# **Birthweight Smoking Analysis**

**Initial Setup** 

```
import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
import statsmodels.api as sm

# Converted to csv since pandas wasn't playing nice
smoking_df = pd.read_csv("birthweight_smoking_1.csv")
# Reorder according to the documentation order
smoking_df = smoking_df[["birthweight", "smoker", "age", "educ", "unmarried"
display(smoking_df)
```

	birthweight	smoker	age	educ	unmarried	alcohol	drinks	tripre1	tripre2	triķ
0	4253	1	27	12	1	0	0	1	0	
1	3459	0	24	16	0	0	0	0	1	
2	2920	1	23	11	0	0	0	1	0	
3	2600	0	28	17	0	0	0	1	0	
4	3742	0	27	13	0	0	0	1	0	
•••		•••		•••		•••	•••	•••	•••	
2995	2520	0	42	12	0	0	0	0	1	
2996	3062	0	27	17	0	0	0	1	0	
2997	3799	0	28	12	0	0	0	0	1	
2998	2070	1	21	11	0	0	0	1	0	
2999	2948	0	23	14	0	0	0	1	0	

3000 rows x 12 columns

1. Get to know your data. Make histograms and summary statistics of your data to get a sense of distributions.

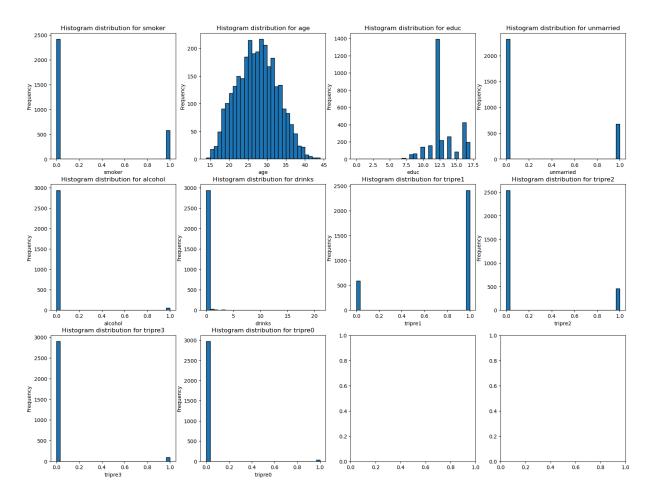
```
In [28]: # Summary statistics for all numerical columns
    print(smoking_df.describe())

# Create histograms for relevant variables
    fig, axes = plt.subplots(3, 4, figsize=(20, 15))
    axes = axes.flatten()

predictor_vars = list(smoking_df.columns)[1:11]
    for i, predictor_var in enumerate(predictor_vars):
        axes[i].hist(smoking_df[predictor_var], bins=30, edgecolor="black")
```

```
axes[i].set_title(f"Histogram distribution for {predictor_var}")
     axes[i].set_xlabel(predictor_var)
     axes[i].set_ylabel("Frequency")
 plt.show()
       birthweight
                          smoker
                                                        educ
                                                                unmarried
                                                                            \
                                           age
count
       3000.000000
                     3000.000000
                                   3000.000000
                                                3000.000000
                                                              3000.000000
mean
       3382,933667
                        0.194000
                                     26.889000
                                                  12.907000
                                                                 0.226667
std
        592.162889
                        0.395495
                                     5.362487
                                                   2.166699
                                                                 0.418745
        425.000000
                        0.000000
min
                                     14.000000
                                                   0.000000
                                                                 0.000000
25%
       3062.000000
                        0.000000
                                     23.000000
                                                  12.000000
                                                                 0.000000
50%
       3420.000000
                        0.000000
                                     27.000000
                                                  12.000000
                                                                 0.000000
75%
       3750.000000
                        0.000000
                                     31.000000
                                                  14.000000
                                                                 0.000000
                                                  17.000000
max
       5755.000000
                        1.000000
                                     44.000000
                                                                 1.000000
                                       tripre1
                                                    tripre2
                                                                  tripre3
           alcohol
                          drinks
count
       3000.000000
                     3000.000000
                                   3000.000000
                                                3000.000000
                                                              3000.000000
          0.019333
                        0.058333
                                      0.804000
                                                   0.153000
                                                                 0.033000
mean
std
          0.137717
                        0.687814
                                      0.397035
                                                   0.360048
                                                                 0.178666
min
          0.000000
                        0.000000
                                      0.000000
                                                   0.000000
                                                                 0.000000
25%
          0.000000
                        0.000000
                                      1.000000
                                                   0.000000
                                                                 0.000000
50%
          0.000000
                        0.000000
                                      1.000000
                                                   0.000000
                                                                 0.000000
75%
          0.000000
                        0.000000
                                      1.000000
                                                   0.000000
                                                                 0.000000
                       21.000000
max
          1.000000
                                      1.000000
                                                   1.000000
                                                                 1.000000
```

	tripre0	nprevist
count	3000.000000	3000.000000
mean	0.010000	10.991667
std	0.099515	3.672069
min	0.000000	0.000000
25%	0.000000	9.000000
50%	0.000000	12.000000
75%	0.000000	13.000000
max	1.000000	35.000000



A. What is the average value of birthweight for mothers who smoke? For mothers who don't smoke?

smoker\_avg\_birthweight=3178.831615120275
non\_smoker\_avg\_birthweight=3432.0599669148055

