

# Econometrics HW4

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## 1 Question 1

1.1 Show that  $\bar{X}_{n+1} = (n\bar{X}_n + X_{n+1})/(n+1)$ .

$$\begin{aligned}\bar{X}_{n+1} &= \frac{1}{n+1} \sum_{i=1}^{n+1} X_i = \frac{1}{n+1} \left( \left( \sum_{i=1}^n X_i \right) + X_{n+1} \right) \\ &= \frac{1}{n+1} (n\bar{X}_n + X_{n+1}).\end{aligned}$$

1.2 Show that  $s_{n+1}^2 = ((n-1)s_n^2 + (n/(n+1))(X_{n+1} - \bar{X}_n)^2)/n$ .

$$\begin{aligned}s_{n+1}^2 &= \frac{1}{n} \sum_{i=1}^{n+1} (X_i - \bar{X}_{n+1})^2 = \frac{1}{n} \sum_{i=1}^{n+1} (X_i - \bar{X}_n + \bar{X}_n - \bar{X}_{n+1})^2 \\ &= \frac{1}{n} \left( (n-1) \sum_{i=1}^n (X_i - \bar{X}_{n+1})^2 + n(\bar{X}_n - \bar{X}_{n+1})^2 + 2 \left( \sum_{i=1}^n (X_i - \bar{X}_n)(\bar{X}_n - \bar{X}_{n+1}) \right) + (X_{n+1} - \bar{X}_{n+1})^2 \right) \\ &= \frac{1}{n} \left( (n-1) \sum_{i=1}^n (X_i - \bar{X}_{n+1})^2 + n(\bar{X}_n - \bar{X}_{n+1})^2 + (X_{n+1} - \bar{X}_{n+1})^2 \right) \\ &= \frac{1}{n} ((n-1)s_n^2 + n\bar{X}_n^2 - 2n\bar{X}_n\bar{X}_{n+1} + n\bar{X}_{n+1}^2 + X_{n+1}^2 - 2X_{n+1}\bar{X}_{n+1} + \bar{X}_{n+1}^2) \\ &= \frac{1}{n} \left( (n-1)s_n^2 + \frac{n}{n+1}(X_{n+1} - \bar{X}_n)^2 \right).\end{aligned}$$

## 2 Question 2

Define  $\hat{\mu}_k := \frac{1}{n} \sum_{i=1}^n X_i^k$ . We will show that this is unbiased.

$$\begin{aligned}E[\hat{\mu}_k] &= E \left[ \frac{1}{n} \sum_{i=1}^n X_i^k \right] = \frac{1}{n} \sum_{i=1}^n E[X_i^k] \\ &= \frac{1}{n} \sum_{i=1}^n \mu_k \\ &= \mu_k.\end{aligned}$$

Thus,  $\hat{\mu}_k$  is an unbiased estimator for  $\mu_k$ .

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\*I worked on this assignment with my study group: Alex von Hafften, Andrew Smith, and Ryan Mather. I have also discussed problem(s) with Emily Case, Sarah Bass, and Danny Edgel.

### 3 Question 3

Define  $\hat{m}_k := \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^k$ . This estimator is biased, which can be seen from the fact that  $\hat{m}_2 = \hat{\sigma}^2 \neq s_n^2$ , so  $\hat{m}_2$  is not an unbiased estimator for  $m_2 = \sigma^2$ . There exists no general formula for an unbiased estimator of  $m_k, k > 3$  to the best of my knowledge.

### 4 Question 4

$$\begin{aligned} \text{Var}(\hat{\mu}_k) &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i^k) = \frac{1}{n} (E[X_i^{2k}] - E[X_i^k]^2) \\ &= \frac{1}{n} (\mu_{2k} - \mu_k^2). \end{aligned}$$

### 5 Question 5

Note that  $f(x) = x^2$  is convex. By Jensen's inequality,

$$E[s_n]^2 \leq E[s_n^2] = \sigma^2,$$

so  $E[s_n] \leq \sigma$  as both are nonnegative.

