

CS 107, Probability, Spring 2020

Lecture 37

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AUA

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- Simple Concentration Inequalities
 - Markov's Inequality
 - Chebyshev's Inequality
- Limit Theorems
 - The Law of Large Numbers
 - The Central Limit Theorem

Simple Concentration Inequalities: Markov and Chebyshev inequalities

Markov and Chebyshev Inequalities: Intro

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First, we will consider the Markov's Inequality, assuming we know the Expectation $\mathbb{E}(X)$.

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Now, it is easy to see that

$$Y \leq X$$

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In the general case, if X is any r.v., we can apply the Markov's Inequality to $|X|$ and obtain

$$\mathbb{P}(|X| \geq a) \leq \frac{\mathbb{E}(|X|)}{a}.$$

Markov and Chebyshev Inequalities

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Note: Also, we can rewrite the Markov's Inequality as: for any r.v. $X \geq 0$, and for any $a > 0$,

$$\mathbb{P}(X < a) \geq 1 - \frac{\mathbb{E}(X)}{a}.$$

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Proof: Indeed, by the Markov's Inequality,

$$\mathbb{P}(X \geq 10\mathbb{E}(X)) \leq \frac{\mathbb{E}(X)}{10\mathbb{E}(X)} = \frac{1}{10}.$$

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From Wikipedia: by the Markov Inequality, no more than $1/5$ of the population can have more than 5 times the average income.

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$$\mathbb{P}\left(|X - \mathbb{E}(X)| \geq a \cdot SD(X)\right) \leq \frac{1}{a^2}.$$

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Note: Another way to represent the Chebyshev's Inequality is

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Say, if the Variance of X is small, then the probability that X will be far from $\mathbb{E}(X)$ is small.

E.g., the probability that X will be more than $3SD(X)$ -away from $\mathbb{E}(X)$ is:

$$\mathbb{P}\left(|X - \mathbb{E}(X)| \geq 3 \cdot SD(X)\right) \leq \frac{1}{3^2} = \frac{1}{9}.$$

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And this works for ANY r.v. X !

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- If X is ANY r.v. with $\mathbb{E}(X) = \mu$ and $SD(X) = \sigma$, then, by Chebyshev Inequality,

$$\mathbb{P}(|X - \mu| < 3\sigma) \geq 1 - \frac{1}{9} = \frac{8}{9}.$$

Chebyshev Inequality, Example

Example 37.1: Assume X is a r.v. with $\mathbb{E}(X) = 9$ and $\text{Var}(X) = 0.2$. Estimate

- a. $\mathbb{P}(8 < X < 10)$;
- b. $\mathbb{P}(8.1 < X < 10)$.

Chebyshev Inequality, Example

Example 37.2: Assume X_1, X_2, \dots, X_n are IID r.vs with finite $\mu = \mathbb{E}(X_k)$ and $\sigma^2 = \text{Var}(X_k)$, and let

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}.$$

Prove that

$$\mathbb{P}(|\bar{X}_n - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{n \cdot \varepsilon^2}.$$

Chebyshev Inequality, Example

Example 37.3: Assume X is a r.v. and $g(t)$ is a nonnegative strictly increasing function.

a. Prove that

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}(g(X))}{g(a)};$$

b. Prove that

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}(e^X)}{e^a}.$$

Limit Theorems

General Facts about the Sum and Average of IID r.v.s

IID Sequence of r.v.s

Recall that we defined X_1, X_2, \dots, X_n are IID if

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Note: Please note that since all X_k -s have the same Distribution, then they have the same Expected values, the same Variances and all other characteristics, i.e.

$$\mathbb{E}(X_1) = \mathbb{E}(X_2) = \dots = \mathbb{E}(X_n),$$

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In the rest, we will consider an infinite sequence of IID r.v.s X_1, X_2, X_3, \dots

The Question

The Questions we consider here are:

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Assume X_1, X_2, \dots, X_n are IID r.v.s.

Q1 What is the Distribution of

$$S_n = X_1 + X_2 + \dots + X_n?$$

Q2 What is the Distribution of

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As you remember, calculation of the Distribution of the sum $X + Y$ is not an easy job (one needs to calculate Convolutions), so calculation of the exact Distribution of S_n and \bar{X}_n is not an easy job, in general.

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$$S_n = X_1 + \dots + X_n \sim \text{Binom}(n, p);$$

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- If $X_k \sim \text{Binom}(m, p)$, $k = 1, \dots, n$, are independent, then

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- If $X_k \sim \mathcal{N}(\mu, \sigma^2)$, $k = 1, \dots, n$, are Independent, then

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$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right).$$

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- If $X_k \sim \text{Pois}(\lambda)$, $k = 1, \dots, n$, are independent, then

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- If $X_k \sim \text{Geom}(p)$, $k = 1, \dots, n$, are independent, then

$$S_n = X_1 + \dots + X_n \sim \text{NBinom}(n, p).$$

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Folklore: Diversification in one sentence: Do not put all your eggs into one basket!

