Survival analysis and regression – Part 2

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BCCDC Biostats Session
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Session overview

- In this session we will discuss
 - continue exploring regression models for survival data
 - an example of a Cox proportional hazards regression

Background



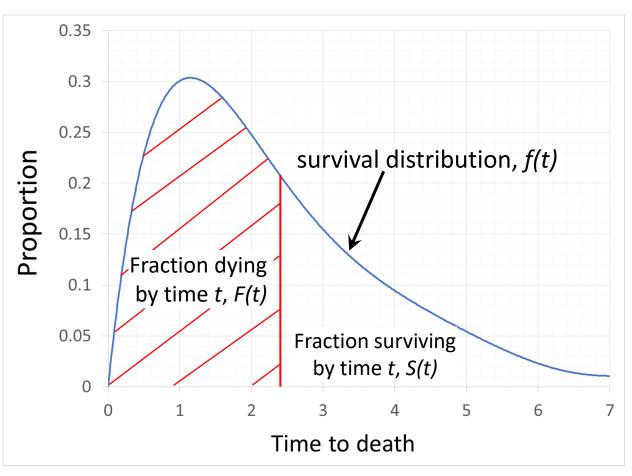
- Simply put, 'survival analysis' is the analysis of longitudinal event data, specifically the <u>time-to-event</u>
- Often, and historically, these analyses focussed on the survival, or time-to-death, of people
- But, the same models apply to the time to injury, illness, admission, readmission, recovery, or any definable health or disease state, and even the time to failure of machines!

Survival analysis

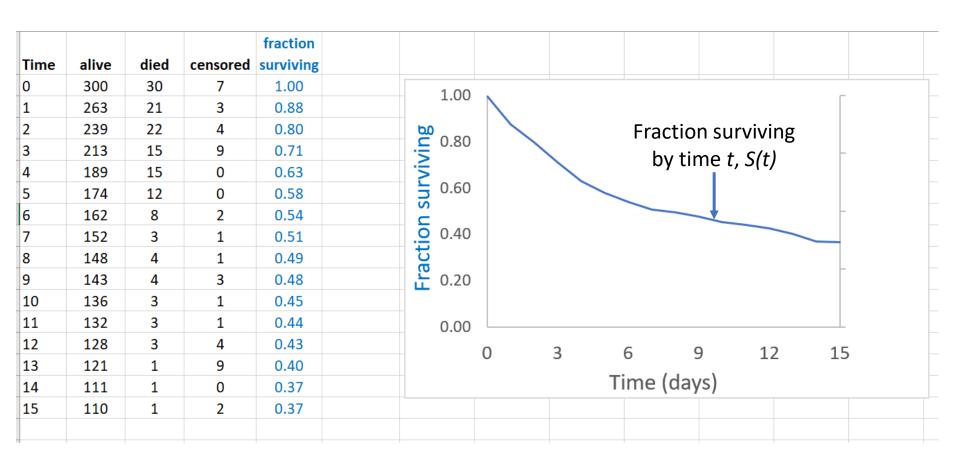


- Define event of interest, time zero, time scale and how participants exit
 - Consideration of censoring
- 2. Descriptive analysis: univariate modeling
 - KM curves and descriptive statistics
- 3. Inferential analysis: multivariate modeling
 - Cox regression (semi-parametric)