### 6 Question 6

$$\begin{aligned} \hat{\sigma}^2 &= \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 = \frac{1}{n} \sum_{i=1}^n (X_i^2 - 2X_i\bar{X}_n + \bar{X}_n^2) \\ &= \frac{1}{n} \left( \sum_{i=1}^n X_i^2 - 2\bar{X}_n \sum_{i=1}^n X_i + \bar{X}_n^2 \sum_{i=1}^n 1 \right) \\ &= \frac{1}{n} \left( \sum_{i=1}^n X_i^2 - 2n\bar{X}_n^2 + n\bar{X}_n^2 \right) = \frac{1}{n} \left( \sum_{i=1}^n X_i^2 - n\bar{X}_n^2 \right). \\ \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - (\bar{X}_n - \mu)^2 &= \frac{1}{n} \sum_{i=1}^n X_i^2 - 2X_i\mu + \mu^2 - \bar{X}_n^2 + 2\mu\bar{X}_n - \mu^2 \\ &= \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}_n^2 + 2\mu(\bar{X}_n - X_i) \\ &= \frac{1}{n} \left( \sum_{i=1}^n X_i^2 - \bar{X}_n^2 \sum_{i=1}^n (1) + \sum_{i=1}^n 2\mu(\bar{X}_n - X_i) \right) \\ &= \frac{1}{n} \left( \sum_{i=1}^n X_i^2 - n\bar{X}_n^2 \right) = \hat{\sigma}^2. \end{aligned}$$

## 7 Question 7

$$\begin{aligned}
Cov(\hat{\sigma}^2, \bar{X}_n) &= E[\hat{\sigma}^2(\bar{X}_n - E\bar{X}_n)] \\
&= E\left[\left(\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - (\bar{X}_n - \mu)^2\right)(\bar{X}_n - \mu)\right] \\
&= \frac{1}{n^2} E\left[\left(\sum_{i=1}^n (X_i - \mu)^2\right)\left(\sum_{i=1}^n X_i - n\mu\right)\right] - E[(\bar{X}_n - \mu)^3] \\
&= \frac{1}{n^2} \left(\sum_{i=1}^n E[(X_i - \mu)^3]\right) + \frac{1}{n^2} \left(\sum_{i \neq j} E[(X_i - \mu)^2(X_j - \mu)]\right) - E[(\bar{X}_n - \mu)^3] \\
&= \frac{1}{n} E[(X_i - \mu)^3] - E[(\bar{X}_n - \mu)^3] \\
&= \frac{1}{n} E[(X_i - \mu)^3] \\
&\quad - \frac{1}{n^3} \left(\sum_{i=1}^n E[(X_i - \mu)^3] + 3 \sum_{i \neq j} E[(X_i - \mu)^2(X_j - \mu)] + 3 \sum_{i \neq j \neq k} E[(X_i - \mu)(X_j - \mu)(X_k - \mu)]\right) \\
&= \frac{1}{n} E[(X_i - \mu)^3] - \frac{1}{n^2} E[(X_i - \mu)^3] \\
&= \left(\frac{1}{n} - \frac{1}{n^2}\right) E[(X_i - \mu)^3]
\end{aligned}$$

This quantity will be 0 when  $X_i$  has no skewness.

## 8 Question 8

8.1 Find  $E[\bar{X}_n]$

$$\begin{aligned}
E[\bar{X}_n] &= E\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n E[X_i] \\
&= \frac{1}{n} \sum_{i=1}^n \mu_i.
\end{aligned}$$

8.2 Find  $Var(\bar{X}_n)$ .

$$\begin{aligned}
Var(\bar{X}_n) &= \frac{1}{n^2} Var\left(\sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n Var(X_i) \\
&= \frac{1}{n^2} \sum_{i=1}^n \sigma_i^2.
\end{aligned}$$

