## EE3731C Pattern Recognition 1

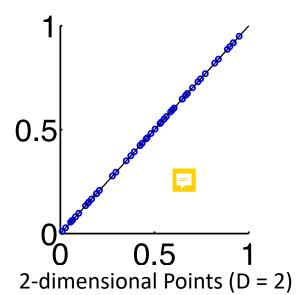
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#### **Curse of Dimensionality**

- Data dimensionality = number of "free" parameters in data / model
  - -1M-pixel image  $\implies 1M$  dimensions
  - Text document has N words  $\implies N$  dimensions
- Dimensionality depends on problem definition
  - Word analysis of text document  $\implies$  dimensionality = # words
  - Character analysis of text document  $\implies$  dimensionality = # characters
- Curse of dimensionality: pattern recognition harder for big dimensions
  - Joint distribution of binary variables  $p(x_1, \dots, x_N)$  has  $2^N 1$  parameters
  - Need a lot of data to estimate so many parameters

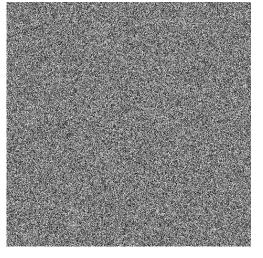
#### **Intrinsic Dimensions**

- ullet Given data has D dimensions, hope "intrinsic" data dimensions d less than D
- "Random" images do not look like "real" images  $\implies$  image pixels not independent  $\implies$  true dimensions less than number of pixels



with 1"intrinsic" dimension (d = 1)

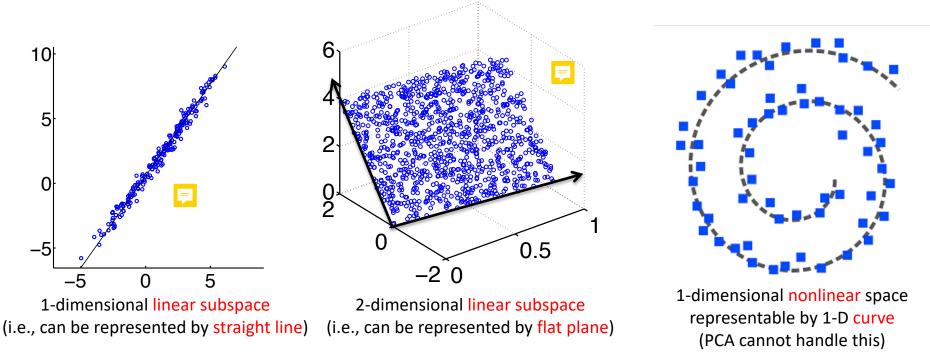
"Real" Image



"Random" Image (Each pixel independent with uniform U[0, 1] distribution)

### Dimensionality Reduction

- Dimensionality reduction: map high-dimensional data into lower dimensions
  - Also known as feature extraction or feature reduction
  - Different assumptions lead to different mapping
- Principal component analysis (PCA) assumes data lies on linear subspace



