-correlation) time scale (i.e., the "color" of noise). It is a building block.

Ornstein-Uhlenbeck process - (my) definition

- is a continuous-time, continuous-valued stochastic process (of time)

- I define it (abusing the mathematical notation of o.d.e.) as

$$\tau \frac{d}{dt}x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t) \quad \text{with } x(t_0) = x_0$$

(deterministically)

- It is an abuse because $\xi(t)$ is not a differentiable "function" (it not at all a function of time), but ...hey, I am a bioengineer!
- Beware: there are slightly different definitions in the literature. You will understand in a moment why I like this one.

Ornstein-Uhlenbeck process - properties

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

- $m(t) = \langle x(t) \rangle = x_0 e^{-(t-t_0)/\tau} + \mu (1 - e^{-(t-t_0)/\tau}) \rightarrow \mu,$
- $s(t)^2 = \langle [x(t) - m(t)]^2 \rangle = \sigma^2 (1 - e^{-2(t-t_0)/\tau}) \rightarrow \sigma^2,$
- $Cov_x(t, T) = \langle [x(t) - m(t)][x(t+T) - m(t+T)] \rangle =$
 $= \sigma^2 e^{-|T|/\tau} (1 - e^{-2(t-t_0)/\tau}) \rightarrow \sigma^2 e^{-|T|/\tau}$

Ornstein-Uhlenbeck process
and ...the Leaky Integrate-and-Fire?!??!?

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

$$C \frac{dV}{dt} = G (E - V) + I_{ext}$$

They look the same!

How do I... simulate it?

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

$$\xrightarrow{\hspace{10cm}} t$$
$$\hat{x}(t) \quad \hat{x}(t + \Delta t) \quad \hat{x}(t + 2\Delta t) \quad \hat{x}(t + 3\Delta t) \dots$$



Ornstein-Uhlenbeck process - discrete time version

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

$$\hat{x}(t + \Delta t) = e^{-\Delta t/\tau} \hat{x}(t) + \mu(1 - e^{-\Delta t/\tau}) + \sigma \sqrt{1 - e^{-2\Delta t/\tau}} \Phi$$

- where $x(t)$ and $\hat{x}(t)$ have the same statistical properties (PDF, mean, var,...)
- where Φ is a (discrete sequence of) uncorrelated Gauss-distributed RVs, with zero-mean and unitary variance.
- Note: Δt is NOT required to be infinitesimal!

Ornstein-Uhlenbeck process - **wrong** discrete time ver.

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

$$\hat{x}(t + \Delta t) = e^{-\Delta t/\tau} \hat{x}(t) + \mu(1 - e^{-\Delta t/\tau}) + \sigma \sqrt{1 - e^{-2\Delta t/\tau}} \Phi$$

is **NOT** the Euler version of the o.d.e!!!

$$\hat{x}(t + \Delta t) \approx \hat{x}(t) + \frac{\Delta t}{\tau} (\mu - \hat{x}(t)) + \frac{\Delta t}{\tau} \sigma \sqrt{2\tau} \Phi$$

- With Euler, $x(t)$ and $\hat{x}(t)$ do **NOT** have the same statistical properties.
- Making Δt smaller and smaller, things should **NOT** change, right?
- With that Euler *abomination* things would depend on how you choose Δt !

Ornstein-Uhlenbeck process - discrete time version

$$\hat{x}(t + \Delta t) = e^{-\Delta t/\tau} \hat{x}(t) + \mu(1 - e^{-\Delta t/\tau}) + \sigma \sqrt{1 - e^{-2\Delta t/\tau}} \Phi$$

When Δt is set to be very small, by Taylor expansion... $e^{-\Delta t/\tau} \approx 1 - \frac{\Delta t}{\tau}$

$$\hat{x}(t + \Delta t) \approx \hat{x}(t) + \frac{\Delta t}{\tau} (\mu - \hat{x}(t)) + \sigma \sqrt{2 \frac{\Delta t}{\tau}} \Phi$$

$$\hat{x}(t + \Delta t) \approx \hat{x}(t) + \frac{\Delta t}{\tau} (\mu - \hat{x}(t)) + \frac{\sqrt{\Delta t}}{\tau} \sigma \sqrt{2\tau} \Phi$$

$$\hat{x}(t + \Delta t) \approx \hat{x}(t) + \frac{\Delta t}{\tau} (\mu - \hat{x}(t)) + \frac{\Delta t}{\tau} \sigma \sqrt{2\tau} \Phi$$

Proof - O.U. process - properties

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

- write the analytical solution for $x(t)$, assuming that $\xi(t)$ is a normal function: sum of two terms... or the convolution integral... bla bla bla bla...)

