

# **Computational and Statistical Methodology for Highly-Structured Data**

Ph.D. Thesis Defense: 2020-09-10

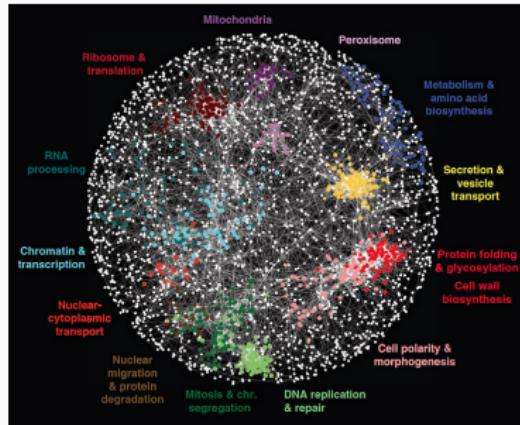
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Michael Weylandt

Slides Available Online at <https://tinyurl.com/WeylandtDefense>

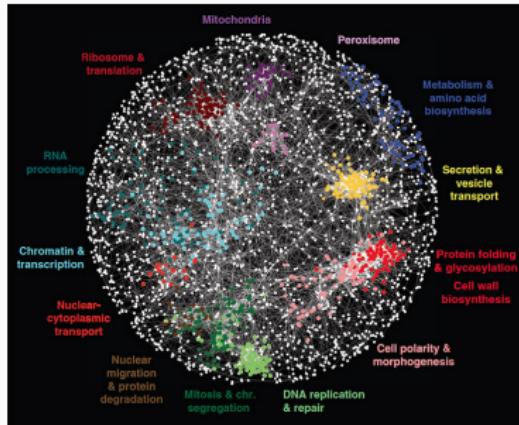
Rice University, Department of Statistics

# Big-Data



“Big Data” enables “Big Models”

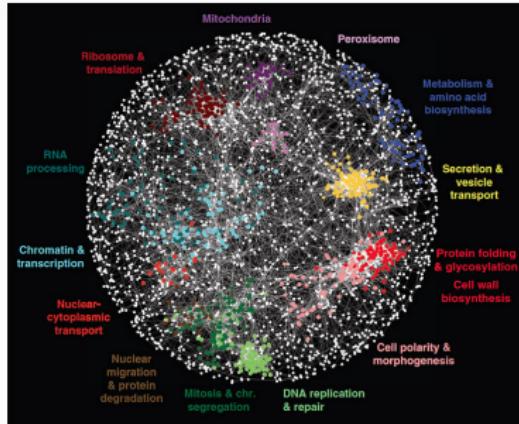
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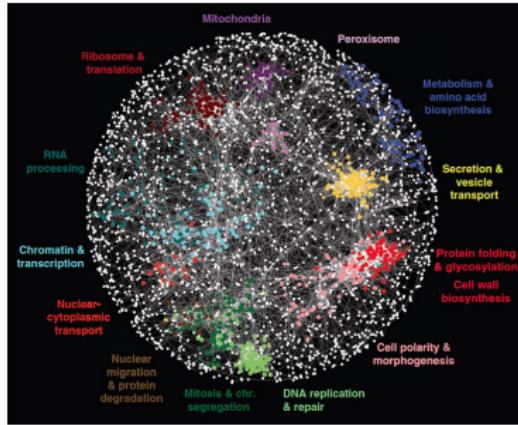


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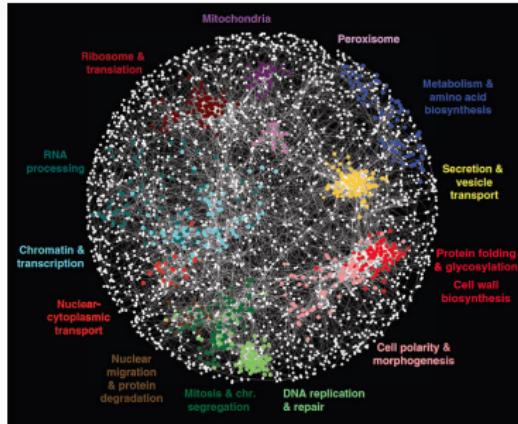
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Highly-structured data requires flexible but powerful models to reflect  
and capture dependencies in data

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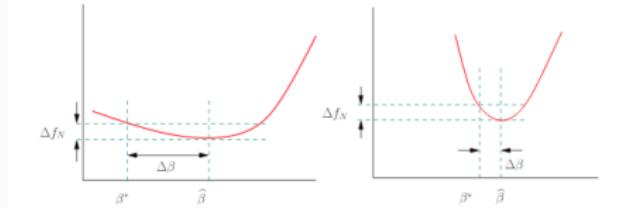
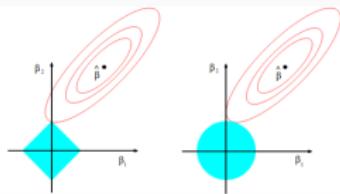
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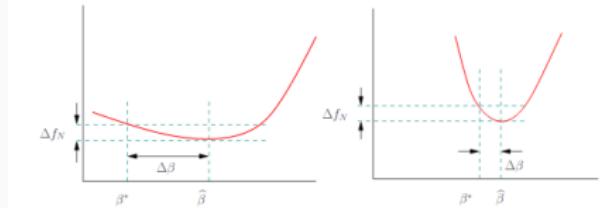
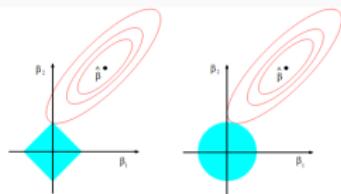
Big data allows us to fit such models

# Convex Revolution in Statistical Machine Learning



Biggest advance in 21<sup>st</sup> c. statistics – convex analysis and optimization:

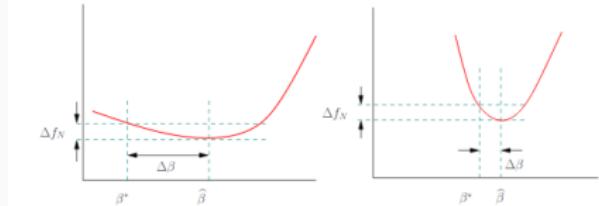
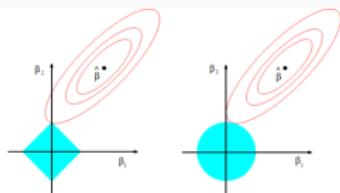
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- Development of novel regularized estimation schemes

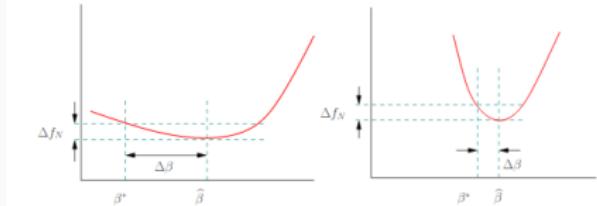
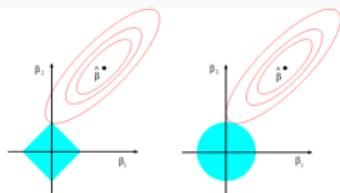
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- Algorithms that efficiently scale to enormous data sets

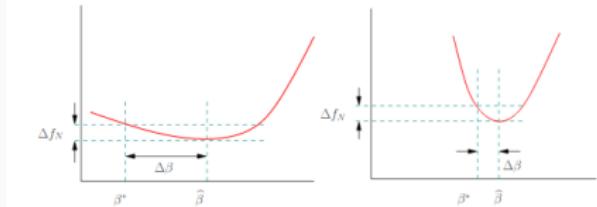
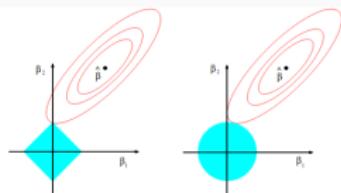
# Convex Revolution in Statistical Machine Learning



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- Algorithms that efficiently scale to enormous data sets
- Theoretical advances based on powerful convex analysis

# Convex Revolution in Statistical Machine Learning



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**Develop Methodology for Big Highly-Structured Data  
Built on Powerful Convex Analysis and Optimization**

# Agenda

Splitting Methods for Clustering

Multi-Rank Regularized PCA

Multivariate Models for Gas Markets

Complex Convex Analysis

Conclusion & Discussion

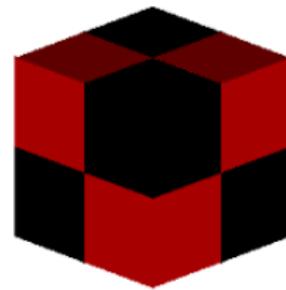
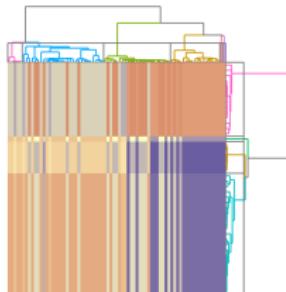
## **Splitting Methods for Clustering**

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# Tensor Co-Clustering

Co-Clustering:

- Simultaneous clustering along all faces of a tensor
- Discover “checkerboard” patterns in data
- “Cluster Heatmap” for 2-tensors
- Manifold learning for  $K$ -tensors (Mishne *et al.*, 2019)



## Convex Bi-Clustering

Convex formulation of co-clustering: Chi *et al.* (2017) and Chi *et al.* (2018)

- Frobenius norm loss  $\implies$  approximate observed data
- Convex fusion penalty  $\implies$  encourages clustering

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Matrix (2-tensor) case:

$$\hat{\mathbf{U}} = \arg \min_{\mathbf{U} \in \mathbb{R}^{n \times p}} \frac{1}{2} \|\mathbf{X} - \mathbf{U}\|_F^2 + \lambda \left( \sum_{\substack{i,j=1 \\ i \neq j}}^n w_{ij} \|\mathbf{U}_{i \cdot} - \mathbf{U}_{j \cdot}\|_q + \sum_{\substack{k,l=1 \\ k \neq l}}^p \tilde{w}_{kl} \|\mathbf{U}_{\cdot k} - \mathbf{U}_{\cdot l}\|_q \right)$$

