AUA CS108, Statistics, Fall 2020 Lecture 43

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07 Dec 2020

Contents

▶ Intro to Linear Regression

Linear Regression

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- to find a good Point Estimator and Estimate;
- ▶ to find a CI for θ of given CL;
- ▶ to Test a Hypothesis about θ , say, is it likely that $\theta = 3.1415926$ or not.

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Recall that, in the Descriptive Statistics part, we considered two Datasets, and defined the Sample Covariance and Correlation Coefficients, to measure the Linear Relationship between that Datasets. That was defined for two **Numerical Dataset**, without any assumptions behind the Process generating that Datasets. Now, if we assume that that Datasets are coming from some Distribution, we are at the stage of doing a Statistical Inference, Statistical Analysis.

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Note: The coordinates of **X** are called **Explanatory**, **Predictor**, **Independent Variables or Features**, and *Y* is called the **Dependent**, **Response Variable or the Label**.

Example: We want to find a relationship between the Stat Total Grade, STG, and the Mean Weekly Hours Spent on Statistics (during the Semester), MWHSS.

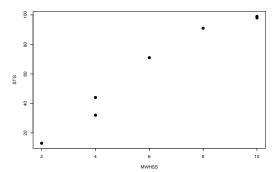
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Say, see Wiki page for BMI.

The simplest Regression Model is the Linear Model: in the Linear Regression Problem, we assume

$$Y = \beta_0 + \beta^T \cdot \mathbf{X} + \varepsilon,$$

where ε is a r.v., for each value of **X**, with $\mathbb{E}(\varepsilon) = 0$.

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$$\beta = \left[\begin{array}{c} \beta_1 \\ \vdots \\ \beta_d \end{array} \right],$$

so our Model, in the expanded way, is

$$Y = \beta_0 + \beta_1 \cdot X^1 + \beta_2 \cdot X^2 + \dots + \beta_d \cdot X^d + \varepsilon.$$

Simple Linear Regression

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This is called the **Simple Linear Regression Model**.

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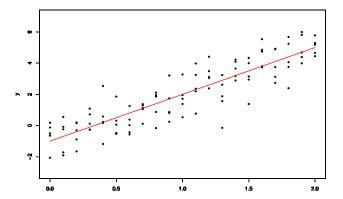
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Again, let us assume we have a Simple Linear Regression Model:

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where ε_k -s are Independent, and our aim is to find good Estimators for β_0 and β_1 .

Now, the idea of the Ordinary Least Squares Method for Estimating the Parameters β_0, β_1 is the following: Find

$$(\hat{\beta}_0, \hat{\beta}_1) = \underset{(\beta_0, \beta_1) \in \mathbb{R}^2}{\operatorname{argmin}} \sum_{k=1}^n \left(Y_k - \beta_0 - \beta_1 \cdot X_k \right)^2.$$

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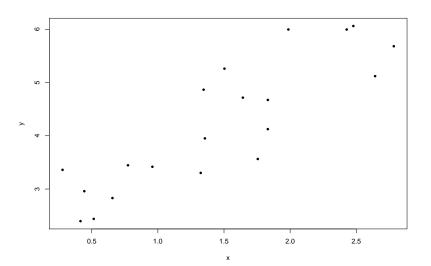
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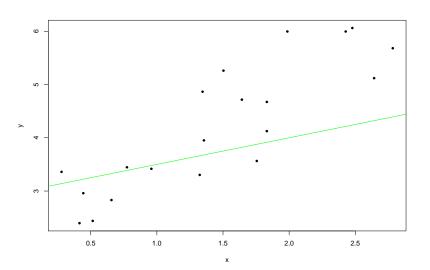
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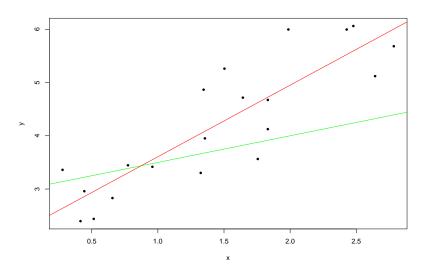
Here, the solution $(\hat{\beta}_0, \hat{\beta}_1)$ will be a pair of r.v.s, since Y_k -s are r.v.s. So we will obtain *Estimators* for β_0 and β_1 . If we will have an Observation (y_k, x_k) , k = 1, ..., n, then we will solve

