

Applied Data Analysis for Public Policy Studies

Simple Linear Regression

Michele Fioretti SciencesPo Paris 2020-09-11

Recap from past weeks

- R basics, importing data
- Exploratory data analysis:
 - Summary statistics: *mean*, *median*, *variance*, *standard deviation*
 - Data visualization: base R and ggplot2
 - Data wrangling: dplyr



Recap from past weeks

- R basics, importing data
- Exploratory data analysis:
 - Summary statistics: *mean*, *median*, *variance*, *standard deviation*
 - Data visualization: base R and ggplot2
 - Data wrangling: dplyr

Today - Real 'metrics finally 送

- Introduction to the **Simple Linear Regression Model** and **Ordinary Least Squares** *estimation*.
- Empirical application: class size and student performance
- Keep in mind that we are interested in uncovering **causal** relationships



Class size and student performance

- What policies *lead* to improved student learning?
- Class size reduction has been at the heart of policy debates for *decades*.



Class size and student performance

- What policies *lead* to improved student learning?
- Class size reduction has been at the heart of policy debates for decades.
- We will be using data from a famous paper by Joshua Angrist and Victor Lavy (1999), obtained from Raj Chetty and Greg Bruich's course.
- Consists of test scores and class/school characteristics for fifth graders (10-11 years old) in Jewish public elementary schools in Israel in 1991.
- National tests measured *mathematics* and (Hebrew) *reading* skills. The raw scores were scaled from 1-100.

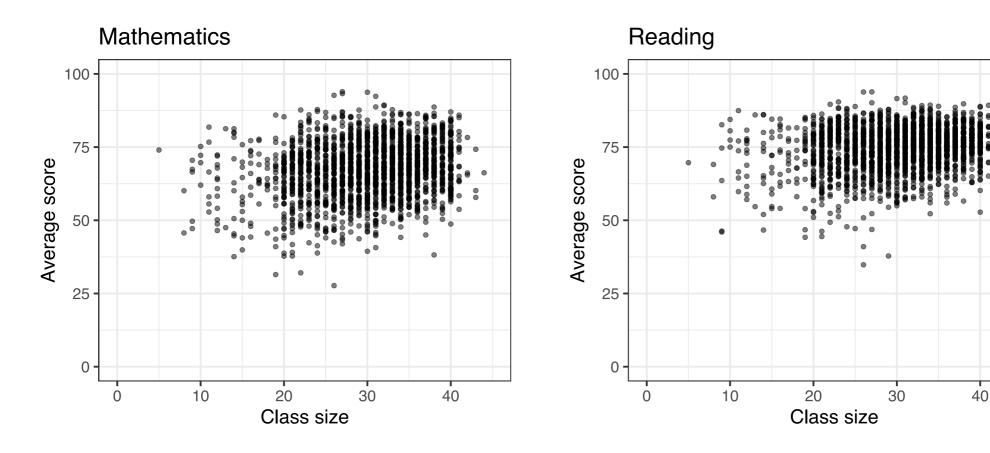


Task 1: Getting to know the data (7 minutes)

- 1. Load the data from here. You need to find the function that enables importing .dta files. (FYI: .dta is the extension for data files used in *Stata*)
- 2. Describe the dataset:
 - What is the unit of observations, i.e. what does each row correspond to?
 - How many observations are there?
 - What variables do we have? View the dataset to see what the variables correspond to.
 - What do the variables avgmath and avgverb correspond to?
 - Use the skim function from the skimr package to obtain common summary statistics for the variables classize, avgmath and avgverb.
 - Hint: use dplyr to select the variables and then simply pipe (%>%) skim().
- 3. Do you have any priors about the actual (linear) relationship between class size and student achievement? What would you do to get a first insight?
- 4. Compute the correlation between class size and math and verbal scores. Is the relationship positive/negative, strong/weak?

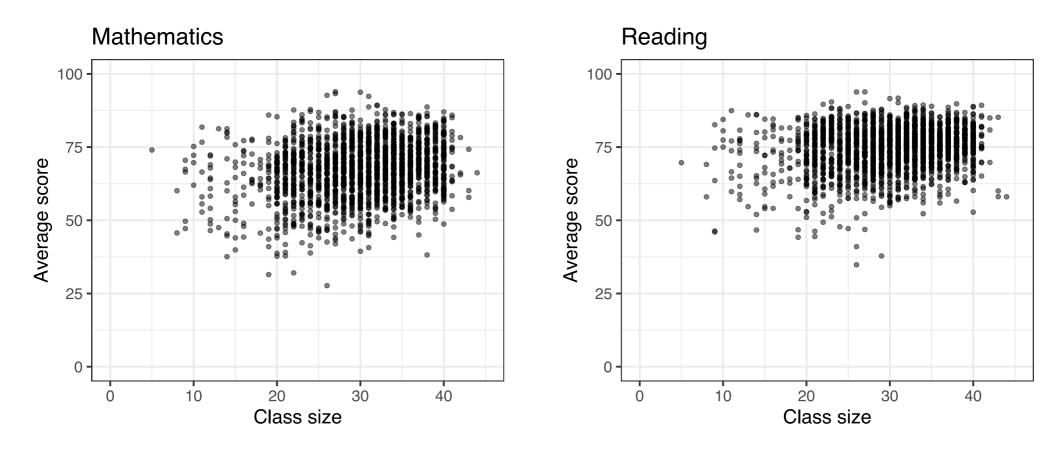


Class size and student performance: Scatter plot





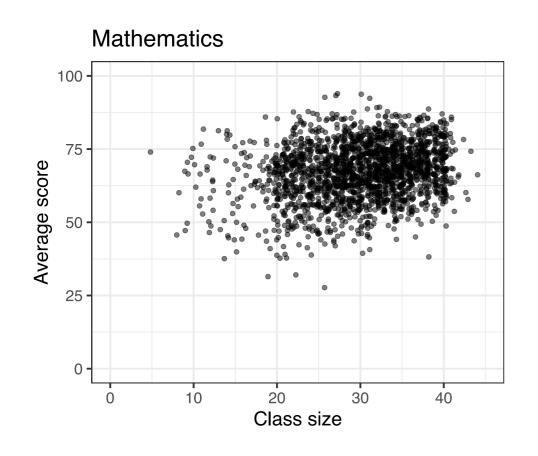
Class size and student performance: Scatter plot

