

Econometrics HW2

Michael B. Nattinger*

February 12, 2021

1 Question 1

1.1 Part i

$$\begin{aligned} E[ZX'] &= E \left[\begin{pmatrix} Z_1 \\ X_2 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}' \right] \\ &= \begin{pmatrix} E[Z_1X_1] & E[Z_1X_2'] \\ E[X_2X_1] & E[X_2X_2'] \end{pmatrix} \\ E[ZZ'] &= E \left[\begin{pmatrix} Z_1 \\ X_2 \end{pmatrix} \begin{pmatrix} Z_1 \\ X_2 \end{pmatrix}' \right] \\ &= \begin{pmatrix} E[Z_1^2] & E[Z_1X_2'] \\ E[X_2Z_1] & E[X_2X_2'] \end{pmatrix} \end{aligned}$$

Note that $E[X_2X_2']$ must be invertible for either $E[ZX']$ or $E[ZZ']$ to be invertible.¹

Block inversion implies that $E[ZX']$ is invertible iff $E[Z_1X_1] - E[Z_1X_2']E[X_2X_2']^{-1}E[X_2X_1] \neq 0$, and similarly $E[ZZ']$ is invertible iff $E[Z_1^2] - E[Z_1X_2']E[X_2X_2']^{-1}E[X_2Z_1] \neq 0$. We can rewrite these expressions as follows: $E[\hat{Z}_1X_1] \neq 0$, $E[\hat{Z}_1^2] \neq 0$ for $\hat{Z}_1 := Z_1 - X_2'E[X_2X_2']^{-1}E[X_2Z_1]$. From FWL, for $\pi_1 = E[\tilde{Z}_1X_1]$. Together, $E[\tilde{Z}_1^2] \neq 0$ and $\pi_1 \neq 0$ imply that $E[\tilde{Z}_1X_1] \neq 0$, and the reverse direction comes from Cauchy-Schwarz:

$$0 < E[Z_1X_1]^2 \leq E[\hat{Z}_1^2]E[X_1^2].$$

1.2 Part ii

Under homoskedasticity, $\Omega = \sigma_u^2 E[ZX']^{-1}E[ZZ']E[XZ']^{-1}$. We again go back to block inversion and find that:

$$E[ZX']^{-1} = E[\tilde{Z}_1]^{-1} \begin{pmatrix} 1 & -E[Z_1X_2']E[X_2X_2']^{-1} \\ \dots & \dots \end{pmatrix},$$

*I worked on this assignment with my study group: Alex von Hafften, Andrew Smith, and Ryan Mather. I have also discussed problem(s) with Emily Case, Sarah Bass, Katherine Kwok, and Danny Edgel.

¹ $E[X_2X_2']$ not being invertible implies the existence of some t such that $E[X_2X_2']t = 0 \Rightarrow E[(X_2't)^2] = 0 \Rightarrow E[ZX'](0, t')' = E[XZ'](0, t')' = 0$ so $E[ZX']$ and $E[ZZ']$ are not invertible.

where the second block row does not enter into the upper left entry of Ω . We then have the following:

$$\begin{aligned}\Omega_{1,1} &= \frac{\sigma_U^2}{E[Z_1 X_1']^2} (E[Z_1]^2 - E[Z_1 X_2']^{-1} E[X_2 Z_1]) \\ &= \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[Z_1 X_1']^2} \\ &= \frac{\sigma_U^2}{E[\tilde{Z}_1^2] \pi_1^2}\end{aligned}$$

1.3 Part iii

π is the population projection (regression) coefficients mapping Z into X_1 , and \tilde{Z} is Z residualized at the population level with respect to X_2 .

