# **Principal Component Analyis**

### Literature

#### Motivation

Getting an idea how PCA works and what kind of problems can be solved.

## **Background**

- 1. A measurement is defined as a collection of K items.
- 2. In general there are N measurement
- 3. The j-th measurement has K measured items. These items are arranged into a column vector denoted  $\mathbf{d}_{j}$ .

$$\mathbf{d}_j = egin{bmatrix} d_1[j] \ dots \ d_k[j] \ dots \ d_K[j] \end{bmatrix} = egin{bmatrix} d_{1,\,j} \ dots \ d_{k,\,j} \ dots \ d_{K,\,j} \end{bmatrix}$$

 $d_k[j] = d_{k\,j}$  denotes the  $\,$  j-th  $\,$  measurement / data of the  $\,$  k-th  $\,$  item.

### centering the data set

The mean value of the data set is computed from all measurements  $\mathbf{d}_j$ :  $j=1,\ldots,N$  by taking the *element-wise* average of each measurement item. The mean value is defined as a column vector:

$$E(\mathbf{d}) = egin{bmatrix} E(d_1) \ dots \ E(d_k) \ dots \ E(d_K) \end{bmatrix} = egin{bmatrix} rac{1}{N} \sum_{j=1}^N d_{1,j} \ dots \ rac{1}{N} \sum_{j=1}^N d_{k,j} \ dots \ rac{1}{N} \sum_{j=1}^N d_{K,j} \end{bmatrix}$$

The **centered** data vector are denoted  $\mathbf{x}_j$ :  $j=1,\ldots,N$ . They are computed from data vectors  $\mathbf{d}_j$  by removing the mean value.

$$\mathbf{x}_j = egin{bmatrix} d_1[j] - E(d_1) \ dots \ d_k[j] - E(d_k) \ dots \ d_K[j] - E(d_K) \end{bmatrix} = egin{bmatrix} d_{1,\,j} - E(d_1) \ dots \ d_{k,\,j} - E(d_k) \ dots \ d_{K,\,j} - E(d_K) \end{bmatrix} = egin{bmatrix} x_{1,\,j} \ dots \ x_{k,\,j} \ dots \ x_{K,\,j} \end{bmatrix}$$

Now we define a unit vector  $\mathbf{w}: \in \mathbf{R}^K$ ;  $||\mathbf{w}||=1$ . This vector shall be to determine the component of the  $\mathbf{j}$ -th measurement in the direction of vector  $\mathbf{w}$ . This vector is denoted  $\mathbf{p}_j$ . It is the projection of vector  $\mathbf{x}_j$  onto  $\mathbf{w}$ . The residual vector  $\mathbf{r}_j$  is orthogonal to vector  $\mathbf{p}_j$ .

To summarise (projection vector & residual vector):

$$\mathbf{p}_j = \left(\mathbf{x}_j^T \cdot \mathbf{w}\right) \cdot \mathbf{w} \tag{1}$$

$$\mathbf{r}_{i} = \mathbf{x}_{i} - \mathbf{p}_{i} \tag{3}$$

$$\mathbf{r}_{j} = \mathbf{x}_{j} - \left(\mathbf{x}_{j}^{T} \cdot \mathbf{w}\right) \cdot \mathbf{w} \tag{4}$$

The squared norm of the residual  $\mathbf{r}_i$  is computed from:

$$\left|\left|\mathbf{r}_{j}\right|\right|^{2}=\mathbf{r}_{j}^{T}\cdot\mathbf{r}_{j}=\left(\mathbf{x}_{j}-\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)\cdot\mathbf{w}\right)^{T}\cdot\left(\mathbf{x}_{j}-\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)\cdot\mathbf{w}\right)$$
 (8)

$$\mathbf{x}_{j}^{T} \cdot \mathbf{x}_{j} - \mathbf{x}_{j}^{T} \cdot \left(\mathbf{x}_{j}^{T} \cdot \mathbf{w}\right) \cdot \mathbf{w} - \left(\mathbf{x}_{j}^{T} \cdot \mathbf{w} \cdot \mathbf{w}\right)^{T} \cdot \mathbf{x}_{j} + \left(\mathbf{x}_{j}^{T} \cdot \mathbf{w} \cdot \mathbf{w}\right)^{T} \cdot \mathbf{x}_{j}^{T} \cdot \mathbf{w} \cdot \mathbf{w} \quad (6)$$

$$\mathbf{x}_{j}^{T}\cdot\mathbf{x}_{j}-2\cdot\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)^{2}+\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)^{2}$$
 (7)

$$=\mathbf{x}_{j}^{T}\cdot\mathbf{x}_{j}-\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)^{2}=\left|\left|\mathbf{x}_{j}\right|\right|^{2}-\left(\mathbf{x}_{j}^{T}\cdot\mathbf{w}\right)^{2}$$
(8)

We have used the unit vector property of vector  ${\bf w}$  namely  $||{\bf w}||=1$ .