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Simultaneous clustering of **rows** and **columns**:

- Rows are clustered together if  $\hat{\mathbf{U}}_{i \cdot} = \hat{\mathbf{U}}_{j \cdot}$
- Columns are clustered together if  $\hat{\mathbf{U}}_{\cdot k} = \hat{\mathbf{U}}_{\cdot l}$
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$\lambda$  controls the number of co-clusters smoothly

# Splitting Methods for Convex Bi-Clustering

Simplified form:

$$\hat{\mathbf{U}} = \arg \min_{\mathbf{U} \in \mathbb{R}^{n \times p}} \frac{1}{2} \|\mathbf{X} - \mathbf{U}\|_F^2 + \lambda \left( \underbrace{\|\mathbf{D}_{\text{row}} \mathbf{U}\|_{\text{row},q}}_{P_{\text{row}}(\mathbf{D}_{\text{row}} \mathbf{U})} + \underbrace{\|\mathbf{U} \mathbf{D}_{\text{col}}\|_{\text{col},q}}_{P_{\text{col}}(\mathbf{U} \mathbf{D}_{\text{col}})} \right)$$

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Current state of the art:

COBRA - Dykstra-Like Proximal Algorithm

(Bauschke and Combettes, 2008; Chi and Lange, 2015)

- Alternating row- and column-wise convex clustering

Convex clustering subproblems are still slow, so COBRA doesn't scale

# Splitting Methods for Convex Bi-Clustering

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Can we develop a fast splitting approach?

# Splitting Methods for Convex Bi-Clustering

$$\hat{\mathbf{U}} = \arg \min_{\mathbf{U} \in \mathbb{R}^{n \times p}} \frac{1}{2} \|\mathbf{X} - \mathbf{U}\|_F^2 + \lambda \left( \underbrace{\|D_{\text{row}} \mathbf{U}\|_{\text{row},q}}_{P_{\text{row}}(D_{\text{row}} \mathbf{U})} + \underbrace{\|\mathbf{U} D_{\text{col}}\|_{\text{col},q}}_{P_{\text{col}}(\mathbf{U} D_{\text{col}})} \right)$$

Can we develop a fast splitting approach?

Davis and Yin (2017) three-block ADMM:

1.  $\mathbf{U}^{(k+1)} = \mathbf{X} - D_{\text{row}}^T Z_{\text{row}}^{(k)} - Z_{\text{col}}^{(k)} D_{\text{col}}^T$
- 2(a).  $V_{\text{row}}^{(k+1)} = \text{prox}_{\lambda/\rho P_{\text{row}}}(\cdot)(D_{\text{row}} \mathbf{U}^{(k+1)} + Z_{\text{row}}^{(k)})$
- 2(b).  $V_{\text{col}}^{(k+1)} = \text{prox}_{\lambda/\rho P_{\text{col}}}(\cdot)(\mathbf{U}^{(k+1)} D_{\text{col}} + Z_{\text{col}}^{(k)})$
- 3(a).  $Z_{\text{row}}^{(k+1)} = Z_{\text{row}}^{(k)} + \rho(D_{\text{row}} \mathbf{U}^{(k+1)} - V_{\text{row}}^{(k+1)})$
- 3(b).  $Z_{\text{col}}^{(k+1)} = Z_{\text{col}}^{(k)} + \rho(\mathbf{U}^{(k+1)} D_{\text{col}} - V_{\text{col}}^{(k+1)})$

Equivalent to AMA and to prox-gradient on the dual - very slow!

(Tseng, 1991)

# Splitting Methods for Convex Bi-Clustering

Why not apply ADMM directly?

# Splitting Methods for Convex Bi-Clustering

Why not apply ADMM directly? Lifted problem:

$$\arg \min_{\substack{\mathbf{U} \in \mathbb{R}^{n \times p} \\ (\mathbf{V}_{\text{row}}, \mathbf{V}_{\text{col}}) \in \mathbb{R}^{\#\text{row} \times p} \times \mathbb{R}^{n \times \#\text{col}}} \frac{1}{2} \|\mathbf{X} - \mathbf{U}\|_F^2 + \lambda (P_{\text{row}}(\mathbf{V}_{\text{row}}) + P_{\text{col}}(\mathbf{V}_{\text{col}}))$$

subject to

$$\mathcal{L}_1 \mathbf{U} - (\mathbf{V}_{\text{row}}, \mathbf{V}_{\text{col}}) = 0 \text{ where } \mathcal{L}_1 \mathbf{U} = (\mathbf{D}_{\text{row}} \mathbf{U}, \mathbf{U} \mathbf{D}_{\text{col}})$$

Isomorphic, but much better computationally!

$\mathbf{V}, \mathbf{Z}$  updates as before (separable penalties + Cartesian structure)  
 $\mathbf{U}$  more complicated

# Splitting Methods for Convex Bi-Clustering

$\mathbf{U}$ -subproblem:

$$\arg \min_{\mathbf{U} \in \mathbb{R}^{n \times p}} \frac{1}{2} \|\mathbf{X} - \mathbf{U}\|_F^2 + \frac{\rho}{2} \|\mathbf{D}_{\text{row}} \mathbf{U} - \mathbf{V}_{\text{row}}^{(k)} + \rho^{-1} \mathbf{Z}_{\text{row}}^{(k)}\|_F^2 + \frac{\rho}{2} \|\mathbf{U} \mathbf{D}_{\text{col}} - \mathbf{V}_{\text{col}}^{(k)} + \rho^{-1} \mathbf{Z}_{\text{col}}^{(k)}\|_F^2$$

Stationary condition - Sylvester equation:

$$\mathbf{X} + \mathbf{D}_{\text{row}}^T (\mathbf{V}_{\text{row}}^{(k)} - \rho^{-1} \mathbf{Z}_{\text{row}}^{(k)}) + (\mathbf{V}_{\text{col}}^{(k)} - \rho^{-1} \mathbf{Z}_{\text{col}}^{(k)}) \mathbf{D}_{\text{col}}^T = \mathbf{U} + \rho \mathbf{D}_{\text{row}}^T \mathbf{D}_{\text{row}} \mathbf{U} + \rho \mathbf{U} \mathbf{D}_{\text{col}} \mathbf{D}_{\text{col}}^T$$

# Splitting Methods for Convex Bi-Clustering

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Alternative: add quadratic term to make  $\mathbf{U}$ -subproblem easier to solve  
(Deng and Yin, 2016)

$$\arg \min_{\mathbf{U} \in \mathbb{R}^{n \times p}} \dots + \alpha \|\mathbf{U}\|_F^2 - \rho \|\mathcal{L}_1 \mathbf{U}\|^2 \text{ where } \mathcal{L}_1 \mathbf{U} = (\mathbf{D}_{\text{row}} \mathbf{U}, \mathbf{U} \mathbf{D}_{\text{col}})$$

$$\begin{aligned} \mathbf{U}^{(k+1)} = & \left( \alpha \mathbf{U}^{(k)} + \mathbf{X} + \rho \mathbf{D}_{\text{row}}^T (\mathbf{V}^{(k)} - \rho^{-1} \mathbf{Z}_{\text{row}}^{(k)} - \mathbf{D}_{\text{row}} \mathbf{U}^{(k)}) \right. \\ & \left. + \rho (\mathbf{V}_{\text{col}}^{(k)} - \rho^{-1} \mathbf{Z}_{\text{col}}^{(k)} - \mathbf{U}^{(k)} \mathbf{D}_{\text{col}}) \mathbf{D}_{\text{col}}^T \right) / (1 + \alpha) \end{aligned}$$

# Results

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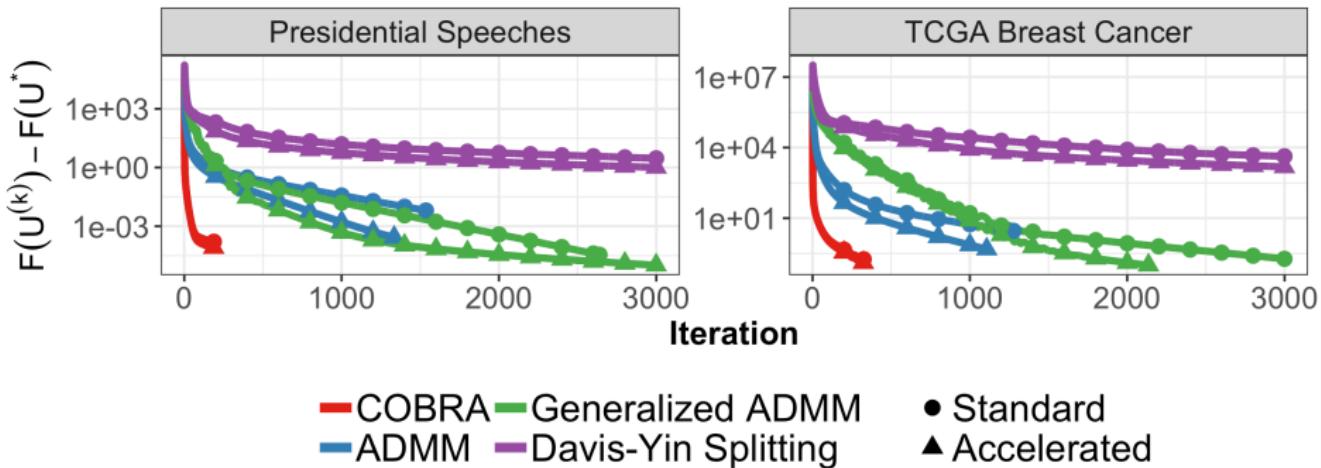
Compare:

- ADMM
- Generalized ADMM
- Davis-Yin Three-Block ADMM
- COBRA  $\implies$  alternating row- and column-clustering sub-problems