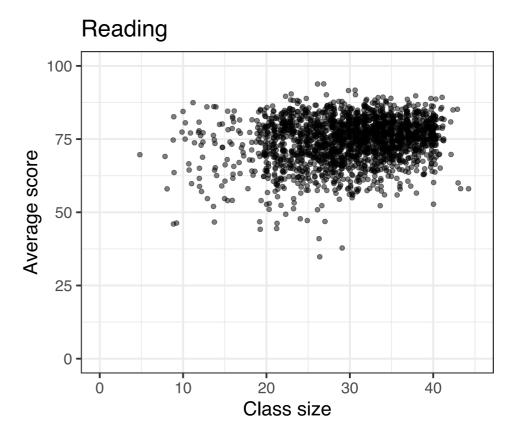




• Hard to see much because all the data points are aligned vertically. Let's add a bit of jitter to disperse the data slightly.

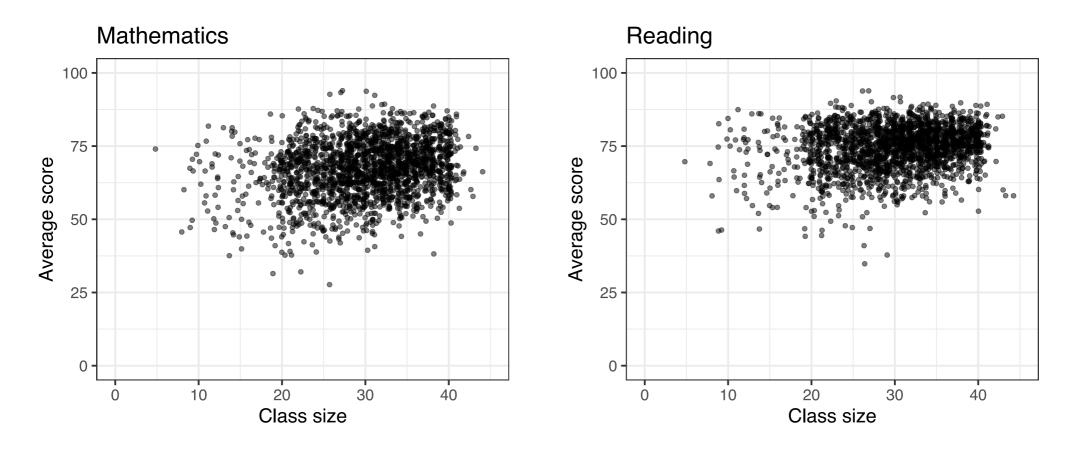
Class size and student performance: jitter scatter plot







Class size and student performance: jitter scatter plot





• Somewhat positive association as suggested by correlations. Let's compute the average score by class size to see things more clearly!

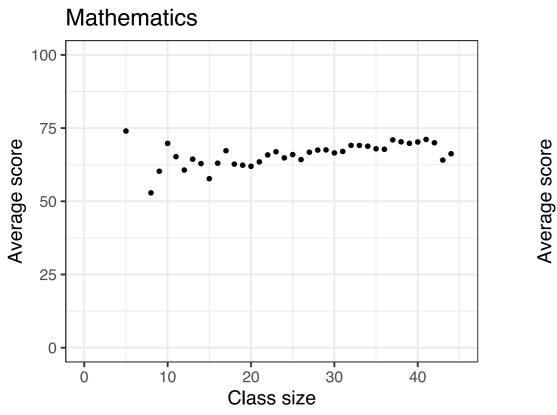
Task 2: Binned scatter plot (7 minutes)

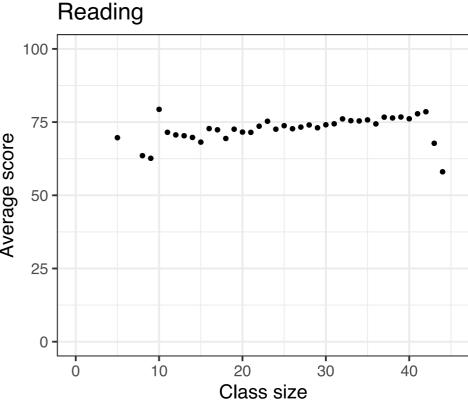
- 1. Create a new dataset (grades_avg_cs) where math and verbal scores are averaged by class size. Let's call these new average scores avgmath_cs and avgverb_cs.

 N.B.: the "raw" scores are already averages at the class level. Here we average these averages by class size.
- 2. Redo the same plots as before. Is the sign of the relationship more apparent?
- 3. Compute the correlation between class size and the new aggreagated math and verbal scores variables. Why is the (linear) association so much stronger?



