

# A Bayesian Approach to Predicting NFL Quarterback Scores in Fanduel Tournaments

*STAT 578, Fall 2017, Team 5: Aaron Ray, Kiomars Nassiri, Michael Chan*

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## Project Description

The National Football League (NFL), being one of the major professional sports leagues in North America, has a wide audience. participates in the NFL craze by competing in fantasy football tournaments organized by the daily fantasy site, “FanDuel.com”. Participants in a **Fantasy Football** game act as the managers of a virtual football team and try to maximize their points by picking up the best line-up. Points are given based on actual performance of players in real-world competition. For the purpose of this project we have chosen to work with the data gathered from the **FanDuel** internet company. We will leverage a Hierarchical Bayesian approach with the Markov Chain Monte Carlo method to predict the fantasy points likely to be scored by an NFL quarterback in any given game. The goal is to predict the points scored by each player given certain prior conditions and predictor variables that will assist our model in providing credible posterior prediction intervals.

The analysis is inspired by the study presented in the article, **Bayesian Hierarchical Modeling Applied to Fantasy Football Projections for Increased Insight and Confidence**, by Scott Rome.

## Team Members

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## Dataset Description

Team has set up a process to gather the historical data from the RotoGuru website. The following is the code used to get the data from RotoGuru:

```
# Scrape rotoguru1 site for weekly FanDuel stats and bind each week's data to the
# pre-defined dataframe, 'd'.

for(year in 2014:2017){
  for(week in 1:16){
    page = read_html(
      gsub(" ", "", ,
           paste("http://rotoguru1.com/cgi-bin/fyday.pl?week=", week, "&year=",
                 year, "&game=fd&scsv=1"))
    )
    dtext = page %>% html_nodes("pre") %>% html_text(trim = TRUE)
    dtable = read.table(text=dtext, sep = ";", header=TRUE, col.names = cnames,
```

```

    quote=NULL)
d = rbind(d,dtable)
}
}

```

Data cleaning is performed using R routines. Some data cleaning tasks are needed to calculate Player rank.

## Response Variables

- **FanDuelPts**: Points position at the end of a single game

## Predictor Variables

- **AvgPts5Wks**: The 5 game average points of the player
- **AvgOppPAP7Wks** : The 7 game average Opposing Points Allowed to Position (OppPAP) by the current player's opposing defense. For example, if the Buffalo Bills defense allowed a total of 30 points per game to wide receivers for six games straight, then this number would equal to the average of 30 for any wide receiver facing the Bills defense.
- **Position**: The position the player plays
- **HomeGame**: Whether it is home game.
- **Rank**: The rank of a player based on recent performance

## Analysis Ideas

### Model

At the lowest level, we model the performance (**FanDuelPts**) as normally-distributed around a true value:

$$y|\alpha, \beta_{defense}, \beta_{home}, \beta_{away}, \sigma_r^2 \sim N(\alpha + X_{defense} \cdot \beta_{defense} + X_{home} \cdot \beta_{home} + X_{away} \cdot \beta_{away}, \sigma_y^2 I)$$

where

$\alpha$  = The average fan duel point of the previous 5 weeks of the player, **AvgPts5Wks**

$\beta_{defense,p}$  = defense coefficient against team t for position p

$\beta_{home,p,r}$  = home coefficient for position p and a rank r player

$\beta_{away,p,r}$  = Away coefficient for position p and a rank r player

$y$  = **FanDuelPts**

$x_p$  = interaction indicator term for opposing team score allowed by position p

$x_{home,p,r}$  = interaction indicator term for rank r, position p, and whether it is home game

At higher level, we model the defense effect,  $\beta_{defense}$ , as how good(bad) a particular team's defense is against the player's position. We pool the effect based on the position of the player. That is, the defense coefficient is normally distributed from the same position specific distribution.

$$\beta_{defense,p} \sim N(\delta_p, \sigma_\delta^2)$$

where  $\sigma_\delta$  is constant = 1000

For the home and away game effect,  $\beta_{home}$  and  $\beta_{away}$ , we model the effect for player of the same rank has the same distribution. We model the home and away game effect to be the same for players of the same position.

$$\beta_{home,p,r} \sim N(\eta_r, \sigma_\eta^2)$$

$$\beta_{away,p,r} \sim N(\rho_r, \sigma_\rho^2)$$

where  $\sigma_\eta, \sigma_\rho$  are constant = 1000

We will approximate non informative prior using:

$$\sigma_y \sim Inv-gamma(0.0001, 0.0001)$$

$$\delta \sim N(0, 10000^2)$$

$$\eta \sim N(0, 10000^2)$$

$$\rho \sim N(0, 10000^2)$$

Here is the JAGS model:

```
#sink("fdp.bug")
#cat("
model {
  for (i in 1:length(y)) {
    y[i] ~ dnorm(alpha[i] + inprod(X.defense[i, ], beta.defense)
                  + inprod(X.home[i, ], beta.home)
                  + inprod(X.away[i, ], beta.away), sigmasqinv)
  }

  # The entry of the beta.defense corresponds to Opponent:Position
  # In our model, we pool the beta.defense based on position.
  # i.e. All defense effects of the same position are drawn from the same distribution
  for (p in 1:Num.Position) {
    beta.defense[p] ~ dnorm(delta[p], 1/1000^2)
    delta[p] ~ dnorm(0, 1/100000^2)
  }

  # The entry of the beta.home and beta.away corresponds to Rank:Position
  # In our model, we pool the beta.home/away based on rank
  for (r in 1:Num.Rank) {
    for (t in 1:Num.Position) {
      beta.home[(t-1) * Num.Rank + r] ~ dnorm(eta[r], 1/1000^2)
      beta.away[(t-1) * Num.Rank + r] ~ dnorm(rho[r], 1/1000^2)
    }
    eta[r] ~ dnorm(0, 1/100000^2)
    rho[r] ~ dnorm(0, 1/100000^2)
  }

  sigmasqinv ~ dgamma(0.0001, 0.0001)
  sigmasq <- 1/sigmasqinv
}

#      ",fill = TRUE)
#sink()

library(knitr)
```