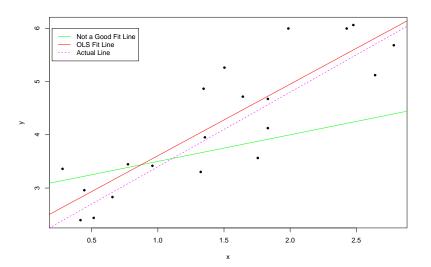
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Calculation of the best fit line is easy: we just define

$$\varphi(\beta_0,\beta_1) = \sum_{k=1}^n \left(Y_k - \beta_0 - \beta_1 \cdot X_k \right)^2, \qquad (\beta_0,\beta_1) \in \mathbb{R}^2,$$

and using our Calc 3 knowledge, find the Minimum Point of φ by solving the System

$$\begin{cases} \frac{\partial \varphi}{\partial \beta_0} = 0 \\ \frac{\partial \varphi}{\partial \beta_1} = 0 \end{cases}$$

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and

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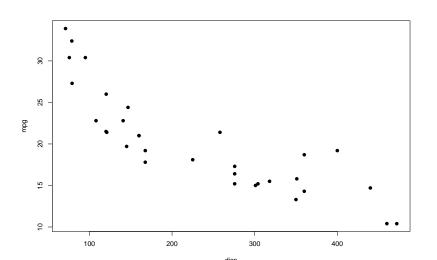
Note: So the cor(X, Y) is not the Slope of the Regression Line, but

$$cor(X, Y) \cdot \frac{sd(Y)}{sd(X)}$$

is. Recall our Descriptive Statistics part!

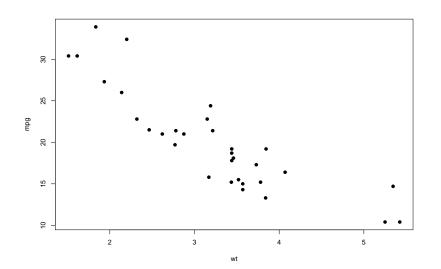
Example: We will use the mtcars Dataset from **R**:

```
plot(mpg ~ disp, data = mtcars, pch = 19)
```



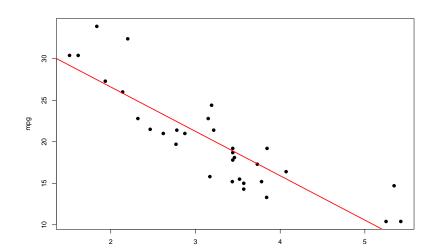
Example, Cont'd

```
plot(mpg ~ wt, data = mtcars, pch = 19)
```



Example, Cont'd

```
plot(mpg ~ wt, data = mtcars, pch = 19)
model <- lm(mpg ~ wt, data = mtcars)
abline(model, col = "red", lwd = 2)</pre>
```



```
Example, Cont'd

model <- lm(mpg ~ wt, data = mtcars)
summary(model)

##

## Call:
## lm(formula = mpg ~ wt, data = mtcars)
##

## Residuals:
## Min 1Q Median 3Q Max</pre>
```

(Intercept) 37.2851 1.8776 19.858 < 2e-16 ***

Estimate Std. Error t value Pr(>|t|)

-5.3445 0.5591 -9.559 1.29e-10 ***

Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' '

Residual standard error: 3.046 on 30 degrees of freedom
Multiple R-squared: 0.7528, Adjusted R-squared: 0.7446
F-statistic: 91.38 on 1 and 30 DF, p-value: 1.294e-10

-4.5432 -2.3647 -0.1252 1.4096 6.8727

##

##

##

wt.