- OR ... follow this short-cut: multiply both sides by $e^{t/\tau}$ and then integrate both sides as $\int_0^t dt'$

$$\tau e^{t/\tau} \frac{d}{dt} x(t) = \mu e^{t/\tau} - x e^{t/\tau} + \sigma \sqrt{2\tau} e^{t/\tau} \xi(t)$$

$$\tau e^{t/\tau} \frac{d}{dt} x(t) + x e^{t/\tau} = \mu e^{t/\tau} + \sigma \sqrt{2\tau} e^{t/\tau} \xi(t)$$

$$\tau \frac{d}{dt} (e^{t/\tau} x(t)) = \mu e^{t/\tau} + \sigma \sqrt{2\tau} e^{t/\tau} \xi(t)$$

$$\tau (e^{t/\tau} x(t) - x_0) = \mu \tau (e^{t/\tau} - 1) + \sigma \sqrt{2\tau} \int_0^t e^{t'/\tau} \xi(t') dt'$$

Proof - O.U. process - properties

$$\tau (e^{t/\tau} x(t) - x_0) = \mu \tau (e^{t/\tau} - 1) + \sigma \sqrt{2\tau} \int_0^t e^{t'/\tau} \xi(t') dt'$$

$$x(t) = x_0 e^{-(t-t_0)/\tau} + \mu (1 - e^{-(t-t_0)/\tau}) + \sigma \sqrt{2\tau} \int_0^t e^{-(t-t')/\tau} \xi(t') dt'$$

- Now you can (simply!) use this expression for $x(t)$ to evaluate the mean, the variance, and the covariance.
- You remember that $\xi(t)$ has very specific statistical properties (mean, variance, covariance, PDF) and remember that $\langle \dots \rangle$ is a *linear* operator and therefore you can swap it with the sign of integral and with the *product by a deterministic quantity*.

Stein process - definition

- is a continuous-time, continuous-valued stochastic process (of time)
- I define it, through an ordinary differential equation as

$$\tau \frac{d}{dt} y(t) = -y + \tau J \sum_k \delta(t - t_k) \quad \text{with } y(t_0) = y_0$$

(deterministically)

- There is an (exponentially) decaying term and a *shot-noise* term (deja-vu?!)
- By using that specific definition, *around* t_k , the *jump* is $y(t_k^+) = y(t_k^-) + J$

Stein process - definition

- $\{t_k\}$ are sequences of RVs in time (called technically a **point process**).
- I require that *renewal* property is satisfied, i.e. the occurrence of an event is independent on the previous history:

$$\text{Prob}\{\text{an event in } [t; t + dt]\} = \lambda dt + O(dt)$$

λ as unit of 1/time, and it represents the *rate* of occurrence of the events.

- When λ does not change over time, then $\{t_k\}$ is called **homogeneous** Poisson Point process.
- The RVs defined as the *inter-event* intervals $\{T_k\}$, $T_k = t_k - t_{k-1}$ have a PDF that is an *exponential* function $\lambda e^{-\lambda T}$.

Stein process
and ...the eq. for tot. synaptic currents?!??!?

$$\tau \frac{d}{dt} y(t) = -y + \tau J \sum_k \delta(t - t_k)$$

$$\frac{dI_{syn\ i}}{dt} \approx -\beta I_{syn\ i} + \sum_j C_{ij} W_{ij} \sum_{k_j} \delta(t - t_{k_j})$$

They look the same!

Stein - Let's integrate both sides of the o.d.e

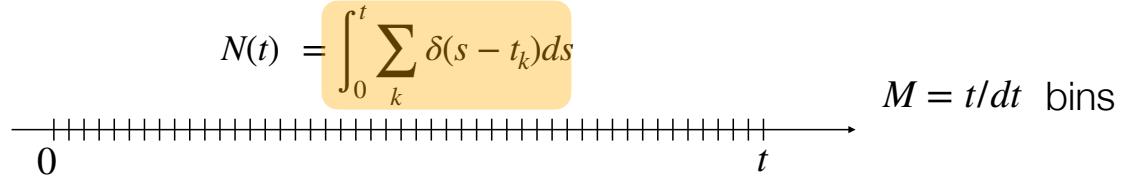
$$\begin{aligned} \tau \frac{d}{dt} y(t) &= -y + \tau J \sum_k \delta(t - t_k) \\ \tau \int_0^t \frac{d}{dt} y(t) dt &= - \int_0^t y dt + \tau J \int_0^t \sum_k \delta(s - t_k) ds \\ y(t) &= y_0 - \frac{1}{\tau} \int_0^t y dt + J \int_0^t \sum_k \delta(s - t_k) ds \end{aligned}$$

$$N(t) = \int_0^t \sum_k \delta(s - t_k) ds$$

At every moment in time, $N(t)$ is an integer number, a RV distributed according to a Poisson distribution.

It counts how many events occurred since $t = 0$.