2. Consider associations. Plot each predictor (variables 2 through 11 in the pdf data description) against the response (birthweight). You could also do a quick line fit or get its correlation. Correlation is with "cor()". A line fit can be achieved using the linear model function. Two commands, model<-lm(response~predictor), followed by summary(model) will give you least squares result for an individual predictor (e.g., smoker) against the response. This will give you a rough idea of what might be important. Try for regressions 2 through 11.

```
In []: predictors = list(smoking_df.columns)[1:11]
    fig, axes = plt.subplots(3, 4, figsize=(20, 25))
    axes = axes.flatten()

correlations = {}
    regression_results = {}
```

```
for i, predictor in enumerate(predictors):
   # Calculate correlation
   correlation = smoking_df[predictor].corr(smoking_df["birthweight"])
   correlations[predictor] = correlation
   # Perform simple linear regression
   intercept = sm.add_constant(smoking_df[predictor])
   model = sm.OLS(smoking df["birthweight"], intercept).fit()
   regression_results[predictor] = model
   # Plot predictor vs birthweight
   axes[i].scatter(smoking_df[predictor], smoking_df["birthweight"])
   # Add regression line
   x_range = np.linspace(smoking_df[predictor].min(), smoking_df[predictor]
   y_pred = model.params[0] + model.params[1] * x_range
   axes[i].plot(x_range, y_pred, "r--", linewidth=2)
   axes[i].set_xlabel(predictor)
   axes[i].set_ylabel('birthweight')
   axes[i].set_title(f'{predictor} vs birthweight\nCorr: {correlation:.3f},
   print(f"{predictor} {model.summary()=}")
plt.show()
```

smoker model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

				JLS F	kegress 	10n ke	SULTS 		
=======================================	=====	====	======	=====	=====	=====	========	=======	=======
Dep. Variab	le:		bi	rthwe	eight	R-squ	ared:		0.0
Model: 28					0LS	Adj.	R-squared:		0.0
Method: 28			Leas <sup>.</sup>	t Squ	iares	F-sta	tistic:		88.
Date:			Sun, 10	Aug	2025	Prob	(F—statistic	:	1.09e-
20 Time:				05:2	23:37	Log-L	ikelihood:		-2336
4. No. Observa	tions:				3000	AIC:			4.673e+
04 Df Residual	s:				2998	BIC:			4.674e+
04 Df Model:					1				
Covariance	Type:		ı	nonro	bust				
=======================================	=====	====		====					
5]		coef	std 	err 		t 	P> t	[0.025	0.97
 const	3432.	0600	11	.871	289	.115	0.000	3408.784	3455.3
36	34321	0000		.071	203	.113	01000	54001704	545515
smoker 83	-253 <b>.</b>	2284	26	.951	-9	.396	0.000	-306.074	-200.3
======================================	=====	====	======	 473	===== 3.891	Durhi	======= n-Watson:	=======	1.9
73	- \ -								
Prob(Omnibus	S):				0.000		e-Bera (JB):		1247.4
Skew: 71				-0	858	Prob(	JB):		1.30e-2
Kurtosis: 64				5	652	Cond.	No.		2.
=======================================	=====	====	======	====	=====	=====	========	=======	=======
Notes: [1] Standard ctly specification		rs a	ssume tl	hat t	the cov	arianc	e matrix of	the errors	is corre
age model.s	ummary	( )=<	class '	stats	smodels	.iolib	.summary.Sum	nmary'>	
			(	OLS F	Regress	ion Re	sults		
	=====	====	======	====	=====	=====	========	=======	=======
Dep. Variab	le:		bi	rthwe	eight	R-squ	ared:		0.0

OLS Adj. R-squared:

0.0

Model:

06 Method:		Least Squar	es	F–sta	tistic:		19.
35 Date:	Sun	, 10 Aug 20	)25	Prob	(F—statisti	c):	1.13e-
05 Time:		, 05:23:			ikelihood:		-2339
8. No. Observation	nc:		000	AIC:			4.680e+
04	131						
Df Residuals:		28	998	BIC:			4.681e+
Df Model: Covariance Type	e:	nonrobu	1 ıst				
=======================================	=======	========	:====	=====	========	=======	=======
5]	coef	std err		t	P> t	[0.025	0.97
const 314	15.1747	55.119	57	.061	0.000	3037.099	3253.2
age	8.8422	2.010	4	.398	0.000	4.901	12.7
84 ========	-=====	=======	====	=====	========	=======	
== Omnibus:		459.5	571	Durbi	n-Watson:		1.9
77 Prob(Omnibus):		0.0	000	Jarqu	e-Bera (JB)	:	1196.2
54 Skew:		-0.8	38	Prob(.	JB):		1.73e-2
60 Kurtosis:		5.6	500	Cond.	No.		14
0.							
==							
Notes: [1] Standard Enctly specified.		me that the	e cov	arianc	e matrix of	the errors	is corre
educ model.sumn	mary()= <cl< td=""><td>ass 'statsm</td><td>nodel</td><td>s.ioli</td><td>b.summary.Su</td><td>ummary'&gt;</td><td></td></cl<>	ass 'statsm	nodel	s.ioli	b.summary.Su	ummary'>	
		OLS Reg					
==							
Dep. Variable: 11		birthweig		•			0.0
Model: 11		C	)LS	Adj. I	R-squared:		0.0
Method: 56		Least Squar	es	F–sta	tistic:		33.
Date: 09	Sun	, 10 Aug 20	)25	Prob	(F—statisti	c):	7 <b>.</b> 65e-
Time:		05:23:	37	Log-L	ikelihood:		-2339
No. Observation	ns:	30	000	AIC:			4.679e+

