Variational Bayesian Inference for a Mixture of Mixtures

A Finite Mixture of Finite Gaussian Mixtures

Let $\mathbf{y} = (y_1, y_2, \dots, y_N)$ be a sample from a finite mixture of finite Gaussian mixtures. We use shared kernels by letting $\mu_{bk} = \mu_k$ and $\sigma_{bk}^2 = \sigma_k^2$ for all $k = 1, 2, \dots, K$ and $b = 1, 2, \dots, B$. Then their joint density is given by

$$p(\mathbf{y} \mid \boldsymbol{\mu}, \boldsymbol{\sigma^2}, \{\mathbf{p}\}_{b=1}^B, \boldsymbol{\phi}) = \prod_{i=1}^N \sum_{b=1}^B \phi_b \sum_{k=1}^K p_{bk} N(y_i \mid \mu_{bk}, \sigma_{bk}^2)$$
$$= \prod_{i=1}^N \sum_{b=1}^B \phi_b \sum_{k=1}^K p_{bk} N(y_i \mid \mu_k, \sigma_k^2).$$

where $\phi = (\phi_1, \phi_2, \dots, \phi_B)'$ with $\sum_{b=1}^B \phi_b = 1$ and $\{\mathbf{p}\}_{b=1}^B = \{\mathbf{p}_1, \dots, \mathbf{p}_B\}$ and $\mathbf{p}_b = (p_{b1}, p_{b2}, \dots, p_{bK})'$ with $\sum_{k=1}^K p_{bk} = 1$.

Let the weights ϕ and \mathbf{p}_b have a Dirichlet $(\alpha_{\phi}\mathbf{1})$ and Dirichlet $(\alpha_{p}\mathbf{1})$ prior respectively. For each y_i we can introduce auxiliary variable vectors $\mathbf{w} = (\mathbf{w}_1, \mathbf{w}_2, \dots, \mathbf{w}_N)$ for the "parent" distribution where $\sum_{b=1}^{B} w_{ib} = 1$ and $p(w_{ib} = 1) = \phi_b$, Then given the "parent" distribution we can introduce another set of auxiliary variable vectors $\mathbf{z}_b = (\mathbf{z}_{b1}, \mathbf{z}_{b2}, \dots, \mathbf{z}_{bN})$ for the "child" distribution where $\sum_{k=1}^{K} z_{ibk} = 1$ and $p(z_{ibk} = 1) = p_{bk}$. Then, $\mathbf{w} \mid \phi \sim \text{Multinomial}(N, \phi)$ and $\mathbf{z}_b \mid \mathbf{w}, p_b \sim \text{Multinomial}(N, \mathbf{p}_b)$.

Now the joint density can be written as

$$p(\mathbf{y} \mid \boldsymbol{\mu}, \boldsymbol{\sigma^2}, \mathbf{w}, \{\mathbf{z}_b\}_{b=1}^B) = \prod_{i=1}^N \prod_{b=1}^B \prod_{k=1}^K p(y_i \mid \mu_k, \sigma_k^2)^{w_{ib} z_{ibk}} p(\mathbf{z}_b \mid \mathbf{w}, \boldsymbol{p}_b) p(\mathbf{w} \mid \boldsymbol{\phi})$$

A Finite Mixture of Infinite Gaussian Mixtures

Extending this to a finite mixture of infinite Gaussian mixtures model using the Stick-breaking representation of a Dirichlet Process.

$$p(\mathbf{y} \mid \boldsymbol{\mu}, \boldsymbol{\sigma^2}, \mathbf{w}, \{\mathbf{z}_b\}_{b=1}^B) = \prod_{i=1}^N \prod_{b=1}^M \prod_{k=1}^\infty p(y_i \mid \mu_k, \sigma_k^2)^{w_{ib}z_{ibk}} p(\mathbf{z}_b \mid \mathbf{w}, \boldsymbol{p}_b) p(\mathbf{w} \mid \boldsymbol{\phi})$$

$$p(\mathbf{z}_b \mid \mathbf{w}, \boldsymbol{p}_b) = \text{Multinomial}(N, \mathbf{p}_b) \quad \text{where} \quad \mathbf{p}_b = (p_{b1}, p_{b2}, \dots)$$

$$p_{bk} = V_{bk} \prod_{k=1}^{k-1} (1 - V_{bk}) \quad \text{where} \quad V_{bk} \sim \text{Beta}(1, \beta_b) \quad b = 1, 2, \dots, B$$

