# STA9715 - Test 3 - Formula Sheet

#### **Advanced Inequalities**

- Chernoff Gaussian: if  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $\mathbb{P}(|X \mu| > t) \leq 2e^{-t^2/2\sigma^2}$  and  $\mathbb{P}(X > \mu + t) \leq e^{-t^2/2\sigma^2}$
- Chernoff Bounded: if X takes values in the range [a,b] with mean  $\mu$ , then  $\mathbb{P}(|X-\mu|>t)\leq 2e^{-2t^2/(b-a)^2}$  and  $\mathbb{P}(X>\mu+t)\leq e^{-2t^2/(b-a)^2}$

## **Moment Generating Functions**

- Moment Generating Function:  $\mathbb{M}_X(t) = \mathbb{E}[e^{tX}]$
- MGF to Moments:  $\mathbb{E}[X^k] = \mathbb{M}_X^{(k)}(0)$
- MGF of linear transforms:  $\mathbb{M}_{aX+b} = \mathbb{E}[e^{t(aX+b)}] = e^{tb}\mathbb{M}_X(ta)$
- MGF of sums of independent RVs:  $\mathbb{M}_{X+Y}(t) = \mathbb{M}_X(t)\mathbb{M}_Y(t)$
- If  $\mathbb{M}_X(t) = \mathbb{M}_Y(t)$ , then X and Y have the same distribution.

### Limit Theory

- Convergence in Probability:  $X_n \xrightarrow{P} X_*$  means  $\mathbb{P}(|X_n X_*| > \epsilon) \xrightarrow{n \to \infty} 0$
- Convergence in Distribution:  $X_n \xrightarrow{d} X_*$  means  $\mathbb{E}[f(X_n)] \xrightarrow{n \to \infty} \mathbb{E}[f(X_*)]$  for 'reasonable' functions  $f(\cdot)$ .
- Law of Large Numbers: if  $X_1, X_2, \ldots$  are IID random variables with mean  $\mu$  and finite variance,  $\overline{X}_n \stackrel{P}{\to} \mu$  for  $\overline{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ .
- Central Limit Theorem: if  $X_1, X_2, \ldots$  are IID random variables with mean  $\mu$  and finite variance  $\sigma^2$ , then  $\sqrt{n} \left( \frac{\overline{X}_n \mu}{\sigma} \right) \stackrel{d}{\to} \mathcal{N}(0, 1)$ . More usefully:  $\overline{X}_n \stackrel{d}{\to} \mathcal{N}(\mu, \sigma^2/n)$
- Delta Method: if  $X_n \stackrel{d}{\to} \mathcal{N}(\mu, \sigma^2)$ , then  $g(X_n) \stackrel{d}{\to} \mathcal{N}(g(\mu), \sigma^2 g'(\mu)^2)$  for any differentiable  $g(\cdot)$
- Glivenko-Cantelli Theorem ('Fundamental Theorem of Statistics'): the sample CDF  $\hat{F}_n(\cdot)$  converges to the true CDF  $F_X(\cdot)$  at all points where  $F_X(\cdot)$  is continuous
- Massart's Inequality:  $\mathbb{P}(\max_x |\hat{F}_n(x) F_X(x)| > \epsilon) \le 2e^{-2n\epsilon^2}$  for any distribution and any  $\epsilon > 0$

### **Key Statistical Distributions**

- Gaussian / Normal See above. Standardizing, CLT, Delta Method.
- $\chi_k^2$  sum of squares of k IID Standard Normals. Arises from 'goodness of fit' type statistics (e.g., SSE in OLS)
- $t_k$  (Student's) t distribution with k degrees of freedom.  $t_k \stackrel{d}{=} Z/\sqrt{\chi_k^2/k}$  where  $Z \perp \!\!\! \perp \chi_k^2$ .  $t_1$  is a Cauchy;  $t_\infty$  is a standard normal. Can replace Z with other normal distributions. Arises in testing with unknown variance.
- $\chi_2^2$  is an Expo(1/2) distribution with mean 2
- Gamma distribution: sum of n exponential distributions with mean  $1/\theta$  is  $\Gamma(n,\theta)$  distributed.  $\Gamma(n/2,2) \xrightarrow{d} \chi_n^2$
- Beta distribution = Gamma ratio.  $X \sim \Gamma(\alpha, \theta), Y \sim \Gamma(\beta, \theta) \implies X/(X+Y) \sim B(\alpha, \beta)$ . Support on [0, 1]
- F distribution:  $\frac{\chi_{k_1}^2/k_1}{\chi_{k_2}^2/k_2}$

Name	Parameters	Density	Mean	Variance	MGF
Standard Normal	None	$\phi(z) = e^{-z^2/2} / \sqrt{2\pi}$	0	1	$e^{t^2/2}$
Normal	Mean $\mu$ , StdDev $\sigma$	$e^{-(x-\mu)^2/2\sigma^2}/\sqrt{2\pi\sigma^2}$	$\mu$	$\sigma^2$	$e^{\mu t + \sigma^2 t^2/2}$
$\chi^2$	Degrees of freedom $k$	$\propto x^{k/2-1}e^{-x/2}$	k	2k	$(1-2t)^{-k/2}$
Standard Student's $t$	Degrees of freedom $k$		0	k/(k-2)	NA
Gamma	Shape $k$ , Scale $\theta$	$\propto x^{k-1}e^{-x/2}$	$k\theta$	$k heta^2$	$(1-\theta t)^{-k}$
Beta	Shapes $\alpha, \beta$	$\propto x^{\alpha-1}(1-x)^{\beta-1}$	$\alpha/(\alpha+\beta)$	$(\alpha\beta)(\alpha+\beta)^{-2}(\alpha+\beta+1)^{-1}$	Hard
F	Deg. Freedom $k_1, k_2$	Hard	$k_2/(k_2-2)$	Hard	NA

Sampling (Probability Integral Transform): If X has CDF  $F_X$ ,  $F_X^{-1}(U) \stackrel{d}{X}$  for  $U \sim \mathcal{U}([0,1])$ 

Sampling (Box-Mueller): Let  $R^2 \sim \chi^2_2$  and  $\Theta \sim \mathcal{U}([0,2\pi])$ ; then  $X = R\cos\Theta, Y = R\sin\Theta$  are independent  $Z_1, Z_2$