04 Df Residuals: 04 Df Model: Covariance Type:			1	BIC:			4.680e+
	======		====	======	=======	========	=======
==	coef	std err		t	P> t	[0.025	0.97
5]							
const 3011 91	. 8137	64.963	46	362	0.000	2884.437	3139.1
	.7534	4.964	5	793	0.000	19.021	38.4
======================================		 441 <sub>-</sub> 1	77	Durhin			1.9
73							
Prob(Omnibus): 87		0.0	000	Jarque	-Bera (JB):	1	1139.8
Skew:		-0.8	809	Prob(J	B):		3.00e-2
48 Kurtosis: 9.5		5.5	49	Cond.	No.		7
=========	======	=======	====	======	=======	========	=======
==							
Notes: [1] Standard Errctly specified.							
unmarried model.	summary(	)= <class 's<="" td=""><td>tats</td><td>smodels.</td><td>iolib.summa</td><td>ary.Summary</td><td>'&gt;</td></class>	tats	smodels.	iolib.summa	ary.Summary	'>
			,	sion Res			
=======================================	======	=======		-=====		========	
Dep. Variable: 41		birthweig	ht	R-squa	red:		0.0
Model:		C	LS	Adj. R	-squared:		0.0
41 Method:		Least Squar	es	E c+a+	istic:		
9.2		Ecast Squa.	-	F-51 a1			12
Data:	_					,	12
Date: 29	Sur	n, 10 Aug 20	25			c):	12 2.46e-
29 Time:	Sur	05:23:		Prob (		c):	
29		05:23:		Prob (	F—statistio	c):	2.46e-
<pre>29 Time: 4. No. Observations 04</pre>		05:23: 30	37 100	Prob ( Log-Li AIC:	F—statistio	c):	2.46e- -2334 4.669e+
<pre>29 Time: 4. No. Observations 04 Df Residuals: 04</pre>		05:23: 30	37	Prob ( Log-Li	F—statistio	c):	2.46e- -2334
29 Time: 4. No. Observations 04 Df Residuals: 04 Df Model:	:	05:23: 30	37 900 998	Prob ( Log-Li AIC:	F—statistio	c):	2.46e- -2334 4.669e+
<pre>29 Time: 4. No. Observations 04 Df Residuals: 04</pre>	:	05:23: 30 29 nonrobu	37 900 998 1	Prob ( Log-Li AIC: BIC:	F—statistion		2.46e- -2334 4.669e+

const 85	3448.0780	12.040	286.3	396 0.0	000	3424.471	3471.6
unmarried 18	-287.4015	25.288	-11.3	365 0.	000	-336.985	-237.8
=======================================		=======	======	=======	======	=======	======
Omnibus: 75		426.	713 I	Durbin-Wats	on:		1.9
Prob(Omnibu	ıs):	0.	000 .	Jarque-Bera	(JB):		1112.4
Skew:		-0.	782	Prob(JB):			2.77e-2
Kurtosis: 54		5.	540	Cond. No.			2.
========		=======	=====	=======	======	=======	======
==							

# Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.  $\hfill \Box$ 

alcohol model.summary()=<class 'statsmodels.iolib.summary.Summary'>

111111	IIII									
		0LS	Regres	sion Re	sults					
=======================================	========	========	=====	======	=======	========	=======			
Dep. Varia	ble:	birthw	eight	R-squ	ared:		0.0			
01 Model:			01.5	Adi. I	R-squared:		0.0			
01			020		· squarear		0.0			
Method: 98		Least Sq	uares	F-sta	tistic:		3.3			
Date:		Sun, 10 Aug	2025	Prob	(F-statist	ic):	0.06			
54		25					22.40			
Time: 6.		05:	23:3/	Log-L:	ikelihood:		-2340			
No. Observ	ations:		3000	AIC:			4.682e+			
04 Df Residua	1c•		2998	BIC:			4.683e+			
04			2990	DIC.			4:0036+			
Df Model:	_		1							
Covariance	Type:	nonr	obust =====	======	=======	========	=======			
==										
5]	coef	f std err		t	P> t	[0.025	0.97			
 const	3385.7308	3 10.913	31	0.246	0.000	3364.333	3407.1			
29										
alcohol 13	-144.6791	1 78 <b>.</b> 486	_	1.843	0.065	-298 <b>.</b> 571	9.2			
=======	========		=====	======	=======	========	=======			

==							
Omnibus:		448	589	Durbi	n-Watson:		1.9
84 Prob(Omnibus):		0	.000	Jarqu	e-Bera (JB):		1145.1
53 Skew:		-0	826	Prob(	JB):		2.15e-2
49 Kurtosis:		5.	.537	Cond.	No.		7.
27							
=======================================	======	=======	=====	=====	========		======
Notes: [1] Standard Erroctly specified.							is corre
drinks model.summ	nary()=<	class 'sta	atsmode	els.io	lib.summary.	Summary'>	
		OLS Re	egress	ion Re	sults		
=======================================		:=======	======	=====	========	========	=======
Dep. Variable: 01		birthwe	ight	R-squ	ared:		0.0
Model:			0LS	Adj.	R-squared:		0.0
01 Method:		Least Squa	ares	F-sta	tistic:		3.1
52 Date:	Sun	, 10 Aug 2	2025	Prob	(F—statistic	:):	0.07
59 Time:		05:23	3:37	Log-L	ikelihood:		-2340
<pre>6. No. Observations:</pre>	:	3	3000	AIC:			4.682e+
04 Df Residuals: 04		2	2998	BIC:			4.683e+
Df Model:			1				
Covariance Type:					========	:=======	=======
==					B 111	[0.025	0.07
5]	соет	sta err		τ	P> t	[0.025	0.97
const 3384.	5613	10.846	312	.048	0.000	3363.294	3405.8
drinks -27.	.9022	15.715	-1.	.775	0.076	-58.716	2.9
=======================================	======	:======	======	=====	========	:======	=======
== Omnibus:		449	492	Durbi	n-Watson:		1.9
83 Prob(Omnibus):		0	.000	Jarqu	e-Bera (JB):		1150.3
03 Skew: 50		-0	826	Prob(	JB):		1.64e-2
Kurtosis:		5	544	Cond.	No.		1.