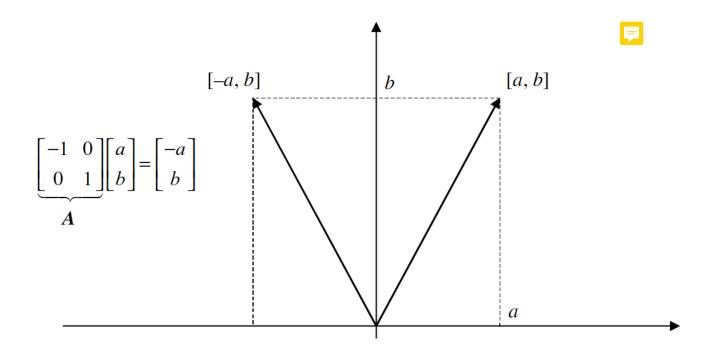
## Linear System Revision

#### Linear System (or Transformation)

- Suppose system T maps input  $x \in \mathbb{R}^n$  to output  $y \in \mathbb{R}^m$ 
  - If system is linear, then  $y_1 = T(x_1)$  and  $y_2 = T(x_2)$  implies  $T(ax_1 + bx_2) = aT(x_1) + bT(x_2) = ay_1 + by_2$
  - All linear systems can be written in matrix format, i.e., y = Ax for some  $m \times n$  matrix
- Example:  $Ax = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \\ a_{31} & a_{32} \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} a_{11}x_1 + a_{12}x_2 \\ a_{21}x_1 + a_{22}x_2 \\ a_{31}x_1 + a_{32}x_2 \end{bmatrix}$
- Example:  $Ax = [\vec{a}_1 \ \vec{a}_2] \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = x_1\vec{a}_1 + x_2\vec{a}_2$ , where  $\vec{a}_i$  is *i*-th column of A
  - Output is linear combination of columns of A

#### Linear Transformation: Example 1

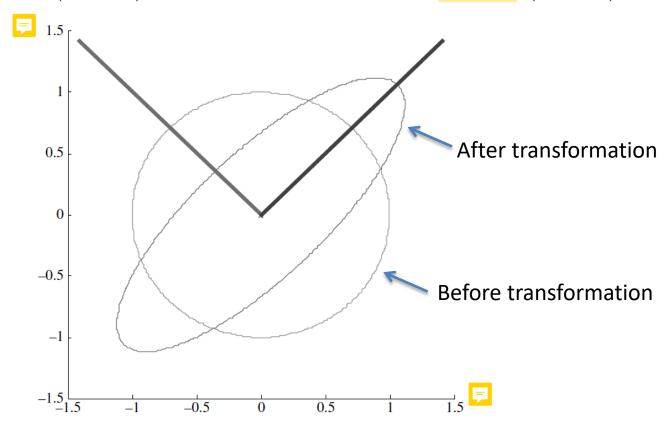
• 
$$A = \begin{bmatrix} -1 & 0 \\ 0 & 1 \end{bmatrix}$$
 flips vector about vertical axis



#### Linear Transformation: Example 2

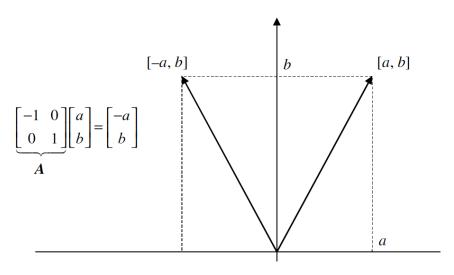
$$\bullet \ A = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix} \implies \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} x_1 + 0.5x_2 \\ 0.5x_1 + x_2 \end{bmatrix}$$

• Circle of points (below) mapped by A onto an ellipse (below)

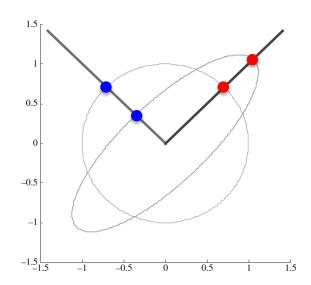


#### **Eigenvectors & Eigenvalues**

- Consider linear system y = Ax, where  $A = M \times M$  matrix
- Suppose x is eigenvector of A
  - By definition of eigenvectors,  $Ax = \lambda x$  (for some number  $\lambda$ )
  - Therefore direction of x preserved under transformation A



$$x = \begin{bmatrix} 0 \\ 1 \end{bmatrix}$$
 is eigenvector with  $\lambda = 1$   $\begin{bmatrix} \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{bmatrix}$  &  $\begin{bmatrix} -\frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{bmatrix}$  eigenvectors with eigenvalues 1.5 and 0.5



$$\begin{bmatrix} \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{bmatrix} \& \begin{bmatrix} -\frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{bmatrix}$$