Class size and student performance: Binned scatter plot

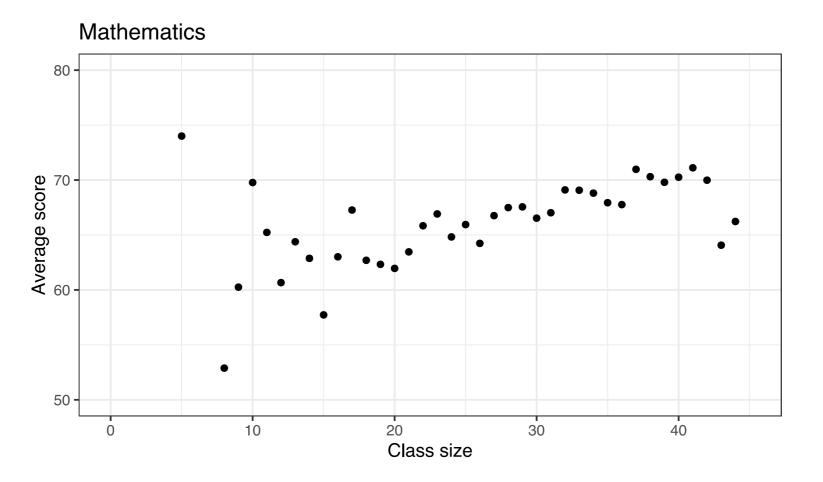






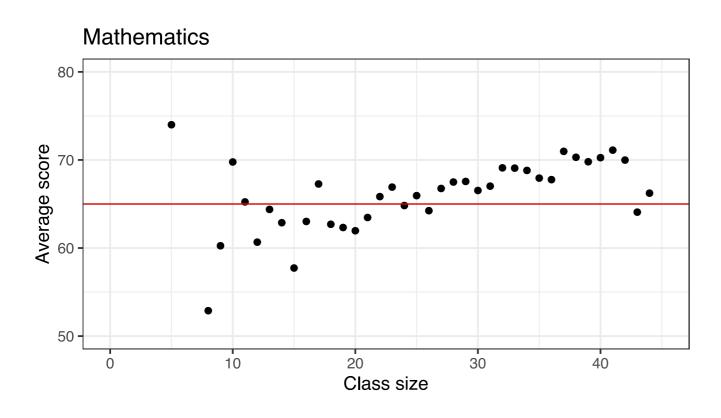
Class size and student performance: Binned scatter plot

• We'll first focus on the mathematics scores and for visual simplicity we'll zoom in

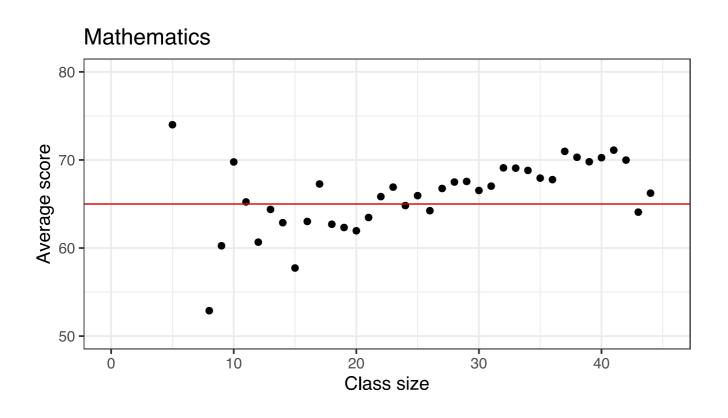






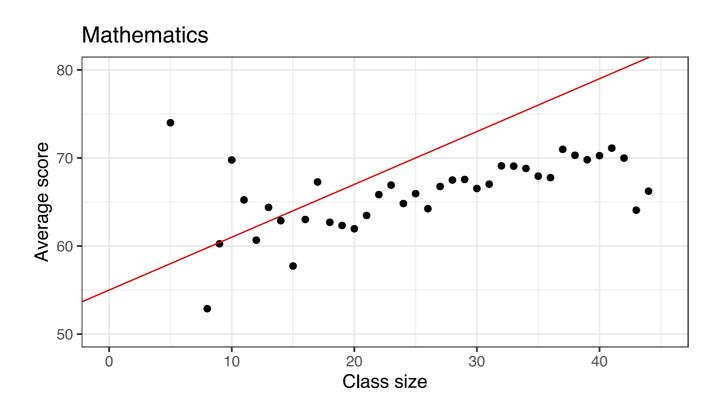




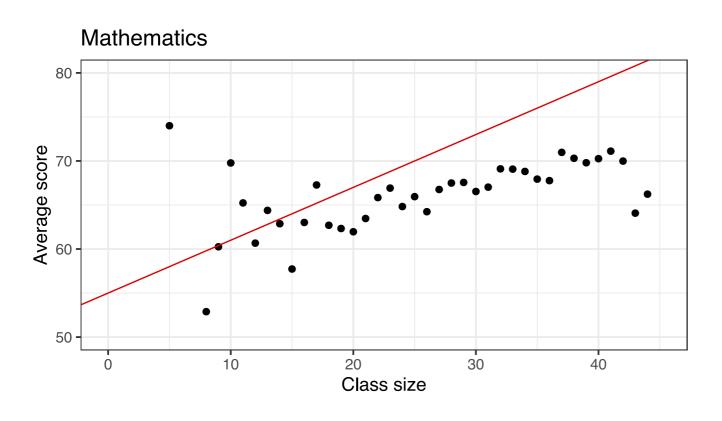


- A line! Great. But which line? This one?
- That's a *flat* line.
 But average
 mathematics score
 is somewhat
 increasing with
 class size









- That one?
- Slightly better! Has a slope and an intercept
- We need a rule to decide!



Let's formalise a bit what we are doing so far.

• We are interested in the relationship between two variables:



- We are interested in the relationship between two variables:
 - \circ an **outcome variable** (also called **dependent variable**): average mathematics score (y)



- We are interested in the relationship between two variables:
 - \circ an **outcome variable** (also called **dependent variable**): average mathematics score (y)
 - \circ an **explanatory variable** (also called **independent variable** or **regressor**): class size (x)



- We are interested in the relationship between two variables:
 - an outcome variable (also called dependent variable): average mathematics score (y)
 - \circ an **explanatory variable** (also called **independent variable** or **regressor**): class size (x)
- For each class i we observe both x_i and y_i , and therefore we can plot the *joint distribution* of class size and average mathematics score.



- We are interested in the relationship between two variables:
 - an outcome variable (also called dependent variable): average mathematics score (y)
 - \circ an **explanatory variable** (also called **independent variable** or **regressor**): class size (x)
- For each class i we observe both x_i and y_i , and therefore we can plot the *joint distribution* of class size and average mathematics score.
- We summarise this relationship with a line (for now). The equation for such a line with an intercept b_0 and a slope b_1 is:

$${\hat y}_i = b_0 + b_1 x_i$$



Let's formalise a bit what we are doing so far.