The Law of Large Numbers

Intro to LLN and CLT

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The two famous Limit Theorems in Probability Theory,

- The Law of the Large Numbers (LLN)
- The Central Limit Theorem (CLT)

help us to get an information about the **asymptotic** (i.e., limiting, or, for large n) properties of \bar{X}_n and S_n , in the general case.

The Weak LLN

The Weak Law of Large Numbers, WLLN

If $X_1, X_2, \dots, X_n, \dots$ are IID, with finite $\mathbb{E}(X_1)$ and Variance $\text{Var}(X_1)$, then

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n} \xrightarrow{\mathbb{P}} \mathbb{E}(X_1), \quad n \rightarrow +\infty,$$

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Note: This means that for any $\varepsilon > 0$, the chances that \bar{X}_n is far from $\mathbb{E}(X_1)$ more than ε , is very small, if n is large.

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The rigorous proof is by using the Chebyshev Inequality: OTB

The Strong LLN

The Strong LLN Says that the above convergence holds also in the Strong Sense, under less restrictive settings:

The Strong Law of Large Numbers, SLLN, Kolmogorov

If $X_1, X_2, \dots, X_n, \dots$ are IID, with finite $\mathbb{E}(X_1)$, then

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where *a.s.* over the arrow sign means the convergence is in Almost Sure sense:

$$\mathbb{P} \left(\lim_{n \rightarrow +\infty} \frac{X_1 + X_2 + \dots + X_n}{n} = \mathbb{E}(X_1) \right) = 1.$$

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Note: Assume $g : \mathbb{R} \rightarrow \mathbb{R}$ is a continuous function. Then one can prove that, under the SLLN(WLLN) conditions,

$$g\left(\frac{X_1 + X_2 + \dots + X_n}{n}\right) \xrightarrow{a.s. (\mathbb{P})} g(\mathbb{E}(X_1)), \quad n \rightarrow +\infty.$$

The Strong LLN

Note: Sometimes we are required to calculate limits of the form:

$$\lim_{n \rightarrow +\infty} \frac{g(X_1) + g(X_2) + \dots + g(X_n)}{n}$$

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in the Probability or a.s. sense, for some nice function g . Clearly, under the condition that $\mathbb{E}(g(X_1))$ and $\text{Var}(g(X_1))$ are finite (for the WLLN), or just $\mathbb{E}(g(X_1))$ is finite (for the SLLN), we will have, for $n \rightarrow +\infty$,

$$\frac{g(X_1) + g(X_2) + \dots + g(X_n)}{n} \xrightarrow{a.s. (\mathbb{P})} \mathbb{E}(g(X_1)).$$

W(S)LLN, Example

Example 37.4: Assume $X_1, X_2, \dots, X_n \sim \text{Unif}[-1, 2]$ are IID. Calculate, in the \mathbb{P} and a.s. sense,

- a. $\lim_{n \rightarrow +\infty} \frac{X_1 + X_2 + \dots + X_n}{n};$
- b. $\lim_{n \rightarrow +\infty} \frac{X_1^2 + X_2^2 + \dots + X_n^2}{n};$
- c. $\lim_{n \rightarrow +\infty} \frac{e^{X_1} + e^{X_2} + \dots + e^{X_n}}{n}.$

W(S)LLN, Example

Example 37.5: Assume we have a coin, for which the probability of having Heads is $p \in (0, 1)$. We are tossing that coin many times, independently. We calculate the proportion of the Heads for that tosses. What is the limit of that proportion, almost surely, if we repeat tossing infinitely many times?

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Note: This example is the base of Frequentists interpretation of the Probability: "the Probability of an Event is p " can be understood as the following: we repeat the Experiment many times, independently, and calculate the proportion of times we will have that Event appearing. Then the limit of that proportion is exactly p .

W(S)LLN, Example

Example 37.6: Assume $g : [a, b] \rightarrow \mathbb{R}$ is a Continuous function, and we want to calculate the integral

$$\int_a^b g(x) dx.$$

- a. Prove that if X_1, X_2, \dots are IID with $X_k \sim \text{Unif}[a, b]$, then

$$\frac{g(X_1) + g(X_2) + \dots + g(X_n)}{n} \rightarrow \frac{1}{b-a} \cdot \int_a^b g(x) dx;$$

- b. Calculate, in \mathbf{R} , the integral

$$\int_0^2 \sin(x) \cdot e^{-x^2} dx.$$

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Question: Why the idea of a Casino works?

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Consider, say, the American Roulette game: we have 38 pockets (numbers from 0 to 36, and 00), 0 and 00 are in green, 18 numbers are in red, and 18 are in black. One of the betting methods is to bet on the color - either black or red. Say, a person is winning 1\$ if he/she correctly guessed the color, and, otherwise, is losing 1\$.