#### Computing Eigenvectors/Eigenvalues

•  $u \neq 0$  is an eigenvector of  $M \times M$  matrix A if  $Au = \lambda u$  for some  $\lambda \models$ 

$$Au = \lambda u$$
 
$$(A - \lambda I)u = 0$$
 
$$\det(A - \lambda I) = 0$$
 u is non-zero => A –  $\lambda$ I is not full rank

• For example,  $A = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$ , then

$$\det \begin{pmatrix} a - \lambda & b \\ c & d - \lambda \end{pmatrix} = 0$$

$$(a - \lambda)(d - \lambda) - bc = 0$$

$$\lambda^2 - (a + d)\lambda + (ad - bc) = 0$$

- Quadratic equations  $\implies$  up to two  $\lambda$ s (in general up to M eigenvalues for  $M \times M$  matrix A)
- Solve for  $\lambda s$ , and then eigenvectors
- In real applications, use software to solve (e.g., 'eig' function in matlab)

#### Matrix Eigen-Decomposition

• For  $M \times M$  matrix A with M eigenvalues  $\lambda_1, \dots, \lambda_M$  and corresponding eigenvectors  $u_1, \dots, u_M$ , where  $u_m^T u_m = 1$  (unit length)

$$A[u_1 \ u_2 \ \cdots \ u_M] = [\lambda_1 u_1 \ \lambda_2 u_2 \ \cdots \ \lambda_M u_M]$$

$$= [u_1 \ u_2 \ \cdots \ u_M] \begin{bmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \lambda_M \end{bmatrix}$$

$$AU = U\Lambda,$$

where  $\Lambda$  is a diagonal matrix which is all 0 except the diagonals which correspond to eigenvalues and U is a matrix whose columns are eigenvectors

• A symmetric  $\Longrightarrow M$  (not necessarily unique) eigenvalues, and M orthogonal eigenvectors, i.e.,  $u_i^T u_j = \delta(i-j) \Longrightarrow U^{-1} = U^T$ 

$$AU = U\Lambda \implies A = U\Lambda U^T = \lambda_1 u_1 u_1^T + \lambda_2 u_2 u_2^T \cdots + \lambda_M u_M u_M^T = \sum_m \lambda_m u_m u_m^T = \sum_m \lambda_m u_m u_m^T$$

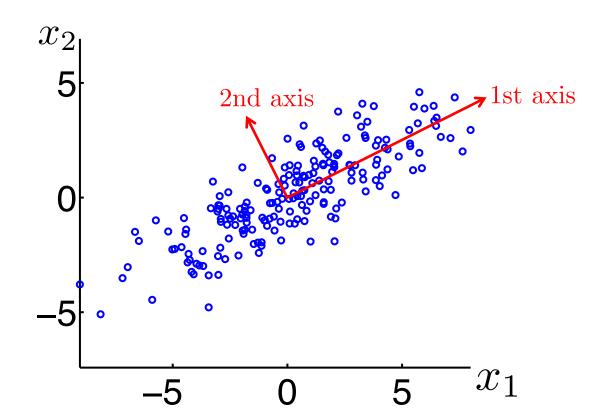
• A positive semidefinite, i.e.,  $x^T A x \ge 0$  for all  $x \implies \lambda_m \ge 0$  for all m.

## Principal Component Analysis (PCA)

#### Principal Component Analysis (PCA)

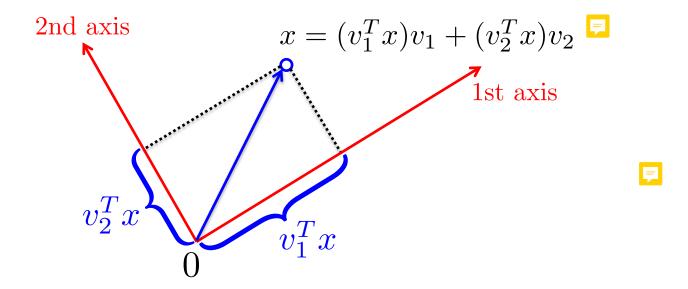
- PCA = new coordinate system with orthogonal axes
  - Assume mean of input data =  $\vec{0}$
  - 1st axis direction of largest variability in data
  - 2nd axis direction of largest variability orthogonal to 1st axis
  - 3rd axis direction of largest variability orthogonal to 1st and 2nd axes

- . . .