Coefficients:

Example, Cont'd

Now, we predict the value of mpg for a new values of a wt Variable:

```
pred <- predict(model, data.frame(wt=4.7))
pred</pre>
```

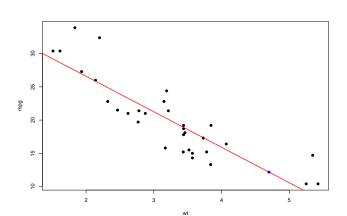
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## 12.16611
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pred</pre>
```

```
## 1
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```



```
y <- runif(100, 2, 10)
z <- 2.7 - 1.7*x + 13.5*y + rnorm(100)
head(x)

## [1] 0.5271618 -2.0473435 0.4363159 -0.1708086 -0.7298618 -1
head(y)

## [1] 3.977900 9.480366 2.867116 6.922705 7.797675 7.675757
head(z)

## [1] 55.39792 133.41440 42.64337 95.71013 109.41316 108.627
```

 $x \leftarrow rnorm(100, mean = -1, sd = 1)$

Call:

##

```
mod1 \leftarrow lm(z \sim x); summary(mod1)
```

$lm(formula = z \sim x)$

```
## Residuals:
      Min 1Q Median 3Q
##
                                  Max
## -58.755 -27.454 2.906 28.465 50.215
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 84.135 4.202 20.022 <2e-16 ***
              -3.130 2.917 -1.073 0.286
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' '
##
## Residual standard error: 31.23 on 98 degrees of freedom
## Multiple R-squared: 0.01161, Adjusted R-squared: 0.00152
```

F-statistic: 1.152 on 1 and 98 DF, p-value: 0.2859

Example, Cont'd

Residuals:

Call:

##

 $mod2 \leftarrow lm(z \sim x + y); summary(mod2)$

$lm(formula = z \sim x + y)$

```
10 Median
                           30
##
      Min
                                    Max
## -2.25835 -0.69422 -0.04329 0.72539 2.12851
##
## Coefficients:
##
            Estimate Std. Error t value Pr(>|t|)
## (Intercept) 2.92180 0.29018 10.07 <2e-16 ***
## x -1.50675 0.09194 -16.39 <2e-16 ***
          ## v
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' '
##
## Residual standard error: 0.9827 on 97 degrees of freedom
## Multiple R-squared: 0.999, Adjusted R-squared: 0.999
## F-statistic: 5.001e+04 on 2 and 97 DF, p-value: < 2.2e-16
```

Properties of the Estimators: \hat{eta}_0 and \hat{eta}_1

Here we assume that $\varepsilon \sim \mathcal{N}(0, \sigma^2)$.

Properties of the Estimators: $\hat{\beta}_0$ and $\hat{\beta}_1$

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Fact 1: Estimators $\hat{\beta}_0$ and $\hat{\beta}_1$ are UnBiased:

$$\mathbb{E}(\hat{\beta}_0) = \beta_0, \qquad Var(\hat{\beta}_0) = \frac{\sigma^2}{n} \cdot \frac{\sum_{k=1}^n X_k^2}{\sum_{k=1}^n (X_k - \overline{X})^2}$$

$$\mathbb{E}(\hat{\beta}_1) = \beta_1, \qquad Var(\hat{\beta}_1) = \frac{\sigma^2}{\sum_{k=1}^n (X_k - \overline{X})^2}$$

Properties of the Estimators: $\widehat{\sigma^2}$

Fact 2: Assume σ^2 is unknown.

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is an UnBiased Estimator for σ^2 , and

$$\widehat{\sigma^2} = \frac{1}{n} \cdot \sum_{k=1}^{n} (\hat{\varepsilon}_k)^2$$

is the MLE for σ^2 .

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is the MLE for σ^2 . Here

$$\hat{\varepsilon}_k = Y_k - \hat{\beta}_0 - \hat{\beta}_1 \cdot X_k$$

is the k-th **residual**.

Fsyo!

Thank You!