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### Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.  $\footnote{\colored}$ 

tripre1 model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

=======	======	=====	======	=====	======	========	:=======	=======
== Dep. Varia	ble:		birthwe	ight	R-squa	ared:		0.0
13 Model:				0LS	۸ ط <del>۱</del> ۱	R-squared:		0.0
12				ULS	Auj. I	K-Squareu:		0.0
Method: 34		l	_east Squ	ares	F-sta	tistic:		38.
Date: 10		Sun,	, 10 Aug	2025	Prob	(F—statistic	:):	6.74e-
Time:			05:2	3:37	Loa-L:	ikelihood:		-2338
9.					5			
No. Observ	ations:			3000	AIC:			4.678e+
04 Df Residua	ls:			2998	BIC:			4.679e+
04								
Df Model:	_			1				
Covariance			nonro			========		
==								
	C	oef	std err		t	P> t	[0.025	0.97
5]								
const	3248.1	.786	24.270	133	8.837	0.000	3200.592	3295.7
66 tripre1	167 6	050	27.067	6	102	0.000	114.534	220.6
77	107.0	סכשו	27.007	U	192	0.000	114.334	220.0
=======================================	======	======	=======	=====	=====	=======	:=======	=======
Omnibus:			429	.414	Durbin	n-Watson:		1.9
82 Prob(Omnib	s) •		ρ	.000	larque	e-Bera (JB):		1092.1
92	u3/:		U	. 000	Jarque	. Bela (JD).		1037 1
Skew: 38			-0	.795	Prob(	JB):		6.81e-2
Kurtosis: 31			5	<b>.</b> 492	Cond.	No.		4.
=======================================	======	======	======	=====	:=====	========	:=======	=======

#### Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

tripre2 model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

		0LS R	Regress	sion Re	sults 		
==							
Dep. Variable: 04		birthwe	eight	R-squ	ared:		0.0
Model:			0LS	Adj.	R-squared:		0.0
04 Method:		Least Squ	iares	F-sta	tistic:		13.
40	Cum	10 1	2025	Dunah	/F atatioti	۵).	0 0002
Date: 56	Sun	, 10 Aug	2025	Prob	(F—statisti	C):	0.0002
Time:		05:2	23:37	Log-L	ikelihood:		-2340
<pre>1. No. Observations:</pre>			3000	AIC:			4.681e+
04 Df Residuals:			2998	BIC:			4.682e+
04 Df Model:			1				
Covariance Type:		nonro					
=======================================	======	======	=====	=====	=======	=======	======
r1	coef	std err		t	P> t	[0.025	0.97
5] 							
 const 3399.	7214	11.723	290	0.002	0.000	3376.735	3422.7
08							
tripre2 –109. 58	7235	29.971	-3	8.661	0.000	-168.489	-50.9
=======================================	======	=======	=====	=====	========	=======	======
Omnibus:		447	.961	Durbi	n-Watson:		1.9
81 Prob(Omnibus):		0	.000	Jarqu	e-Bera (JB)	:	1161.0
36 Skew:		-0	.820	Prob(	JB):		7.66e-2
53 Kurtosis:		-	E60	Cond.	No		2
85		3	5.569	Cona.	NO.		2.
=======================================	======	======	=====		=======	=======	======
Notes: [1] Standard Erroctly specified.	ors assu	me that t	the cov	/arianc	e matrix of	the errors	is corre
tripre3 model.sum	nmary()=	<class 's<="" td=""><td>tatsmo</td><td>dels.i</td><td>olib.summar</td><td>y.Summary'&gt;</td><td></td></class>	tatsmo	dels.i	olib.summar	y.Summary'>	
		OLS R	Regress	sion Re	sults		
	======	======	=====	=====	=======	========	======
== Dep. Variable:		birthwe	eight	R-squ	ared:		0.0

OLS Adj. R-squared:

02 Model:

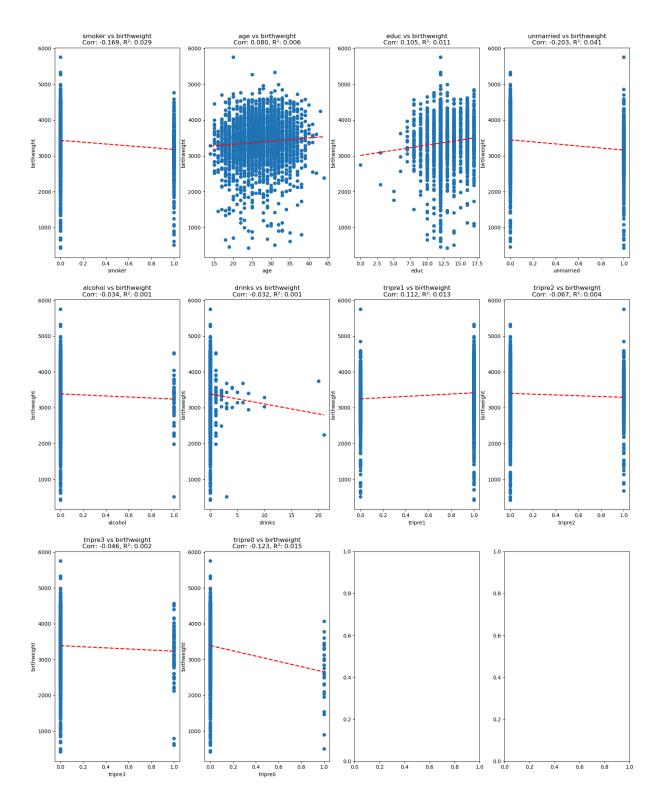
0.0

02							
Method:		Least Squa	res	F-sta	atistic:		6.4
95 Date:	Sun	, 10 Aug 2	025	Prob	(F-statisti	c):	0.01
09 Time:		05:23	:37	Loa-l	_ikelihood:		-2340
4.							
No. Observations 04	5:	3	000	AIC:			4.681e+
Df Residuals: 04		2	998	BIC:			4.682e+
Df Model:			1				
Covariance Type:			ust =====			========	=======
==	coef	std err		+	P> t	[0.025	0.97
5]	2021	Std Cil			17 [6]	[01023	0137
const 3388 56	3.0190	10.984	308	3.444	0.000	3366.482	3409.5
tripre3 -154 40	1.0998	60.466	-2	2.549	0.011	-272.659	-35.5
=======================================		=======	====	=====		=======	=======
Omnibus:		443.	461	Durbi	in-Watson:		1.9
86 Prob(Omnibus):		0.	000	Jarqı	ue-Bera (JB)	:	1123.6
46 Skew:		_0	819				1.01e-2
44							
Kurtosis: 60		5.	511	Cond.	. No.		5.
=======================================	======	=======	====	======	=======	========	=======
Notes:							
<pre>[1] Standard Err ctly specified.</pre>	ors assu	me that th	ie cov	/ariand	ce matrix of	the errors	is corre
tripre0 model.su	ummary()=	<class 'st<="" td=""><td>atsmo</td><td>odels.i</td><td>iolib.summar</td><td>y.Summary'&gt;</td><td></td></class>	atsmo	odels.i	iolib.summar	y.Summary'>	
		0LS Re	gress	sion Re	esults		
		=======	=====	======		=======	
== Dep. Variable: 15		birthwei	.ght	R-squ	uared:		0.0
Model:			0LS	Adj.	R-squared:		0.0
15 Method:		Least Squa	res	F-sta	atistic:		46.
43 Date:	Sun	, 10 Aug 2	025	Prob	(F—statisti	c):	1.14e-
11 Time:		05:23			_ikelihood:		-2338
5.					TING CTHOOD.		
No. Observations	S:	3	000	AIC:			4.677e+

04 Df Residuals: 04 Df Model: Covariance Type:	299 nonrobus	1			4.679e+
coef	std err	t	P> t	[0.025	0.97
 const 3390.2825 28 tripre0 -734.8825 26		314.368 -6.814	0.000	3369 <b>.</b> 137 -946 <b>.</b> 339	3411.4 -523.4
======================================	425.98 0.00 -0.79	0 Jarque		:	1.9 1072.7 1.14e-2
33 Kurtosis: 0.1 ===================================	5.46			========	1

#### Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.  $\footnote{\colored{A}}$ 



A. What does a regression of birthweight on the binary variable smoker suggest about the relationship between maternal smoking and infant birthweight?

Looking at the regression of birthweight on smoker, there seems to be a negative correlation between maternal smoking and infant birthweight, that is, smoking seems to have a negative effect and reduce the infant birthweight.

B. Do you think the regression above accurately captures the impact of smoking on birthweight? (Consider the assumptions of the linear regression model and whether

they are met. Hint: do you think smoking is uncorrelated with other factors that cause low birthweight?)

The regression above likely does not solely accurately capture the impact of smoking on birthweight, as some of the key assumptions of the linear regression model include normality and independence, both of which are likely correlated to other factors that cause low birthweight due to lifestyle choices or tightly coupled backgrounds.

C. Regress birthweight on smoker, alcohol, and nprevist. Explain why the exclusion of these variables could lead to a biased regression coefficient in (a) above. Is the estimated effect of smoking on birthweight substantially different from the regression in (a) above?

```
In [49]: x_multi = sm.add_constant(smoking_df[["smoker", "alcohol", "nprevist"]])
    multi_model = sm.OLS(smoking_df["birthweight"], x_multi).fit()
    print(f"{multi_model.summary()=}")

pct_change = abs((-253.2284 - (-217.5801)) / -253.228)
    print(f"{pct_change=}")
```

### OLS Regression Results

========	========		=====	=====	=====	=========		=======
== Dep. Varia	hle:	h	irthu	eight	P_ca	uared:		0.0
73	D.C.	L	TI CIIW	zigiit	IN-34	uai eu.		0.0
Model:				0LS	Adj.	R-squared:		0.0
72 Method:		Loo	Cau	12505	E c+	atictic.		78.
47		Lea	ist squ	iares	r-5 t	atistic:		/0.
Date:		Sun, 1	.0 Aug	2025	Prob	(F-statistic)	:	7.31e-
49								
Time: 4.			05:5	66:49	Log-	Likelihood:		-2329
No. Observ	ations:			3000	AIC:			4.660e+
04								
Df Residua	ls:			2996	BIC:			4.662e+
04 Df Model:				3				
	Type:		nonro					
	:======:		=====	-====	=====	=========		=======
==		£ 6+	.d orr		+	P> t	[0 025	0.07
5]	coe	ı St	.a err		ι	P> L	[0.025	0.97
	2051 240		14 016	0.0	701	0.000	2004 552	2117 0
const 46	3051.2480	0 3	4.010	89	./01	0.000	2984.552	3117.9
	-217.580	1 2	6.680	-8	.155	0.000	-269.892	-165.2
68								
alcohol 85	-30.4913	3 7	6.234	-0	.400	0.689	-179.968	118.9
	34.0699	9	2.855	11	.933	0.000	28.472	39.6
68						01000		3313
=======	=======	=====	=====		=====	=========	=======	=======
== Omnibus:			37/	1.095	Durh	in-Watson:		1.9
74			375	+.093	Duib	in-watson.		1.9
Prob(Omnib	us):		(	0.000	Jarq	ue-Bera (JB):		869.2
20				700		(30)		4 70 4
Skew: 89			-(	729	Prob	(JB):		1.78e-1
Kurtosis:				5 <b>.</b> 197	Cond	. No.		8
5.2								
=======	=======	=====	=====	=====	=====	=========	======	=======
==								

#### Notes:

### pct\_change=0.1407755066580315

The exclusion of these variables could lead to a biased regression other variables like alcohol and nprevist could also have an effect on birthweight, so smoking could appear

to be more harmful than it actually is on its own. The coefficient seems to have changed by  $\sim$ 14%, which is a relatively large amount to have shifted just due to two variables, so the impact is indeed somewhat substantially different.

