

Adjusting Standard ANOVA Methods to Account for Heterogeneous Variances With an Application to Turfgrass Management

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Section 1

Introduction

Background

- Slides available on my GitHub [here](#)
- PhD in Statistics from Oregon State University (2020)
- Research statistician at the Environmental Protection Agency

Background

- OSU Statistics Consulting Practicum
 - Encourage you to sign up!
 - Long format vs drop-in
 - Faculty are encouraged too - separate process
- Worked on several turfgrass projects with Alec Kowalewski and Clint Mattox
- Use Analysis of Variance (ANOVA) to study designed experiments
 - Are there *statistically significant differences* among treatment effects?
- One common problem: unequal variance / standard deviation within treatment groups
 - How can we use ANOVA to best understand our data when there is unequal variance?

Experiment Roadmap

- ① Formulate a hypothesis
- ② Choose an experimental design
- ③ Choose an analysis method
- ④ **Randomize** treatments
- ⑤ Collect data
- ⑥ Analyze data using ANOVA
 - $Y_i = \mu + \alpha_i + \epsilon_i$ (focus on one-way ANOVA)
 - Estimate treatment effects from the data
 - Do these estimates ($\hat{\alpha}$) suggest *statistically significant differences* among the true treatment effects (α)?
- ⑦ Report results

Section 2

Why ANOVA?

Properties

ANOVA has several attractive properties:

- ① Estimates of treatment effects equal the true treatment effects *on average*
 - But we only get to run the experiment once!
- ② Treatment effect confidence intervals are as small as possible
- ③ Hypothesis tests have well known forms

But 2 and 3 rely on specific **assumptions** on the errors, ϵ

Assumptions on ϵ

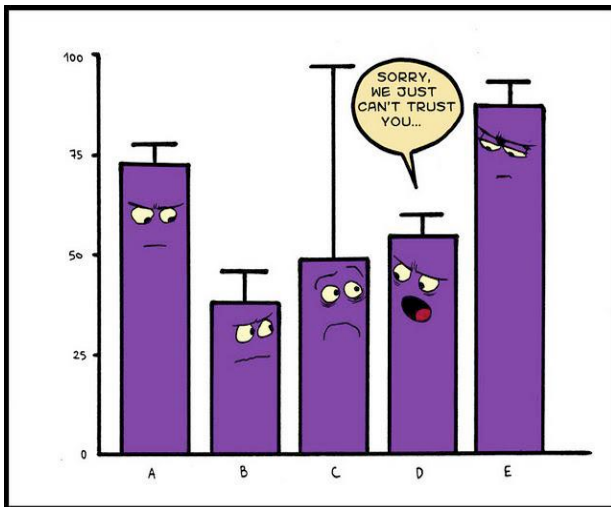
- ① Independence
 - Unit A does not depend on unit B
 - Dice roll, coin flip
- ② Normality
- ③ Constant Variance
 - Variance homogeneity
 - This presentation focuses on 3.

Assumptions on ϵ

When the constant variance assumption is violated, the ANOVA

- Estimates of treatment effects still equal the true treatment effects *on average*
 - Great!
- Treatment effect confidence intervals are too large
 - Inefficient use of resources
- Hypothesis tests don't have well known forms
 - Incorrect p-values → misleading conclusions, poor policy decisions

Assumptions



How Do I Know?

- Graphics! If it looks off, it probably is
- Ratio of largest and smallest variances
 - Suggestions of cutoff range from 1.5 to 9
- Statistical tests for constant variance
 - Levene's, Brown-Forsythe, several others
 - Come with their own assumptions

What Now?

So I know my data does not have constant variance, what now?

- Could transform the response, Y , so that the transformed Y satisfy standard assumptions
 - Can be very useful
 - Generally require a specific mean / variance relationship, $\log_e(Y)$ often used

Poses practical considerations:

- Challenging to find an appropriate transformation
- Difficult to interpret on original scale (usually of interest)
 - Significant difference between treatments on transformed scale
DOES NOT imply the same on the original scale
- What else?

Section 3

GV-ANOVA

What is GV-ANOVA?

- Can use Generalized Variance ANOVA (GV-ANOVA) to directly model variances within groups
 - Separate variance for each treatment level
 - Does not require a mean / variance relationship
 - No transformation requirement
 - Requires the estimation of more variance parameters than when using a transformation
- Goal of this talk is to expose you to another possible way to handle the non constant variance problem
- Important to be aware of both approaches

Section 4

Application

Example

```
anova_mod <- gls(response ~ trt, data = data)
anova(anova_mod)
```

```
## Denom. DF: 28
##               numDF  F-value p-value
## (Intercept)      1 10358.47 <.0001
## trt              3   30.27 <.0001
emmeans(anova_mod, list(pairwise ~ trt), adjust = "bonferroni")
```

```
## $`emmeans of trt`
##   trt emmean   SE df lower.CL upper.CL
## A    47.0 1.07 28    44.8    49.2
## B    52.3 1.07 28    50.1    54.5
## C    58.1 1.07 28    55.9    60.2
## D    60.0 1.07 28    57.8    62.2
##
## Degrees-of-freedom method: df.error
## Confidence level used: 0.95
##
## $`pairwise differences of trt`
##   contrast estimate   SE df t.ratio p.value
## A - B      -5.24 1.51 28  -3.473 0.0102
## A - C     -11.03 1.51 28  -7.301 <.0001
## A - D     -12.97 1.51 28  -8.589 <.0001
## B - C      -5.78 1.51 28  -3.829 0.0040
## B - D      -7.73 1.51 28  -5.116 0.0001
## C - D      -1.94 1.51 28  -1.287 1.0000
##
## Degrees-of-freedom method: df.error
## P value adjustment: bonferroni method for 6 tests
```


Example

```
gvanova_mod <- gls(response ~ trt, weights = varIdent(form = ~ 1|trt), data = data)
anova(gvanova_mod)
```

```
## Denom. DF: 28
##          numDF    F-value p-value
## (Intercept)      1 230058.47 <.0001
## trt              3    67.27 <.0001
emmeans(gvanova_mod, list(pairwise ~ trt), adjust = "bonferroni")
```

```
## $`emmeans of trt`
##   trt emmean      SE    df lower.CL upper.CL
##   A    47.03 0.9674  6.97    44.74    49.32
##   B    52.27 1.8105  7.01    47.99    56.55
##   C    58.05 0.5745  7.00    56.69    59.41
##   D    60.00 0.1288  7.00    59.69    60.30
##
## Degrees-of-freedom method: satterthwaite
## Confidence level used: 0.95
##
## $`pairwise differences of trt`
##   contrast estimate      SE    df t.ratio p.value
##   A - B        -5.24 2.053 10.72  -2.555 0.1636
##   A - C       -11.03 1.125 11.37  -9.799 <.0001
##   A - D       -12.97 0.976  7.21 -13.289 <.0001
##   B - C        -5.78 1.900  8.40  -3.044 0.0905
##   B - D        -7.73 1.815  7.08  -4.256 0.0220
##   C - D        -1.94 0.589  7.70  -3.302 0.0686
##
## Degrees-of-freedom method: satterthwaite
## P value adjustment: bonferroni method for 6 tests
```

Section 5

Conclusions