## Sample Data

```
fdp <- read.csv("fdpfinal.csv", sep = ',', header = TRUE)

head(fdp)

##   Position Year YearWeek Opponent Week PlayerId          Name      Team
## 1      QB 2015     201513  Steelers  13    1060 Hasselbeck, Matt  Colts
## 2      QB 2015     201514  Jaguars  14    1060 Hasselbeck, Matt  Colts
## 3      QB 2015     201515  Texans   15    1060 Hasselbeck, Matt  Colts
## 4      QB 2015     201516 Dolphins  16    1060 Hasselbeck, Matt  Colts
## 5      QB 2015     201501  Ravens   1    1081 Manning, Peyton Broncos
## 6      QB 2015     201502 Chiefs    2    1081 Manning, Peyton Broncos
##   HomeGame FanDuelPts FanDuelSalary AvgOppPAP7Wks SdOppPAP7Wks OallAvgPAP
## 1          0       6.86        6500      20.95      8.045    17.44
## 2          0       9.08        6600      24.16      6.738    17.44
## 3          1       8.98        6400      16.36      9.690    17.44
## 4          0       3.96        6000      20.46      6.002    17.44
## 5          1       5.90        9100      18.99     11.697    17.44
## 6          0      21.24        8200      13.85      4.342    17.44
##   OallStddevPAP AvgPts5Wks StdevPts5Wks OffRnk5Wks DefRnk7Wks
## 1      2.909     14.85      4.824    Rank4     Rank1
## 2      2.909     14.82      4.897    Rank4     Rank1
## 3      2.909     13.56      5.490    Rank4     Rank3
## 4      2.909     12.11      5.566    Rank4     Rank2
## 5      2.909     14.96      8.496    Rank3     Rank1
## 6      2.909     10.53      5.011    Rank4     Rank4

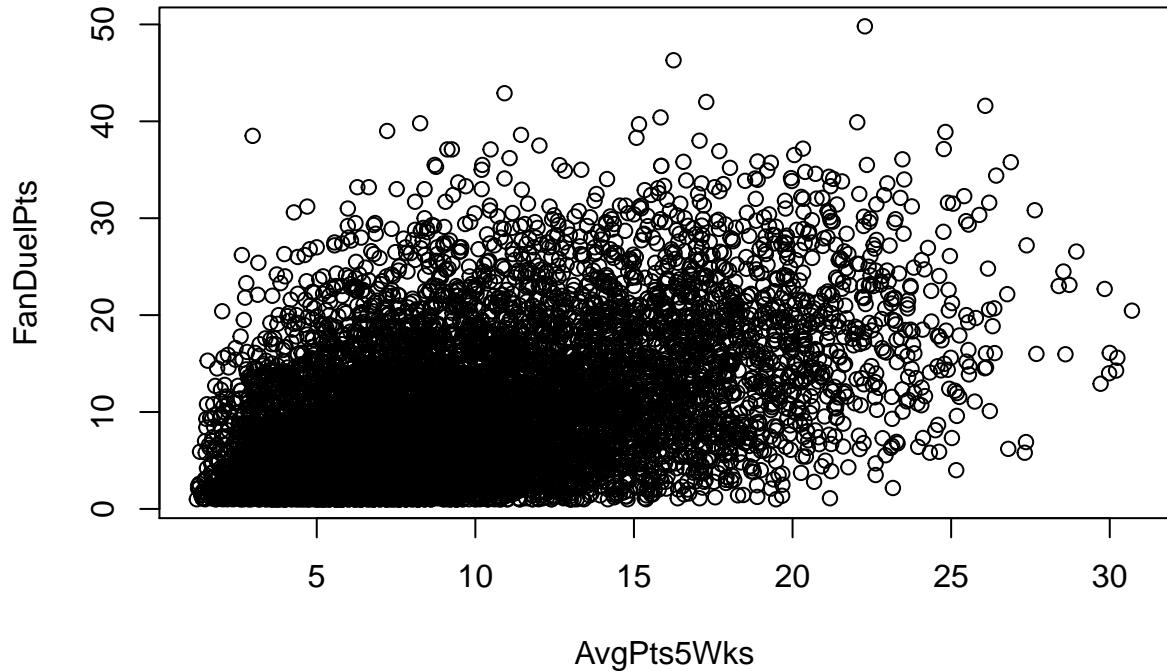
fdp['Rank'] = fdp$OffRnk5Wks
fdp['Locality'] = 'Away'
fdp[fdp$HomeGame == 1, 'Locality'] = 'Home'
```

## Simple Ideas

$$y|\alpha \sim N(\alpha, \sigma_y^2 I)$$

```
mod.classic = lm(FanDuelPts ~ AvgPts5Wks, data = fdp)

plot(FanDuelPts ~ AvgPts5Wks, data = fdp)
```



All over the place, let's add the team defense

X.defense is the indicator matrix

\*\* Set up train data \*\*

```
#fdp_train=fdp[fdp$Year == 2015, ]
fdp_train=fdp[fdp$Year == 2016 & ((fdp$Position == "QB"& fdp$FanDuelSalary > 6500 & !is.na(fdp$FanDuelSalary)) | (fdp$Position == "RB"& fdp$FanDuelSalary > 6500 & !is.na(fdp$FanDuelSalary)))
#fdp_train=fdp[fdp$Year == 2015 & (fdp$Position == "QB" | fdp$Position == "RB") & fdp$FanDuelSalary > 6500 & !is.na(fdp$FanDuelSalary)]

fdp_test = fdp_train[fdp_train$YearWeek >= 201615, ]
fdp_train=fdp_train[fdp_train$YearWeek < 201615, ]
fdp_train = droplevels(fdp_train)
```

```
Use.Rank = TRUE
Num.Opponent = length(unique(fdp_train[, "Opponent"]))
Num.Position = length(unique(fdp_train[, "Position"]))
if (Use.Rank) {
  Num.Rank = length(unique(fdp_train[, "Rank"]))
  Num.HomeAwayInit = Num.Rank
  Model.File.Ext = ""
} else {
  Num.HomeAwayInit = 1
  Model.File.Ext = ".norank"
}
```

```

if (Num.Position == 1) {
  #X.offense = model.matrix(~ 0 + AvgPts5Wks, data=fdp_train)
  X.defense = model.matrix(~ 0 + AvgOppPAP7Wks, data=fdp_train)
  if (Use.Rank) {
    X.home = model.matrix(~ 0 + Rank , data=fdp_train)
    X.away = model.matrix(~ 0 + Rank , data=fdp_train)
  } else {
    X.home = rep(1, nrow(fdp_train))
    X.away = rep(1, nrow(fdp_train))
  }
} else {
  #X.offense = model.matrix(~ 0 + AvgPts5Wks:Position, data=fdp_train)
  X.defense = model.matrix(~ 0 + AvgOppPAP7Wks:Position, data=fdp_train)
  if (Use.Rank) {
    X.home = model.matrix(~ 0 + Rank:Position , data=fdp_train)
    X.away = model.matrix(~ 0 + Rank:Position , data=fdp_train)
  } else {
    X.home = model.matrix(~ 0 + Position , data=fdp_train)
    X.away = model.matrix(~ 0 + Position , data=fdp_train)
  }
}

X.home = X.home * fdp_train$HomeGame
X.away = X.away * (1- fdp_train$HomeGame)
X = cbind(X.defense, X.home, X.away)

```

```

library(rjags)
set.seed(20171008)

```

```

# Initialization List for the 4 chains
jags.inits=list(
  list( sigmasqinv= 0.01, delta = rep(-100000, Num.Position),
        eta = c(100000, -100000, 100000, -100000)[1:Num.HomeAwayInit],
        rho = c(-100000, 100000, -100000, 100000)[1:Num.HomeAwayInit],
        .RNG.name = "base::Mersenne-Twister", .RNG.seed = 20171008 ),
  list( sigmasqinv= 0.01, delta = rep(100000, Num.Position),
        eta = c(100000, -100000, -100000, 100000)[1:Num.HomeAwayInit],
        rho = c(-100000, 100000, 100000, -100000)[1:Num.HomeAwayInit],
        .RNG.name = "base::Mersenne-Twister", .RNG.seed = 20171008 + 1 ),
  list( sigmasqinv=0.000001, delta = rep(-100000, Num.Position),
        eta = c(-100000, 100000, -100000, 100000)[1:Num.HomeAwayInit],
        rho = c(100000, -100000, 100000, -100000)[1:Num.HomeAwayInit],
        .RNG.name = "base::Mersenne-Twister", .RNG.seed = 20171008 + 2 ),
  list( sigmasqinv=0.000001, delta = rep(100000, Num.Position),
        eta = c(-100000, 100000, 100000, -100000)[1:Num.HomeAwayInit],
        rho = c(100000, -100000, -100000, 100000)[1:Num.HomeAwayInit],
        .RNG.name = "base::Mersenne-Twister", .RNG.seed = 20171008 + 3 )
)

data.jags <- list(
  y= fdp_train$FanDuelPts,
  alpha = fdp_train$AvgPts5Wks,