D. Jane smoked during her pregnancy, did not drink alcohol, and had 8 prenatal care visits. Use the regression in (c) to predict the birthweight of Jane's infant.

```
In [52]: smoker_jane = 1
    alcohol_jane = 0
    nprevist_jane = 8

intercept, smoker_coeff, alcohol_coeff, nprevisit_coeff = multi_model.params
birthweight_jane = intercept + smoker_coeff * smoker_jane + alcohol_coeff *
    print(f"{birthweight_jane=}")
```

birthweight\_jane=3106.227800368578

3. An alternative way to control for prenatal visits is to use binary variables tripre0 through tripre3. Regress birthweight on smoker, alcohol, tripre0, tripre2, and tripre3.

```
In [57]: x_tri_multi = sm.add_constant(smoking_df[["smoker", "alcohol", "tripre0", "t
x_tri_multi_model = sm.OLS(smoking_df["birthweight"], x_tri_multi).fit()
print(f"{x_tri_multi_model.summary()=}")
```

x\_tri\_multi\_model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

========			=====	======			=======
== Dep. Varia 46	able:	birthwei	.ght	R-squa	red:		0.0
Model: 45			0LS	Adj. F	R-squared:		0.0
Method: 18		Least Squa	res	F-stat	istic:		29.
Date: 29	Sı	un, 10 Aug 2	025	Prob (	F-statistic	c):	5.20e-
Time:		06:04	:13	Log-Li	kelihood:		-2333
No. Observ	/ations:	3	000	AIC:			4.668e+
Df Residua 04	als:	2	994	BIC:			4.672e+
Df Model: Covariance	e Type:						
5]		std err					0.97
 const	3454.5493	12.650	273	.077	0.000	3429.745	3479.3
54 smoker	-228.8476	27.165	-8	.424	0.000	-282.111	-175.5
	-15.1000	77.541	-0	.195	0.846	-167.138	136.9
•	-697.9687	106.876	-6	.531	0.000	-907.526	-488.4
	-100.8373	29.619	-3	.404	0.001	-158.913	-42.7
62 tripre3 31	-136.9553	59.581	-2	.299	0.022	-253.780	-20.1
=======================================	========			======		========	
Omnibus: 76	,	443.			n-Watson:		1.9
Prob(Omnik	ous):		000		e-Bera (JB):	:	1157.6
Skew: 52			811	Prob(J			4.20e-2
Kurtosis: 0.5 =======		5 <b>.</b> 	575 =====	Cond.	No.		1

# ==

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

A. Why is tripre1 excluded from the model? What happens if you include it in the regression?

```
In [59]: x_tri_full_multi = sm.add_constant(smoking_df[["smoker", "alcohol", "tripre@x_tri_full_multi = sm.OLS(smoking_df["birthweight"], x_tri_multi).fit()
print(f"{x_tri_full_multi.summary()=}")
```

x\_tri\_full\_multi.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

========		-====	======	======	=====	========	=======	=======
•	able:		birthwe	eight	R-sq	uared:		0.0
46 Model:				0LS	Adj.	R-squared:		0.0
45 Method:		Le	ast Squ	uares	F-st	atistic:		29.
18 Date:		Sun,	10 Aug	2025	Prob	(F-statistic	:):	5.20e-
29 Time:			06:0	04:44	Log-	Likelihood:		-2333
No. Observ	/ations:			3000	AIC:			4.668e+
04 Df Residua	als:			2994	BIC:			4.672e+
04 Df Model:	. T			5				
	e Type: 		nonro				.=======	=======
==		_						
5]	coe	r s	td err		t	P> t	[0.025	0.97
const 44	2576.4872	2	25.835	99.	.727	0.000	2525.830	2627.1
	-228.8476	5	27.165	-8.	424	0.000	-282.111	-175.5
	-15.1000	ð	77.541	-0.	195	0.846	-167.138	136.9
	180.0934	1	85.926	2.	096	0.036	11.614	348.5
	878.0622	l	26.598	33.	012	0.000	825.909	930.2
tripre2 60	777.2249	9	32.505	23.	911	0.000	713.490	840.9
tripre3 87	741.1069	9	51.552	14.	376	0.000	640.027	842.1
==	=======	=====				 · ,, ,	=======	
Omnibus: 76			443	3.968	Durb	in-Watson:		1.9
Prob(Omnib	ous):		(	0.000	Jarq	ue-Bera (JB):		1157.6
Skew: 52			-6	0.811	Prob	(JB):		4.20e-2
Kurtosis: 16			:=====	5.575 ======	Cond	. No.	:=======	1.42e+

# ==

# Notes:

[1] Standard Errors assume that the covariance matrix of the errors is corre

ctly specified.

[2] The smallest eigenvalue is 2.63e-29. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.

It seems that including tripre1 drastically affects the coefficients of the tripre variables by turning all of them to positive whereas they were previously all negative. It is likely that the results become meaningless because these tripre variables add up perfectly to 1, and with including these all we cannot do an accurate regression since we cannot find unique solutions for the coefficients of the regression.

# B. The estimated coefficient on tripre0 is large and negative. What does this coefficient measure? Interpret its value.

This coefficient measures the difference in average birthweight of babies to mothers who did not have any prenatal care visits compared to mothers who had their prenatal care visit in the first trimester. The ~-698 value indicates that the average birthweight of babies with no prenatal care visits is approximately 698 grams less than those born to mothers with their first prenatal care visit in the first trimester.

### C. Interpret the value of the estimated coefficients on tripre2 and tripre3.

- The coefficient of tripre2, ~-101, means that babies born to mothers with their first prenatal care visit in the 2nd trimester had birthweights of 101 grams less than babies born to mothers with their first prenatal care visit in the first trimester, with everything else in the model held constant.
- The coefficient of tripre3, ~-137, means that babies born to mothers with their first prenatal care visit in the 2nd trimester had birthweights of 137 grams less than babies born to mothers with their first prenatal care visit in the first trimester, with everything else in the model held constant.

# D. Does the regression in (3) explain a larger fraction of the variance in birthweight than the regression in (2c)? (Hint: consider R2.)

The regression in (3) has R<sup>2</sup> value of 0.046 while the regression in part (2C) has R<sup>2</sup> value of 0.073. We can therefore determine that with the smaller R<sup>2</sup> value, it actually explains a **lower fraction** of the variance.

