

# STA9715 - Test 3 - Formula Sheet

## Advanced Inequalities

- Chernoff - Gaussian: if  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $\mathbb{P}(|X - \mu| > t) \leq 2e^{-t^2/2\sigma^2}$  and  $\mathbb{P}(X > \mu + t) \leq e^{-t^2/2\sigma^2}$
- Chernoff - Bounded: if  $X$  takes values in the range  $[a, b]$  with mean  $\mu$ , then  $\mathbb{P}(|X - \mu| > t) \leq 2e^{-2t^2/(b-a)^2}$  and  $\mathbb{P}(X > \mu + t) \leq e^{-2t^2/(b-a)^2}$

## Moment Generating Functions

- Moment Generating Function:  $\mathbb{M}_X(t) = \mathbb{E}[e^{tX}]$
- MGF to Moments:  $\mathbb{E}[X^k] = \mathbb{M}_X^{(k)}(0)$
- MGF of linear transforms:  $\mathbb{M}_{aX+b} = \mathbb{E}[e^{t(aX+b)}] = e^{tb}\mathbb{M}_X(ta)$
- MGF of sums of independent RVs:  $\mathbb{M}_{X+Y}(t) = \mathbb{M}_X(t)\mathbb{M}_Y(t)$
- If  $\mathbb{M}_X(t) = \mathbb{M}_Y(t)$ , then  $X$  and  $Y$  have the same distribution.

## Limit Theory

- Convergence in Probability:  $X_n \xrightarrow{P} X_*$  means  $\mathbb{P}(|X_n - X_*| > \epsilon) \xrightarrow{n \rightarrow \infty} 0$
- Convergence in Distribution:  $X_n \xrightarrow{d} X_*$  means  $\mathbb{E}[f(X_n)] \xrightarrow{n \rightarrow \infty} \mathbb{E}[f(X_*)]$  for ‘reasonable’ functions  $f(\cdot)$ .
- Law of Large Numbers: if  $X_1, X_2, \dots$  are IID random variables with mean  $\mu$  and finite variance,  $\bar{X}_n \xrightarrow{P} \mu$  for  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ .
- Central Limit Theorem: if  $X_1, X_2, \dots$  are IID random variables with mean  $\mu$  and finite variance  $\sigma^2$ , then  $\sqrt{n} \left( \frac{\bar{X}_n - \mu}{\sigma} \right) \xrightarrow{d} \mathcal{N}(0, 1)$ . More usefully:  $\bar{X}_n \xrightarrow{d} \mathcal{N}(\mu, \sigma^2/n)$
- Delta Method: if  $X_n \xrightarrow{d} \mathcal{N}(\mu, \sigma^2)$ , then  $g(X_n) \xrightarrow{d} \mathcal{N}(g(\mu), \sigma^2 g'(\mu)^2)$  for any differentiable  $g(\cdot)$
- Glivenko-Cantelli Theorem (‘Fundamental Theorem of Statistics’): the sample CDF  $\hat{F}_n(\cdot)$  converges to the true CDF  $F_X(\cdot)$  at all points where  $F_X(\cdot)$  is continuous
- Massart’s Inequality:  $\mathbb{P}(\max_x |\hat{F}_n(x) - F_X(x)| > \epsilon) \leq 2e^{-2n\epsilon^2}$  for any distribution and any  $\epsilon > 0$

## Key Statistical Distributions

- Gaussian / Normal - See above. Standardizing, CLT, Delta Method.
- $\chi_k^2$  sum of squares of  $k$  IID Standard Normals. Arises from ‘goodness of fit’ type statistics (*e.g.*, SSE in OLS)
- $t_k$  - (Student’s)  $t$  distribution with  $k$  degrees of freedom.  $t_k \stackrel{d}{=} Z/\sqrt{\chi_k^2/k}$  where  $Z \perp \chi_k^2$ .  $t_1$  is a Cauchy;  $t_\infty$  is a standard normal. Can replace  $Z$  with other normal distributions. Arises in testing with unknown variance.
- $\chi_2^2$  is an Expo(1/2) distribution with mean 2
- Gamma distribution: sum of  $n$  exponential distributions with mean  $1/\theta$  is  $\Gamma(n, \theta)$  distributed.  $\Gamma(n/2, 2) \xrightarrow{d} \chi_n^2$
- Beta distribution = Gamma ratio.  $X \sim \Gamma(\alpha, \theta), Y \sim \Gamma(\beta, \theta) \implies X/(X+Y) \sim B(\alpha, \beta)$ . Support on  $[0, 1]$
- $F$  distribution:  $\frac{\chi_{k_1}^2/k_1}{\chi_{k_2}^2/k_2}$

Name	Parameters	Density	Mean	Variance	MGF
Standard Normal	None	$\phi(z) = e^{-z^2/2}/\sqrt{2\pi}$	0	1	$e^{t^2/2}$
Normal	Mean $\mu$ , StdDev $\sigma$	$e^{-(x-\mu)^2/2\sigma^2}/\sqrt{2\pi\sigma^2}$	$\mu$	$\sigma^2$	$e^{\mu t + \sigma^2 t^2/2}$
$\chi^2$	Degrees of freedom $k$	$\propto x^{k/2-1}e^{-x/2}$	$k$	$2k$	$(1-2t)^{-k/2}$
Standard Student’s $t$	Degrees of freedom $k$	$\propto (1+x^2/k)^{-(k+1)/2}$	0	$k/(k-2)$	NA
Gamma	Shape $k$ , Scale $\theta$	$\propto x^{k-1}e^{-x/\theta}$	$k\theta$	$k\theta^2$	$(1-\theta t)^{-k}$
Beta	Shapes $\alpha, \beta$	$\propto x^{\alpha-1}(1-x)^{\beta-1}$	$\alpha/(\alpha+\beta)$	$(\alpha\beta)(\alpha+\beta)^{-2}(\alpha+\beta+1)^{-1}$	Hard
$F$	Deg. Freedom $k_1, k_2$	Hard	$k_2/(k_2-2)$	Hard	NA

Sampling (Probability Integral Transform): If  $X$  has CDF  $F_X$ ,  $F_X^{-1}(U) \stackrel{d}{=} X$  for  $U \sim \mathcal{U}([0, 1])$

Sampling (Box-Mueller): Let  $R^2 \sim \chi_2^2$  and  $\Theta \sim \mathcal{U}([0, 2\pi])$ ; then  $X = R \cos \Theta, Y = R \sin \Theta$  are independent  $Z_1, Z_2$