```

```

X.defense = X.defense,
X.home = X.home,
X.away = X.away,
Num.Position=Num.Position,
#Num.Opponent=Num.Opponent,
Num.Rank=Num.Rank
)

burnAndSample = function(m, N.burnin, N.iter, show.plot, mon.col) {
  update(m, N.burnin) # burn-in

  x <- coda.samples(m, mon.col, n.iter=N.iter)

  if(show.plot) {
    plot(x, smooth=FALSE)
  }

  gelman.R = gelman.diag(x, autoburnin=FALSE, multivariate = FALSE)
  print(gelman.R)

  result <- list(
    coda.sam = x,
    gelman.R.max=max(gelman.R$psrf[, 1])
  )

  return(result)
}

runModel=TRUE
runSample=TRUE

mon.col <- c("delta", "eta", "rho", "beta.defense", "beta.home", "beta.away", "sigmasq")

NSim = 30000
if (runModel) {
  m <- jags.model("fdp.bug", data.jags, inits = jags.inits, n.chains=4, n.adapt = 1000)
  save(file=paste("fdp.jags.model.init", Model.File.Ext, ".Rdata", sep=""), list="m")
} else {
  load(paste("fdp.jags.model.init", Model.File.Ext, ".Rdata", sep=""))
  m$recompile()
}

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
##   Observed stochastic nodes: 746
##   Unobserved stochastic nodes: 29
##   Total graph size: 18344
##
## Initializing model