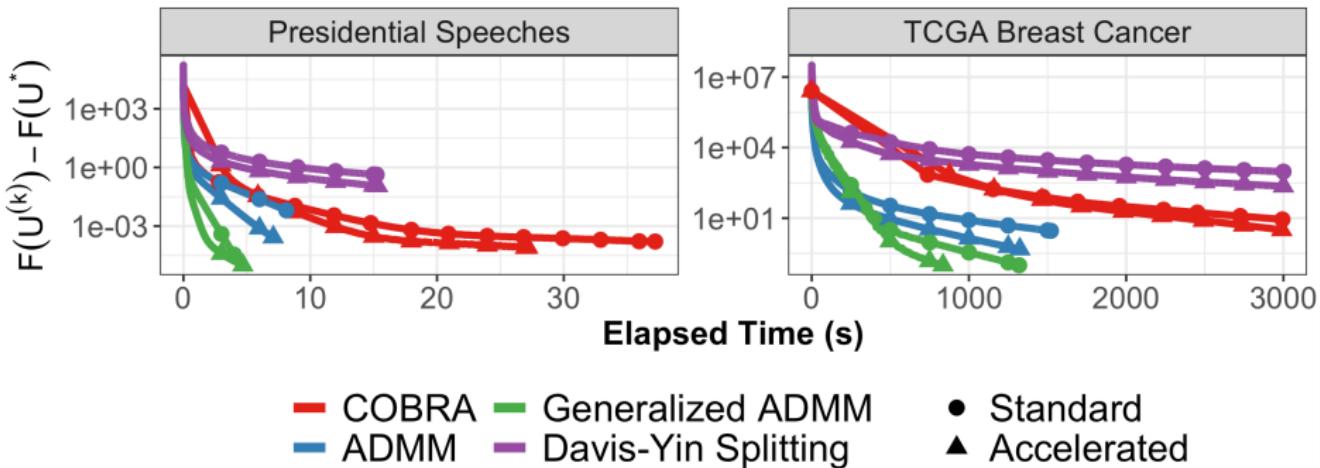
Data:

- Presidents  $\in \mathbb{R}^{44 \times 75}$
- TCGA Breast Cancer  $\in \mathbb{R}^{438 \times 353}$

## Results: Iteration Count



## Results: Elapsed Time



## Higher-Order Extensions

$$\hat{\mathcal{U}} = \arg \min_{\mathcal{U}} \frac{1}{2} \|\mathcal{X} - \mathcal{U}\|_F^2 + \lambda \sum_{j=1}^J \|\mathcal{U} \times_j \mathcal{D}_j\|_{j,q}$$

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Same “lifting” approach works for Generalized ADMM and Davis-Yin:

$$\mathcal{U}_{\text{Gen-ADMM}}^{(k+1)} = \frac{\alpha}{1+\alpha} \mathcal{U}^{(k)} + \frac{\mathcal{X}}{1+\alpha} + \frac{\rho}{1+\alpha} \sum_{j=1}^J (\mathcal{V}_j^{(k)} - \rho^{-1} \mathcal{Z}_j^{(k)} - \mathcal{U}^{(k)} \times_j \mathcal{D}_j) \times_j \mathcal{D}_j^T$$

$$\mathcal{U}_{\text{DY / AMA}}^{(k+1)} = \mathcal{X} - \sum_{j=1}^J \mathcal{Z}_j^{(k)} \times_j (\mathcal{D}_j)^T$$

$$\mathcal{V}_j^{(k+1)} = \underset{\lambda/\rho \|\cdot\|_{j,q}}{\text{prox}} \left( \mathcal{U}^{(k+1)} \times_j \mathcal{D}_j + \rho^{-1} \mathcal{Z}_j^{(k)} \right) \quad \forall j \in \{1, \dots, J\}$$

$$\mathcal{Z}_j^{(k+1)} = \mathcal{Z}_j^{(k)} + \rho(\mathcal{U}^{(k+1)} \times_j \mathcal{D}_j - \mathcal{V}_j^{(k+1)}) \quad \forall j \in \{1, \dots, J\}$$

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Standard ADMM  $\implies$  tensor Sylvester equation:

$$\mathcal{X} + \rho \sum_{j=1}^J (\mathcal{V}_j^{(k)} - \rho^{-1} \mathcal{Z}_j^{(k)}) \times_j \mathcal{D}_j^T = \mathcal{U}_{\text{ADMM}} + \rho \sum_{j=1}^J \mathcal{U}_{\text{ADMM}} \times_j \mathcal{D}_j \times_j \mathcal{D}_j^T.$$

# Implications

Efficient Convex Clustering Algorithm

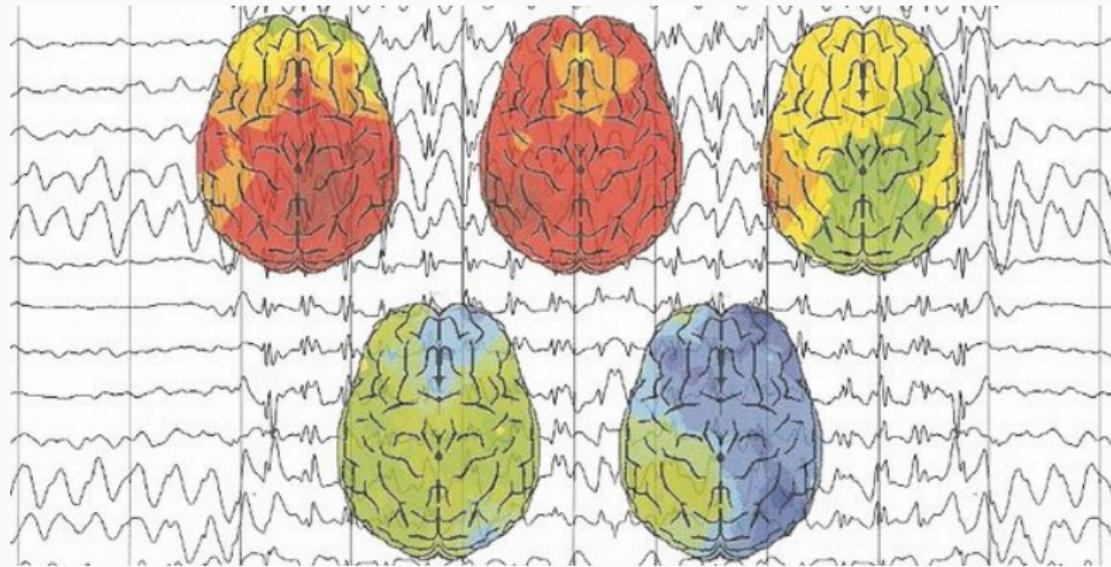
Embed in More Complex Schemes (e.g., 2D trend filtering)

“Lifting” Trick Useful for Multiply-Regularized Problems

## Multi-Rank Regularized PCA

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# Motivation



Principal Components Analysis:

- Exploratory Data Analysis
- Pattern Recognition
- Dimension Reduction
- Data Visualization

# Regularization in PCA

Low-rank model for PCA - estimate low-rank mean of  $\mathbf{X}$ :

$$\mathbf{X} = \mathbf{u}\mathbf{v}^T + \mathbf{E} \text{ where } \mathbf{E} \stackrel{\text{IID}}{\sim} \mathcal{N}(0, \sigma^2)$$

$$\arg \min_{\mathbf{u}, \mathbf{v}, d} \|\mathbf{X} - d\mathbf{u}\mathbf{v}^T\|_F^2 \Leftrightarrow \arg \max_{\mathbf{u}, \mathbf{v}} \mathbf{u}^T \mathbf{X} \mathbf{v} \quad \text{subject to } \|\mathbf{u}\| = \|\mathbf{v}\| = 1$$

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- Identify patterns in rows and columns of  $\mathbf{X}$
- $\mathbf{u}, \mathbf{v}, d$  calculated using SVD of  $\mathbf{X}$

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Convergence is slow: RMT asymptotics ( $p/n \rightarrow c$ ) more relevant

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- $\mathbf{u}, \mathbf{v}, d$  calculated using SVD of  $\mathbf{X}$

PCA is consistent under standard ( $n \rightarrow \infty$ ) asymptotics (Anderson, 1963)

Convergence is slow: RMT asymptotics ( $p/n \rightarrow c$ ) more relevant

PCA in high-dimensions is inconsistent (Johnstone and Lu, 2009)

# Regularization in PCA

Low-rank model for PCA - estimate low-rank mean of  $\mathbf{X}$ :

$$\mathbf{X} = \mathbf{u}\mathbf{v}^T + \mathbf{E} \text{ where } \mathbf{E} \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$$

$$\arg \min_{\mathbf{u}, \mathbf{v}, d} \|\mathbf{X} - d\mathbf{u}\mathbf{v}^T\|_F^2 \Leftrightarrow \arg \max_{\mathbf{u}, \mathbf{v}} \mathbf{u}^T \mathbf{X} \mathbf{v} \quad \text{subject to } \|\mathbf{u}\| = \|\mathbf{v}\| = 1$$

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**Regularization Needed**

# Sparse and Functional PCA

Sparse and Functional PCA: Allen and W., (DSW 2019)

$$\arg \max_{\mathbf{u} \in \overline{\mathbb{B}}_{S_u}^n, \mathbf{v} \in \overline{\mathbb{B}}_{S_v}^p} \mathbf{u}^T \mathbf{X} \mathbf{v} - \lambda_u P_u(\mathbf{u}) - \lambda_v P_v(\mathbf{v})$$

where

$$\overline{\mathbb{B}}_{S_u}^n = \left\{ \mathbf{x} \in \mathbb{R}^n : \mathbf{x}^T \mathbf{S}_u \mathbf{x} = \mathbf{x}^T (\mathbf{I} + \alpha_u \boldsymbol{\Omega}_u) \mathbf{x} \leq 1 \right\}$$

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Sparse and Functional PCA:

- Smoothness in  $\mathbf{u}$  - structure  $\mathbf{S}_u$  + strength  $\alpha_u$
- Sparsity in  $\mathbf{u}$  - structure  $P_u$  + strength  $\lambda_u$
- Smoothness in  $\mathbf{v}$  - structure  $\mathbf{S}_v$  + strength  $\alpha_v$
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Well-posed and non-degenerate

## Computing SPCA

$$\text{SPCA: } \arg \max_{\mathbf{u} \in \overline{\mathbb{B}}_{S_u}^n, \mathbf{v} \in \overline{\mathbb{B}}_{S_v}^p} \mathbf{u}^T \mathbf{X} \mathbf{v} - \lambda_u P_u(\mathbf{u}) - \lambda_v P_v(\mathbf{v})$$

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Projection + (accelerated) proximal gradient to solve sub-problems

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## Theorem

1. The  $\mathbf{u}$ -update converges to a solution of

$$\arg \min_{\mathbf{u} \in \overline{\mathbb{B}}_{S_u}^n} \frac{1}{2} \|\mathbf{X} \mathbf{v} - \mathbf{u}\|_2^2 + \lambda_u P_u(\mathbf{u}) + \frac{\alpha_u}{2} \mathbf{u}^T \Omega_u \mathbf{u}$$

2.  $\mathbf{u}$ -update finds global optimum for fixed  $\mathbf{v}$
3. Converges to block-coordinate-wise global optima (Nash points)

Fast!