## 9 Question 9

$$\begin{aligned}
 E[Q] &= E \left[ \sum_{i=1}^r X_i^2 \right] = \sum_{i=1}^r E[X_i^2] \\
 &= \sum_{i=1}^r \mu_X^2 + \sigma_X^2 = \sum_{i=1}^r 0^2 + 1^2 \\
 &= r.
 \end{aligned}$$

$$\begin{aligned}
 Var(Q) &= E[Q^2] - E[Q]^2 = E \left[ \left( \sum_{i=1}^r X_i^2 \right)^2 \right] - r^2 \\
 &= \sum_{i=1}^r E[X_i^4] + 2 \sum_{1 \leq i < j \leq r} E[X_i^2 X_j^2] - r^2 \\
 &= \sum_{i=1}^r 3 + 2 \sum_{1 \leq i < j \leq r} (1)(1) - r^2 \\
 &= 3r + 2 \left( \frac{r!}{2(r-2)!} \right) - r^2 = 3r + r(r-1) - r^2 \\
 &= 2r.
 \end{aligned}$$

Note: we calculated  $E[X^4] = 3$  for  $X \sim N(0, 1)$  in the previous problem set.

## 10 Question 10

We will first show that a sum of independent normal random variables is normal. Let  $Z_i \sim N(\mu_i, \sigma_i^2) \forall i \in [1, \dots, n]$  for some  $n \in \mathbb{N}$ . Then the MGF of  $Z_i$  is  $M_{Z_i}(t) = \exp(\mu_i t + \frac{\sigma_i^2 t^2}{2})$ .<sup>1</sup> Then,<sup>2</sup>

$$M_{\sum_{i=1}^n Z_i}(t) = \prod_{i=1}^n M_{Z_i}(t) = \exp \left( \sum_{i=1}^n \left( \mu_i t + \frac{\sigma_i^2 t^2}{2} \right) \right) = \exp \left( \left( \sum_{i=1}^n \mu_i \right) t + \left( \sum_{i=1}^n \sigma_i^2 \right) \frac{t^2}{2} \right).$$

So  $\sum_{i=1}^n Z_i$  is of the form of a normal random variable with mean  $\sum_{i=1}^n \mu_i$  and variance  $\sum_{i=1}^n \sigma_i^2$ .

### 10.1 Find $E[\bar{X}_n - \bar{Y}_n]$

From the above,  $\bar{X}_n, \bar{Y}_n$  are of the form of normal variables with means  $\mu_X, \mu_Y$  and variances  $\frac{1}{n_1} \sigma_X^2, \frac{1}{n_2} \sigma_Y^2$ . Then,  $\bar{X}_n - \bar{Y}_n$  is also normal with mean  $\mu_X - \mu_Y$ .

### 10.2 Find $Var[\bar{X}_n - \bar{Y}_n]$

From the above,  $\bar{X}_n - \bar{Y}_n$  is normal with variance  $\frac{1}{n_1} \sigma_X^2 + \frac{1}{n_2} \sigma_Y^2$ .

### 10.3 Find the distribution of $\bar{X}_n - \bar{Y}_n$

From the above,  $\bar{X}_n - \bar{Y}_n$  is of the form of a normal random variable with mean  $\mu_X - \mu_Y$  and variance  $\frac{1}{n_1} \sigma_X^2 + \frac{1}{n_2} \sigma_Y^2$ .

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<sup>1</sup>  $M_{Z_i}(t) = \int_{-\infty}^{\infty} \exp(xt) \frac{1}{\sqrt{2\pi\sigma_i^2}} \exp(-(x - \mu_i)^2/(2\sigma_i^2)) dx$ . Define  $z = (x - \mu_i)/\sigma_i$ . Then  $M_{Z_i}(t) = \exp(\mu_i t) \int_{-\infty}^{\infty} \exp(z\sigma_i t) \frac{1}{\sqrt{2\pi\sigma_i^2}} \exp(-z^2/2) (dx/dz) dz = \exp(\mu_i t) \int_{-\infty}^{\infty} \exp(z\sigma_i t) \frac{1}{\sqrt{2\pi}} \exp(-z^2/2) dz = \exp(\mu_i t) \exp(\sigma_i^2 t^2/2) = \exp(\mu_i t + \frac{\sigma_i^2 t^2}{2})$ .

<sup>2</sup>  $M_{\sum_{i=1}^n Z_i}(t) = E[\exp(\sum_{i=1}^n t Z_i)] = E[\prod_{i=1}^n \exp(t Z_i)] = \prod_{i=1}^n E[\exp(t Z_i)] = \prod_{i=1}^n M_{Z_i}(t)$