```

```

load.module("dic")

## module dic loaded

N.Retry.Loop = 1
if (runSample) {
  N.burnin=2500/2
  for (loopIdx in 1:N.Retry.Loop) {
    (start_time <- Sys.time())
    (N.burnin = N.burnin * 2)
    result = burnAndSample(m, N.burnin, NSim, show.plot=FALSE, mon.col = mon.col)
    (end_time <- Sys.time())
    (result$gelman.R.max)
  }
  save(file=paste("fdp.jags.samples.", N.burnin, Model.File.Ext, ".Rdata", sep=""), list="result")
  save(file=paste("fdp.jags.model.", N.burnin, Model.File.Ext, ".Rdata", sep=""), list="m")
} else {
  N.burnin=2500/2 * (2**N.Retry.Loop)
  load(paste("fdp.jags.samples.", N.burnin, Model.File.Ext, ".Rdata", sep=""))
  load(paste("fdp.jags.model.", N.burnin, Model.File.Ext, ".Rdata", sep=""))

  gelman.diag(result$coda.sam, autoburnin=FALSE, multivariate = FALSE)
}

## Potential scale reduction factors:
##          Point est. Upper C.I.
## beta.away[1]      1     1.00
## beta.away[2]      1     1.00
## beta.away[3]      1     1.00
## beta.away[4]      1     1.00
## beta.away[5]      1     1.01
## beta.away[6]      1     1.00
## beta.away[7]      1     1.00
## beta.away[8]      1     1.01
## beta.defense[1]   1     1.00
## beta.defense[2]   1     