## adding up the squared residual

The mean squared error is the expectation of the squared residuals for all N measurements.

$$MSE(\mathbf{w}) = rac{1}{N} \sum_{j=1}^{N} \left|\left|\mathbf{x}_{j}
ight|
ight|^{2} - rac{1}{N} \sum_{j=1}^{N} \left(\mathbf{x}_{j}^{T} \cdot \mathbf{w}
ight)^{2}$$

To make  $MSE(\mathbf{w})$  as small as possible the term

$$V = rac{1}{N} \sum_{i=1}^{N} \left( \mathbf{x}_{j}^{T} \cdot \mathbf{w} 
ight)^{2}$$

must be maximised by appropriate choice of vector  $\mathbf{w}$ . Re-writing V yields:

$$V = rac{1}{N} \sum_{j=1}^{N} \mathbf{w}^T \cdot \left( \mathbf{x}_j \cdot \mathbf{x}_j^T 
ight) \cdots \mathbf{w} = \mathbf{w}^T \cdot \left( rac{1}{N} \sum_{j=1}^{N} \mathbf{X}_j 
ight) \cdots \mathbf{w} = \mathbf{w}^T \cdot \mathbf{Z}$$

 $\mathbf{X}_j = \mathbf{x}_j \cdot \mathbf{x}_j^T : \in \mathbb{R}^{K \times K}$  is a square matrix computed as the outer product of centered data vector  $\mathbf{x}_j$ .

 $\mathbf{X}_j$  has elements  $x_{l,\,m}[j]$ .

$$x_{l,m}[j] = x_l[j] \cdot x_m[j] = (d_l[j] - E(d_l)) \cdot (d_m[j] - E(d_m))$$
(9)

$$x_{l,m}[j] = d_l[j] \cdot d_m[j] - d_l[j] \cdot E(d_m) - d_m[j] \cdot E(d_l) + E(d_l) \cdot E(d_m)$$
 (10)

The sum of all matrices  $X_i$  is matrix X:

$$\mathbf{X} = rac{1}{N} \sum_{j=1}^{N} \mathbf{X}_{j}$$

with elements  $x_{l,m}$ :

$$x_{l,m} = E(d_l \cdot d_m) - E(d_m) \cdot E(d_l) - E(d_m) \cdot E(d_l) + E(d_l) \cdot E(d_m)$$
 (11)

$$(12)$$

$$x_{l,m} = E(d_l \cdot d_m) - E(d_m) \cdot E(d_l) \tag{13}$$

For l=m matrix elements  $x_{l,l}$  are just the variance:

$$x_{l,\,l} = E(d_l \cdot d_l) - E(d_l)^2 = E(d_l^2) - E(d_l)^2$$

For  $l \neq m$  matrix elements  $x_{l,\,m}$  are the covariance.

Hence matrix  $\mathbf{X}$  is also known as the covariance matrix.

#### summary

It has been demonstrated that by appropriate choice of projection vector  $\mathbf{w}$  the maximisation of

$$V = \mathbf{w}^T \cdot \mathbf{X} \cdot \mathbf{w}$$

minimises the mean squared error  $MSE(\mathbf{w})$ .

The next step is to determine the optimum vector  ${\bf w}$  with the constraint that  ${\bf w}$  has unit length ( $||{\bf w}||=1$ ).

## Finding the vector minimising the MSE

The optimum vector must be computed with the unit length constraint.

Using the Lagrange multiplier method, the objective function  $F(\mathbf{w},\lambda)$  may be expressed like this:

$$F(\mathbf{w}, \lambda) = \mathbf{w}^T \cdot \mathbf{X} \cdot \mathbf{w} - \lambda \cdot (\mathbf{w}^T \cdot \mathbf{w} - 1)$$

### **ToDo**

Get a solid understanding of constrained optimisation with Lagrange multipliers. (in this notebook I just copied the method without having understood how it works)

The solution vector is found by setting derivatives

$$\frac{\partial}{\partial \mathbf{w}} F(\mathbf{w}, \lambda) = 0 \tag{14}$$

$$\frac{\partial}{\partial \lambda} F(\mathbf{w}, \lambda) = 0 \tag{15}$$

$$\frac{\partial}{\partial \mathbf{w}} F(\mathbf{w}, \lambda) = 2 \cdot \mathbf{X} \cdot \mathbf{w} - 2 \cdot \lambda \cdot \mathbf{w} = \mathbf{0}$$

$$rac{\partial}{\partial \lambda} F(\mathbf{w}, \lambda) = \mathbf{w}^T \cdot \mathbf{w} - 1$$

Leading to

$$\mathbf{X} \cdot \mathbf{w} = \lambda \cdot \mathbf{w} \tag{16}$$

$$\mathbf{w}^T \cdot \mathbf{w} = 1 \tag{17}$$

From the first equation we conclude that the optimum vector  $\mathbf{w}$  has been found as an **eigenvector** of the covariance matrix.