$$p(\mathbf{w} \mid \boldsymbol{\phi}) = \text{Multinomial}(N, \boldsymbol{\phi})$$

$$p(\boldsymbol{\phi}) = \text{Dirichlet}(\alpha_{\boldsymbol{\phi}} \mathbf{1})$$

$$p(\mu_k) = N(\mu_0, \sigma_0^2)$$

$$p(\sigma_k^2) = \text{inverse-gamma}(A, B)$$

Posterior

$$p(\boldsymbol{\mu}, \boldsymbol{\sigma^2}, \mathbf{w}, \{\mathbf{z}_b\}_{b=1}^B \mid \mathbf{y}) \propto \prod_{i=1}^N \prod_{b=1}^B \prod_{k=1}^\infty p(y_i \mid \mu_k, \sigma_k^2)^{w_{ib} z_{ibk}} p(z_{ibk} \mid \mathbf{w}, \boldsymbol{p}_b) p(w_{ib} \mid \boldsymbol{\phi}) p(p_{bk}) p(\phi_b) p(\mu_k) p(\sigma_k^2)$$

Full Conditionals

$$\begin{split} p(w_{ib} = 1 \mid \cdot) &\propto \phi_b \prod_{k=1}^{\infty} p_{bk} N(y_i \mid \mu_k, \sigma_k^2) \quad \text{for} \quad b = 1, 2, \dots, B, \quad i = 1, 2, \dots, N \\ p(z_{ibk} = 1 \mid \cdot) &\propto p_{bk} N(y_i \mid \mu_k, \sigma_k^2) \prod_{b=1}^{B} \phi_b \quad \text{for} \quad k = 1, 2, \dots, \infty, \quad i = 1, 2, \dots, N \\ p(V_{bk} \mid \cdot) &= \text{Beta} \Big(1 + n_{bk}, \beta + \sum_{j=k+1}^{\infty} n_{bj} \Big) \quad \text{where} \quad n_{bk} = \sum_{i=1}^{N} z_{ibk} w_{ib} \quad k = 1, 2, \dots \\ p(\phi \mid \cdot) &= \text{Dirichlet}(\alpha) \quad \alpha_j = \alpha_{\phi_j} + \sum_{i=1}^{N} y_i w_{ib} \\ p(\mu_k \mid \cdot) &= N(a^{-1}b, a^{-1}) \quad \text{where} \quad a = \frac{1}{\sigma_0^2} + \frac{\sum_{b=1}^{B} n_{bk}}{\sigma_k^2} \quad \text{and} \quad b = \frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^{N} \sum_{b=1}^{B} y_i z_{ibk} w_{ib}}{\sigma_k^2} \quad k = 1, 2, \dots \\ p(\sigma_k^2 \mid \cdot) &= \text{inverse-gamma} \Big(A + \frac{\sum_{b=1}^{B} n_{bk}}{2}, B + \frac{\sum_{i=1}^{N} \sum_{b=1}^{B} (y_i - \mu_k)^2 z_{ibk} w_{ib}}{2} \Big) \quad k = 1, 2, \dots \end{split}$$

Optimal Densities

The optimal densities are given by

$$q_i^*(\theta_i) \propto \exp\{E_{-\theta_i} \log p(\theta_i \mid \cdot)\}.$$

Optimal density for w_{ib}

$$\log q_w^*(w_{ib}) \propto E_{\phi,\mu_k,\sigma_k^2} \log p(w_{ik} = 1 \mid \cdot)$$

$$= E_{\phi}[\log \phi_b] + \sum_{k=1}^{\infty} \left[E_{\mathbf{p}}[\log p_{bk}] - \frac{1}{2} \log(2\pi) - \frac{1}{2} E_{\sigma_k^2}[\log \sigma_k^2] - \frac{1}{2E_{\sigma_k^2}[\sigma_k^2]} E_{\mu_k}[(y_i - \mu_k)^2] \right]$$

$$= \log \tilde{\phi}_{ib_{q(\phi)}}$$

It follows that

$$q_w^*(w_i) = \text{categorical}(\phi_{q(\phi)}) \quad \text{where} \quad \phi_{ib_{q(\phi)}} = \frac{\tilde{\phi}_{ib_{q(\phi)}}}{\sum_{b=1}^B \tilde{\phi}_{ib_{q(\phi)}}}$$

