

# Econometrics HW2

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February 16, 2021

## 1 Question 1

### 1.1 Part i

$$\begin{aligned} E[ZX'] &= E \left[ \begin{pmatrix} Z_1 \\ X_2 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}' \right] \\ &= \begin{pmatrix} E[Z_1X_1] & E[Z_1X_2'] \\ E[X_2X_1] & E[X_2X_2'] \end{pmatrix} \\ E[ZZ'] &= E \left[ \begin{pmatrix} Z_1 \\ X_2 \end{pmatrix} \begin{pmatrix} Z_1 \\ X_2 \end{pmatrix}' \right] \\ &= \begin{pmatrix} E[Z_1^2] & E[Z_1X_2'] \\ E[X_2Z_1] & E[X_2X_2'] \end{pmatrix} \end{aligned}$$

Note that  $E[X_2X_2']$  must be invertible for either  $E[ZX']$  or  $E[ZZ']$  to be invertible.<sup>1</sup>

Block inversion implies that  $E[ZX']$  is invertible iff  $E[Z_1X_1] - E[Z_1X_2']E[X_2X_2']^{-1}E[X_2X_1] \neq 0$ , and similarly  $E[ZZ']$  is invertible iff  $E[Z_1^2] - E[Z_1X_2']E[X_2X_2']^{-1}E[X_2Z_1] \neq 0$ . We can rewrite these expressions as follows:  $E[\tilde{Z}_1X_1] \neq 0$ ,  $E[\tilde{Z}_1^2] \neq 0$  for  $\tilde{Z}_1 := Z_1 - X_2'E[X_2X_2']^{-1}E[X_2Z_1]$ . From FWL, for  $\pi_1 = E[\tilde{Z}_1X_1]E[\tilde{Z}_1^2]^{-1}$ . Together,  $E[\tilde{Z}_1^2] \neq 0$  and  $\pi_1 \neq 0$  imply that  $E[\tilde{Z}_1X_1] \neq 0$ , and the reverse direction comes from Cauchy-Schwarz:

$$0 < E[Z_1X_1]^2 \leq E[\tilde{Z}_1^2]E[X_1^2].$$

### 1.2 Part ii

Under homoskedasticity,  $\Omega = \sigma_u^2 E[ZX']^{-1}E[ZZ']E[XZ']^{-1}$ . We again go back to block inversion and find that:

$$E[ZX']^{-1} = E[\tilde{Z}_1]^{-1} \begin{pmatrix} 1 & -E[Z_1X_2']E[X_2X_2']^{-1} \\ \dots & \dots \end{pmatrix},$$

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\*I worked on this assignment with my study group: Alex von Hafften, Andrew Smith, and Ryan Mather. I have also discussed problem(s) with Emily Case, Sarah Bass, Katherine Kwok, and Danny Edgel.

<sup>1</sup> $E[X_2X_2']$  not being invertible implies the existence of some  $t$  such that  $E[X_2X_2']t = 0 \Rightarrow E[(X_2't)^2] = 0 \Rightarrow E[ZX'](0, t')' = E[ZX'](0, t')' = 0$  so  $E[ZX']$  and  $E[ZZ']$  are not invertible.

where the second block row does not enter into the upper left entry of  $\Omega$ . We then have the following:

$$\begin{aligned}\Omega_{1,1} &= \frac{\sigma_U^2}{E[Z_1 X_1']^2} (E[Z_1]^2 - E[Z_1 X_2']^{-1} E[X_2 Z_1]) \\ &= \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[Z_1 X_1']^2} \\ &= \frac{\sigma_U^2}{E[\tilde{Z}_1^2] \pi_1^2}\end{aligned}$$

### 1.3 Part iii

$\pi$  is the population projection (regression) coefficients mapping  $Z$  into  $X_1$ , and  $\tilde{Z}$  is  $Z$  residualized at the population level with respect to  $X_2$ .