1.4 Part iv

$$E[\tilde{Z}_1 X_2] = E[X_2 Z_1 - X_2 X_2' E[X_2 X_2']^{-1} E[X_2 Z_1]] = 0.$$

The above expression implies the following:

$$\begin{aligned}E[\tilde{Z}_1 X_2 E[X_2 X_2']^{-1} E[X_2 E[X_1|Z]]] &= 0 \\ E[\tilde{Z}_1 X_1] &= E[\tilde{Z}_1 E[X_1|Z]] - E[\tilde{Z}_1 X_2 E[X_2 X_2']^{-1} E[X_2 E[X_1|Z]]] = E[\tilde{Z} Z_*]\end{aligned}$$

We apply Cauchy-Schwarz to achieve the desired inequality:

$$\begin{aligned}\Omega_{1,1} &= \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[\tilde{Z} Z_*]^2} \\ &\geq \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[\tilde{Z}^2] E[Z_*^2]} \\ &= \frac{\sigma_U^2}{E[Z_*^2]}\end{aligned}$$

We will achieve the lower bound when $Z_* = \tilde{Z}_1 \pi_1$. This occurs when $E[X_1|Z] = Z_1 \pi_1 + X_2' \pi_2$.

1.5 Part v

If X_2 is just a constant, $\tilde{Z} = Z - E[Z]$, $E[\tilde{Z}_1^2] = \text{Var}(Z_1)$, $E[\tilde{Z}_1 X_1] = \text{Cov}(Z_1, X_1)$. Thus,

$$\Omega_{1,1} = \frac{\sigma_U^2 \text{Var}(Z_1)}{\text{Cov}(Z_1, X_1)^2}.$$

2 Question 2

2.1 Part i

$$\begin{aligned}E[h(Z)(Y - X\beta)] &= E[h(Z)(X(\beta_1 - \beta)) + U] \\ &= E[h(Z)X](\beta_1 - \beta) + E[h(Z)U].\end{aligned}$$

If exogeneity holds, $E[h(Z)U] = E[h(Z)E[U|Z]] = 0$ so $E[h(Z)(Y - X\beta)] = 0 \iff \beta_1 = \beta$.

2.2 Part ii

Define $\hat{\beta}_1^h$ as the solution to the following:

$$\frac{1}{n} \sum_i (h(Z_i)(Y_i - X_i \hat{\beta}_1^h)) = 0 \quad (1)$$

$$\Rightarrow \frac{1}{n} \sum_i (h(Z_i)(Y_i)) - \frac{1}{n} \sum_i (h(Z_i)X_i) \hat{\beta}_1^h = 0 \quad (2)$$

$$\Rightarrow \hat{\beta}_1^h = \left[\frac{1}{n} \sum_i (h(Z_i)X_i) \right]^{-1} \frac{1}{n} \sum_i (h(Z_i)(Y_i)) \quad (3)$$

2.3 Part iii

By applying the law of large numbers twice and continuous mapping theorem, we reach the following probability limit:

$$\begin{aligned} \hat{\beta}_1^h &\rightarrow_p [E[h(Z)X]]^{-1} E[h(Z)Y] \\ &= [E[h(Z)X]]^{-1} E[h(Z)(X\beta_1 + U)] \\ &= [E[h(Z)X]]^{-1} E[h(Z)X]\beta_1 + [E[h(Z)X]]^{-1} E[h(Z)U] \\ &= \beta_1 + [E[h(Z)X]]^{-1} E[h(Z)E[U|Z]] \\ &= \beta_1. \end{aligned}$$

Now, we have the following:

$$\sqrt{n}(\hat{\beta}_1^h - \beta_1) = \left[\frac{1}{n} \sum_i (h(Z_i)X_i) \right]^{-1} \frac{1}{\sqrt{n}} \sum_i (h(Z_i)(U_i))$$

By LLN, $\frac{1}{n} \sum_i (h(Z_i)X_i) \rightarrow_p E[h(Z)X]$. Also, by the CLT,

$$\frac{1}{\sqrt{n}} \sum_i (h(Z_i)(U_i)) \rightarrow_d N(0, V)$$

where $V = \text{Var}(h(Z)U) = E[(h(Z))^2 U^2] - E[h(Z)U]^2 = E[(h(Z))^2 U^2] - E[h(Z)E[U|Z]]^2 = E[(h(Z))^2 U^2]$. By the continuous mapping theorem,

$$\sqrt{n}(\hat{\beta}_1^h - \beta_1) \rightarrow_d N(0, E[h(Z)X]^{-2} E[(h(Z))^2 U^2])$$