#### **Projecting Onto Principal Axes**

- Coordinates with respective to new axes given by projection onto axes
  - For *D*-dimensional point x, new coordinates =  $(v_1^T x, v_2^T x, \dots, v_K^T x)$ , where  $v_i$  is unit vector in same direction as i-th axis
  - -x now K-dimensional vector: we have achieved dimensionality reduction for K < D
  - $-v_i^T x$  called *i*-th principal component (PC)



### 1st Principal Axis is Direction of Largest Variability

- Consider  $x^{(1)}, \dots, x^{(N)}$  (N data points, each dimension D)
- Assume data centered at origin, i.e.,  $\frac{1}{N} \sum_{n} x^{(n)} = \vec{0}$
- Let  $v_1$  = unit vector in same direction as 1st principal axis, so resulting projected data are  $[v_1^T x^{(1)}, \dots, v_1^T x^{(N)}]$
- Note that projected data has  $\frac{1}{N} \sum_{n} v_1^T x^{(n)} = 0$  mean, so (empirical) variance of projected data are

where  $\Sigma$  is covariance of data samples

$$-\Sigma = \frac{1}{N}XX^T$$
, where X is the  $D \times N$  matrix  $[x^{(1)}, \dots, x^{(N)}]$ 

• Therefore 1st principal axis (direction of data with largest variance):

$$v_1 = \operatorname{argmax} v^T \Sigma v \quad \text{where} \quad v^T v = 1$$

#### Finding Principal Axes

• 1st principal axis (direction of data with largest variance):

$$v_1 = \operatorname*{argmax}_{v} v^T \Sigma v \quad \text{where} \quad v^T v = 1$$

•  $\Sigma$  is symmetric & positive semidefinite, so has D eigenvalues  $\lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_D \geq 0$  with eigenvectors  $u_1, \cdots, u_D$ , where  $u_i^T u_j = \delta(i-j)$ :

$$\Sigma = \lambda_1 u_1 u_1^T + \lambda_2 u_2 u_2^T + \dots + \lambda_D u_D u_D^T$$

• Therefore

$$v^{T} \Sigma v = v^{T} \left( \lambda_{1} u_{1} u_{1}^{T} + \lambda_{2} u_{2} u_{2}^{T} + \dots + \lambda_{D} u_{D} u_{D}^{T} \right) v$$
$$= \lambda_{1} (v^{T} u_{1})^{2} + \lambda_{2} (v^{T} u_{2})^{2} + \dots + \lambda_{D} (v^{T} u_{D})^{2}$$

- Since  $\lambda_1$  largest and  $v_1^T v_1 = 1$ , therefore  $v_1 = u_1$ 
  - $-v_2 = \operatorname{argmax}_v v^T \Sigma v$  where  $v^T v = 1$  and  $v_1^T v_2 = 0 \implies v_2 = u_2$
  - In general,  $v_n = u_n$

#### Principal Component Analysis (PCA)

- Consider  $x^{(1)}, \dots, x^{(N)}$  (N data points, each dimension D)
- Compute mean  $\mu = \frac{1}{N} \sum_{n} x^{(n)}$
- Subtract mean from each sample point  $\bar{x}^{(n)} = x^{(n)} \mu$  (this ensures resulting data are centered at the origin) procedure called "demean"
- Form  $D \times N$  matrix  $X = [\bar{x}^{(1)}, \cdots, \bar{x}^{(N)}]$
- Compute sample covariance matrix  $\Sigma = \frac{1}{N}XX^T$
- Find top K eigenvalues and corresponding eigenvectors  $v_1, \dots, v_K$
- Given new (or old) datapoint x, we can represent x by new coordinates  $[v_1^T(x-\mu) \ v_2^T(x-\mu) \ \cdots \ v_K^T(x-\mu)]$
- Why top K eigenvalues and not bottom K?