1.01
## beta.home[1]      1     1.00
## beta.home[2]      1     1.00
## beta.home[3]      1     1.00
## beta.home[4]      1     1.00
## beta.home[5]      1     1.01
## beta.home[6]      1     1.01
## beta.home[7]      1     1.01
## beta.home[8]      1     1.01
## delta[1]          1     1.00
## delta[2]          1     1.00
## eta[1]            1     1.00
## eta[2]            1     1.00
## eta[3]            1     1.00
## eta[4]            1     1.00
## rho[1]            1     1.00
## rho[2]            1     1.00

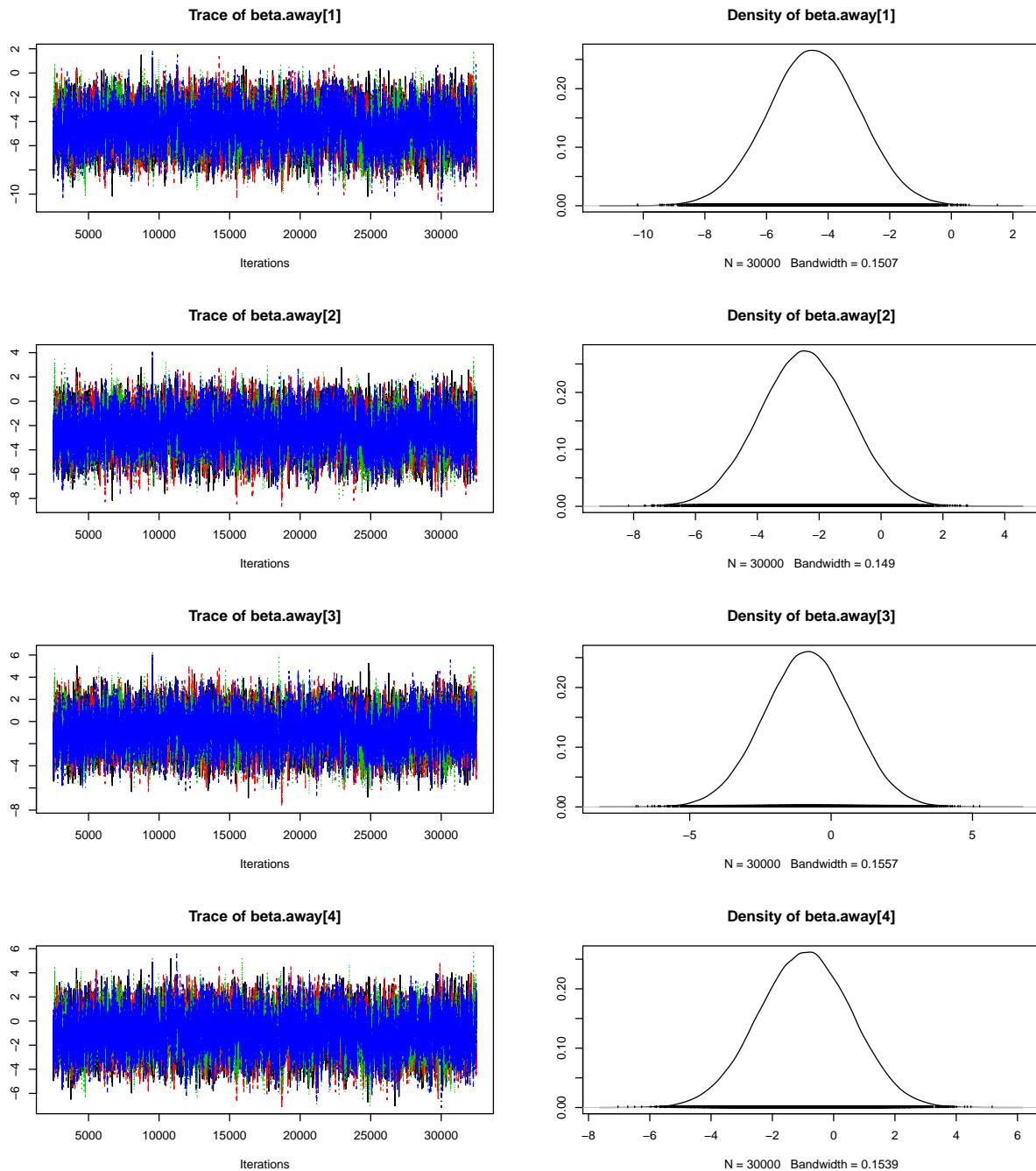
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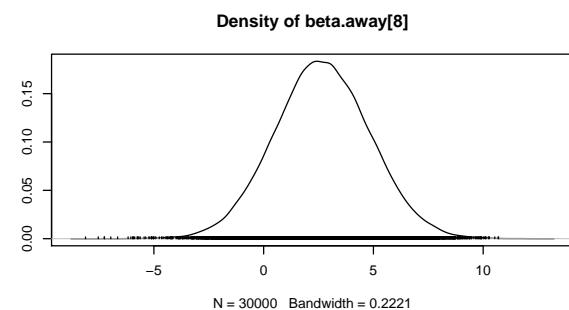
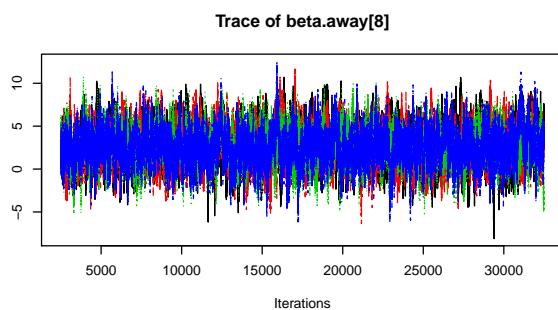
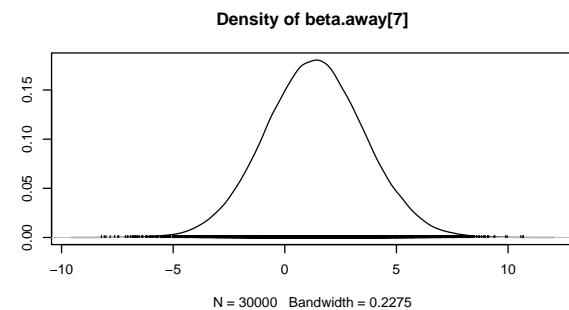
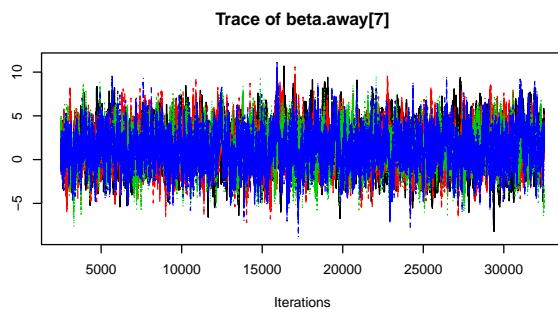
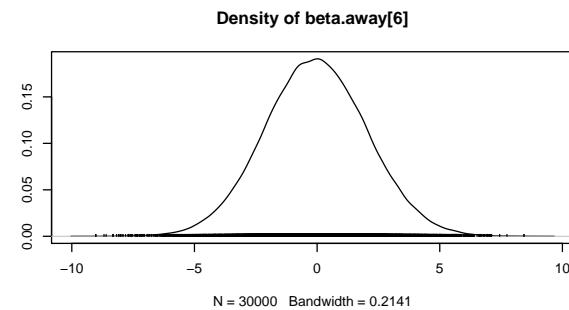
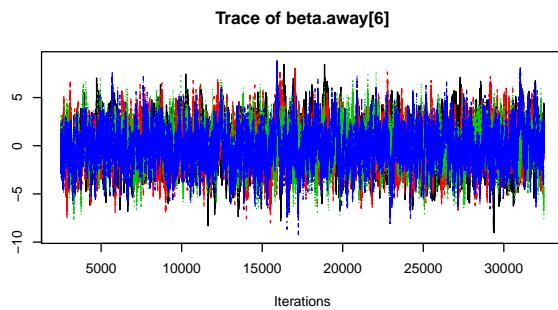
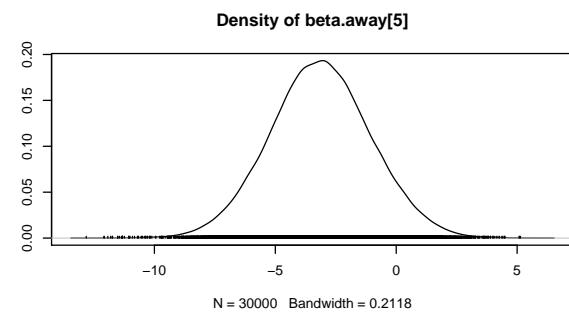
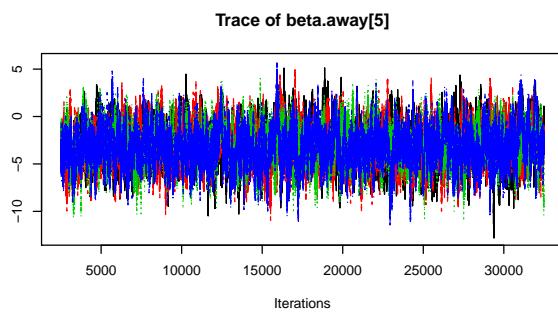
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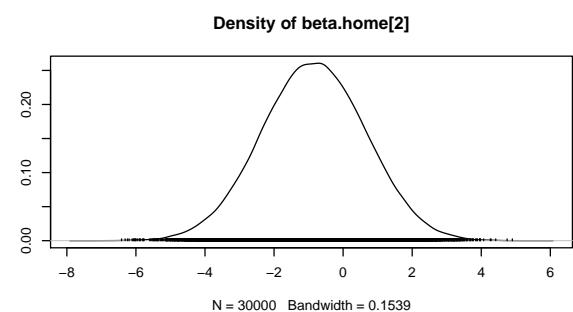
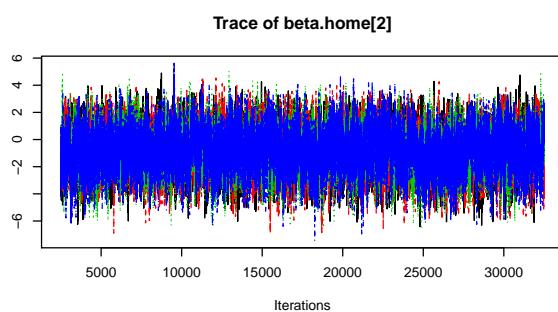
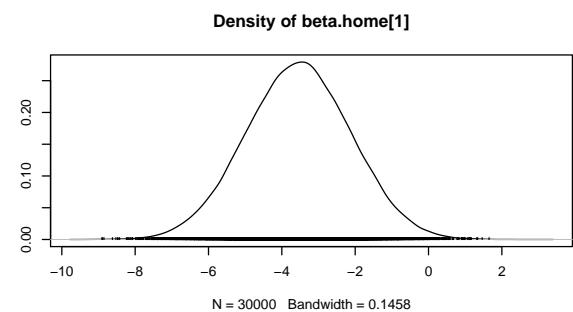
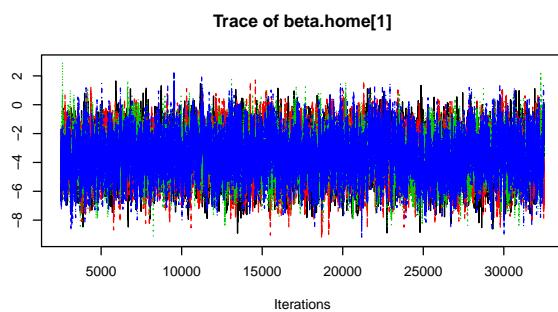
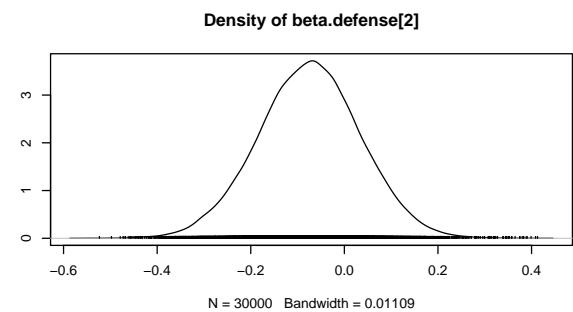
