

Statistics 12, Lab 3

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February 19th, 2019

Exercise 1

a) I ran the following code.

```
linear_model <- lm(soil$lead ~ soil$zinc)
summary(linear_model)
```

This output the following statistics.

Call:

```
lm(formula = soil$lead ~ soil$zinc)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-79.853	-12.945	-1.646	15.339	104.200

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	17.367688	4.344268	3.998	9.92e-05 ***
soil\$zinc	0.289523	0.007296	39.681	< 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 33.24 on 153 degrees of freedom

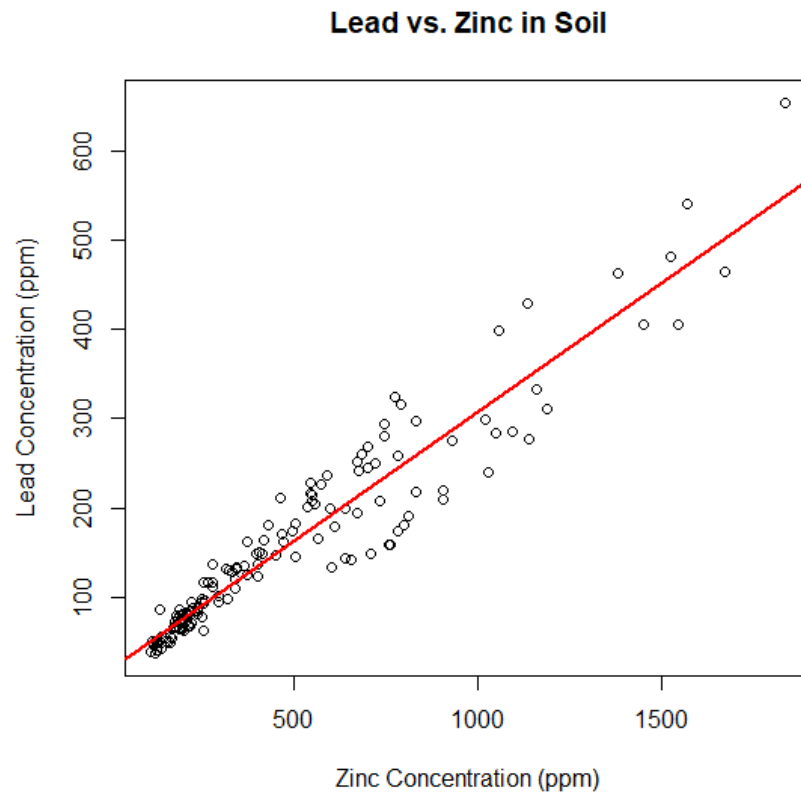
Multiple R-squared: 0.9114, Adjusted R-squared: 0.9109

F-statistic: 1575 on 1 and 153 DF, p-value: < 2.2e-16

b) I ran the following code.

```
plot(soil$lead ~ soil$zinc, xlab="Zinc Concentration (ppm)",  
     ylab="Lead Concentration (ppm)", main="Lead vs. Zinc in Soil")  
abline(linear_model, col="red", lwd=2)
```

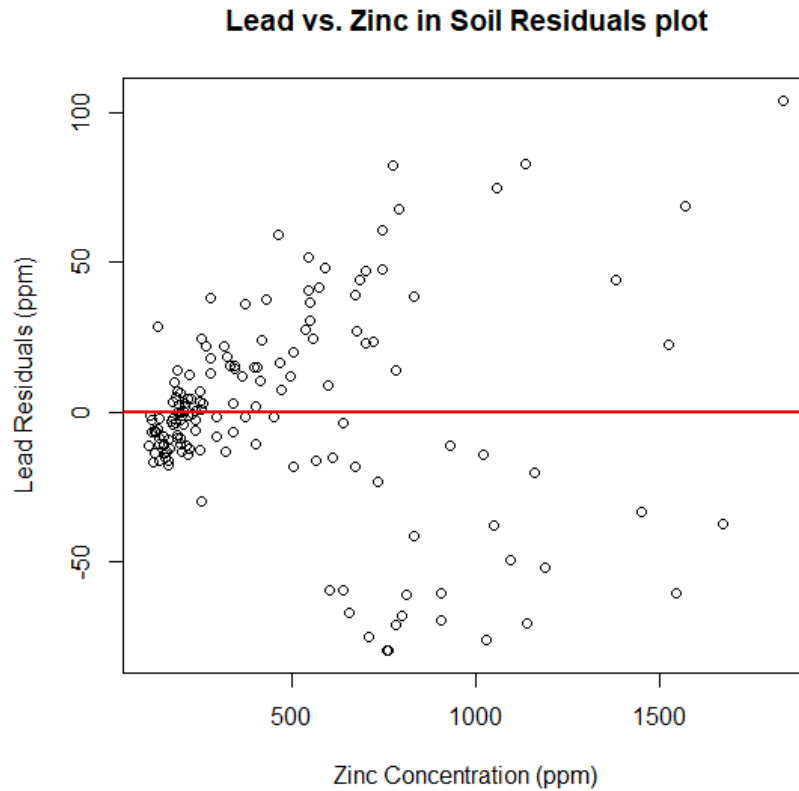
This generated the following plot.



c) I ran the following code