# Principal Component Analysis (PCA) Representation

#### What does PCA representation buy us?

- Consider  $x^{(1)}, \dots, x^{(N)}$  (N data points, each dimension D) with 0 mean
- Perform PCA and represent old datapoint  $x^{(n)}$  by  $\begin{bmatrix} v_1^T x^{(n)} & v_2^T x^{(n)} & \cdots & v_K^T x^{(n)} \end{bmatrix}$
- Can "reconstruct"  $x^{(n)}$  by  $\hat{x}^{(n)} = \sum_{k=1}^{K} (v_k^T x^{(n)}) v_k$
- Consider reconstruction error

$$||x^{(n)} - \hat{x}^{(n)}||^2 \stackrel{\triangle}{=} \sum_{i=1}^{D} (x_i^{(n)} - \hat{x}_i^{(n)})^2$$

2nd axis  $x = (v_1^T x)v_1 + (v_2^T x)v_2$ 1st axis  $v_2^T x$ 

• Total reconstruction error (proof at end of notes):

$$\sum_{n=1}^{N} ||x^{(n)} - \hat{x}^{(n)}||^2 = \sum_{n=1}^{N} x^{(n)T} x^{(n)} - N \sum_{i=1}^{K} v_i^T \Sigma v_i$$

• First term in reconstruction error always positive  $\implies$  to minimize reconstruction error, want second term  $\sum_{i=1}^{K} v_i^T \sum v_i$  as big as possible, which is how PCA was defined  $\implies$  PCA lets us represent our original data with the least amount of error with K coefficients

#### Remarks

- Previous argument still works when  $x^{(1)}, \dots, x^{(N)}$  has non-zero mean  $\mu$
- Just need to subtract  $\mu$  from data before PCA
- When projecting datapoint x onto principal axes, new coordinates are  $[v_1^T(x-\mu), v_2^T(x-\mu), \cdots, v_K^T(x-\mu)]$
- Reconstructed  $\hat{x} = \mu + \sum_{k=1}^{K} (v_k^T(x \mu)) v_k$
- No magic way to select K. One way is the following criterion

$$\frac{\sum_{i=1}^{K} \lambda_i}{\sum_{i=1}^{D} \lambda_i} \ge \text{threshold (e.g., 0.95)}$$

• For  $M \times M$  images, can convert into a vector (i.e.,  $D = M^2$ ) by stacking columns of pixel values into one single column

## Eigenfaces Example

### Eigenfaces

• PCA on 200 face images



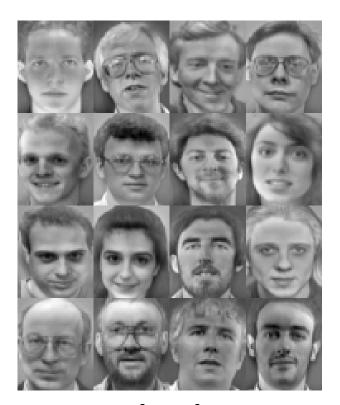
16 of 200 faces



Mean of 200 faces

#### Eigenfaces

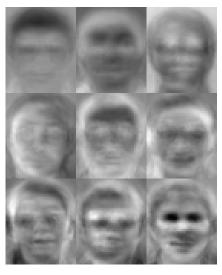
• PCA on 200 face images



16 of 200 faces after de-meaning



Mean of 200 faces (origin of coordinate System)



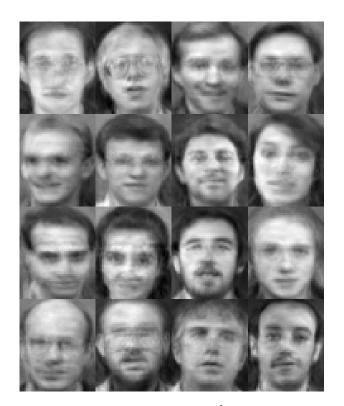
Top 9 Eigenvectors (Eigenfaces)