The second equation just states the fact that  $\mathbf{w}$  has unit length.

We have defined

$$V = \mathbf{w}^T \cdot \mathbf{X} \cdot \mathbf{w}$$

inserting the optimum vector yields:

$$V(\mathbf{w}) = \lambda \mathbf{w}^T \cdot \mathbf{w} = \lambda ||\mathbf{w}||^2 = \lambda$$

The MSE is minimised if we choose the eigenvector  $\mathbf{w} = \mathbf{w}_1$  with the largest eigenvalue  $\lambda = \lambda_1$ .

For the residual vector we get:

$$\mathbf{r}_{j(1)} = \mathbf{x}_j - \left(\mathbf{x}_j^T \cdot \mathbf{w}_1
ight) \cdot \mathbf{w}_1$$

and for the average of squared residuals:

$$\frac{1}{N} \sum_{j=1}^{N} ||\mathbf{r}_{j(1)}||^2 = \frac{1}{N} \sum_{j=1}^{N} ||\mathbf{x}_j||^2 - \mathbf{w}_1^T \cdot \mathbf{X} \cdot \mathbf{w}_1$$
 (18)

$$\frac{1}{N} \sum_{j=1}^{N} ||\mathbf{r}_{j(1)}||^2 = \frac{1}{N} \sum_{j=1}^{N} ||\mathbf{x}_j||^2 - \lambda_1$$
(19)

## Going one step further

We are going to project vector  $\mathbf{r}_{j(1)}$  onto some other vector  $\mathbf{w}_2$  and get a new residual vector  $\mathbf{r}_{j(2)}$ .

$$\mathbf{r}_{j(2)} = \mathbf{r}_{j(1)} - \left(\mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2
ight) \cdot \mathbf{w}_2$$

Then vector  $\mathbf{w}_2$  shall be chosen such as to minimise the sum of the squared residuals:

$$\frac{1}{N}\sum_{j=1}^{N}\left|\left|\mathbf{r}_{j(2)}\right|\right|^{2}=\frac{1}{N}\sum_{j=1}^{N}\mathbf{r}_{j(2)}^{T}\cdot\mathbf{r}_{j(2)}=\frac{1}{N}\sum_{j=1}^{N}\left(\left|\left|\mathbf{r}_{j(1)}\right|\right|^{2}-\left(\mathbf{r}_{j(1)}^{T}\cdot\mathbf{w}_{2}\right)^{2}\right)=\frac{1}{N}\sum_{j=1}^{N}\left|\left|\mathbf{r}_{j(1)}\right|\right|^{2}$$

$$rac{1}{N} \sum_{j=1}^{N} \left| \left| \mathbf{r}_{j(2)} 
ight| 
ight|^2 = (rac{1}{N} \sum_{j=1}^{N} \left| \left| \mathbf{x}_{j} 
ight| 
ight|^2 - \lambda_1) - rac{1}{N} \sum_{j=1}^{N} \left( \mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2 
ight)^2$$

Choose  $\mathbf{w}_2$  to maximise

$$rac{1}{N} \sum_{j=1}^{N} \left( \mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2 
ight)^2$$

$$\mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2 = \left(\mathbf{x}_j - \left(\mathbf{x}_j^T \cdot \mathbf{w}_1\right) \cdot \mathbf{w}_1\right)^T \cdot \mathbf{w}_2 \tag{20}$$

(21)

$$\mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2 = \mathbf{x}_j^T \cdot \mathbf{w}_2 - \mathbf{w}_1^T \cdot \left(\mathbf{x}_j^T \cdot \mathbf{w}_1\right) \cdot \mathbf{w}_2 \tag{22}$$

At this point we postulate the vectors  $\mathbf{w}_1$  and  $\mathbf{w}_2$  are **orthonormal**.

$$\mathbf{r}_{j(1)}^T \cdot \mathbf{w}_2 = \mathbf{x}_j^T \cdot \mathbf{w}_2$$

Hence

$$\frac{1}{N}\sum_{j=1}^{N}\left(\mathbf{r}_{j(1)}^{T}\cdot\mathbf{w}_{2}\right)^{2}=\frac{1}{N}\sum_{j=1}^{N}\mathbf{w}_{2}^{T}\left(\mathbf{x}_{j}\mathbf{x}_{j}^{T}\right)\cdot\mathbf{w}_{2}=\mathbf{w}_{2}^{T}\cdot(\frac{1}{N}\sum_{j=1}^{N}\mathbf{x}_{j}\mathbf{x}_{j}^{T})\cdot\mathbf{w}_{2}=\mathbf{w}_{2}^{T}\cdot\mathbf{X}\cdot\mathbf{w}_{2}$$