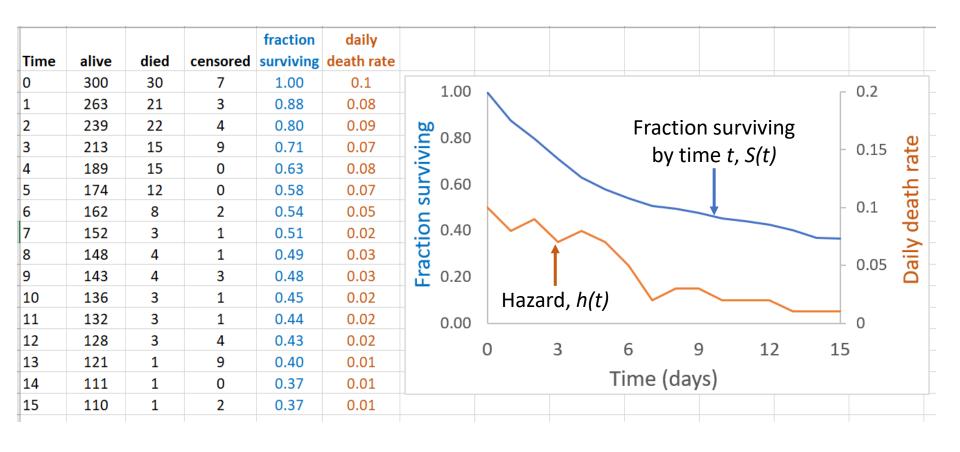
Distribution of survival times



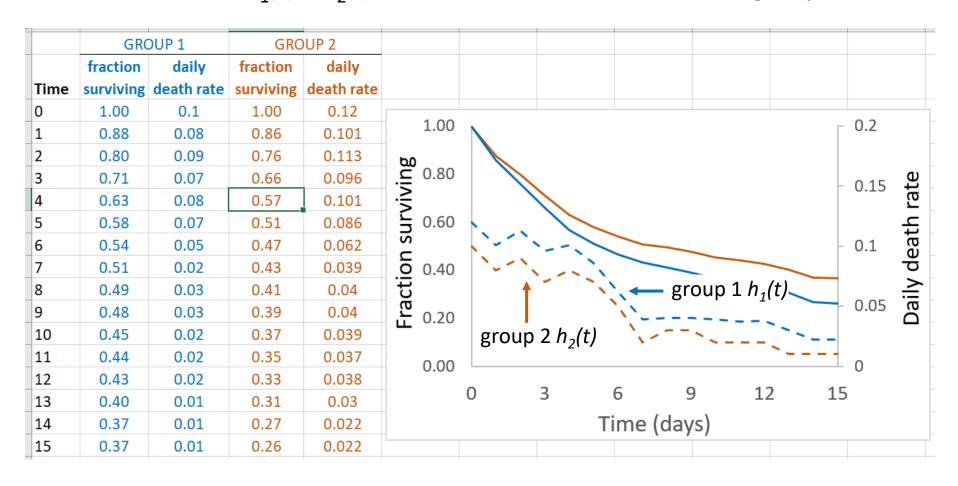
Survival curve, S(t): Fraction (or probability of) surviving by time t



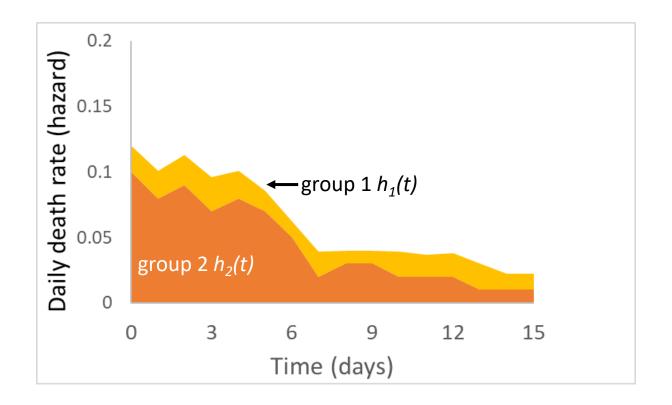
Hazard h(t): risk of death in the next small interval among those still alive



Hazard ratio $h_1(t)/h_2(t)$: ratio of hazards between two groups



If the hazards $h_1(t)$ and $h_2(t)$ remain **proportional** over time, the difference in risk can be properly summarized by a single number, the hazard ratio



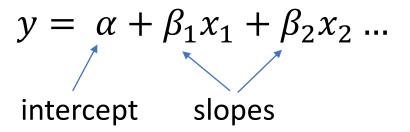


 Recall, most regression models relate observations to a *linear series* of predictors, in the general form

$$y = \alpha + \beta_1 x_1 + \beta_2 x_2 \dots$$



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$$\log(\mu_i) = \alpha + \beta_1 x_1 + \beta_2 x_2 \dots$$
 Poisson model

$$\log\left(\frac{p}{1-\rho}\right) = \alpha + \beta_1 x_1 + \beta_2 x_2 \dots \qquad \text{Logistic model}$$

- In the case of survival (time-to-event) analysis, we model the hazard
- log of the hazard ratio is the link used connect to the linear predictors

$$\log(HR) = \log\left(\frac{h(t)}{h_0(t)}\right) = \beta_1 x_1 + \beta_2 x_2 \dots$$