#### Eigenfaces

• PCA on 200 face images



16 of 200 faces



Reconstructed Faces From Top 50 eigenfaces

See eigenfaces.m on IVLE

#### Summary

- Curse of dimensionality: pattern recognition harder for high dimensions ⇒ useful to reduce dimensions
- PCA finds orthonormal basis (coordinate system) for data
- Principal axes corresponds to eigenvectors of data covariance matrix
- Principal axes sorted in order of importance (based on eigenvalues)
- Can discard axes with lower eigenvalues (hoping they are noise or unimportant for application)
- Reduce data dimensions by projecting onto remaining axes
- PCA optimal in the sense that it minimizes reconstruction error as measured by Euclidean distance
- PCA does not work for nonlinear space

#### **Optional Reading**

• Duda and Hart: Pattern Recognition, Chapter 3.8.1

# Additional Material: PCA minimizes reconstruction error proof

#### What does PCA representation buy us?

- Consider  $x^{(1)}, \dots, x^{(N)}$  (N data points, each dimension D) with 0 mean
- Perform PCA and represent old datapoint  $x^{(n)}$  by  $\begin{bmatrix} v_1^T x^{(n)} & v_2^T x^{(n)} & \cdots & v_K^T x^{(n)} \end{bmatrix}$

2nd axis

 $x = (v_1^T x)v_1 + (v_2^T x)v_2$ 

1st axis

- Can "reconstruct"  $x^{(n)}$  by  $\hat{x}^{(n)} = \sum_{k=1}^{K} (v_k^T x^{(n)}) v_k$
- Consider reconstruction error

$$||x^{(n)} - \hat{x}^{(n)}||^2 \stackrel{\triangle}{=} \sum_{i=1}^{D} (x_i^{(n)} - \hat{x}_i^{(n)})^2$$

$$= (x^{(n)} - \hat{x}^{(n)})^T (x^{(n)} - \hat{x}^{(n)})$$

$$= \left(x^{(n)} - \sum_{k=1}^{K} (v_k^T x^{(n)}) v_k\right)^T \left(x^{(n)} - \sum_{k=1}^{K} (v_k^T x^{(n)}) v_k\right)$$

$$= x^{(n)T} x^{(n)} - 2 \sum_{k=1}^{K} (v_k^T x^{(n)})^2 + \sum_{k=1}^{K} (v_k^T x^{(n)})^2$$

$$= x^{(n)T} x^{(n)} - \sum_{k=1}^{K} (v_k^T x^{(n)})^2$$

#### What does PCA representation buy us?

- From previous slide:  $||x^{(n)} \hat{x}^{(n)}||^2 = x^{(n)T}x^{(n)} \sum_{i=1}^K (v_i^T x^{(n)})^2$
- Total reconstruction error:

$$\sum_{n=1}^{N} ||x^{(n)} - \hat{x}^{(n)}||^2 = \sum_{n=1}^{N} x^{(n)T} x^{(n)} - \sum_{n=1}^{N} \sum_{i=1}^{K} (v_i^T x^{(n)})^2$$

$$= \sum_{n=1}^{N} x^{(n)T} x^{(n)} - \sum_{n=1}^{N} \sum_{i=1}^{K} v_i^T x^{(n)} x^{(n)T} v_i$$

$$= \sum_{n=1}^{N} x^{(n)T} x^{(n)} - \sum_{i=1}^{K} v_i^T \left(\sum_{n=1}^{N} x^{(n)} x^{(n)T}\right) v_i$$

$$= \sum_{n=1}^{N} x^{(n)T} x^{(n)} - N \sum_{i=1}^{K} v_i^T \sum v_i$$

• First term in reconstruction error always positive  $\implies$  to minimize reconstruction error, want second term  $\sum_{i=1}^{K} v_i^T \sum v_i$  as big as possible, which is how PCA was defined  $\implies$  PCA lets us represent our original data with the least amount of error with K coefficients