### 1.4 Part iv

$$E[\tilde{Z}_1 X_2] = E[X_2 Z_1 - X_2 X_2' E[X_2 X_2']^{-1} E[X_2 Z_1]] = 0.$$

The above expression implies the following:

$$\begin{aligned}E[\tilde{Z}_1 X_2 E[X_2 X_2']^{-1} E[X_2 E[X_1|Z]]] &= 0 \\ E[\tilde{Z}_1 X_1] &= E[\tilde{Z}_1 E[X_1|Z]] - E[\tilde{Z}_1 X_2 E[X_2 X_2']^{-1} E[X_2 E[X_1|Z]]] = E[\tilde{Z} Z_*]\end{aligned}$$

We apply Cauchy-Schwarz to achieve the desired inequality:

$$\begin{aligned}\Omega_{1,1} &= \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[\tilde{Z} Z_*]^2} \\ &\geq \frac{\sigma_U^2 E[\tilde{Z}_1^2]}{E[\tilde{Z}^2] E[Z_*^2]} \\ &= \frac{\sigma_U^2}{E[Z_*^2]}\end{aligned}$$

We will achieve the lower bound when  $Z_* = \tilde{Z}_1 \pi_1$ . This occurs when  $E[X_1|Z] = Z_1 \pi_1 + X_2' \pi_2$ .

### 1.5 Part v

If  $X_2$  is just a constant,  $\tilde{Z} = Z - E[Z]$ ,  $E[\tilde{Z}_1^2] = \text{Var}(Z_1)$ ,  $E[\tilde{Z}_1 X_1] = \text{Cov}(Z_1, X_1)$ . Thus,

$$\Omega_{1,1} = \frac{\sigma_U^2 \text{Var}(Z_1)}{\text{Cov}(Z_1, X_1)^2}.$$

## 2 Question 2

### 2.1 Part i

$$\begin{aligned}E[h(Z)(Y - X\beta)] &= E[h(Z)(X(\beta_1 - \beta) + U)] \\ &= E[h(Z)X](\beta_1 - \beta) + E[h(Z)U].\end{aligned}$$

If exogeneity holds,  $E[h(Z)U] = E[h(Z)E[U|Z]] = 0$  so  $E[h(Z)(Y - X\beta)] = 0 \iff \beta_1 = \beta$ .

## 2.2 Part ii

Define  $\hat{\beta}_1^h$  as the solution to the following:

$$\frac{1}{n} \sum_i (h(Z_i)(Y_i - X_i \hat{\beta}_1^h)) = 0 \quad (1)$$

$$\Rightarrow \frac{1}{n} \sum_i (h(Z_i)(Y_i)) - \frac{1}{n} \sum_i (h(Z_i)X_i) \hat{\beta}_1^h = 0 \quad (2)$$

$$\Rightarrow \hat{\beta}_1^h = \left[ \frac{1}{n} \sum_i (h(Z_i)X_i) \right]^{-1} \frac{1}{n} \sum_i (h(Z_i)(Y_i)) \quad (3)$$

## 2.3 Part iii

By applying the law of large numbers twice and continuous mapping theorem, we reach the following probability limit:

$$\begin{aligned} \hat{\beta}_1^h &\rightarrow_p [E[h(Z)X]]^{-1} E[h(Z)Y] \\ &= [E[h(Z)X]]^{-1} E[h(Z)(X\beta_1 + U)] \\ &= [E[h(Z)X]]^{-1} E[h(Z)X]\beta_1 + [E[h(Z)X]]^{-1} E[h(Z)U] \\ &= \beta_1 + [E[h(Z)X]]^{-1} E[h(Z)E[U|Z]] \\ &= \beta_1. \end{aligned}$$

Now, we have the following:

$$\sqrt{n}(\hat{\beta}_1^h - \beta_1) = \left[ \frac{1}{n} \sum_i (h(Z_i)X_i) \right]^{-1} \frac{1}{\sqrt{n}} \sum_i (h(Z_i)(U_i))$$