- We are interested in the relationship between two variables:
 - an outcome variable (also called dependent variable): average mathematics score (y)
 - \circ an **explanatory variable** (also called **independent variable** or **regressor**): class size (x)
- For each class i we observe both x_i and y_i , and therefore we can plot the *joint distribution* of class size and average mathematics score.
- We summarise this relationship with a line (for now). The equation for such a line with an intercept b_0 and a slope b_1 is:

$${\hat y}_i = b_0 + b_1 x_i$$

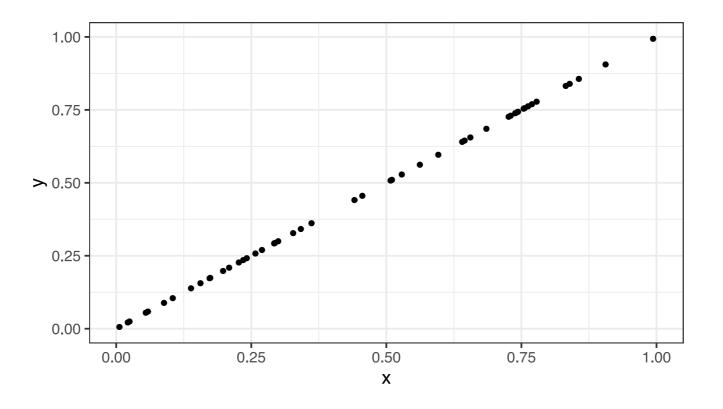


• \hat{y}_i is our *prediction* for y at observation i (y_i) given our model (i.e. the line).

ullet If all the data points were $oldsymbol{ ext{on}}$ the line then $\hat{y}_i=y_i.$

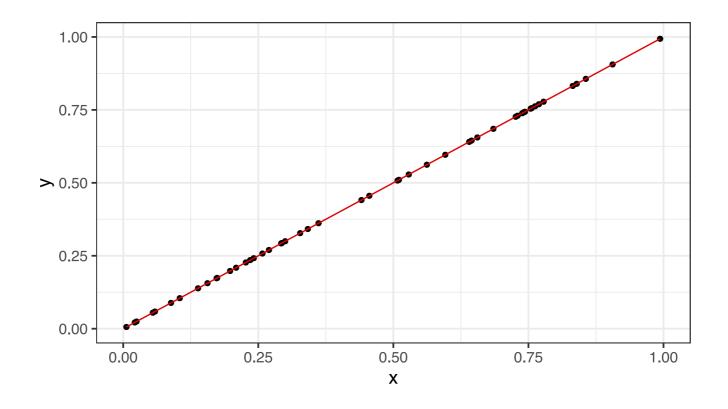


• If all the data points were **on** the line then $\hat{y}_i = y_i$.





• If all the data points were **on** the line then $\hat{y}_i = y_i$.





- If all the data points were **on** the line then $\hat{y}_i = y_i$.
- However, since in most cases the *dependent variable* (y) is not *only* explained by the chosen *independent variable* (x), $\hat{y}_i \neq y_i$, i.e. we make an **error**. This **error** is called the **error term**.



- If all the data points were **on** the line then $\hat{y}_i = y_i$.
- However, since in most cases the *dependent variable* (y) is not *only* explained by the chosen *independent variable* (x), $\hat{y}_i \neq y_i$, i.e. we make an **error**. This **error** is called the **error term**.
- At point (x_i, y_i) , we note this error e_i .



- If all the data points were **on** the line then $\hat{y}_i = y_i$.
- However, since in most cases the *dependent variable* (y) is not *only* explained by the chosen *independent variable* (x), $\hat{y}_i \neq y_i$, i.e. we make an **error**. This **error** is called the **error term**.
- At point (x_i, y_i) , we note this error e_i .
- The actual data (x_i, y_i) can thus be written as prediction + error:

$$y_i=\hat{y}_i+e_i=b_0+b_1x_i+e_i$$



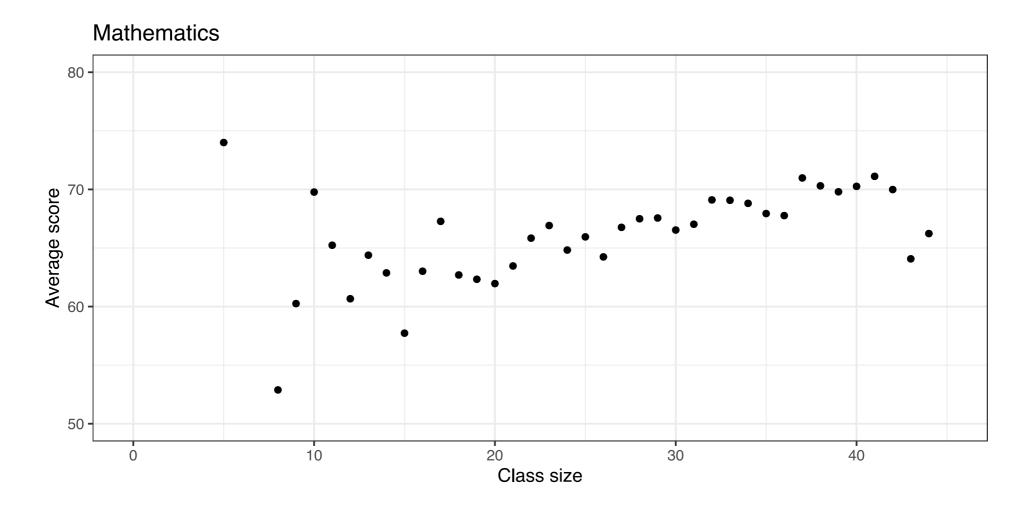
- If all the data points were **on** the line then $\hat{y}_i = y_i$.
- However, since in most cases the *dependent variable* (y) is not *only* explained by the chosen *independent variable* (x), $\hat{y}_i \neq y_i$, i.e. we make an **error**. This **error** is called the **error term**.
- At point (x_i, y_i) , we note this error e_i .
- The actual data (x_i, y_i) can thus be written as prediction + error:

