Econometrics HW2

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March 27, 2021

1 Question 18.2

1.1 Part A

This is essentially just a regression of Y on D with state fixed effects, so the estimator is the following:

$$\hat{\theta} = \frac{\sum_{t=0}^{T} \sum_{i=0}^{N} (D_{it} - \bar{D}_i)(Y_{it} - \bar{Y}_i)}{\sum_{t=i}^{T} \sum_{i=1}^{N} (D_{it} - \bar{D}_i)^2},$$

where $\bar{D}_i = (1/T) \sum_{t=1}^T D_{it}, \bar{Y}_i = (1/T) \sum_{t=1}^T Y_{it}$.

1.2 Part B

For the untreated sample, $D_{it} = 0 \forall t$, so $D_{it} - \bar{D}_i = 0 \forall t$, so $\hat{\theta} = \frac{\sum_{t=0}^{T} \sum_{i=0}^{N} (D_{it} - \bar{D}_i)(Y_{it} - \bar{Y}_i)}{\sum_{t=0}^{T} \sum_{i=0}^{N} (D_{it} - \bar{D}_i)^2} = \frac{\sum_{t=0}^{T} (D_{1t} - \bar{D}_1)(Y_{1t} - \bar{Y}_1)}{\sum_{t=0}^{T} (D_{1t} - \bar{D}_1)^2}$ which is a function only of the treated sample.

1.3 Part C

No, it is only a difference estimator of the treated group. It is not accounting for the control group, as shown from the independence result from Part B.

1.4 Part D

If the time trend is not important then the control group would have no change over time, so the difference estimator would be the same as the difference-in-difference estimator.

2 Question 18.4

For both, if all interaction dummies were included then there would be perfect collinearity and the X'X matrix would not be invertible. Intuitively one of the groups serve as a baseline from which the other groups are compared.

3 Question 18.5

3.1 Part A

The difference in Wisconsin is 16.72 - 15.23 = 1.49. The difference in Minnesota is 18.10 - 16.42 = 1.68. The difference in difference estimate is therefore 1.49 - 1.68 = -0.19.

^{*}I worked on this assignment with my study group: Alex von Hafften, Andrew Smith, and Ryan Mather. I have also discussed problem(s) with Emily Case, Sarah Bass, Katherine Kwok, and Danny Edgel.

3.2 Part B

$$\hat{\beta} = -0.1.$$

3.3 Part C

$$\hat{\gamma} = 15.23 - 16.42 = -1.19.$$

4 Question 17.1

4.1 Part A

We will integrate and make a variable substitution $y = \frac{X_i - x}{h} \Rightarrow h dy = dx$

$$E[X^*] = \int x \frac{1}{nh} \sum_{i=1}^n K((X_i - x)/h) dx$$

$$= \frac{1}{nh} \sum_{i=1}^n \int x K((X_i - x)/h) dx$$

$$= \frac{1}{n} \sum_{i=1}^n \int (X_i - hy) K(y) dy$$

$$= \frac{1}{n} \sum_{i=1}^n X_i \int K(y) dy - \frac{h}{n} \sum_{i=1}^n \int y K(y) dy$$

$$= \frac{1}{n} \sum_{i=1}^n X_i$$

$$= \bar{X}_n$$

where we have used the fact that the kernels integrate to 1 and have mean 0.

4.2 Part B

We will again ust the same substitution.

$$Var(X^*) = E[X^{*2}] - E[X^*]^2$$

$$= \int x^2 \frac{1}{nh} \sum_{i=1}^n K((X_i - x)/h) dx - \bar{X}_n^2$$

$$= \frac{1}{n} \sum_{i=1}^N \int (X_i - hy)^2 K(y) dy - \bar{X}_n^2$$

$$= \frac{1}{n} \sum_{i=1}^N X_i^2 \int K(y) dy - \frac{2h}{n} \sum_{i=1}^N X_i \int y K(y) dy + \frac{h^2}{n} \sum_{i=1}^N \int y^2 K(y) dy - \bar{X}_n^2$$

$$= \frac{1}{n} \sum_{i=1}^N X_i^2 + h^2 - \bar{X}_n^2$$

$$= \hat{\sigma}^2 + h^2.$$

5 Question 17.3

The optimal bandwidth is $h_0 = \left(\frac{R_K}{R(f'')}\right)^{1/5} n^{-1/5}$. Note that, for a uniform distribution, $f'' = 0 \Rightarrow R(f'') = 0$. The optimal bandwidth is the largest feasible bandwidth.

6 Question 17.4

The effective bandwidth would be much larger as the scale was reduced by a factor of 10^6 yet the bandwidth was unchanged. The plot would be much wider, smoother, and lower than it should be.

If the bandwidth was scaled appropriately by 1000000^{-1} when the scale changed, the density plots would have the same shape.

7 Question 19.3

When m(x) is increasing and convex, the bias is positive. When m(x) is increasing and concave, the bias is negative. When m(x) is decreasing and concave, the bias is again negative. When m(x) is decreasing and convex, the bias is positive. The intuition here is that you are locally averaging around a point of interest. This is similar to taking the midpoint of a function over a region of concavity or convexity.

Moreover, asymptotic theory tells us that the bias is $\frac{1}{2}m''(x)h^2$ and m''(x) > 0 for a convex function and m''(x) < 0 for a concave function.