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Intercept slopes

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$$\log h(t) = \log(h_0(t)) + \beta_1 x_1 + \beta_2 x_2 \dots$$

$$h(t) = h_0(t)e^{\beta_1 x_1 + \beta_2 x_2 \dots}$$

- The most common proportional hazards model is the Cox regression
- Sometimes this model is termed semi-parametric -- linear predictor set is parametric, but no assumptions are made about baseline hazard $h_0(t)$ (often written as $\lambda_0(t)$)

$$h(t) = \lambda_0(t)e^{\beta_1 x_1 + \beta_2 x_2 \dots}$$

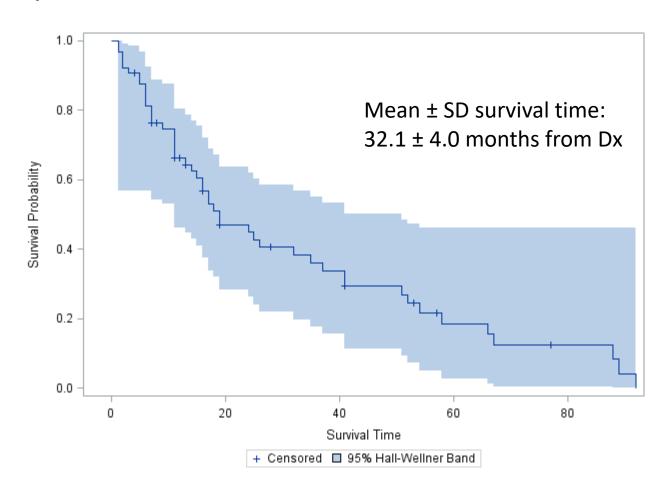
An example

- Multiple myeloma study (see Krall et al, 1975)
 - 65 patients undergoing treatment (48 died during study)
 - Analysis of survival time from diagnosis
 - Identifying factors associated with survival

| Time! | months) | IDesd July | ies en Hemi | gelobin | | white | ploode | 6/92/43 | cells tron | ninurine Seri |
|-------|---------|------------|-------------------|----------|-----|---------|--------|---------|---------------|------------------|
| Time | Status | LogBUN | HGB | Platelet | Age | LogWBC | Frac | LogPBM | Protein | SCalc |
| 1.25 | 1 | 2.2175 | 9.4 | 1 | 67 | 3.6628 | 1 | 1.9542 | 12 | 10 |
| 1.25 | 1 | 1.9395 | 12 | 1 | 38 | 3.9868 | 1 | 1.9542 | 20 | 18 |
| 2.00 | 1 | 1.5185 | 9.8 | 1 | 81 | 3.8751 | 1 | 2 | 2 | 15 |
| 2.00 | 1 | 1.7482 | 11.3 | 0 | 75 | 3.8062 | 1 | 1.2553 | 0 | 12 |
| 2.00 | 1 | 1.301 | 5.1 | 0 | 57 | 3.7243 | 1 | 2 | 3 | 9 |
| 3.00 | 1 | 1.5441 | 6.7 | 1 | 46 | 4.4757 | 0 | 1.9345 | 12 | 10 |
| 4.00 | 0 | 1.9542 | 10.2 | 1 | 59 | 4.0453 | 0 | 0.7782 | 12 | 10 |
| 4.00 | 0 | 1.9243 | 10 | 1 | 49 | 3.959 | 0 | 1.6232 | 0 | 13 |
| 5.00 | 1 | 2.2355 | 10.1 | 1 | 50 | 4.9542 | 1 | 1.6628 | 4 | 9 |
| 5.00 | 1 | 1.6812 | 6.5 | 1 | 74 | 3.7324 | 0 | 1.7324 | 5 | 9 |
| 6 00 | 1 | 1 2617 | ۵ | 1 | 77 | 2 5///1 | Λ | 1 /67/ | 1 | Q |

