

1.

(a)

Given mean  $\mu = 2$  with the exponential distribution, we can directly plug in:

$$P(X > (s - t)) = e^{-\lambda \cdot (s-t)}$$

$$P(X > 2) = e^{-\frac{1}{\mu} \cdot 2}$$

$$P(X > 2) = e^{-\frac{1}{2} \cdot 2}$$

$$P(X > 2) = e^{-1} = \boxed{0.36787944117}$$

(b)

We set up the equivalent conditional property expression:

$$P(X > 5 | X > 3) = \frac{P(X > 5 \cap X > 3)}{P(X > 3)}$$

Logically, we know that if  $P(X > 5)$ ,  $P(X > 3)$  for sure, so we can remove the latter part, i.e.  $P(X > 5 \cap X > 3) = P(X > 5)$ . Plugging back in:

$$P(X > 5 | X > 3) = \frac{P(X > 5)}{P(X > 3)}$$

And now using the exponential pdf:

$$P(X > 5 | X > 3) = \frac{e^{-5\lambda}}{e^{-3\lambda}}$$

$$P(X > 5 | X > 3) = \frac{e^{-5 \cdot \frac{1}{2}}}{e^{-3 \cdot \frac{1}{2}}}$$

$$P(X > 5 | X > 3) = \frac{e^{-\frac{5}{2}}}{e^{-\frac{3}{2}}}$$

$$P(X > 5 | X > 3) = e^{-1} = \boxed{0.36787944117}$$

2.

Given  $X \sim \text{Exp}(\lambda)$ , we aim to prove that  $P(X \geq a + t | X \geq a) = P(X \geq t)$ :

$$\begin{aligned}
 P(X \geq a + t | X \geq a) &= \frac{P(X \geq a + t \cap X \geq a)}{P(X \geq a)} && \text{Conditional probability} \\
 &= \frac{P(X \geq a + t)}{P(X \geq a)} && \text{As above, } X \geq a + t \text{ implies } X \geq a \\
 &= \frac{e^{-\lambda \cdot (a+t)}}{e^{-\lambda a}} && \text{Definition of exponential distribution} \\
 &= \frac{e^{-\lambda a} e^{-\lambda t}}{e^{-\lambda a}} \\
 &= e^{-\lambda t} \\
 &= P(X \geq t)
 \end{aligned}$$

as we sought to show.

3.

4.

5.

(a)

We are given that  $N(5) = 2$ , and Poisson process  $\lambda = 10$ . As the arrival times are uniformly distributed,  $N(t) = n$  in time  $[0, s]$  follows the Binomial with  $p = \frac{s}{t}$ . We therefore can directly use the formula to calculate  $P(k = 2)$ :

$$\begin{aligned}
 P(N(s) = l | N(t) = k) &= \binom{k}{l} \left(\frac{s}{t}\right)^l \left(1 - \frac{s}{t}\right)^{k-l} \\
 P(N(s) = 2 | N(5) = 2) &= \binom{2}{2} \left(\frac{2}{5}\right)^2 \left(1 - \frac{2}{5}\right)^{2-2} \\
 P(N(s) = 2 | N(5) = 2) &= \binom{2}{2} \left(\frac{2}{5}\right)^2 \left(\frac{3}{5}\right)^0 \\
 P(N(s) = 2 | N(5) = 2) &= 1 \cdot \frac{4}{25} \cdot 1 = \boxed{\frac{4}{25}}
 \end{aligned}$$

(b)

The probability that at least one customer arrived in the first two minutes is easier found by taking its complement, i.e. 1 - the probability that no customers arrived in the first two minutes:

$$\begin{aligned}P(N(s) = l | N(t) = k) &= \binom{k}{l} \left(\frac{s}{t}\right)^l \left(1 - \frac{s}{t}\right)^{k-l} \\P(N(2) = 0 | N(5) = 2) &= \binom{2}{0} \left(\frac{2}{5}\right)^0 \left(1 - \frac{2}{5}\right)^{2-0} \\P(N(2) = 0 | N(5) = 2) &= 1 \cdot 1 \cdot \left(\frac{3}{5}\right)^2 \\P(N(2) = 0 | N(5) = 2) &= \frac{9}{25}\end{aligned}$$

Taking the complement, we see that

$$P(N(2) \geq 1 | N(5) = 2) = 1 - \frac{9}{25} = \boxed{\frac{16}{25}}$$

6.