```
plot(linear_model$residuals ~ soil$zinc, xlab="Zinc Concentration (ppm)",  
     ylab="Lead Residuals (ppm)", main = "Lead vs. Zinc in Soil Residuals plot")  
abline(a=0, b=0, col="red", lwd=2)
```

This generated the following plot.



d) The equation of the regression line is the following.

$$y = 0.289523x + 17.367688$$

e) We would expect the lead concentration to be 306.89 ppm.

f) We would expect the lead concentration to be 28.95 ppm higher.

g) The R-squared value is 0.9114. In this context it means that 91.14% of the variation in the lead concentration can be determined by the variation in zinc concentration.

h) I believe that linearity, symmetry, and are met for this data. The data appears very straight and the residuals look symmetrical. However, the condition for equal variance is not met. The residuals appear to become more spread out as the zinc concentration becomes greater. This would indicate that our model does not capture all the details of our system. Perhaps of the logarithms of both zinc and lead would be more appropriate to compare with each other.

Exercise 2

a) I ran the following code.

```
linear_model2 <- lm(ice$Extent ~ ice$Date)
summary(linear_model2)
```

This output the following statistics.

Call:

```
lm(formula = ice$Extent ~ ice$Date)
```

Residuals:

Min	1Q	Median	3Q	Max
-9.445	-5.439	1.442	5.599	7.564

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	1.011e+01	1.558e+00	6.486	4.11e-10 ***
ice\$Date	1.438e-04	1.411e-04	1.019	0.309

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.654 on 273 degrees of freedom

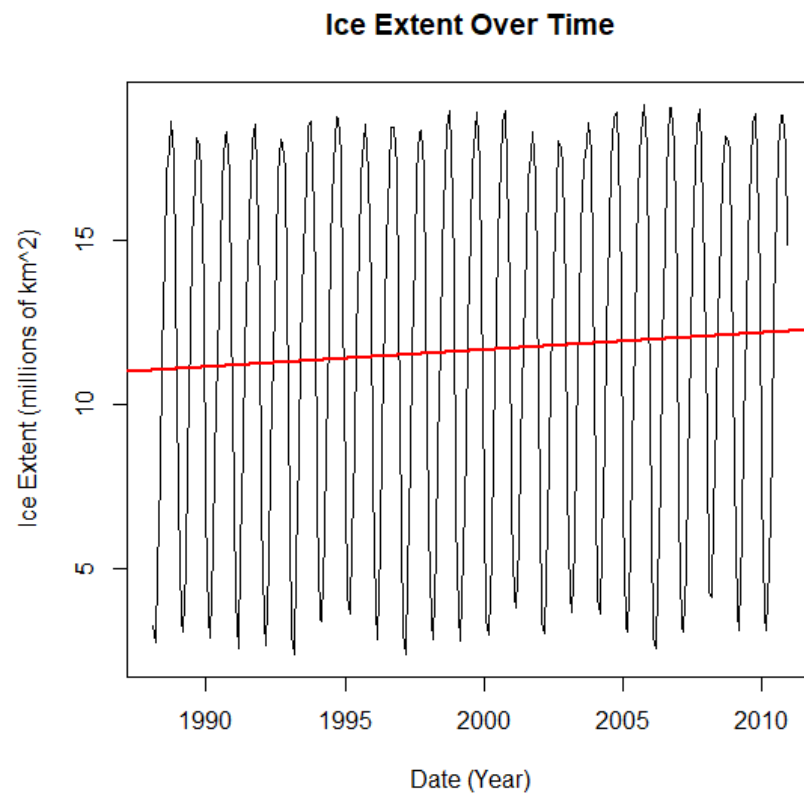
Multiple R-squared: 0.003787, Adjusted R-squared: 0.0001377

F-statistic: 1.038 on 1 and 273 DF, p-value: 0.3093

b) I ran the following code.