Optimal density for z_{ibk}

$$\begin{split} \log q_z^*(z_{ibk}) &\propto E_{\pmb{p},\mu_k,\sigma_k^2} \log p(z_{ibk} = 1 \mid \cdot) \\ &= E_{\mathbf{p}}[\log p_{bk}] - \frac{1}{2} \log(2\pi) - \frac{1}{2} E_{\sigma_k^2}[\log \sigma_k^2] - \frac{1}{2E_{\sigma_k^2}[\sigma_k^2]} E_{\mu_k}[(y_i - \mu_k)^2] + \sum_{b=1}^B E_{\pmb{\phi}} \log(\phi_b) \\ &= \log \tilde{p}_{ibk_{g(p)}} \end{split}$$

It follows that

$$q_z^*(z_{ib}) = \text{categorical}(\boldsymbol{p}_{q(p)}) \quad \text{where} \quad p_{ibk_{q(p)}} = \frac{\tilde{p}_{ibk_{q(p)}}}{\sum_{k=1}^{\infty} \tilde{p}_{ibk_{q(p)}}}$$

Optimal density for V_{bk}

$$q_V^*(V_{bk}) = \text{Beta}\left(\alpha_{q(V_{bk})}, \beta_{q(V_{bk})}\right) \text{ where } \alpha_{q(V_{bk})} = 1 + E_{w,z}[n_{bk}], \quad \beta_{q(V_{bk})} = \beta + \sum_{j=k+1}^{\infty} E_{w,z}[n_{bj}]$$

Optimal density for ϕ

$$q_{\phi}^*(\phi) = \text{Dirichlet}(\boldsymbol{\alpha}_{q(\phi)}) \text{ where } \alpha_{q(\phi_j)} = \alpha_{\phi_j} + \sum_{i=1}^N y_i E_w[w_{ib}]$$

Optimal density for μ_k

$$\begin{split} q_{\mu_k}^*(\mu_k) &= N(\mu_{q(\mu_k)}, \sigma_{q(\mu_k)}^2) \\ \sigma_{q(\mu_k)}^2 &= \left(\frac{1}{\sigma_0^2} + \frac{E_{w,z}[n_{bk}]}{E_{\sigma_k^2}[\sigma_k^2]}\right)^{-1} \\ \mu_{q(\mu_k)} &= \sigma_{q(\mu_k)}^2 * \left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^N \sum_{b=1}^B y_i E_z[z_{ibk}] E_w[w_{ib}]}{E_{\sigma_k^2}[\sigma_k^2]}\right) \end{split}$$

Optimal density for σ_k^2

$$\begin{split} q_{\sigma_k^2}^*(\sigma_k^2) &= \text{inverse-gamma}\Big(A_{q(\sigma_k^2)}, B_{q(\sigma_k^2)}\Big) \\ A_{q(\sigma_k^2)} &= A + \frac{E_{w,z}[n_{bk}]}{2}, \\ B_{q(\sigma_k^2)} &= B + \frac{\sum_{i=1}^N \sum_{b=1}^B E_{w,z,\mu_k}[(y_i - \mu_k)^2 z_{ibk} w_{ib}]}{2} \end{split}$$

Expectations

$$\begin{split} E_{\phi}[\log \phi_b] &= \psi(\alpha_{q(\phi_b)}) + \psi\left(\sum_{b=1}^{B} \alpha_{q(\phi_b)}\right) \\ E_{p}[\log p_{bk}] &= E_{V}[\log V_{bk}] + \sum_{j=1}^{k-1} E_{V}[\log(1 - V_{bj})] \\ &= \psi(\alpha_{q(V_{bk})}) - \psi(\alpha_{q(V_{bk})} + \beta_{q(V_{bk})}) + \sum_{j=1}^{k-1} [\psi(\beta_{q(V_{bj})}) - \psi(\alpha_{q(V_{bj})} + \beta_{q(V_{bj})})] \\ E_{\sigma_k^2}[\log \sigma_k^2] &= \log(B_{q(\sigma_k^2)}) - \psi(A_{q(\sigma_k^2)}) \\ E_{\sigma_k^2}[\sigma_k^2] &= \frac{B_{q(\sigma_k^2)}}{A_{q(\sigma_k^2)}} \\ E_{\mu_k}[\mu_k] &= \mu_{q(\mu_k)} \\ E_{\mu_k}[(y_i - \mu_k)^2] &= (y_i - \mu_{q(\mu_k)})^2 + \sigma_{q(\mu_k)}^2 \end{split}$$