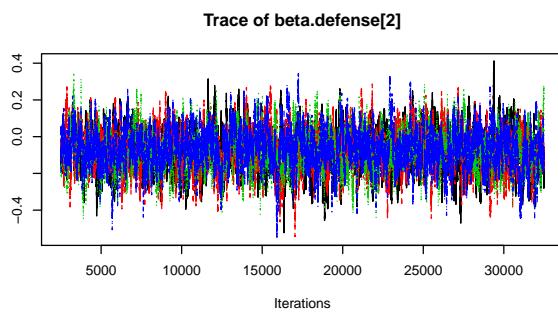
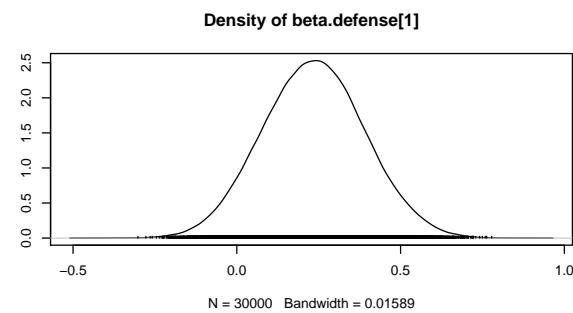
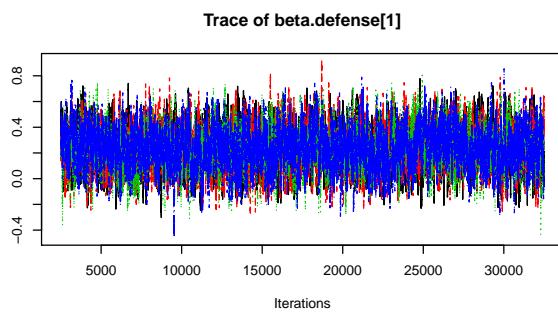
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## rho[4]          1      1.00
## sigmasq         1      1.00

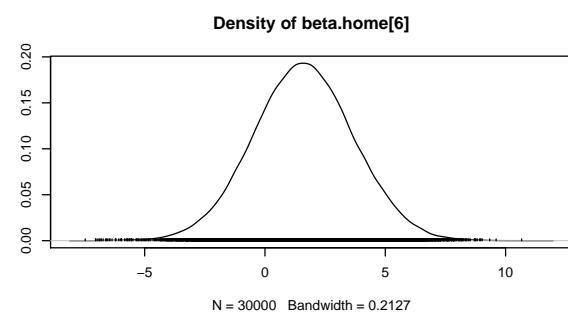
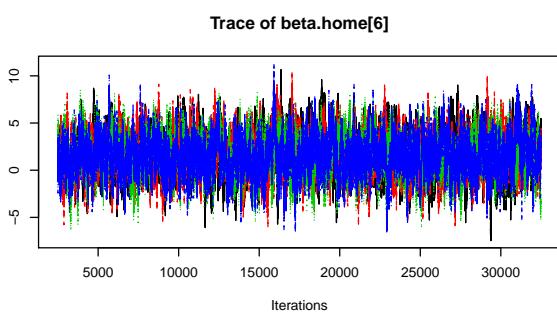
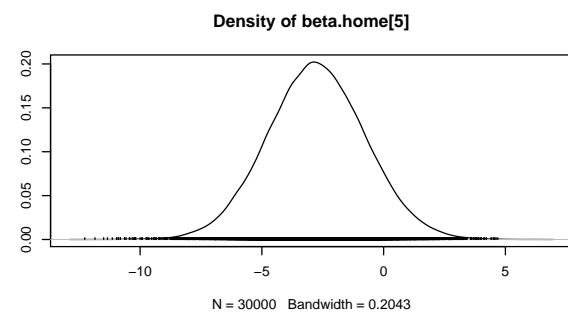
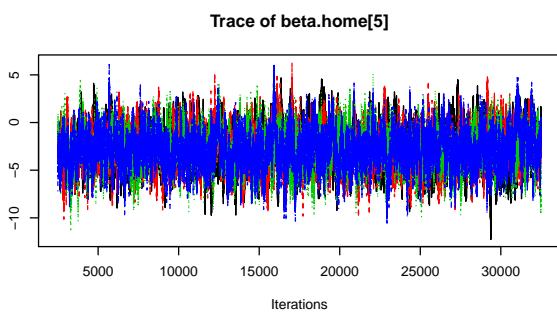
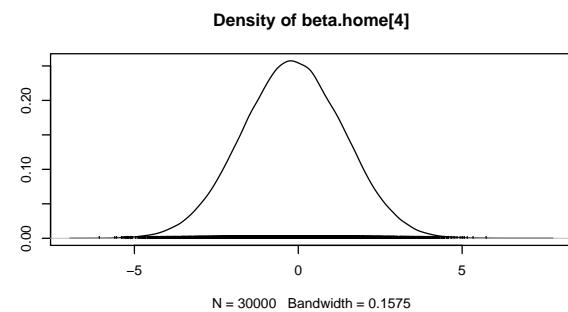
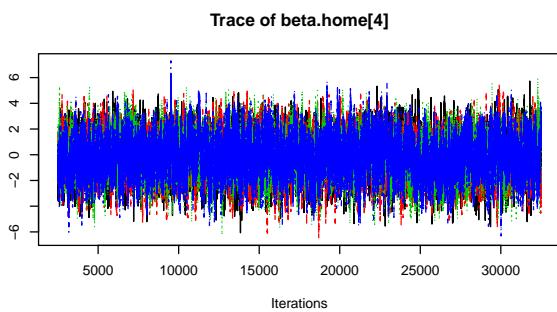
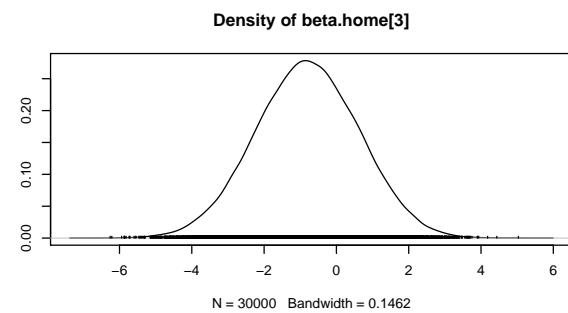
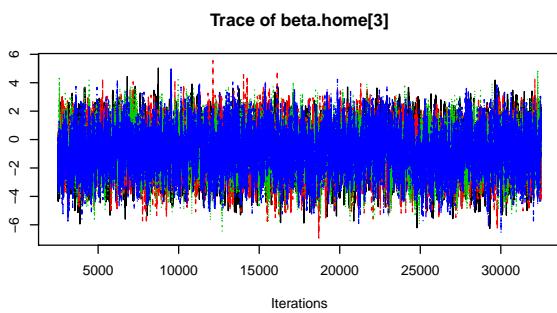
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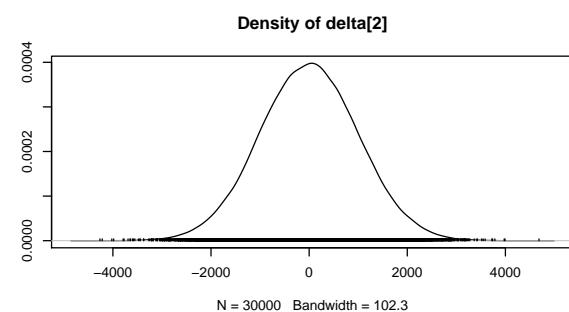
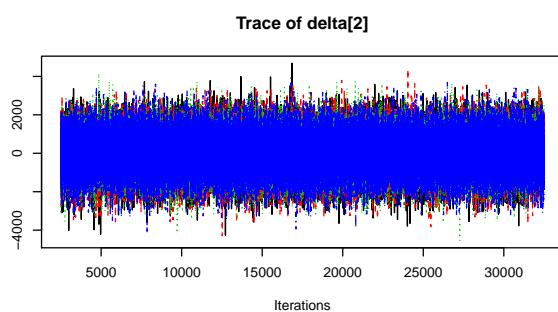
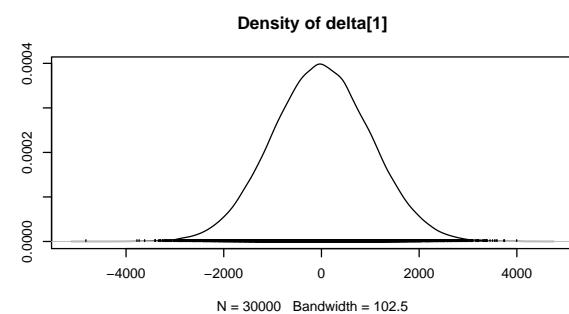
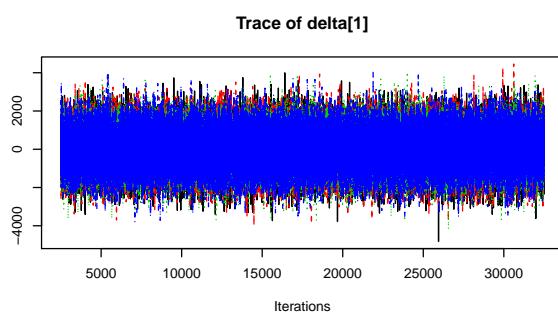
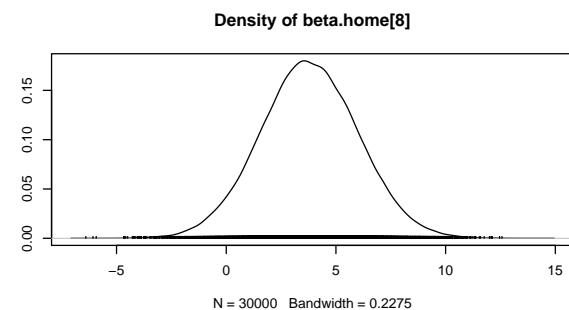
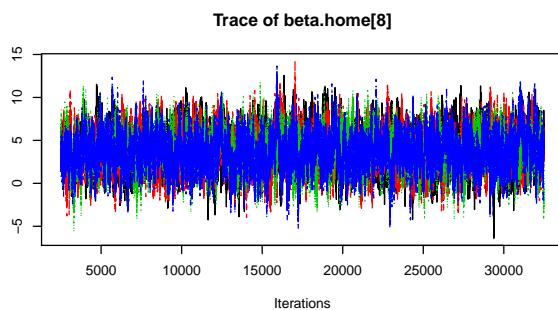
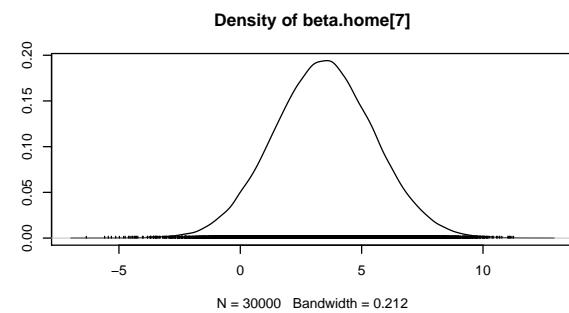
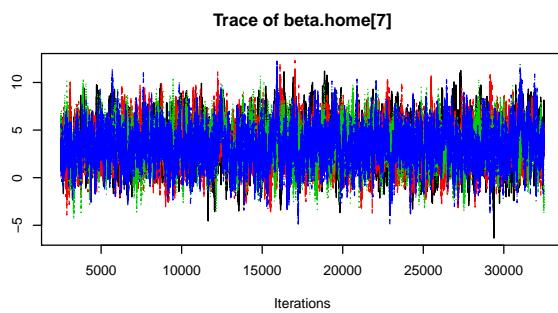
```
plot(result$coda.sam, smooth=FALSE)
```

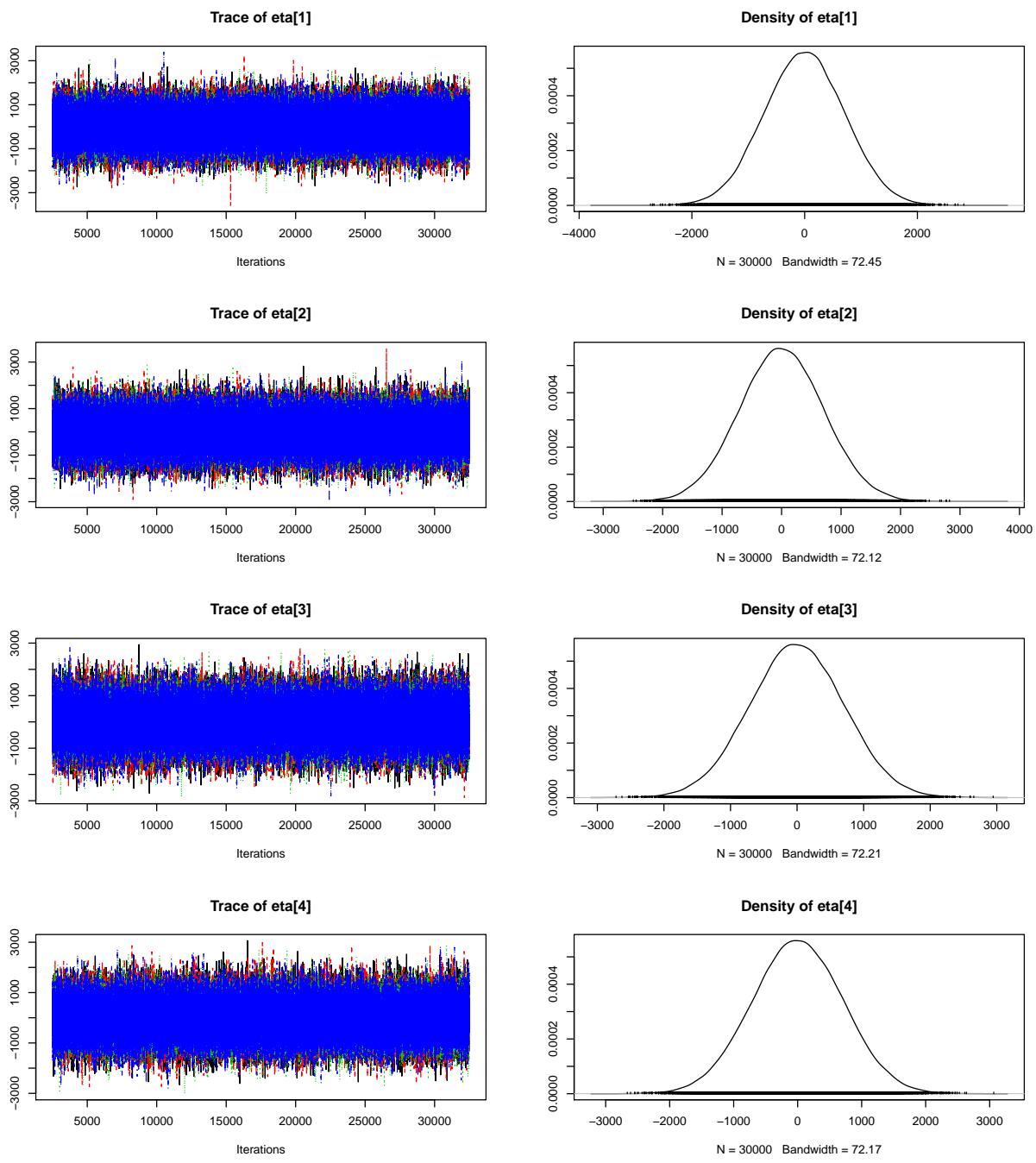


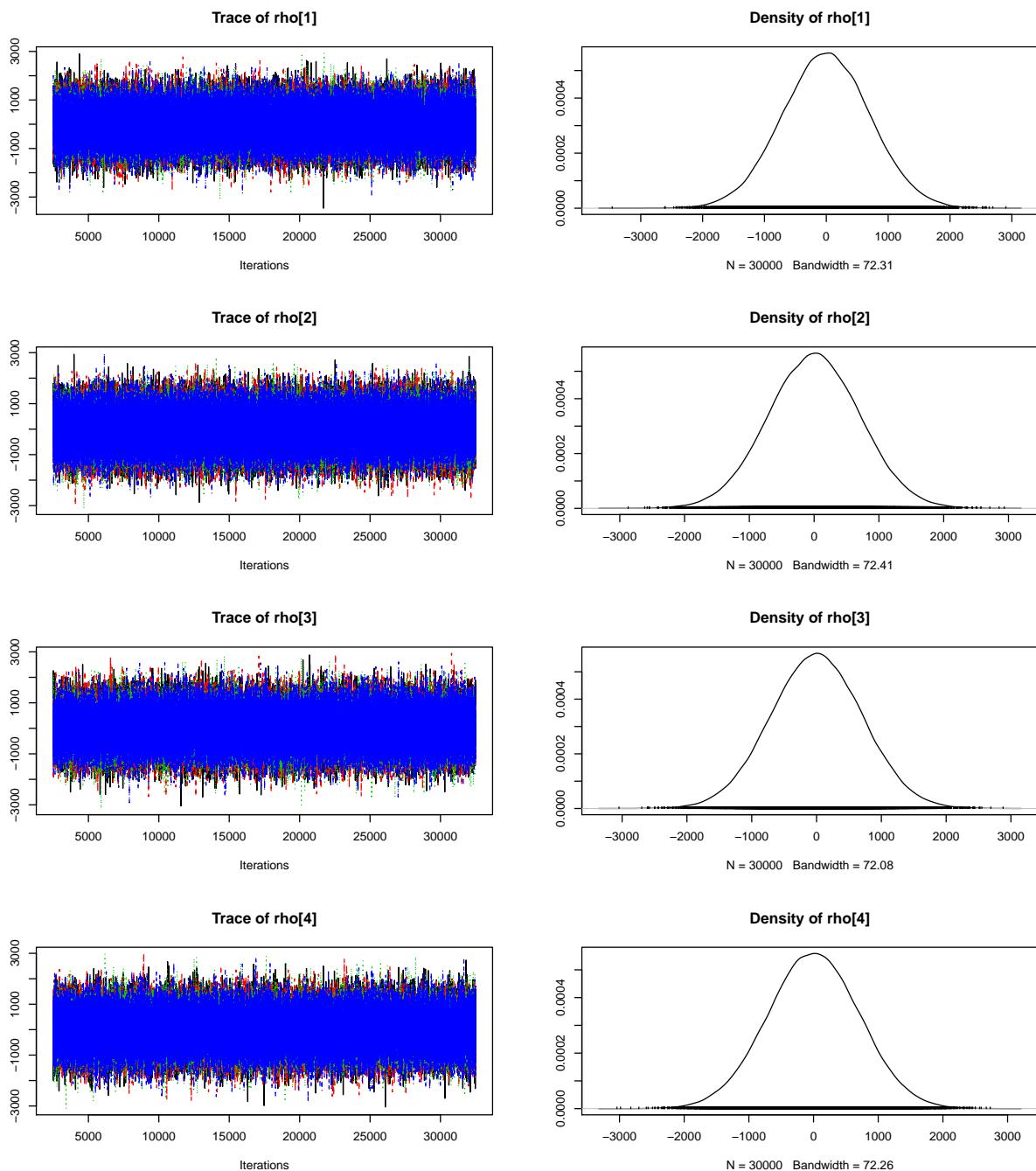


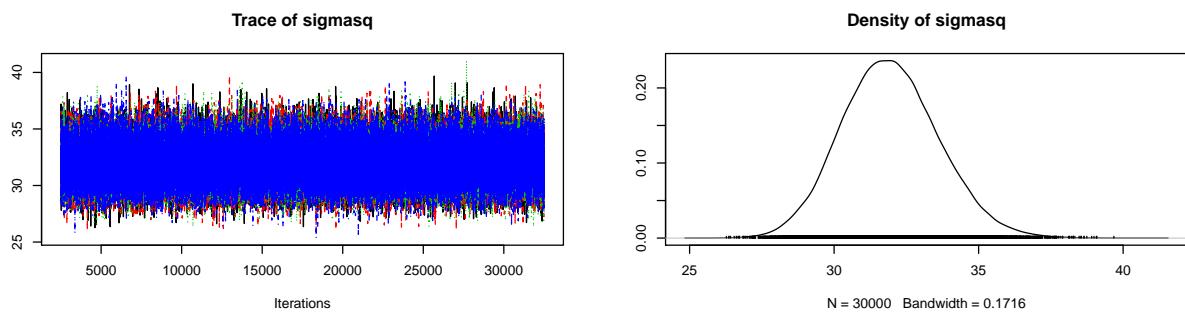












Converged as `gelman.R.max` = 1.002 < 1.1 and the plot also looks good.