Example:

We have calculated that the Expected winning of that person is $\mathbb{E}(W) = -\frac{1}{19}$, or, which is the same, from the point of view of our Casino, the Expected Gain of Casino will be $\mathbb{E}(G) = \frac{1}{19}$.

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Assume G_1 is the gain of Casino from the 1st person, G_2 - from the second one etc. So G_k -s are random, IID. It can happen, of course that, in some scenario, $G_1 = G_2 = \dots = G_{50} = -1$.

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So, even sometimes Casino will pay to players, in the long run, for many players, Casino will win: it will have almost $\frac{n}{19}$ \$ from n persons.

The Central Limit Theorem

Reminder about Standard Normal r.vs

Let me first recall that if Z is Standard Normal, i.e.

$$Z \sim \mathcal{N}(0, 1),$$

then we denote by $\Phi(x)$ its CDF and $\varphi(x)$ its PDF:

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}}, \quad x \in \mathbb{R}$$

and

$$\Phi(x) = \int_{-\infty}^x \varphi(t) dt, \quad x \in \mathbb{R}.$$

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And we know that

$$\mathbb{P}(a \leq Z \leq b) = \Phi(b) - \Phi(a),$$

and we can calculate the values of Φ either by some mathematical software, or by using Standard Normal tables.

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And these hold **for any IID sequence X_k , from any Distribution.**

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Now, it is easy to see that the Standardization of S_n and \bar{X}_n yields to the same r.v.:

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Hence, the above two versions of CLT are the same, just one is in terms of S_n , the other one is in terms of \bar{X}_n .

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We use this in the following form: for large n (in most of the cases, $n \geq 30$ will suffice)

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Now, let me give more precisely how to calculate Probabilities using the CLT.

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We need to have Standardized r.v. for the CLT, so first we write the above inequality in the following form:

$$\mathbb{P}(\bar{X}_n \leq b) = \mathbb{P}\left(\frac{\bar{X}_n - \mathbb{E}(\bar{X}_n)}{\sqrt{\text{Var}(\bar{X}_n)}} \leq \frac{b - \mathbb{E}(\bar{X}_n)}{\sqrt{\text{Var}(\bar{X}_n)}}\right).$$

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And we can calculate $\mathbb{E}(\bar{X}_n)$, $\text{Var}(\bar{X}_n)$, and, finally, the last Probability.

CLT, Example

Example 37.7: Assume I have a Piggy Bank. I am collecting coins worth 500 AMD in my Piggy. The mean weight of a 500 AMD coin is 5 gr with Standard Deviation 0.1 gr. I have collected 250 coins. Assume W is the weight, in grams, of all that 250 coins.

- a. Calculate, approximately, $\mathbb{P}(W > 1248)$;
- b. Calculate, approximately, $\mathbb{P}(1240 \leq W < 1260)$.

CLT, Example

Example 37.8: The lengths of rods produced by a machine have a mean 100 cm and standard deviation 5 cm. Find the probability that if 60 rods are randomly chosen from the machine, the mean length of the sample will be at most 101 cm.

CLT, Roughly

Let us go back again to our CLT statement. In a non-rigorous way, we can write, for large n (here \approx means approximately distributed as):

$$\frac{S_n - \mathbb{E}(S_n)}{\sqrt{\text{Var}(S_n)}} \approx \mathcal{N}(0, 1) \quad \text{and} \quad \frac{\bar{X}_n - \mathbb{E}(\bar{X}_n)}{\sqrt{\text{Var}(\bar{X}_n)}} \approx \mathcal{N}(0, 1).$$

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Using $\mu = \mathbb{E}(X_1)$ and $\sigma^2 = \text{Var}(X_1)$, we can write

$$\frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma} \approx \mathcal{N}(0, 1) \quad \text{and} \quad \frac{\bar{X}_n - \mu}{\sigma / \sqrt{n}} \approx \mathcal{N}(0, 1).$$

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or

$$S_n \approx \mathcal{N}(n\mu, n\sigma^2) \quad \text{and} \quad \bar{X}_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right).$$

CLT, Roughly -2

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- If X_k -s are independent, have the Mean $\mathbb{E}(X_k) = \mu$ and $Var(X_K) = \sigma^2$, and **are Normally Distributed**, i.e., $X_k \sim \mathcal{N}(\mu, \sigma^2)$, then

CLT, Roughly -2

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so we know the **exact Distributions** of S_n and \bar{X}_n .

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CLT, Roughly -2

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- If X_k -s are independent, have the Mean $\mathbb{E}(X_k) = \mu$ and $\text{Var}(X_K) = \sigma^2$, and are **from any Distribution**, then

$$S_n \approx \mathcal{N}(n\mu, n\sigma^2) \quad \text{and} \quad \bar{X}_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right);$$

and we know the **asymptotic Distributions**
(approximate Distributions for large n) of S_n and \bar{X}_n .