(a)

We are given Poisson  $N(t)$  with  $\lambda = 2$ . To find  $P(N(2) = 1, N(3) = 4, N(5) = 5)$ , we can break this up into independent periods with how many events occur within each period, and then multiply. Let's begin:

1 event from  $[0, 1]$ :

$$\begin{aligned}P[N(s) = n] &= \frac{e^{-\lambda s} \cdot (\lambda s)^n}{n!} \\P[N(2) = 1] &= \frac{e^{-2 \cdot 2} \cdot (2 \cdot 2)^1}{1!} \\P[N(2) = 1] &= \frac{e^{-4} \cdot 4^1}{1} = 4e^{-4}\end{aligned}$$

4-1 = 3 events from  $[2, 3]$ :

$$P[N(3-2) = 3] = \frac{e^{-2 \cdot (3-2)} \cdot (2 \cdot (3-2))^3}{3!}$$

$$P[N(1) = 3] = \frac{e^{-2} \cdot 2^3}{6} = \frac{4e^{-2}}{3}$$

5-4 = 1 event from [3, 5]:

$$P[N(5-3) = 1] = \frac{e^{-2 \cdot (5-3)} \cdot (2 \cdot (5-3))^1}{1!}$$

$$P[N(2) = 1] = \frac{e^{-2 \cdot 2} \cdot (2 \cdot 2)^1}{1} = 4e^{-4}$$

So finally,

$$\begin{aligned} P(N(2) = 1, N(3) = 4, N(5) = 5) &= P(N(2) = 1) \cdot P(N(3) - N(2) = 3) \cdot P(N(5) - N(3) = 1) \\ &= 4e^{-4} \cdot \frac{4e^{-2}}{3} \cdot 4e^{-4} = \frac{4 \cdot 4 \cdot 4}{3} e^{-4-2-4} \\ &= \frac{64}{3} e^{-10} = \boxed{0.00096853183} \end{aligned}$$

(b)

To find  $P(N(4) = 3 | N(2) = 1, N(3) = 2)$ , we can first remove the  $N(2) = 1$  part because we know that  $N(3) = 2$  is the only relevant part given the independent time increments. Therefore, we just need to find  $P(N(4) = 3 | N(3) = 2)$ :

$$P(N(4) = 3 | N(3) = 2) = P[N(4-3) = 3-2]$$

$$P[N(s) = n] = \frac{e^{-\lambda s} \cdot (\lambda s)^n}{n!}$$

$$P[N(1) = 1] = \frac{e^{-2 \cdot 1} \cdot (2 \cdot 1)^1}{1!}$$

$$P[N(1) = 1] = \frac{e^{-2} \cdot 2^1}{1}$$

$$P[N(1) = 1] = 2e^{-2} = \boxed{0.27067056647}$$

(c)

$$E(N(4)|N(2) = 2)$$

Given

$$E(N(4) - N(2) + N(2)|N(2) = 2) \quad N(4) \text{ in independent increments}$$

$$E[N(4) - N(2)|N(2) = 2] + E[N(2)|N(2) = 2] \quad \text{Linearity of expectation}$$

$$E[N(4) - N(2)|N(2) = 2] + 2 \quad \text{Definition of expectation}$$

$$E[N(4) - N(2)] + 2 \quad \text{Independent increments}$$

$$E[N(4 - 2)] + 2 \quad \text{Definition of Poisson process}$$

$$E[N(2)] + 2$$

$$\lambda \cdot 2 + 2 \quad \text{Definition of Poisson process}$$

$$2 \cdot 2 + 2$$

$$4 + 2 = \boxed{6}$$

7.

8.

9.

(a)

(b)