```
summary(result$coda.sam)
```

```
##
## Iterations = 2501:32500
## Thinning interval = 1
## Number of chains = 4
## Sample size per chain = 30000
##
```

```

## 1. Empirical mean and standard deviation for each variable,
## plus standard error of the mean:
##
##          Mean        SD Naive SE Time-series SE
## beta.away[1] -4.4636 1.474 0.004256 0.02372
## beta.away[2] -2.4624 1.458 0.004210 0.02317
## beta.away[3] -0.8332 1.524 0.004399 0.02341
## beta.away[4] -0.9295 1.506 0.004348 0.02333
## beta.away[5] -3.1336 2.092 0.006039 0.05109
## beta.away[6] -0.0485 2.095 0.006047 0.04977
## beta.away[7] 1.3464 2.235 0.006451 0.05245
## beta.away[8] 2.6680 2.180 0.006292 0.05058
## beta.defense[1] 0.2348 0.155 0.000449 0.00285
## beta.defense[2] -0.0750 0.110 0.000316 0.00286
## beta.home[1] -3.5465 1.427 0.004119 0.02297
## beta.home[2] -0.8547 1.505 0.004346 0.02281
## beta.home[3] -0.8048 1.430 0.004129 0.02218
## beta.home[4] -0.1613 1.541 0.004447 0.02484
## beta.home[5] -2.7594 2.013 0.005810 0.04794
## beta.home[6] 1.6027 2.088 0.006028 0.05020
## beta.home[7] 3.4434 2.079 0.006002 0.04957
## beta.home[8] 3.7885 2.228 0.006430 0.05187
## delta[1] 0.5124 1002.753 2.894698 2.90765
## delta[2] -0.5926 1000.900 2.889350 2.90995
## eta[1] -4.4295 708.899 2.046415 2.04643
## eta[2] -0.1967 705.636 2.036996 2.08127
## eta[3] -2.6215 706.540 2.039605 2.05043
## eta[4] -0.0937 706.160 2.038509 2.03851
## rho[1] -2.1788 707.476 2.042306 2.05159
## rho[2] 0.1973 708.476 2.045195 2.03842
## rho[3] 1.1455 705.234 2.035836 2.02869
## rho[4] 0.0402 706.971 2.040849 2.03517
## sigmasq 31.9444 1.679 0.004846 0.00504
##
## 2. Quantiles for each variable:
##
##          2.5%      25%      50%      75%     97.5%
## beta.away[1] -7.3597 -5.458 -4.4632 -3.46216 -1.597
## beta.away[2] -5.3210 -3.447 -2.4598 -1.47056 0.381
## beta.away[3] -3.8368 -1.859 -0.8279 0.19686 2.137
## beta.away[4] -3.8975 -1.949 -0.9238 0.09688 1.993
## beta.away[5] -7.2200 -4.526 -3.1435 -1.74877 1.016
## beta.away[6] -4.1384 -1.459 -0.0593 1.35282 4.092
## beta.away[7] -3.0286 -0.151 1.3423 2.83241 5.763
## beta.away[8] -1.5715 1.214 2.6541 4.12679 6.974
## beta.defense[1] -0.0662 0.129 0.2348 0.33967 0.541
## beta.defense[2] -0.2929 -0.147 -0.0740 -0.00212 0.139
## beta.home[1] -6.3444 -4.514 -3.5429 -2.58185 -0.756
## beta.home[2] -3.8255 -1.873 -0.8525 0.16951 2.075
## beta.home[3] -3.6191 -1.773 -0.8028 0.16886 1.979
## beta.home[4] -3.1849 -1.202 -0.1595 0.88608 2.834
## beta.home[5] -6.6870 -4.100 -2.7698 -1.42217 1.224
## beta.home[6] -2.4973 0.204 1.5951 2.99237 5.717
## beta.home[7] -0.6236 2.051 3.4406 4.83075 7.546

```