# 4. Consider adding an additional regressor: Regress birthweight on smoker, alcohol, nprevist, and unmarried.

```
In []: x_unmarried_multi = sm.add_constant(smoking_df[["smoker", "alcohol", "nprevi
x_unmarried_multi_model = sm.OLS(smoking_df["birthweight"], x_unmarried_mult
print(f"{x_unmarried_multi_model.summary()=}")
```

x\_unmarried\_multi\_model.summary()=<class 'statsmodels.iolib.summary.Summar
y'>

# OLS Regression Results

========		======	=====	=====	=====	========	=======	=======
== Dep. Variak 89	ole:	bi	rthwei	ght	R-squ	ared:		0.0
Model:				0LS	Adj. I	R-squared:		0.0
87 Method:		Leas	t Squa	res	F–sta	tistic:		72.
79 Date:		Sun, 10	Aug 2	025	Prob	(F—statistic	):	6.12e-
59 Time:			19:07	:08	Log-L	ikelihood:		-2326
8. No. Observa	ations:		3	000	AIC:			4.655e+
04 Df Residua	ls:		2	995	BIC:			4.658e+
04 Df Model:				4				
Covariance			nonrob 					
==						P> t		0 07
5]	COE	s tu	CII		·	r> t	[0:023	0.97
 const	3134.4000	a 25	656	27	907	0.000	3064.487	3204.3
13					472		-228.511	
smoker 43								
alcohol 64					279		-169.331	
nprevist 86					213		23.920	
unmarried 39	-187 <b>.</b> 1332	2 26	.007	<b>-7.</b>	195	0.000	-238 <b>.</b> 128	-136.1
=======================================	=======	======	=====	=====	:====:	=======	=======	=======
Omnibus:			369.	861	Durbi	n-Watson:		1.9
Prob(Omnibu	ıs):		0.	000	Jarqu	e-Bera (JB):		880.8
Skew:			-0.	714	Prob(.	JB):		5.27e-1
92 Kurtosis: 5.2				238 =====	Cond.	No.	=======	8

#### Notes:

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[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.  $\hfill {\tt min}$ 

A. Compare the coefficient on smoker in this regression to the coefficients on smoker in regressions (2a) and (2c). What is the estimated effect of smoking on birthweight in each regression?

The coefficient on smoker in this regression is ~-175, which indicates a 175 gram decrease in birthweight for mothers who smoke relative to those who do not in this model. This is lower than the coefficients from (2A) and (2C), which are ~-253 and ~-218 respectively, which indicates a much larger decrease in birthweight (253 and 218 grams respectively) attributed to the fact of whether the mother of the child is a smoker or not.

## B. Interpret differences in estimated effects.

The smaller absolute value of the smoking coefficient in this regression indicates that the regression with fewer variables is less accurate, and controlling for more variables leads to a model that is closer to modeling the actual effect of smoking on birthweight, with the previously simpler regressions in question 2 overestimating its effect and possibly having some correlation that is obscured with other variables. With more variables, we can further isolate and decorrelate these variables, so the true value is likely closer to the regression we just did, with a smaller effect on birthweight truly attributed solely to smoking.

C. Interpret the estimated effect of marital status on birthweight. Is the coefficient on unmarried statistically significant? Is the magnitude of the coefficient large?

The unmarried coefficient is ~-187 with an extremely small p-value (that displays as 0.000 when rounded to the thousandths), which in turn is below our threshold of considering it to be statistically significant, so the effect is indeed statistically significant. The effect seems to be that a mother being unmarried is associated with an approximately 187 gram decrease in child birthweight, while holding other variables in the model constant. The magnitude seems to be fairly large -- it is larger than the magnitude of all other variables within the model, including smoking, which we previously found to be a large factor in the other models.

D. A family advocacy group notes that the large coefficient suggests that public policies that encourage marriage will lead, on average, to healthier babies. Do you agree? (Hint: consider some of the various factors that unmarried may be controlling for and how this affects the interpretation of this coefficient).

It makes more sense to disagree with the family advocacy group. In this model, the coefficient indicates that being unmarried while controlling for the other variables not present in the model such as age and education, while the coefficient itself reflects the impact of being unmarried while combined with other factors like the behaviors of smoking or drinking during the pregnancy, so it cannot be taken just at face value as an impactful factor in isolation.

# 5. Consider the other coefficients in this data set. Which do you think should be included in the regression?

All the other factors that are left in the data set should honestly be included -- they are all relevant to the attributes of the mother and it would be more impactful to include them so that we can also better isolate and understand the impacts of each factor individually. So we should add age and education, which we haven't yet used, to the regression.

A. Try adding in some of these additional variables. Share your findings and conclusions.

x\_age\_multi\_model.summary()=<class 'statsmodels.iolib.summary.Summary'>

### OLS Regression Results

========			=====	=====	========		=======
	ble:	birthwe	ight	R-squ	ared:		0.0
89 Model: 87			0LS	Adj.	R-squared:		0.0
Method: 51		Least Squ	ares	F-sta	tistic:		58.
Date: 58	S	Sun, 10 Aug	2025	Prob	(F-statistio	c):	2.88e-
Time: 8.		20:0	8:12	Log-L	ikelihood:		-2326
No. Observa	ations:		3000	AIC:			4.655e+
Df Residua 04	ls:		2994	BIC:			4.658e+
Df Model: Covariance	Type:	nonro 					
5]		std err					0.97
	3201.4887	68.100	47	.011	0.000	3067.960	3335.0
	-177.1410	27.140	-6	.527	0.000	-230.356	-123.9
26 alcohol 54	-14.6826	75.806	-0	.194	0.846	-163.319	133.9
	29.7909	2.903	10	.263	0.000	24.099	35.4
	-199.4890	28.116	-7	.095	0.000	-254.617	-144.3
age 11	-2.4597	2.127	-1	.156	0.248	-6.631	1.7
=========	========		=====			=======	=======
Omnibus: 69			<b>.</b> 192	Durbi	n-Watson:		1.9
Prob(Omnib	us):		.000		e-Bera (JB):	:	872.2
Skew: 90		-0	.708	Prob(	JB):		3.89e-1
Kurtosis: 8.	=========	5=======	.231 =====	Cond.	No.		21