$$y_i=\hat{y}_i+e_i=b_0+b_1x_i+e_i$$

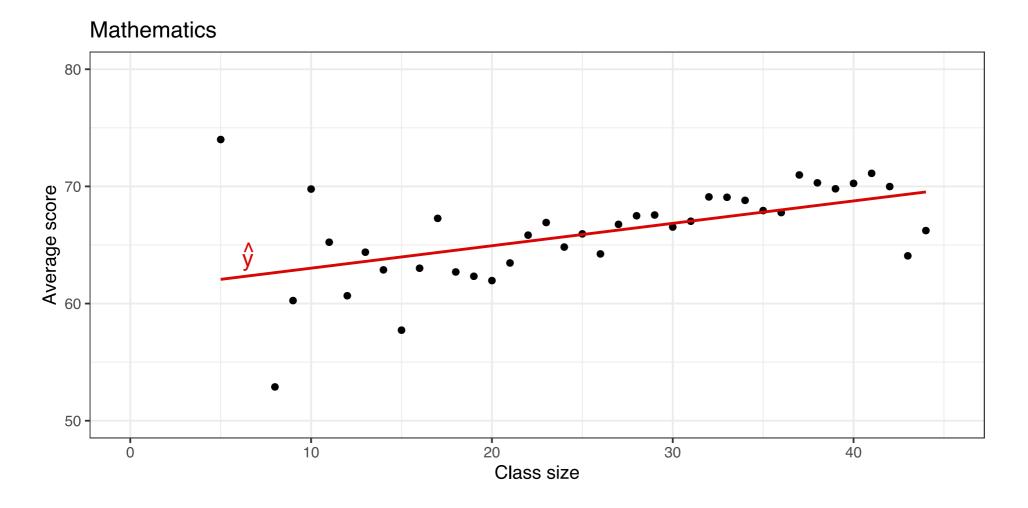
Goals

- 1. Find the values for b_0 and b_1 that make the errors as small as possible,
- 2. Check whether these values give a reasonable description of the data.