```

## beta.home[8]      -0.5856    2.302   3.7748   5.28452   8.170
## delta[1]        -1957.5020 -678.601  -2.3096  675.59568 1974.721
## delta[2]        -1962.0265 -677.623   0.3306  672.01082 1973.059
## eta[1]          -1393.2636 -484.157  -3.8283  475.09084 1380.327
## eta[2]          -1383.1520 -475.871  -1.3895  475.35275 1388.076
## eta[3]          -1390.5465 -476.938  -4.0451  475.59379 1379.797
## eta[4]          -1380.4350 -476.480  -0.7876  478.51307 1384.202
## rho[1]          -1393.2531 -477.516  -1.6231  474.21504 1387.126
## rho[2]          -1388.6359 -478.087   1.1899  478.14215 1389.690
## rho[3]          -1381.9251 -474.961   0.5987  476.30937 1384.538
## rho[4]          -1391.9517 -474.983   0.8282  476.87848 1383.491
## sigmasq         28.8029   30.787   31.8924  33.03684  35.382

```

*Effective Sample Size*

```
(eff.size = effectiveSize(result$coda.sam[, ]))
```

```

##   beta.away[1]   beta.away[2]   beta.away[3]   beta.away[4]   beta.away[5]
##      3879        3980        4257        4184        1681
##   beta.away[6]   beta.away[7]   beta.away[8]   beta.defense[1] beta.defense[2]
##     1775        1821        1863        2987        1469
##   beta.home[1]   beta.home[2]   beta.home[3]   beta.home[4]   beta.home[5]
##      3888        4381        4176        3859        1765
##   beta.home[6]   beta.home[7]   beta.home[8]   delta[1]     delta[2]
##     1733        1766        1847       118955      118320
##   eta[1]         eta[2]         eta[3]         eta[4]       rho[1]
##     120000       115504       118747       120000     118984
##   rho[2]         rho[3]         rho[4]         sigmasq
##     120813       120869       120681       111066

```

The effective sample sizes of all parameters are greater than 400.

*Probabilty of players perform better at home than away*

```

post.samp = as.matrix(result$coda.sam)

beta.home = post.samp[, paste("beta.home[", 1:(Num.Rank*Num.Position), "]")]
beta.away = post.samp[, paste("beta.away[", 1:(Num.Rank*Num.Position), "]")]

prob.home.away = rep(0, Num.Rank * Num.Position)
for (r in 1:Num.Rank) {
  for (p in 1:Num.Position) {
    idx = (p-1) * Num.Rank + r
    prob.home.away[idx] = mean(beta.home[, idx] > beta.away[, idx])
  }
}
prob.home.away

```

```
## [1] 0.8082 0.9165 0.5100 0.7491 0.6302 0.9239 0.9498 0.7928
```

```

prob.home.away.df = data.frame(colnames(X.home))
prob.home.away.df$prob.home.bt.away = prob.home.away

```

```

colnames(prob.home.away.df) = c("Rank:Position", "Prob.home.bt.away")

kable(prob.home.away.df)

```

Rank:Position	Prob.home.bt.away
RankRank1:PositionPK	0.8082
RankRank2:PositionPK	0.9165
RankRank3:PositionPK	0.5100
RankRank4:PositionPK	0.7491
RankRank1:PositionQB	0.6302
RankRank2:PositionQB	0.9239
RankRank3:PositionQB	0.9498
RankRank4:PositionQB	0.7928

### Beta defense

If a player is facing a team which gives up more points to players on average, we expect the player will score more points.

```

beta.defense = post.samp[, paste("beta.defense[", 1:Num.Position, "]",
sep="")]

beta.defense.int.df = data.frame(colnames(X.defense))
beta.defense.int = matrix(rep(0, Num.Position * 4), nrow=Num.Position, ncol = 4)

int.alpha=0.05
for (p in 1:Num.Position) {
  beta.defense.int[p, 1:3] = quantile(beta.defense[, p], c(int.alpha/2, 0.5, 1-int.alpha/2))
  beta.defense.int[p, 4] = mean(beta.defense[, p])
}

beta.defense.int.df$pct025 = beta.defense.int[, 1]
beta.defense.int.df$pct975 = beta.defense.int[, 3]
beta.defense.int.df$median = beta.defense.int[, 2]
beta.defense.int.df$mean = beta.defense.int[, 4]
colnames(beta.defense.int.df) = c("beta.defense.position", "pct025", "pct975", "median", "mean")
kable(beta.defense.int.df)

```

beta.defense.position	pct025	pct975	median	mean
AvgOppPAP7Wks:PositionPK	-0.0662	0.5415	0.2348	0.2348
AvgOppPAP7Wks:PositionQB	-0.2929	0.1386	-0.0740	-0.0750

We observe that the median beta.defense for PK is positive as expected. But for QB, it is negative, that implies QB actually score less against bad defensive team.

*DIC*

```
(dic.samp = dic.samples(m, NSim))
```

```
## Mean deviance: 4700
```

```
## penalty 19
## Penalized deviance: 4719
```

## Alternative Model - no rank

```
#sink("fdp.norank.bug")
#cat("
model {
  for (i in 1:length(y)) {
    y[i] ~ dnorm(alpha[i] + inprod(X.defense[i, ], beta.defense)
                  + inprod(X.home[i, ], beta.home)
                  + inprod(X.away[i, ], beta.away), sigmasqinv)
  }

  # The entry of the beta.defense corresponds to Opponent:Position
  # In our model, we pool the beta.defense based on position.
  # i.e. All defense effects of the same position are drawn from the same distribution
  for (p in 1:Num.Position) {
    beta.defense[p] ~ dnorm(delta[p], 1/1000^2)
    delta[p] ~ dnorm(0, 1/100000^2)
  }

  # The entry of the beta.home and beta.away corresponds to Position
  # In our model, we pool the beta.home/away based on Position
  for (t in 1:Num.Position) {
    beta.home[t] ~ dnorm(eta, 1/1000^2)
    beta.away[t] ~ dnorm(rho, 1/1000^2)
  }
  eta ~ dnorm(0, 1/100000^2)
  rho ~ dnorm(0, 1/100000^2)

  sigmasqinv ~ dgamma(0.0001, 0.0001)
  sigmasq <- 1/sigmasqinv
}
#      ",fill = TRUE)
#sink()
```