# Multi-Rank SPCA

---

Orthogonality of PCs: interpretation and statistical independence

Can we do the same for SPCA?

# Multi-Rank SFPCA

Orthogonality of PCs: interpretation and statistical independence

Can we do the same for SFPCA?

Multi-rank extension of SFPCA:

$$\text{MR-SFPCA: } \arg \max_{\mathbf{U} \in \mathcal{V}_{S_u}^{n \times k}, \mathbf{V} \in \mathcal{V}_{S_v}^{p \times k}} \text{Tr}(\mathbf{U}^T \mathbf{X} \mathbf{V}) - \lambda_u P_u(\mathbf{U}) - \lambda_v P_v(\mathbf{V})$$

where  $\mathcal{V}_{S_u}^{n \times k}$  is the  $k^{\text{th}}$  order generalized Stiefel manifold in  $\mathbb{R}^n$ :

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Generalized Stiefel manifold constraint  $\implies$  manifold optimization (Absil *et al.*, 2007)

As with R1-SFPCA, alternating (partial) maximization

# Manifold Proximal Gradient

Standard SFPCA  $\mathbf{u}$ -subproblem updates:

$$\mathbf{u} := \underset{\frac{\lambda_u}{L_u} P_u(\cdot)}{\text{prox}} \left( \mathbf{u} + L_{\mathbf{u}}^{-1} (\mathbf{X}\hat{\mathbf{v}} - \mathbf{S}_u \mathbf{u}) \right) \quad \hat{\mathbf{u}} := \begin{cases} \mathbf{u} & \|\mathbf{u}\|_{\mathcal{S}_u} \leq 1 \\ \mathbf{u}/\|\mathbf{u}\|_{\mathcal{S}_u} & \text{otherwise} \end{cases}$$

Proximal + projected gradient descent

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Proximal + projected gradient descent

Multi-Rank SFPCA  $\mathbf{U}$ -subproblem updates - Manifold Prox Gradient:  
(Chen et al., 2020a)

$$\begin{aligned} \hat{\mathbf{D}} &= \arg \min_{\mathbf{D} \in \mathbb{R}^{n \times k}} -\langle \mathbf{X}\hat{\mathbf{V}}, \mathbf{D} \rangle_F + \lambda_{\mathbf{U}} P_{\mathbf{U}}(\mathbf{U}^{(k)} + \mathbf{D}) \\ \text{s.t. } \mathbf{D}^T \mathbf{S}_{\mathbf{u}} \mathbf{U}^{(k)} + (\mathbf{U}^{(k)})^T \mathbf{S}_{\mathbf{u}} \mathbf{D} &= \mathbf{0} \\ \mathbf{U}^{(k+1)} &= \text{Retr}_{\mathbf{U}^{(k)}}(\eta \hat{\mathbf{D}}) \end{aligned}$$

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One step of each subproblem  $\implies$  convergence to stationary point:

- Constraint set smooth  $\implies$  Stationary points isolated
- Guaranteed descent at each iteration (Chen et al., 2020b)

# Manifold ADMM

Easier  $\mathbf{U}$ -updates from Manifold ADMM: (Kovnatsky *et al.*, 2016)

$$\mathbf{U}^{(k+1)} = \arg \min_{\mathbf{U} \in \mathcal{V}_{n \times k}^{S_u}} -\text{Tr}(\mathbf{U}^T \mathbf{X} \mathbf{V}) + \frac{\rho}{2} \|\mathbf{U} - \mathbf{W}^{(k)} + \mathbf{Z}^{(k)}\|_F^2$$

$$\begin{aligned}\mathbf{W}^{(k+1)} &= \arg \min_{\mathbf{W} \in \mathbb{R}^{n \times k}} \lambda_U P_U(\mathbf{W}) + \frac{\rho}{2} \|\mathbf{U}^{(k+1)} - \mathbf{W} + \mathbf{Z}^{(k)}\|_F^2 \\ &= \underset{\lambda_U / \rho P_U(\cdot)}{\text{prox}} \left( \mathbf{U}^{(k+1)} + \mathbf{Z}^{(k)} \right)\end{aligned}$$

$$\mathbf{Z}^{(k+1)} = \mathbf{Z}^{(k)} + \mathbf{U}^{(k+1)} - \mathbf{W}^{(k+1)}$$

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$$\begin{aligned}\mathbf{W}^{(k+1)} &= \arg \min_{\mathbf{W} \in \mathbb{R}^{n \times k}} \lambda_{\mathbf{U}} P_{\mathbf{U}}(\mathbf{W}) + \frac{\rho}{2} \|\mathbf{U}^{(k+1)} - \mathbf{W} + \mathbf{Z}^{(k)}\|_F^2 \\ &= \underset{\lambda_{\mathbf{U}}/\rho P_{\mathbf{U}}(\cdot)}{\text{prox}} \left( \mathbf{U}^{(k+1)} + \mathbf{Z}^{(k)} \right)\end{aligned}$$

$$\mathbf{Z}^{(k+1)} = \mathbf{Z}^{(k)} + \mathbf{U}^{(k+1)} - \mathbf{W}^{(k+1)}$$

First step is a *generalized unbalanced Procrustes problem* - analytical solution via SVD of  $\mathbf{S}_u^{-1/2} \mathbf{X} \hat{\mathbf{V}} + \rho \mathbf{S}_u^{1/2} (\mathbf{W}^{(k)} - \mathbf{Z}^{(k)})$

Typically converges quickly and to a good solution (no theory)

## Deflation Techniques

$$\text{Rank-1 SFPCA: } \arg \max_{\mathbf{u} \in \overline{\mathbb{B}}_{S_u}^n, \mathbf{v} \in \overline{\mathbb{B}}_{S_v}^p} \mathbf{u}^T \mathbf{X} \mathbf{v} - \lambda_u P_u(\mathbf{u}) - \lambda_v P_v(\mathbf{v})$$

How to get additional *nested* SFPCA components?

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How to get additional *nested* SFPCA components?

Deflation:

- Hotelling:  $\mathbf{X}_t^{\text{HD}} := \mathbf{X}_{t-1} - \mathbf{U}_t (\mathbf{U}_t^T \mathbf{U}_t)^{-1} \mathbf{U}_t^T \mathbf{X}_{t-1} \mathbf{V}_t (\mathbf{V}_t^T \mathbf{V}_t)^{-1} \mathbf{V}_t^T$
- Projection:  $\mathbf{X}_t^{\text{PD}} := (\mathbf{I}_n - \mathbf{U}_t (\mathbf{U}_t^T \mathbf{U}_t)^{-1} \mathbf{U}_t^T) \mathbf{X}_{t-1} (\mathbf{I}_p - \mathbf{V}_t (\mathbf{V}_t^T \mathbf{V}_t)^{-1} \mathbf{V}_t^T)$
- Schur Complement:  $\mathbf{X}_t^{\text{SD}} := \mathbf{X}_{t-1} - \mathbf{X}_{t-1} \mathbf{V}_t (\mathbf{U}_t^T \mathbf{X}_{t-1} \mathbf{V}_t)^{-1} \mathbf{U}_t^T \mathbf{X}_{t-1}$

HD doesn't fully remove signal and may re-introduce if estimated PCs non-orthogonal

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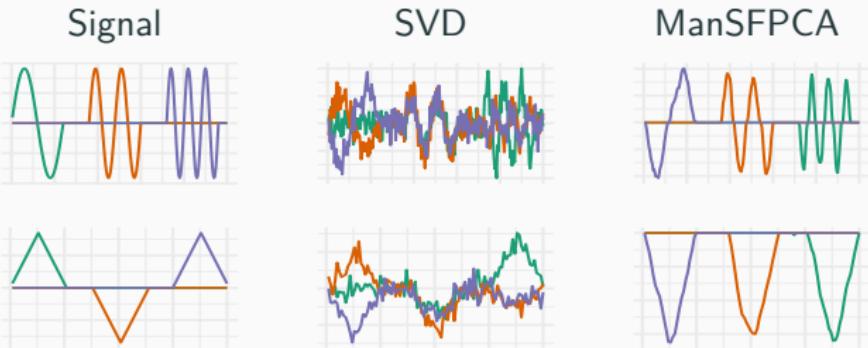
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HD doesn't fully remove signal and may re-introduce if estimated PCs non-orthogonal

| Method     | Two-Way 0-ing<br>$\mathbf{u}_t^T \mathbf{X}_t \mathbf{v}_t = 0$ | One-Way 0-ing<br>$\mathbf{u}_t^T \mathbf{X}_t, \mathbf{X}_t \mathbf{v}_t = 0$ | Subsequent 0-ing ( $\forall s \geq 0$ )<br>$\mathbf{u}_t^T \mathbf{X}_{t+s}, \mathbf{X}_{t+s} \mathbf{v}_t = 0$ | Robust to<br>Scale of $\mathbf{u}_t, \mathbf{v}_t$ |
|------------|---|---|---|--|
| Hotelling  | ✓   | ✗   | ✗   | ✗  |
| Projection | ✓   | ✓   | ✗   | ✗  |
| Schur      | ✓   | ✓   | ✓   | ✓  |