```
plot(ice$Extent ~ ice$Date, xlab="Date (Year)",
     ylab="Ice Extent (millions of km^2)", main="Ice Extent Over Time", type="l")
abline(linear_model2, col="red", lwd=2)
```

This generated the following plot.

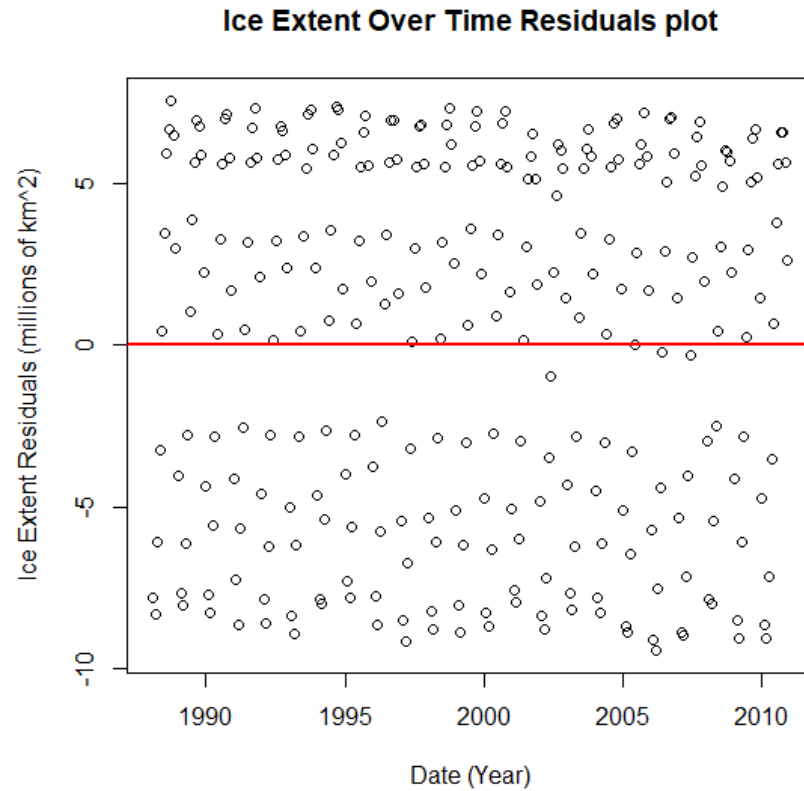


There does seem to be a slight upwards trend in this data as shown by the regression line. However, the main change in sea ice extent seems to come from the change in the seasons. The sea periodically rises and falls with a period of a year.

c) I ran the following code

```
plot(linear_model2$residuals ~ ice$Date, xlab="Date (Year)",  
     ylab="Ice Extent Residuals (millions of km^2)",  
     main = "Ice Extent Over Time Residuals plot")  
abline(a=0, b=0, col="red", lwd=2)
```

This generated the following plot.



The assumption about the model that we should be concerned about is the linearity assumption. The residuals have bands where the density changes, which indicates that our model consistently over and under predicts the sea ice extent within certain time frames. In reality the sea ice extent is probably best modeled by a sinusoidal relationship.

Exercise 3

- a)
- b)
- c)

d)

e)

f)

Exercise 4

a)

b)

c)

d)

e)