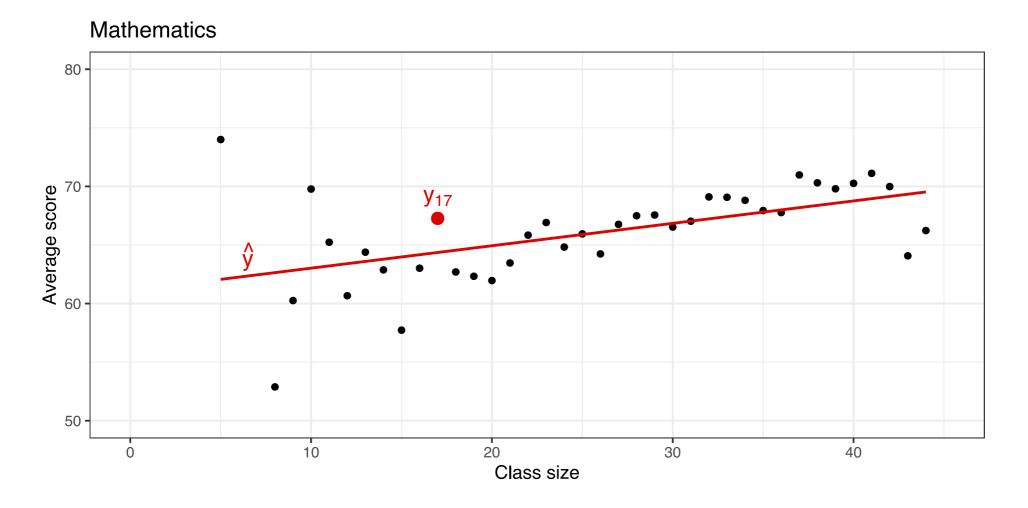




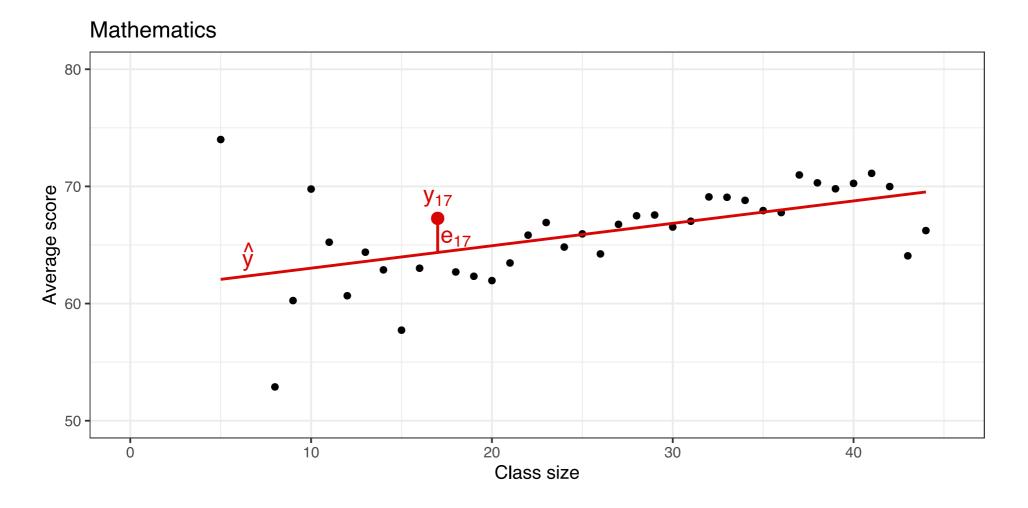




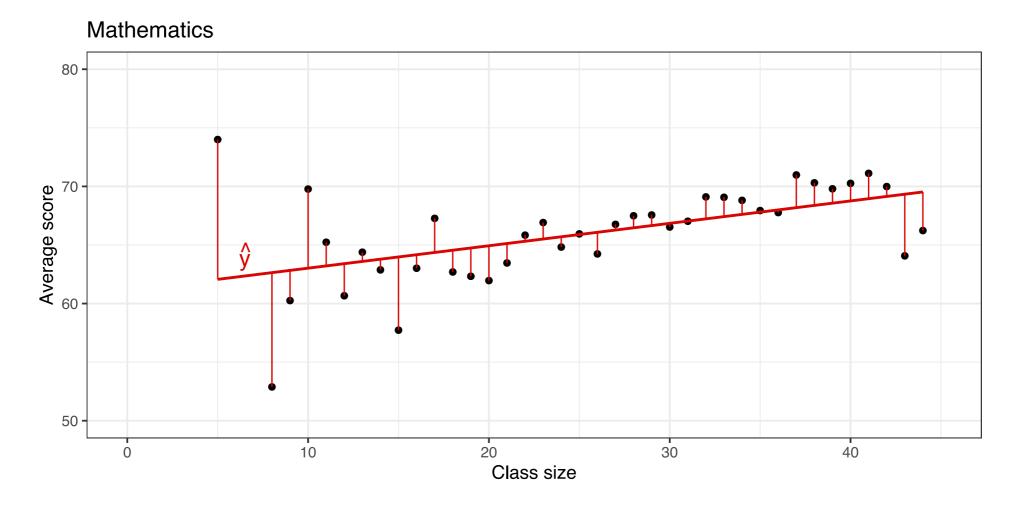














App Time! (5 minutes)

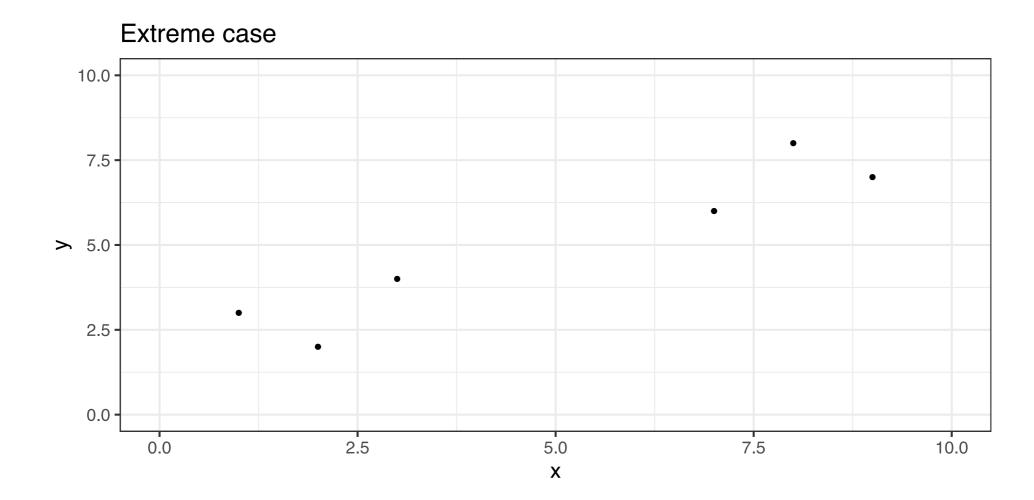
Intuitively one might want to simply minimize the absolute value of the sum of all the errors $|\sum_{i=1}^{n} e_i|$, that is one might want the sum of errors as close to 0 as possible.

Let's try to find the best line by minimizing the absolute value of the sum of errors. The line won't turn green so don't spend all day waiting for it.

```
library(ScPoApps) # load our library
launchApp('reg_simple_arrows')
aboutApp('reg_simple_arrows') # explainer about app
```

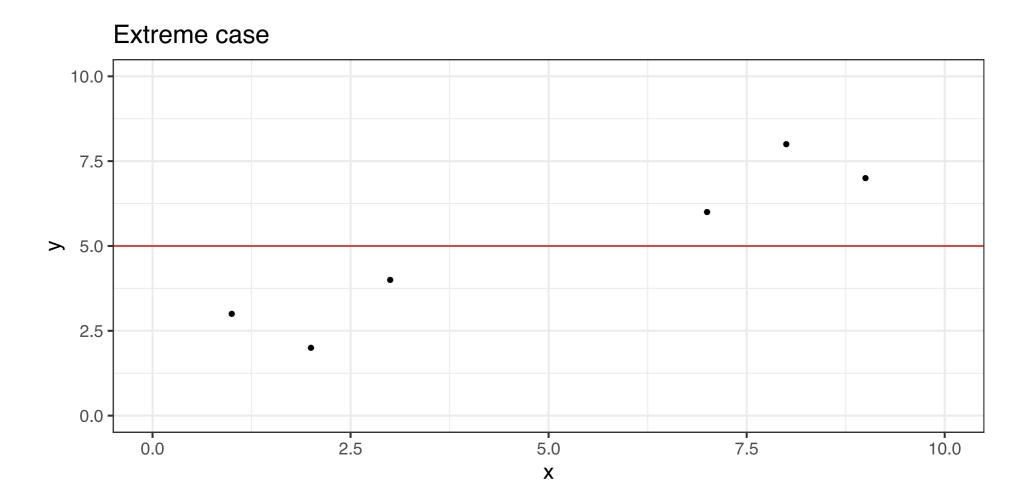


Minimizing the Absolute Value of the Sum of Errors



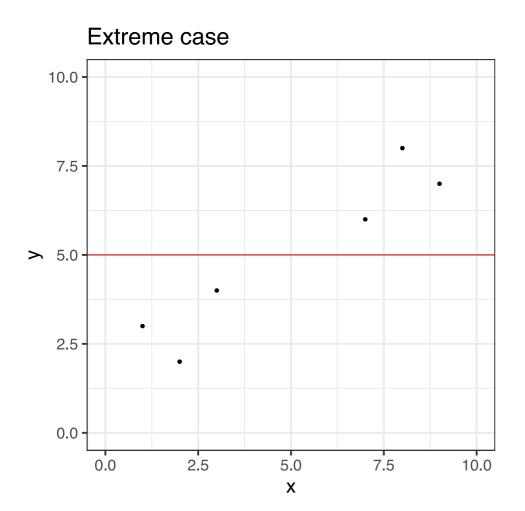


Minimizing the Absolute Value of the Sum of Errors





Minimizing the Absolute Value of the Sum of Errors



- This line minimizes the absolute value of the sum of errors since the data points are symmetric around y=5.
- Yet it clearly does not fit the data well!
- Note also that many other lines would also yield a sum of errors of 0 since the data are symmetric. A unique solution is not guaranteed!