# Simulation: “On Model” Signal Recovery

Scenario 1:  $\mathbf{U}^*$  and  $\mathbf{V}^*$  Orthogonal - SNR  $\approx 1.2$



|           |              | HD      | PD      | SD            | ManSFPSCA     |
|-----------|--------------|---------|---------|---------------|---------------|
| CPVE      | PC1          | 15.92%  | 21.05%  | <b>21.87%</b> |               |
|           | PC2          | 22.21%  | 29.42%  | <b>30.59%</b> | <b>37.12%</b> |
|           | PC3          | 26.80%  | 35.57%  | <b>37.09%</b> |               |
| rSS-Error | $\mathbf{U}$ | 129.54% | 129.55% | 128.35%       | <b>68.66%</b> |
|           | $\mathbf{V}$ | 143.01% | 143.72% | 141.15%       | <b>36.98%</b> |

# Simulation: ‘Off-Model’ Signal Recovery

Scenario 2:  $\mathbf{U}^*$  and  $\mathbf{V}^*$  Not Orthogonal - SNR  $\approx 1.7$



|           |              | HD      | PD      | SD            | ManSFPCA      |
|-----------|--------------|---------|---------|---------------|---------------|
| CPVE      | PC1          | 8.85%   | 19.74%  | <b>29.80%</b> |               |
|           | PC2          | 13.03%  | 28.30%  | <b>39.87%</b> | <b>50.85%</b> |
|           | PC3          | 16.16%  | 34.22%  | <b>46.48%</b> |               |
| rSS-Error | $\mathbf{U}$ | 215.73% | 206.30% | 205.74%       | <b>97.77%</b> |
|           | $\mathbf{V}$ | 211.15% | 207.77% | 204.38%       | <b>78.26%</b> |

# Implications

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Principaled Approach to Multi-Rank PCA:

- Unifies **many** regularized PCA variants
- Extension to multiple PCs without loosing orthogonality
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Additional Multivariate Methods:

- CCA, LDA, PLS, *etc.* all SVD - can all be similarly treated

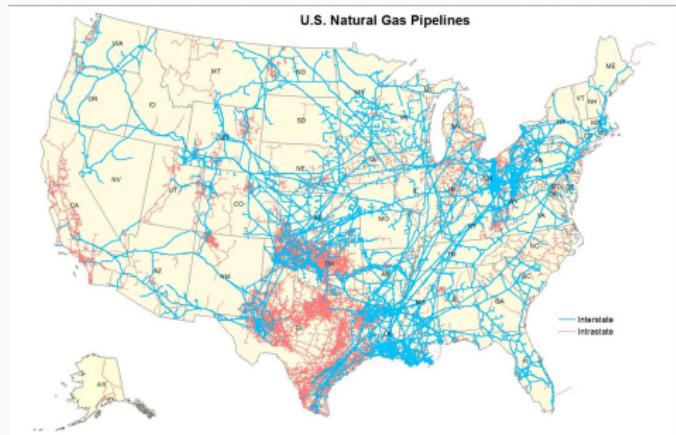
# Multivariate Models for Gas Markets

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# Natural Gas Markets

## LNG Markets:

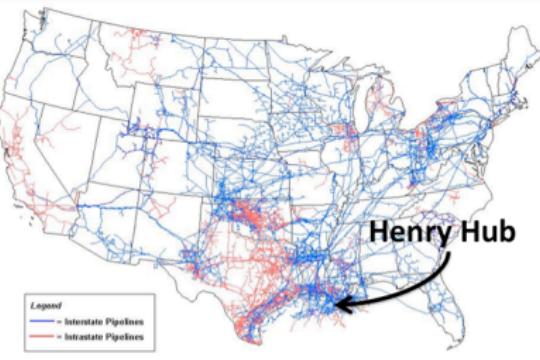
- 32% of all US Electricity  
(1.273 PWh in 2017)
- 3M Miles of NG  
Pipelines
- 150+ Trading Spots



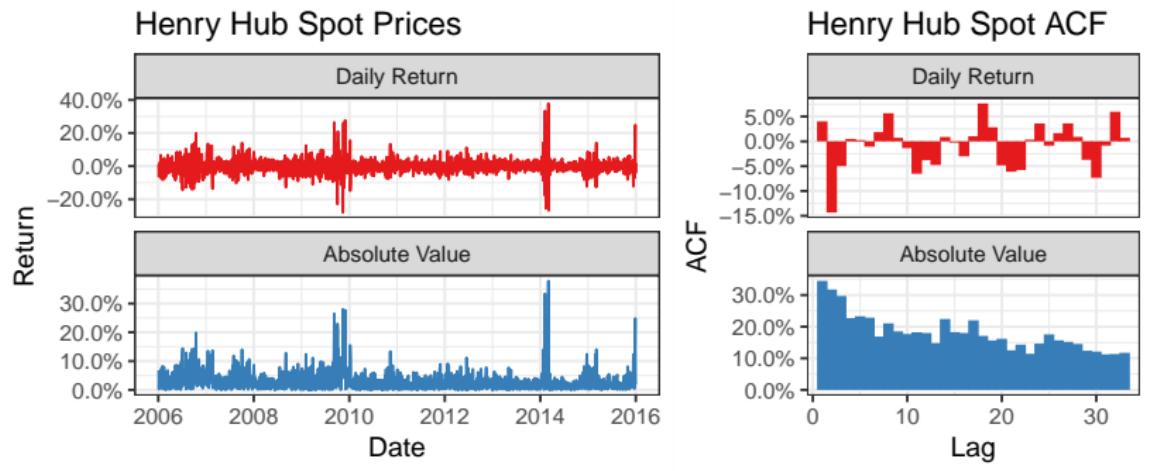
# Natural Gas Markets

Henry Hub:

- \$14B+ Futures Volume Daily
- Common Proxy for Domestic NG Markets Broadly

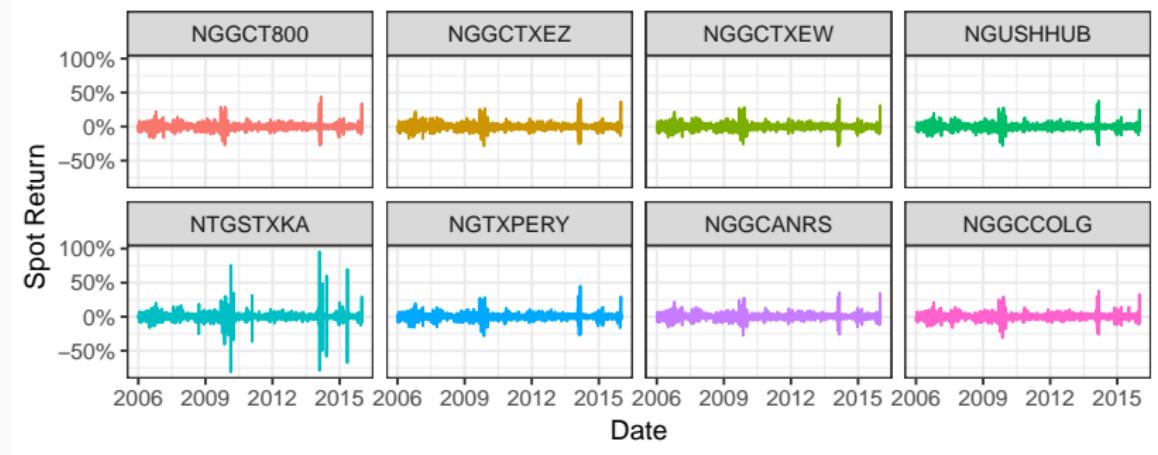


# Natural Gas Markets



Equity-like dynamics: vol clustering, heavy-tails, 2nd moment  
autocorrelation

# Natural Gas Markets



High inter-spot correlation: PC1 (74%) suggests single-factor model

# Natural Gas Markets

We want to capture:

- High-Dimensional Multivariate Time Series
- Irregular Data Availability
  - NG Futures Priced Near-Continuously on Lit Markets
  - NG Spots Traded Over-the-Counter
- GARCH Type Behavior + Single-Factor Structure

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- GARCH Type Behavior + Single-Factor Structure

💡 Realized Beta GARCH Model (Hansen *et al.*, 2014) combining:

- Intra-Day Futures Realized Volatility
- End-of-Day Spot Volatility
- 2<sup>nd</sup> Moment Single-Factor Dynamics

# Bayesian Realized Beta GARCH

- Bi-Variate GARCH Model (Multivariate Skew Normal Specification):
  - “Realized” (High-Frequency) Volatility: improved estimate of  $\sigma_t^2$
  - “Beta” Volatility linkage: Volatility at Henry  $\implies$  volatility in spots

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  - “Beta” Volatility linkage: Volatility at Henry  $\implies$  volatility in spots
- Bayesian Estimation:
  - Priors calibrated to S&P 500 (equity) markets: improve estimation
  - Coherent uncertainty propagation
  - Improved out of sample forecasts

# Application to Tail Forecasting

Does it work?