Let's study ...the monster!



$$\text{Prob}\{\text{an event in}[t; t + dt]\} = \lambda dt + O(dt)$$

$$\text{Prob}\{N \text{ events in}[0; t), \text{i.e. over } M \text{ bins}\} = \binom{M}{N} (\lambda dt)^N (1 - \lambda dt)^{M-N}$$

$$\binom{M}{N} = \frac{M!}{(M-N)!N!}$$

$$\langle N(t) \rangle = M \lambda dt = t\lambda$$

$$\langle N(t)^2 \rangle - (\langle N(t) \rangle)^2 = M \lambda dt (1 - \lambda dt) = t\lambda(1 - \lambda dt)$$

Note: We should anyway remember to go back to continuous time! We exploit the known fact that as $dt \rightarrow 0, M \rightarrow +\infty$, the Binomial distribution \rightarrow Poisson. Now, Poisson \rightarrow Gaussian, but only if $N \rightarrow +\infty$ (i.e. $\lambda \rightarrow +\infty$).

O.U. - Let's integrate both sides of the o.d.e

$$\tau \frac{d}{dt} x(t) = \mu - x + \sigma \sqrt{2\tau} \xi(t)$$

$$\tau \int_0^t \frac{d}{dt} x(t) dt = \int_0^t \mu dt - \int_0^t x dt + \sigma \sqrt{2\tau} \int_0^t \xi(s) ds$$

$$x(t) = x_0 - \frac{1}{\tau} \int_0^t x dt + \frac{\mu}{\tau} t + \frac{\sigma}{\tau} \sqrt{2\tau} \int_0^t \xi(s) ds \quad a + bW(t)$$

$$y(t) = y_0 - \frac{1}{\tau} \int_0^t y dt + J \int_0^t \sum_k \delta(s - t_k) ds$$

$$W(t) = \int_0^t \xi(s) ds$$

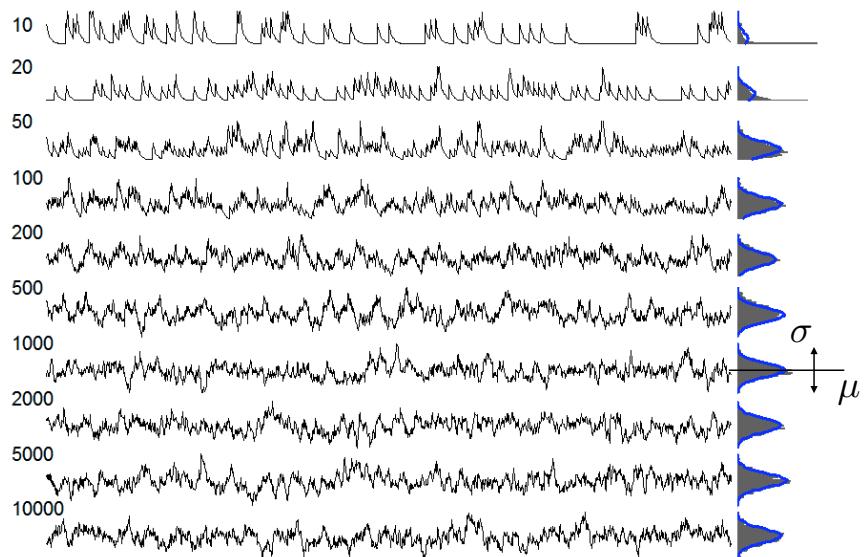
At every moment in time, $W(t)$ is **gaussian** distributed continuous-time, continuous-valued process, called Wiener's process. It has zero mean and variance t .

The diffusion approximation

$$\tau \frac{d}{dt}x(t) = \mu - x + \sigma\sqrt{2\tau} \xi(t)$$

$$\tau \frac{d}{dt}y(t) = -y + \tau J \sum_k \delta(t - t_k)$$

The diffusion approximation



Giugliano et al., 2008, Biological Cybernetics
Arsiero et al., 2007, J Neuroscience
Giugliano et al., 2004, J Neurophysiology

Stein vs O.U.

$$y(t_0) = y_0 = x_0 = x(t_0)$$

$$y(t) = y_0 - \frac{1}{\tau} \int_0^t y dt + J N(t)$$

Poisson

```

graph TD
    A[y(t) = y0 - 1/tau * integral y dt + J N(t)] --> B[J t lambda]
    A --> C[J^2 t lambda]
    B --- D[mean]
    C --- E[variance]
  
```

$$x(t) = x_0 - \frac{1}{\tau} \int_0^t x dt + \frac{\mu}{\tau} t + \frac{\sigma}{\tau} \sqrt{2\tau} W(s)$$

Weiner's process (Gaussian)

```

graph TD
    A[x(t) = x0 - 1/tau * integral x dt + mu/tau * t + sigma/tau * sqrt(2*tau) * W(s)] --> B[mu/tau * t]
    A --> C[2*sigma^2/tau]
    B --- D[mean]
    C --- E[variance]
  
```

$$J \lambda = \frac{\mu}{\tau}$$

$$J^2 \lambda = 2 \frac{\sigma^2}{\tau}$$

$$\mu = J \lambda \tau \quad \sigma^2 = \frac{1}{2} J^2 \lambda \tau$$

Warning: This makes sense only when Poisson \rightarrow Gaussian, but this is only if $N \rightarrow +\infty$ (i.e. $\lambda \rightarrow +\infty$). Therefore only when $J \rightarrow 0$, the term $J\lambda$ remain finite.