2.4 Part iv

We have from part (iii) that $\Omega^h = E[h(Z)X]^{-2} E[(h(Z))^2 U^2]$. By applying the law of iterated expectations and then Cauchy-Schwarz,

$$\begin{aligned} \Omega^h &= \frac{E[(h(Z))^2 U^2]}{E[h(Z)X]^2} \\ &= \frac{E[(h(Z))^2 E[U^2|Z]]}{E \left[\left[h(Z) \sqrt{E[U^2|Z]} \frac{E[X|Z]}{\sqrt{E[U^2|Z]}} \right]^2 \right]} \\ &\geq \frac{E[(h(Z))^2 E[U^2|Z]]}{E[(h(Z))^2 E[U^2|Z]] \left(\frac{E[X|Z]^2}{E[U^2|Z]} \right)} \\ &= \left(\frac{E[X|Z]^2}{E[U^2|Z]} \right)^{-1} \end{aligned}$$

We can find h_* which achieves this lower bound. Let $h_* = E[X|Z](E[U^2|Z])^{-1}$. Then,

$$\begin{aligned}\Omega^{h_*} &= \frac{E[(h_*(Z))^2 U^2]}{E[h_*(Z)X]^2} \\ &= \frac{E[(E[X|Z](E[U^2|Z])^{-1})^2 E[U^2|Z]]}{E[E[X|Z](E[U^2|Z])^{-1} E[X|Z]]^2} \\ &= \frac{E[E[X|Z]^2 (E[U^2|Z])^{-1}]}{E[E[X|Z]^2 (E[U^2|Z])^{-1}]^2} \\ &= \left(\frac{E[X|Z]^2}{E[U^2|Z]} \right)^{-1}.\end{aligned}$$

3 Question 3

| | Results |
|----------|---------|
| Beta hat | 0.10842 |
| SE | 0.01948 |

```
clear; close all; clc
reload = 0;
if reload
xl=readtable('AK91.csv');
y = xl.lwage; ed = xl.educ; n = length(y);
yob = zeros(n,9); sob = zeros(n,50); qob = zeros(n,3);
for t=1:n
    if xl.yob(t)>30&&xl.yob(t)<40
        yob(t,xl.yob(t) - 30) = 1;
    end
    if xl.sob(t)>0 && xl.sob(t)< 51
        sob(t,xl.sob(t)) = 1;
    end
    if xl.qob(t)>1 && xl.qob(t)<5
        qob(t,xl.qob(t)-1) = 1;
    end
end
save 'cleanAK91'
else
load 'cleanAK91'
end
ssob = sum(sob); ssobi = ssob>0;
sob = sob(:,ssobi); % remove columns with no data
x = [ed ones(n,1) yob sob]; z = [qob ones(n,1) yob sob];
% formulas
bhat2sls = (x'*z*inv(z'*z)*z'*x)\(x'*z*inv(z'*z)*z'*y);
e = y-x*bhat2sls; Qzz = z'*z/n; Qxz = x'*z/n;
Om = 0*Qzz;
for i=1:n
    Om = Om + z(i,:)'*z(i,:)*(e(i)^2);
end
Om = Om/n;
minv = inv(Qxz*inv(Qzz)*Qxz');
VhB = minv*(Qxz*inv(Qzz)*Om*inv(Qzz)*Qxz')*minv/n;
tab = table([bhat2sls(1,1); sqrt(VhB(1,1))], 'VariableNames', ...
    {'Results'}, 'RowNames', {'Beta hat' 'SE'});
table2latex(tab, 'ps3.tex')
```