Fitting strategy:

- Fit to 250 window: refit every 50 days
- One-day rolling predictions for out-of-sample

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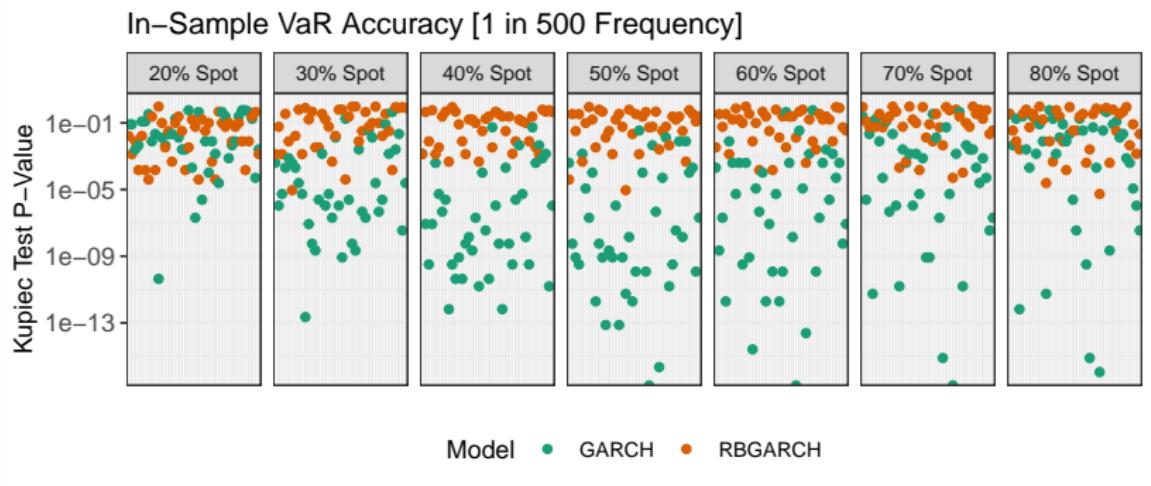
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Measures of Fit:

- In-sample VaR Test (Kupiec, 1995)
  - Binomial test for number of VaR exceedances
- Out-of-sample VaR Test (Kupiec, 1995)
  - Estimated out-of-sample log-likelihood

# Application to Tail Forecasting



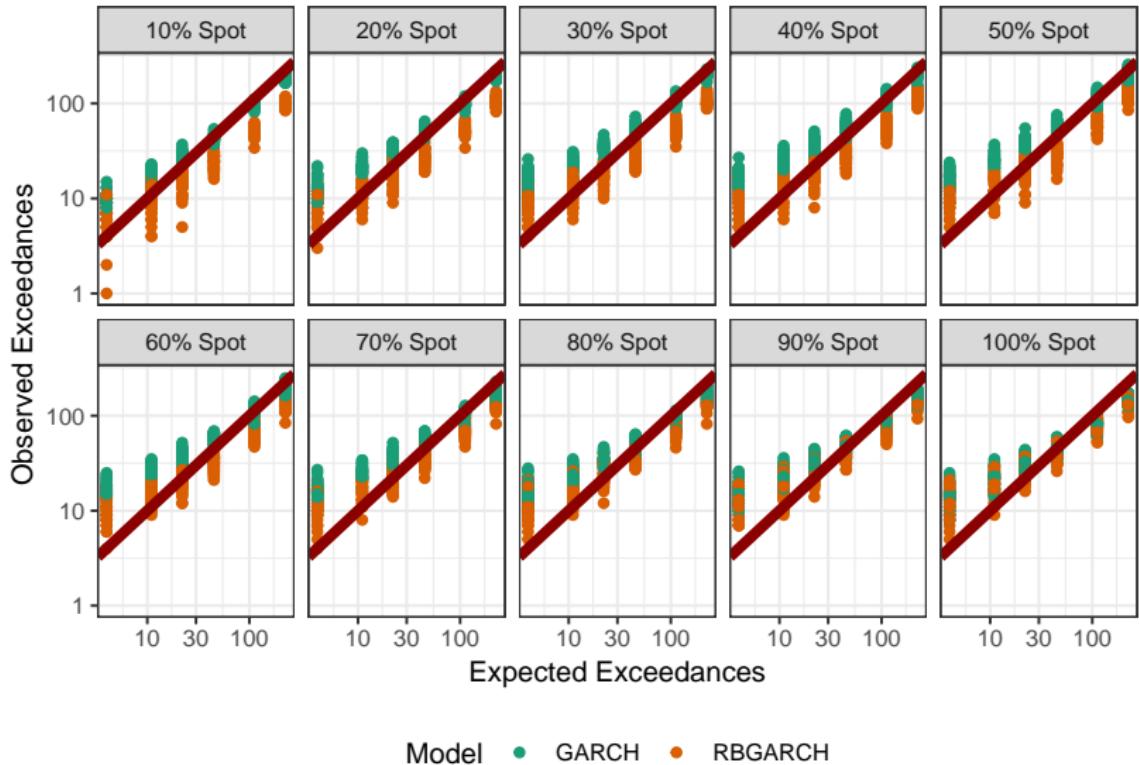
# Application to Tail Forecasting

Out-of-Sample VaR Accuracy [1 in 500 Frequency]



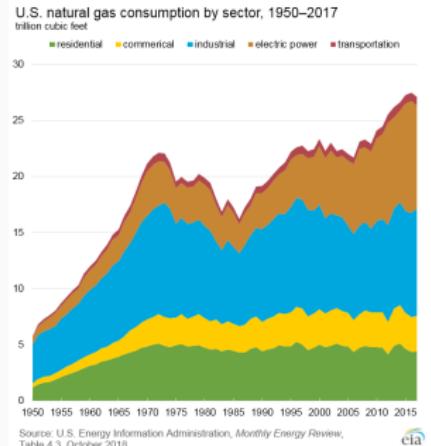
# Application to Tail Forecasting

Out-of-Sample VaR Performance



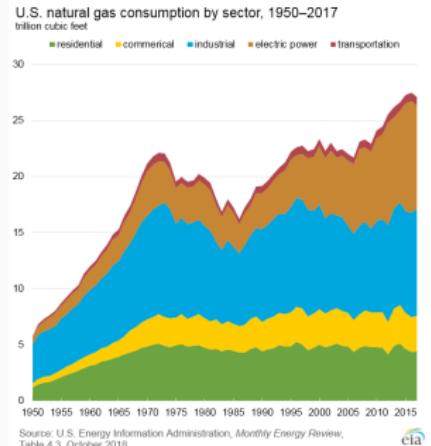
# Implications

- Multivariate and Multi-Resolution Model for NG Volatility
  - Daily and intra-day volatility measures
  - Multivariate Treatment of 50+ NG Trading Hubs



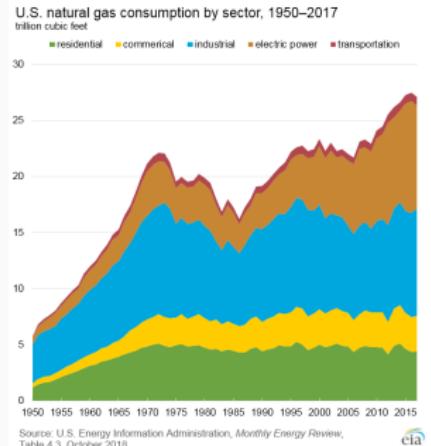
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- Bayesian Approach
  - Market Calibrated Priors



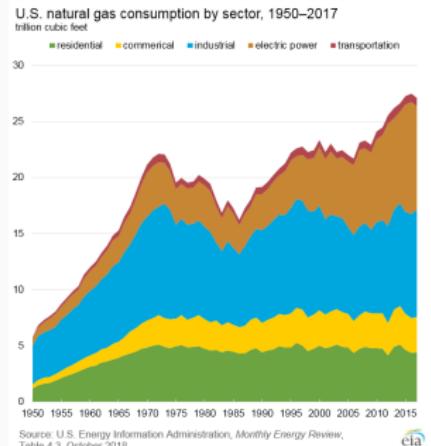
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- Multivariate and Multi-Resolution Model for NG Volatility
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- Amenable to all commodities markets with irregular data availability



# Complex Convex Analysis

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# Motivation

---

Complex-data arise in many domains:

- Signal and radar processing (Schreier and Scharf, 2010; Candès *et al.*, 2015; Mechlenbrauker *et al.*, 2017)
- Neuroscience (Yu *et al.*, 2018; Adrian *et al.*, 2018)
- Geostatistics (de Iaco *et al.*, 2003; Mandic *et al.*, 2009)
- Astronomy (Zechmeister and Kürster, 2009)
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Major sources:

- Spectral analysis (Fourier transforms)
- 2D directional data

# Why Complex?

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Why not treat complex data as real?

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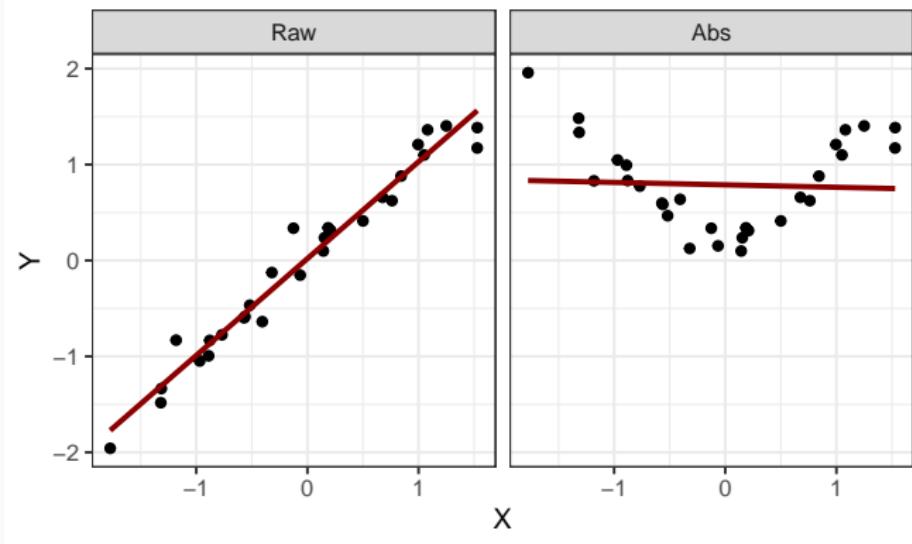
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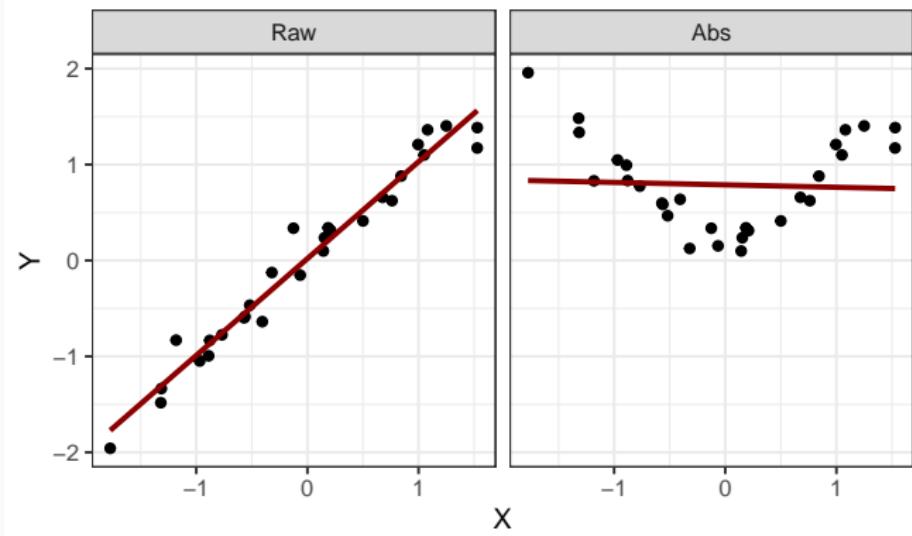
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- Corrupts statistical relationships:



- Natural domain for “spectral” phenomena

## Proper and Improper RVs

Let  $Z$  be a *univariate* complex random variable:

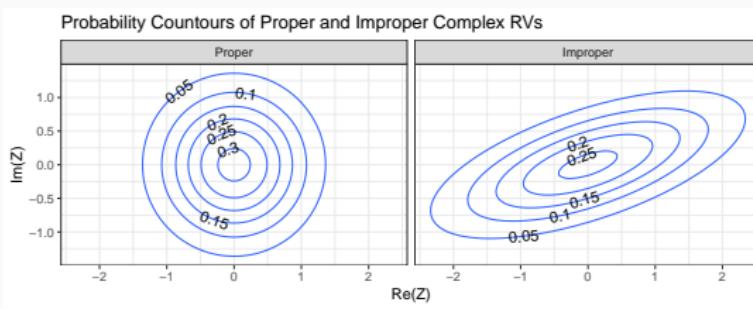
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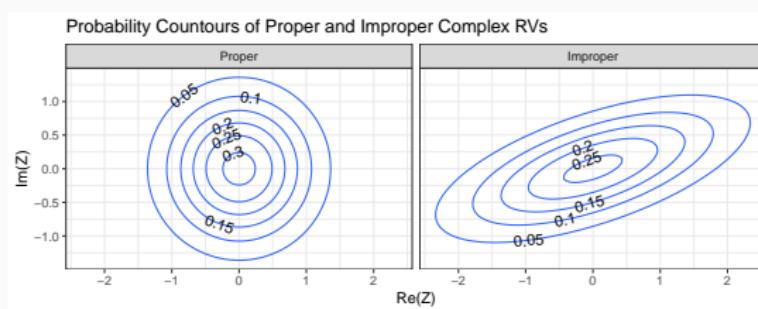


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Complex variables often arise as Fourier transform of *stationary* processes

- Only *relative* phase of multivariate signal matters
- Absolute phase is meaningless

$$\text{Law}[Z] = \text{Law}[e^{i\theta} Z] \implies Z \text{ proper}$$

# Complex Gaussian Distribution

The *complex* Gaussian is a **three** parameter distribution:

- Mean:  $\mu = \mathbb{E}[\mathbf{Z}]$
- Covariance:  $\Sigma = \mathbb{E}[(\mathbf{Z} - \mu)(\mathbf{Z} - \mu)^H] = \mathbb{E}[(\mathbf{Z} - \mu)\overline{(\mathbf{Z} - \mu)}^T]$
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- $\Sigma$  is positive-definite:  $\Sigma_{ij} = \mathbb{E}[z_i \bar{z}_j]$  - non-negative when  $i = j$
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Proper complex normal ( $\Gamma = \mathbf{0}$ ) almost universally assumed in statistics  
(Wooding, 1956; Goodman, 1963; Graczyk *et al.*, 2003)

Non-stationary DSP sometimes uses general case (van den Bos, 1995;  
Schreier and Scharf, 2010; Adali *et al.*, 2011)

# Complex Convex Analysis: Subgradient Analysis

Penalized  $M$ -Estimation Paradigm:

$$\hat{\beta} = \arg \min_{\beta \in \mathbb{R}^p} \mathcal{L}(\beta; \mathbf{X}, \mathbf{y}) + \lambda \mathcal{P}(\beta)$$

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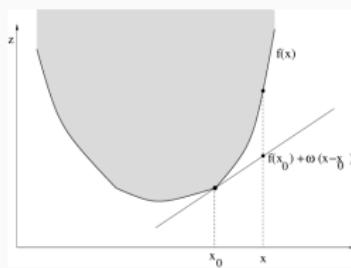
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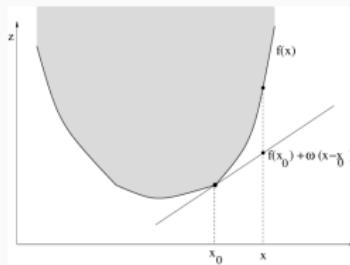
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If  $x, y$  are *complex*, this is ill-defined!



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**GOAL:** rigorously define this derivative and connect it to optimization

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- Can add vectors (elements of  $\mathbb{C}^P$ ) and multiply by  $\mathbb{F}$
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In this work,  $\mathbb{F} = \mathbb{R}$ !

$$\langle \mathbf{a}, \mathbf{b} \rangle = \frac{\mathbf{a}^H \mathbf{b} + \mathbf{a}^T \bar{\mathbf{b}}}{2}$$

Sub-gradient inequality becomes

$$f(\mathbf{w}) \geq f(\mathbf{z}) + \langle \boldsymbol{\gamma}, \mathbf{w} - \mathbf{z} \rangle \quad \text{for all } \mathbf{w} \in \mathbb{C}^p$$

All terms real  $\implies$  well-defined!

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We have freedom to change *algebraic* structure used to analyze problem

Likelihood still defined with “regular” complex multiplication

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Change definition of “multiplication”  
convex analysis still “works!”

## Example: Complex OLS

$$\arg \min_{\beta \in \mathbb{C}^p} \|\mathbf{y} - \mathbf{X}\beta\|_2^2$$

Set Wirtinger derivative to  $\mathbf{0}$ :

$$\begin{aligned} f &= \|\mathbf{y} - \mathbf{X}\beta\|_2^2 \\ &= (\mathbf{y} - \mathbf{X}\beta)^H(\mathbf{y} - \mathbf{X}\beta) \\ &= (\mathbf{y}^H - \bar{\beta}\mathbf{X}^H)(\mathbf{y} - \mathbf{X}\beta) \\ \mathbf{0} &= \frac{\partial f}{\partial \beta} = -\mathbf{y}^H \mathbf{X} + \bar{\beta}^T \mathbf{X}^H \mathbf{X} \\ \implies \bar{\beta} &= (\mathbf{X}^H \mathbf{X})^{-1} \mathbf{X}^H \bar{\mathbf{y}} \\ \beta &= (\mathbf{X}^H \mathbf{X})^{-1} \mathbf{X}^H \mathbf{y} \end{aligned}$$

Intuition from real-domain translates to complex-domain!

## Concentration Inequalities: Regression Noise (W., Lemma 3.9)

Suppose  $Z$  is mean-zero sub-Gaussian with variance proxy  $\sigma^2$ :

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Similar rates for effective noise  $\|\mathbf{X}^H \boldsymbol{\epsilon}\|_\infty$  depending on  $\boldsymbol{\epsilon}$

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$$[\sigma^2 = \max(\Sigma_{ii}^*)]$$

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Better concentration for proper complex  $Z$  than real  $Z$ !

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# The Complex Lasso

Real LASSO: (Tibshirani, 1996; Chen *et al.*, 1998)

$$\hat{\beta} = \arg \min_{\beta \in \mathbb{R}^p} \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_1$$

$\ell_1$ -norm penalty  $\implies \hat{\beta}$  will be *sparse* (have exact zeros):

- Compressed Sensing: can estimate  $\beta^*$  well even with  $p \ll n$  elements
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Rich theoretical literature Fu and Knight (2000), Greenshtein and Ritov (2004), Zhao and Yu (2006), Bickel *et al.* (2009), Zhang and Huang (2008), Bunea *et al.* (2007), Meinshausen and Yu (2009), and van de Geer and Bühlmann (2009) etc.