Ordinary Least Squares (OLS) Estimation

• Errors of different sign (+/-) cancel out, so let's consider **squared residuals**

$$orall i \in [1,N], e_i^2 = (y_i - \hat{y}_i)^2 = (y_i - b_0 - b_1 x_i)^2$$

• Choose (b_0,b_1) such that $\sum_{i=1}^N e_1^2 + \cdots + e_N^2$ is **as small as possible**.



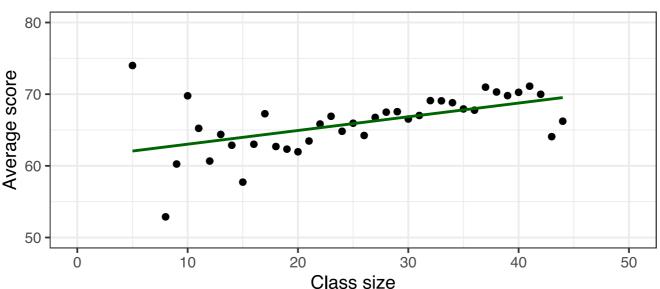
Ordinary Least Squares (OLS) Estimation

• Errors of different sign (+/-) cancel out, so let's consider **squared residuals**

$$orall i \in [1,N], e_i^2 = (y_i - \hat{y}_i)^2 = (y_i - b_0 - b_1 x_i)^2$$

• Choose (b_0,b_1) such that $\sum_{i=1}^N e_1^2 + \cdots + e_N^2$ is **as small as possible**.

Mathematics





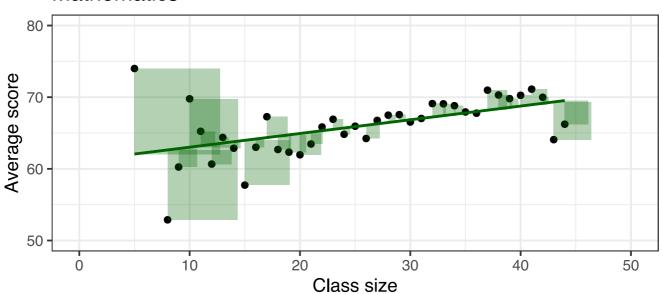
Ordinary Least Squares (OLS) Estimation

• Errors of different sign (+/-) cancel out, so let's consider **squared residuals**

$$orall i \in [1,N], e_i^2 = (y_i - \hat{y}_i)^2 = (y_i - b_0 - b_1 x_i)^2$$

• Choose (b_0,b_1) such that $\sum_{i=1}^N e_1^2 + \cdots + e_N^2$ is **as small as possible**.

Mathematics





App Time! #2 (3 minutes)

Let's minimize some squared errors!

```
launchApp('reg_simple')
aboutApp('reg_simple')
```



- OLS: estimation method consisting in minimizing the sum of squared residuals.
- Yields **unique** solutions to this minization problem.
- So what are the formulas for b_0 (intercept) and b_1 (slope)?



- OLS: estimation method consisting in minimizing the sum of squared residuals.
- Yields unique solutions to this minization problem.
- So what are the formulas for b_0 (intercept) and b_1 (slope)?
- In our single independent variable case:

Slope:
$$b_1^{OLS}=rac{cov(x,y)}{var(x)}$$
 Intercept: $b_0^{OLS}=ar{y}-b_1ar{x}$



- OLS: estimation method consisting in minimizing the sum of squared residuals.
- Yields **unique** solutions to this minization problem.
- So what are the formulas for b_0 (intercept) and b_1 (slope)?
- In our single independent variable case:

Slope:
$$b_1^{OLS}=rac{cov(x,y)}{var(x)}$$
 Intercept: $b_0^{OLS}=ar{y}-b_1ar{x}$

ullet You should know these formulas, especially the one for b_1^{OLS} !



- OLS: estimation method consisting in minimizing the sum of squared residuals.
- Yields **unique** solutions to this minization problem.
- So what are the formulas for b_0 (intercept) and b_1 (slope)?
- In our single independent variable case:

Slope:
$$b_1^{OLS}=rac{cov(x,y)}{var(x)}$$
 Intercept: $b_0^{OLS}=ar{y}-b_1ar{x}$

- ullet You should know these formulas, especially the one for b_1^{OLS} !
- These formulas do not appear from magic. They can be found by solving the minimisation of squared errors. The maths can be found here for those who are interested.



App Time! #3 (3 minutes)

How does OLS actually perform the minimization problem? Some intuition without maths.

```
launchApp('SSR_cone')
aboutApp('SSR_cone') # after
```



For now assume both the dependent variable (y) and the independent variable (x) are numeric.



For now assume both the dependent variable (y) and the independent variable (x) are numeric.

Intercept (b_0) : The predicted value of $y(\hat{y})$ if x=0.



For now assume both the dependent variable (y) and the independent variable (x) are numeric.

```
Intercept (b_0): The predicted value of y(\hat{y}) if x=0.
```

Slope (b_1) : The predicted change, on average, in the value of y associated to a one-unit increase in x.



For now assume both the dependent variable (y) and the independent variable (x) are numeric.

```
Intercept (b_0): The predicted value of y (\hat{y}) if x=0.
Slope (b_1): The predicted change, on average, in the value of y associated to a one-unit increase in x.
```

• A Note that we use the term *associated*, **clearly avoiding interpreting** b_1 **as the causal impact of** x **on** y. To make such a claim, we need some specific conditions to be met. (Next week!)