By LLN,  $\frac{1}{n} \sum_i (h(Z_i)X_i) \rightarrow_p E[h(Z)X]$ . Also, by the CLT,

$$\frac{1}{\sqrt{n}} \sum_i (h(Z_i)(U_i)) \rightarrow_d N(0, V)$$

where  $V = \text{Var}(h(Z)U) = E[(h(Z))^2 U^2] - E[h(Z)U]^2 = E[(h(Z))^2 U^2] - E[h(Z)E[U|Z]]^2 = E[(h(Z))^2 U^2]$ . By the continuous mapping theorem,

$$\sqrt{n}(\hat{\beta}_1^h - \beta_1) \rightarrow_d N(0, E[h(Z)X]^{-2} E[(h(Z))^2 U^2])$$

## 2.4 Part iv

We have from part (iii) that  $\Omega^h = E[h(Z)X]^{-2} E[(h(Z))^2 U^2]$ . By applying the law of iterated expectations and then Cauchy-Schwarz,

$$\begin{aligned} \Omega^h &= \frac{E[(h(Z))^2 U^2]}{E[h(Z)X]^2} \\ &= \frac{E[(h(Z))^2 E[U^2|Z]]}{E \left[ \left[ h(Z) \sqrt{E[U^2|Z]} \frac{E[X|Z]}{\sqrt{E[U^2|Z]}} \right]^2 \right]} \\ &\geq \frac{E[(h(Z))^2 E[U^2|Z]]}{E[(h(Z))^2 E[U^2|Z]] \left( \frac{E[X|Z]^2}{E[U^2|Z]} \right)} \\ &= \left( \frac{E[X|Z]^2}{E[U^2|Z]} \right)^{-1} \end{aligned}$$

We can find  $h_*$  which achieves this lower bound. Let  $h_* = E[X|Z](E[U^2|Z])^{-1}$ . Then,

$$\begin{aligned}\Omega^{h_*} &= \frac{E[(h_*(Z))^2 U^2]}{E[h_*(Z)X]^2} \\ &= \frac{E[(E[X|Z](E[U^2|Z])^{-1})^2 E[U^2|Z]]}{E[E[X|Z](E[U^2|Z])^{-1} E[X|Z]]^2} \\ &= \frac{E[E[X|Z]^2 (E[U^2|Z])^{-1}]}{E[E[X|Z]^2 (E[U^2|Z])^{-1}]^2} \\ &= \left( \frac{E[X|Z]^2}{E[U^2|Z]} \right)^{-1}.\end{aligned}$$

### 3 Question 3

	Results
Beta hat	0.10842
SE	0.01948

```
clear; close all; clc
reload = 0;
if reload
xl=readtable('AK91.csv');
y = xl.lwage; ed = xl.educ; n = length(y);
yob = zeros(n,9); sob = zeros(n,50); qob = zeros(n,3);
for t=1:n
    if xl.yob(t)>30&&xl.yob(t)<40
        yob(t,xl.yob(t) - 30) = 1;
    end
    if xl.sob(t)>0 && xl.sob(t)< 51
        sob(t,xl.sob(t)) = 1;
    end
    if xl.qob(t)>1 && xl.qob(t)<5
        qob(t,xl.qob(t)-1) = 1;
    end
end
save 'cleanAK91'
else
load 'cleanAK91'
end
ssob = sum(sob); ssobi = ssob>0;
sob = sob(:,ssobi); % remove columns with no data
x = [ed ones(n,1) yob sob]; z = [qob ones(n,1) yob sob];
% formulas
bhat2sls = (x'*z*inv(z'*z)*z'*x)\(x'*z*inv(z'*z)*z'*y);
e = y-x*bhat2sls; Qzz = z'*z/n; Qxz = x'*z/n;
Om = 0*Qzz;
for i=1:n
    Om = Om + z(i,:)'*z(i,:)*(e(i)^2);
end
Om = Om/n;
minv = inv(Qxz*inv(Qzz)*Qxz');
VhB = minv*(Qxz*inv(Qzz)*Om*inv(Qzz)*Qxz')*minv/n;
tab = table([bhat2sls(1,1); sqrt(VhB(1,1))], 'VariableNames', ...
    {'Results'}, 'RowNames', {'Beta hat' 'SE'});
table2latex(tab, 'ps3.tex')
```