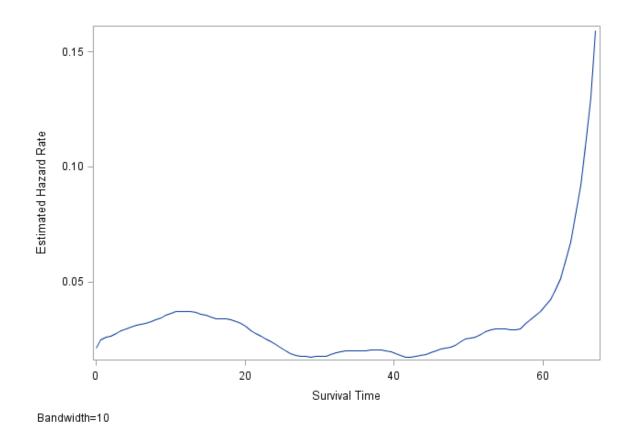
Descriptive analysis

Kaplan-Meier survival curve



Descriptive analysis

Estimated (smoothed) hazard function



Cox proportional hazards regression

$$h(t) = \lambda_0(t)e^{\beta_1 x_1 + \beta_2 x_2 \dots}$$

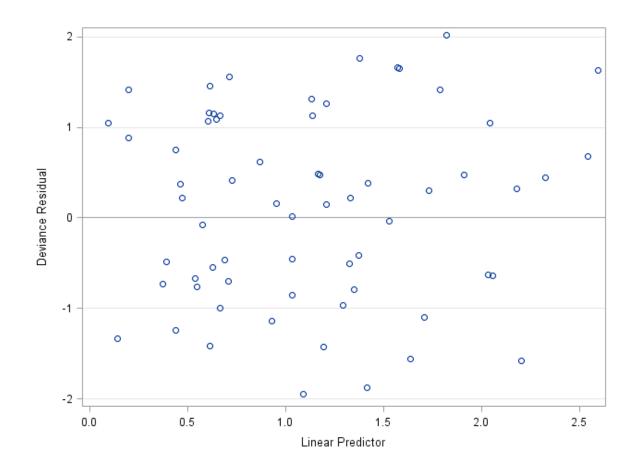
Cox proportional hazards regression

$$h(t) = \lambda_0(t)e^{\beta_1 x_1 + \beta_2 x_2 \dots}$$

| Summ | ary of the | Number of E Values | vent | and Cens | sored | |
|------|------------|---|---------|------------|----------------|--|
| Tota | I Even | t Censo | red | | rcent sored | |
| 6 | 5 48 | 3 | 17 | | 26.15 | |
| Conv | ergence cn | terion (GCON | 1V = 1E | -orsansn | en | |
| | Me | odel Fit Stati | | o y outlon | ou. | |
| | | odel Fit Stati Without Covariates | stics | With | ou. | |
| - | | Without | stics | With | ou. | |
| - | Criterion | Without Covariates | cov | With | ou. | |

| Analysis of Maximum Likelihood Estimates | | | | | | | | | | |
|--|----|-----------------------|-------------------|------------|------------|-----------------|--|--|--|--|
| Parameter | DF | Parameter Estimate | Standard Error | Chi-Square | Pr > ChiSq | Hazard Ratio | | | | |
| LogBUN | 1 | 1.79836 | 0.64833 | 7.6942 | 0.0055 | 6.040 | | | | |
| HGB | 1 | -0.12631 | 0.07183 | 3.0920 | 0.0787 | 0.881 | | | | |
| Platelet | 1 | -0.25059 | 0.50747 | 0.2438 | 0.6214 | 0.778 | | | | |
| Age | 1 | -0.01279 | 0.01948 | 0.4316 | 0.5112 | 0.987 | | | | |
| LogWBC | 1 | 0.35371 | 0.71319 | 0.2460 | 0.6199 | 1.424 | | | | |
| Frac | 1 | 0.33788 | 0.40728 | 0.6883 | 0.4068 | 1.402 | | | | |
| LogPBM | 1 | 0.35893 | 0.48603 | 0.5454 | 0.4602 | 1.432 | | | | |
| Protein | 1 | 0.01307 | 0.02617 | 0.2494 | 0.6175 | 1.013 | | | | |
| SCalc | 1 | 0.12595 | 0.10340 | 1.4837 | 0.2232 | 1.134 | | | | |

Assessing model fit (as usual, with residuals)



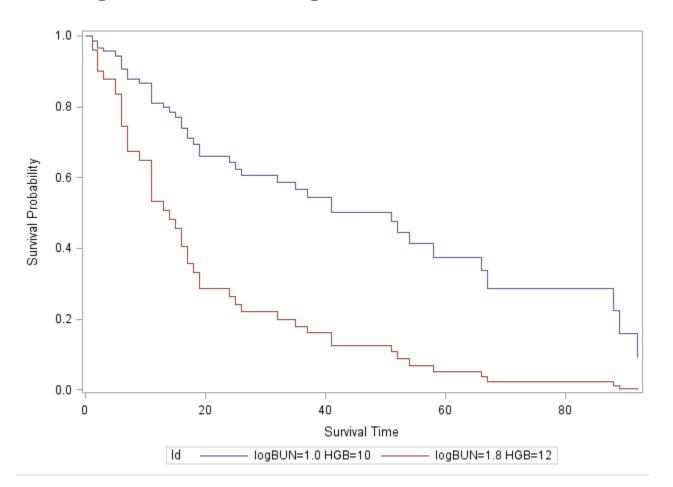
Estimating survival using fitted model

```
data Inrisks;
  length Id $20;
  input LogBUN HGB Id $12-31;
  datalines;

1.00 10.0 logBUN=1.0 HGB=10
  1.80 12.0 logBUN=1.8 HGB=12
;

proc phreg data=Myeloma plots(overlay)=survival;
  model Time*VStatus(0)=LogBUN HGB;
  baseline covariates=Inrisks out=Pred1 survival=_all__ / rowid=Id;
run;
```

Estimating survival using fitted model



References

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