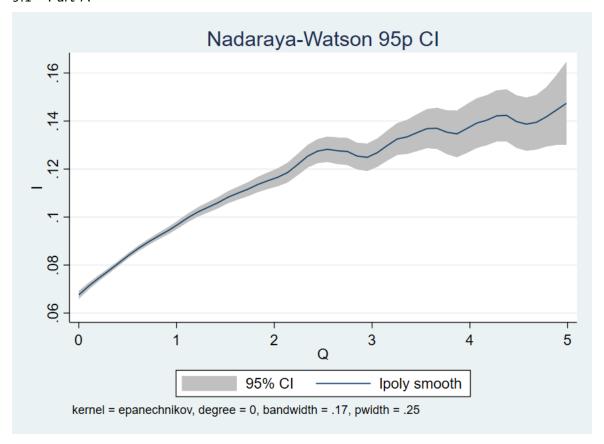
8 Question 19.4

Bias is $B(x) = (1/2)m''(x) + m'(x)f'(x)/f(x) = \beta f'(x)/f(x)$. If $\beta > 0$, then $b(x) < 0 \iff f'(x) < 0$, $b(x) > 0 \iff f'(x) > 0$. In contrast, if $\beta < 0$ then $b(x) < 0 \iff f'(x) > 0$, $b(x) > 0 \iff f'(x) < 0$.

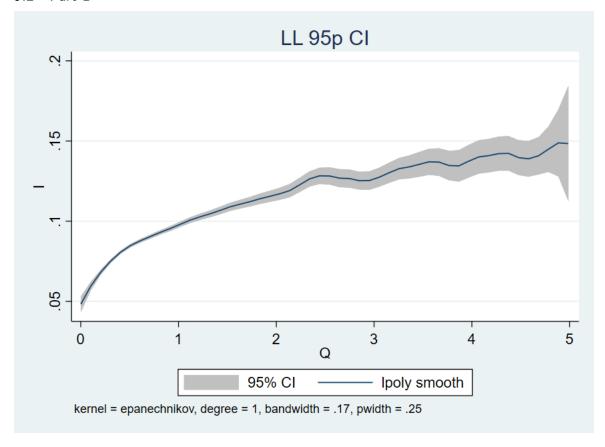
f(x) is the marginal density of X. If f'(x) > 0 then the density of X is increasing at x, so if β is positive (negative), the bias at x is positive because it is influenced more by the mass of X to the right of x, so the positive (negative) influence of this mass on the conditional expectation due to the positive (negative) value of β leads to a positive (negative) bias.

9 Question 19.9

9.1 Part A



9.2 Part B



9.3 Part C

Yes, from the graphs in the preceding parts of this question it appears that there is nonlinearity in the relationship.

```
Stata code:
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use "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\Invest1993",

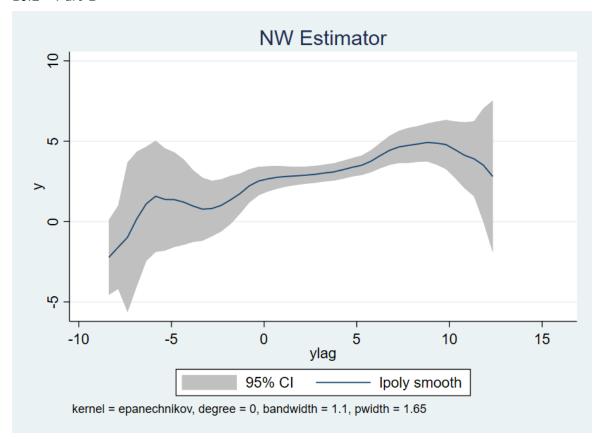
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gen Q = vala
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graph export "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\pings\9p1.png", as(
png) name("Graph")
lpoly I Q if Q<=5, degree(1) ci nosc title("LL 95p CI")
graph export "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\pings\9p2.png", as(
png) name("Graph")
```

10 Question 19.11

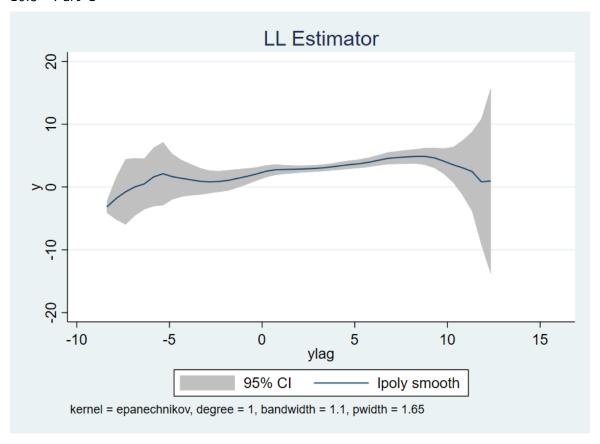
10.1 Part A

Done in stata, see code below the final part of this question.

10.2 Part B



10.3 Part C



10.4 Part D

Yes, from the graphs in the preceding parts of this question it appears that there is nonlinearity in the relationship.

Stata code: pause on use "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\FRED-QD", clear gen y = 100*((gdpc1/L.gdpc1)^4-1) gen ylag = L.y lpoly y ylag, ci nosc title("NW Estimator") graph export "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\pings\11p1.png", as(png) name("Graph") lpoly y ylag, degree(1) ci nosc title("LL Estimator") graph export "C:\Users\micha\OneDrive\Documents\HOMEWORK\Y1S1\Metrics\Q4\PS2\pings\11p2.png", as(png) name("Graph")