# Notes:

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[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

x\_educ\_multi\_model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

========	=======	=======	======	====	======	=======		=======
== Dep. Varia 89	ble:	bi	rthweight	t	R-squar	ed:		0.0
Model: 87			0L9	5	Adj. R-	squared:		0.0
Method: 24		Leas	t Squares	5	F-stati	stic:		58.
Date: 58		Sun, 10	Aug 2025	5	Prob (F	-statistic	:):	5.23e-
Time: 8.			20:08:12	2	Log-Lik	elihood:		-2326
No. Observ	ations:		3000	9	AIC:			4.655e+
04 Df Residua	ls:		2994	4	BIC:			4.658e+
04 Df Model: Covariance			nonrobust	5 t				
=======================================	CO6			====	t	P> t	[0.025	0.97
	2157 01	74	200	42	E00		2012 240	2202 5
const 79					.508	0.000	3012.248	
smoker 31					. 442	0.000	-230.874	
alcohol 41			703		. 261	0.794	-168.229	
nprevist 85					. 165	0.000	24.009	35.4
unmarried 76	-189.80	56 27	.046	<b>-7</b>	.018	0.000	-242.837	-136.7
educ 17	-1.87	54 5	. 198	-0	.361	0.718	-12.068	8.3
=======================================	=======	=======	======	====	======	=======	:=======	=======
Omnibus: 68			369.874	4	Durbin-	Watson:		1.9
Prob(Omnib	us):		0.000	0	Jarque-	Bera (JB):		880.5
Skew: 92			-0.714	4	Prob(JB	):		6.29e-1
92 Kurtosis: 8.	=======	=======	5.237	7	Cond. N	0.	:=======	12

# ==

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

x\_both\_multi\_model.summary()=<class 'statsmodels.iolib.summary.Summary'>

# OLS Regression Results

========	========	========			========		=======
==		1					0.0
Dep. Varia 89	ble:	birth	veignt	R-squ	ared:		0.0
Model: 87			0LS	Adj.	R-squared:		0.0
Method:		Least So	quares	F-sta	tistic:		48.
74 Date:		Sun 10 Aug	1 2025	Proh	(F-statisti	c):	2.32e-
57						C / I	
Time: 8.		20:	08:12	Log-L	ikelihood:		-2326
No. Observ	ations:		3000	AIC:			4.655e+
04 Df Residua	1s:		2993	BIC:			4.659e+
04				DICI			110550
Df Model: Covariance	Tyne:	noni	6 obust				
		=======		======	=======	=======	
==	coe	f std eri	_	t	P> t	[0.025	0.97
5]		. 554 5		-	. 1-1	[01020	
const 30	3199.426	4 83.337	7 3	38.392	0.000	3036.023	3362.8
smoker	-176.958	9 27.474	1 -	-6.441	0.000	-230.828	-123.0
	-14.758	3 75.839	) –	-0.195	0.846	-163.460	133.9
43 nprevist 13	29.775	1 2.926	5 1	10.175	0.000	24.037	35.5
	-199.319	5 28.396	5 -	-7.019	0.000	-254.997	-143.6
	-2.493	5 2.269	) –	-1.099	0.272	-6.942	1.9
educ 05	0.238	0 5.542	2	0.043	0.966	-10.629	11.1
=======================================	=======					=======	
Omnibus: 69		36	66.140	Durbi	n-Watson:		1.9
Prob(Omnib	us):		0.000	Jarqu	e-Bera (JB)	:	872.1
Skew:		-	-0.707	Prob(	JB):		4.04e-1
90 Kurtosis: 6.			5.231	Cond.	No.		26
=======	=======	========	=====		=======	=======	======

# ==

# Notes:

[1] Standard Errors assume that the covariance matrix of the errors is corre

We added in age, education, and then both to the model.

Overall, we find that adding age has a very little effect on the birthweight with single-digit coefficients in all contexts, albeit with a relatively high p-value. Adding in education also did not have a statistically significant effect, even smaller than that of age when it comes to coefficient. Both of these had little impact on the model's R^2 value as well, so it's safe to say they did not contribute much. The effect of the other factors stayed relatively constant, though.

B. The data set includes babies born in Pennsylvania in 1989. Discuss the external validity of your analysis for: (i) California in 1989, (ii) Illinois in 2015, (iii) South Korea in 2014.

- (i) For California in 1989, there is some generalizability and validity. The year is the same, and the country is the same, but the socioeconomic structure of the state, sociocultural norms around education, marriage, smoking, and alcohol, and even other confounding factors such as climate. Still, there is some validity to the data.
- (ii) For Illinois in 2015, there is limited generalizability and validity. The country is still the same, and the state is geographically similar, but there are a lot of differences in the 26 years that have passed when it comes to medical advice and cultural norms around relevant factors.
- (iii) For South Korea in 2014, there is little if any generalizability and validity. The country and demographics and culture are completely different, and completely different medical systems and norms.

C. Overall, explain your conclusions on how maternal smoking impacts birthweight (hint: the regressions you're running should be helping you see that isolating the causal effect of smoking on birthweight is difficult because there are a lot of other confounding variables)

Across the different regressions that we ran, we found that with more controls, the regression with more variables had smoking at significantly different coefficients and impact on birthweight relative to the other regressions that were ran with fewer variables in parts (2A) and (2C) (and even those two regression models had a pretty nontrivial difference in coefficient). This shows what we hoped to determine, which is that isolating the effect that smoking has on birthweight is difficult to isolate and can vary wildly due to other confounding variables that may be interlinked for other reasons. As such, the earlier models with higher absolute value coefficients for maternal smoking's effect seemed to miss confounding variables. We do see that maternal smoking still has some meaningful effect on child birthweight both statistically and magnitude-wise. However, it is still important to not that this is in the context of several other variables that were considered in the model as well.