Results all translate to the complex lasso (CLASSO)!

$$\hat{\beta} = \arg \min_{\beta \in \mathbb{C}^p} \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_1$$

# The Complex Lasso

W., **Theorem 3.3:** Under standard assumptions (Wainwright, 2009), the Complex-Lasso (CLASSO) is model selection consistent with probability

- (a)  $\geq 1 - 2 \exp\{-(\tau - 2)/2 \log(p - s)\}$  for real  $\epsilon$  (real  $\mathbf{X}, \mathbf{y}$ )

$$\epsilon \stackrel{\text{IID}}{\sim} \text{subG}(0, \sigma^2) \quad \lambda_{\min}(\mathbf{X}_S^H \mathbf{X}_S / n) \geq c > 0 \quad \max_{j \in S^c} \|(\mathbf{X}_S^H \mathbf{X}_S)^{-1} \mathbf{X}_S^H \mathbf{x}_j\|_1 \leq 1 - \gamma$$

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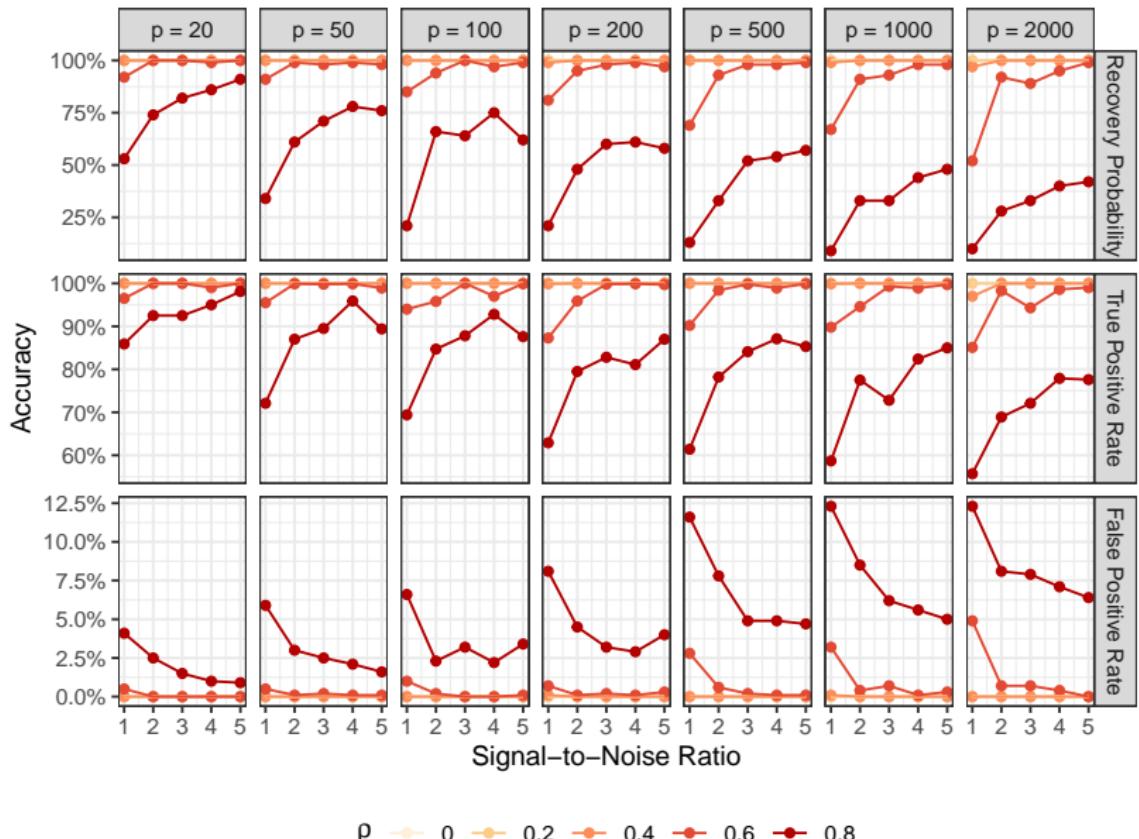
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- (c)  $\geq 1 - 4 \exp\{-(\tau - 16)/16 \log(p - s)\}$  for complex  $\epsilon$
- (d)  $\geq 1 - 8 \exp\{-(\tau - 2)/2 \log(p - s)\}$  for proper complex  $\epsilon$

$$\epsilon \stackrel{\text{IID}}{\sim} \text{subG}(0, \sigma^2) \quad \lambda_{\min}(\mathbf{X}_S^H \mathbf{X}_S / n) \geq c > 0 \quad \max_{j \in S^c} \|(\mathbf{X}_S^H \mathbf{X}_S)^{-1} \mathbf{X}_S^H \mathbf{x}_j\|_1 \leq 1 - \gamma$$

First precise finite sample results for CLASSO: previously studied by Yang and Zhang (2011), Maleki *et al.* (2013), and Mechlenbrauker *et al.* (2017)

# Model Selection Consistency of CLasso



# The Complex Graphical Lasso

Suppose  $\mathbf{Z}$  is drawn from a  $p$ -variate complex Gaussian with precision matrix  $\Theta^* = (\Sigma^*)^{-1}$ . CGLASSO gives a sparse estimate of  $\Theta^*$ :

$$\hat{\Theta} = \arg \min_{\Theta \in \mathbb{C}_{\geq 0}^{p \times p}} -\log \det \Theta + \text{Tr}(\hat{\Sigma}\Theta) + \lambda \|\Theta\|_{1,\text{off-diag}}$$

where  $\hat{\Sigma}$  is the sample covariance (Yuan and Lin, 2007; Friedman *et al.*, 2008; Ravikumar *et al.*, 2011).

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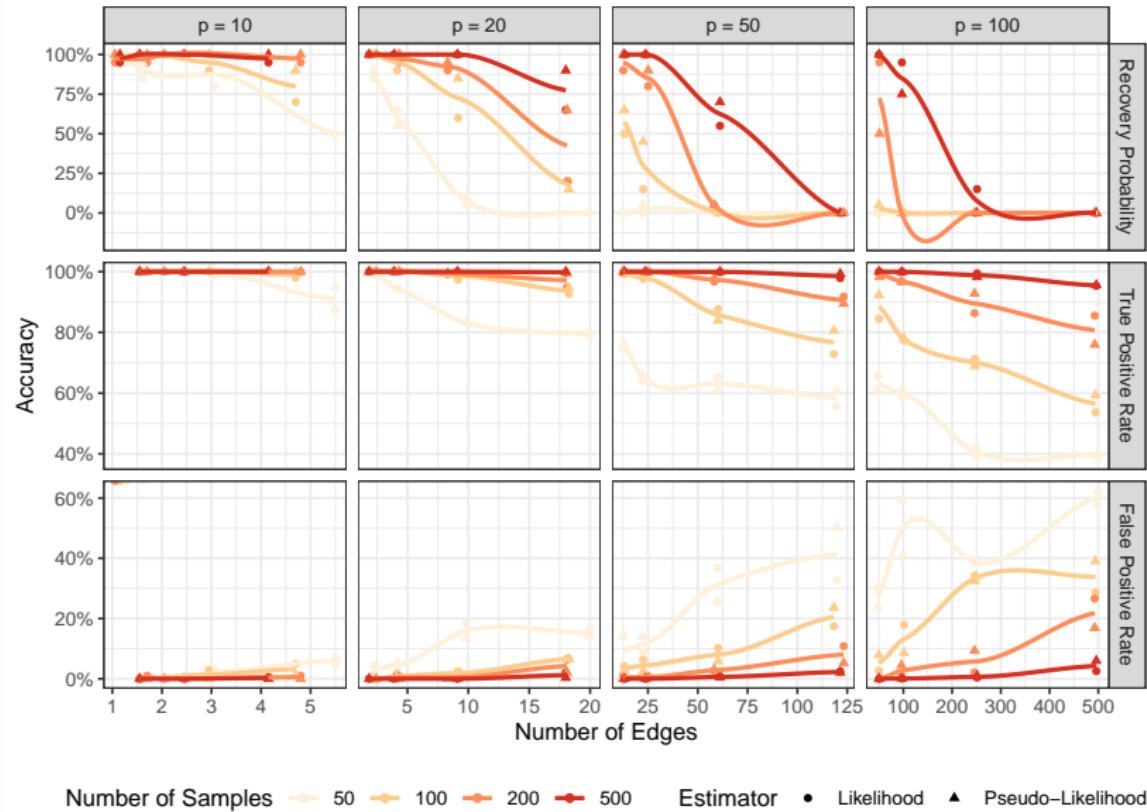
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Similar results for neighborhood selection (regularized pseudo-likelihood) (Meinshausen and Bühlmann, 2006)

First theoretical results for CGLASSO: Tugnait (2018, 2019a, 2019b) gave experimental results and algorithms

# Model Selection Consistency of CGLasso



## Improper and Dependent Observations

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In addition to incoherence of  $\Theta^*$ , CGLASSO requires only that

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## W. Theorem 4.1 + Fiecas et al. (2019) + Dahlhaus (2000)

Suppose  $\mathcal{Z} = \{\mathbf{Z}_t\}_{t=1}^T$  is a stationary Gaussian linear  $p$ -dimensional time series with spectrum  $\Gamma(k) \in \mathbb{C}_{\geq 0}^{p \times p}$ , such that  $\Gamma^{-1}(k)$  satisfies the incoherence conditions at all  $k$ . Let  $\hat{\Gamma}(k)$  be the sample averaged periodogram based on  $T$  observations. Then the graphical model  $\hat{\mathcal{G}} = (\mathcal{V}, \hat{\mathcal{E}})$  with

$$(i, j) \notin \hat{\mathcal{E}} \Leftrightarrow \hat{\Theta}_{ij}^{(k)} = 0 \quad \text{for } \hat{\Theta}^{(k)} = \text{CGLASSO}(\hat{\Gamma}(k)) \quad \text{for all } k < T/2$$

correctly estimates the conditional independence structures in  $\mathcal{Z}$  at all lags with probability  $\geq 1 - Cp^2/T$  for  $T$  sufficiently large.

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- Foundational Optimization and Statistical Theory for *Complex Machine Learning*
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- Exciting Implications for Multivariate Time Series

## **Conclusion & Discussion**

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# Conclusions

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Extensions of classical  $M$ -estimation and convex analysis to Wirtinger functions

# Acknowledgements

M G Z L U K E M I N J I E F M  
M A J F G G L I B B Y U X R I  
A A R O A N D E R S E N N E T  
K U T G H I S G E O R G E D C  
E I G T A A K R I S T E N Y H  
A I M U T R N J U L I A H G T  
N U G A S O E M E W V T G U E  
D U T A T T M T A T A K B R R  
R G I U U T T I R Z K N L E D E  
E R A K Y T E N I L N U N C N  
A F N K O V A O E O U I H O C  
B P Y M E W N M N H C S A F E  
N F I N B T A N E J C M A E G  
I S E Y O O A L A J I I M S S P  
E G F O B C D R J O N E S S E

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**Thank you!**