For now assume both the dependent variable (y) and the independent variable (x) are numeric.

```
Intercept (b_0): The predicted value of y (\hat{y}) if x=0.
Slope (b_1): The predicted change, on average, in the value of y associated to a one-unit increase in x.
```

- A Note that we use the term *associated*, **clearly avoiding interpreting** b_1 **as the causal impact of** x **on** y. To make such a claim, we need some specific conditions to be met. (Next week!)
- Also notice that the units of x will matter for the interpretation (and magnitude!) of b_1 .



OLS with R

- In R, OLS regressions are estimated using the 1m function.
- This is how it works:

lm(formula = dependent variable ~ independent variable, data = data.frame containing the data)



OLS with R

- In R, OLS regressions are estimated using the 1m function.
- This is how it works:

```
lm(formula = dependent variable ~ independent variable, data = data.frame containing the data)
```

Class size and student performance

Let's estimate the following model by OLS: $\operatorname{avgmath_cs}_i = b_0 + b_1 \operatorname{classsize}_i + e_i$

```
# OLS regression of class size on average maths score
lm(avgmath_cs ~ classize, grades_avg_cs)

##
## Call:
## lm(formula = avgmath_cs ~ classize, data = grades_avg_cs)
##
## Coefficients:
## (Intercept) classize
## 61.1092 0.1913
```



Ordinary Least Squares (OLS): Prediction

```
##
## Call:
## lm(formula = avgmath_cs ~ classize, data = grades_avg_cs)
##
## Coefficients:
## (Intercept) classize
## 61.1092 0.1913
```



Ordinary Least Squares (OLS): Prediction

```
##
## Call:
## lm(formula = avgmath_cs ~ classize, data = grades_avg_cs)
##
## Coefficients:
## (Intercept) classize
## 61.1092 0.1913
```

This implies (abstracting the *i* subscript for simplicity):

$$\hat{y} = b_0 + b_1 x$$
 $ext{avgmath_cs} = b_0 + b_1 \cdot ext{classize}$ $ext{avgmath_cs} = 61.11 + 0.19 \cdot ext{classize}$



Ordinary Least Squares (OLS): Prediction

```
##
## Call:
## lm(formula = avgmath_cs ~ classize, data = grades_avg_cs)
##
## Coefficients:
## (Intercept) classize
## 61.1092 0.1913
```

This implies (abstracting the *i* subscript for simplicity):

$$\hat{y} = b_0 + b_1 x$$
 $ext{avgmath_cs} = b_0 + b_1 \cdot ext{classize}$ $ext{avgmath_cs} = 61.11 + 0.19 \cdot ext{classize}$

What's the predicted average score for a class of 26 students? (Using the exact coefficients.)



Task 3: OLS Regression (7 minutes)

- 1. Compute the OLS coefficients using the formulas on slide 28.
- 2. Regress class size (independant variable) on average verbal score (dependent variable).
- 3. Is the slope coefficient similar to the one found for average math score? Was it expected based on the graphical evidence?
- 4. What is the predicted average verbal score when class size is equal to 0? (Does that even make sense?!)
- 5. What is the predicted average verbal score when the class size is equal to 30 students?



OLS variations / restrictions

- All are described in the book. Optional ...
- There is an app for each of them:

type	App	
No Intercept, No regressors	<pre>launchApp('reg_constrained')</pre>	
Centered Regression	<pre>launchApp('demeaned_reg')</pre>	
Standardized Regression	<pre>launchApp('reg_standardized')</pre>	



Predictions and Residuals: Properties

• The average of \hat{y}_i is equal to \bar{y} .

$$egin{aligned} rac{1}{N} \sum_{i=1}^{N} \hat{y}_i &= rac{1}{N} \sum_{i=1}^{N} b_0 + b_1 x_i \ &= b_0 + b_1 ar{x} = ar{y} \end{aligned}$$

• The average (or sum) of errors is 0.

$$egin{aligned} rac{1}{N} \sum_{i=1}^{N} e_i &= rac{1}{N} \sum_{i=1}^{N} (y_i - \hat{y}_i) \ &= ar{y} - rac{1}{N} \sum_{i=1}^{N} \hat{y}_i \ &= 0 \end{aligned}$$



Predictions and Residuals: Properties

• The average of \hat{y}_i is equal to \bar{y} .

$$egin{aligned} rac{1}{N} \sum_{i=1}^{N} \hat{y}_i &= rac{1}{N} \sum_{i=1}^{N} b_0 + b_1 x_i \ &= b_0 + b_1 ar{x} = ar{y} \end{aligned}$$

• The average (or sum) of errors is 0.

$$egin{align} rac{1}{N} \sum_{i=1}^{N} e_i &= rac{1}{N} \sum_{i=1}^{N} (y_i - \hat{y}_i) \ &= ar{y} - rac{1}{N} \sum_{i=1}^{N} \hat{y}_i \ &= 0 \end{split}$$

 Regressor and errors are uncorelated (by definition).

$$Cov(x_i, e_i) = 0$$

Prediction and errors are uncorrelated.

$$egin{aligned} Cov(\hat{y}_i,e_i) &= Cov(b_0+b_1x_i,e_i) \ &= b_1Cov(x_i,e_i) \ &= 0 \end{aligned}$$

Since
$$Cov(a + bx, y) = bCov(x, y)$$
.



Linearity Assumption: Visualize your Data!

- It's important to keep in mind that covariance, correlation and simple OLS regression only measure **linear relationships** between two variables.
- Two datasets with *identical* correlations and regression lines could look *vastly* different.



Linearity Assumption: Visualize your Data!

- It's important to keep in mind that covariance, correlation and simple OLS regression only measure **linear relationships** between two variables.
- Two datasets with *identical* correlations and regression lines could look *vastly* different.

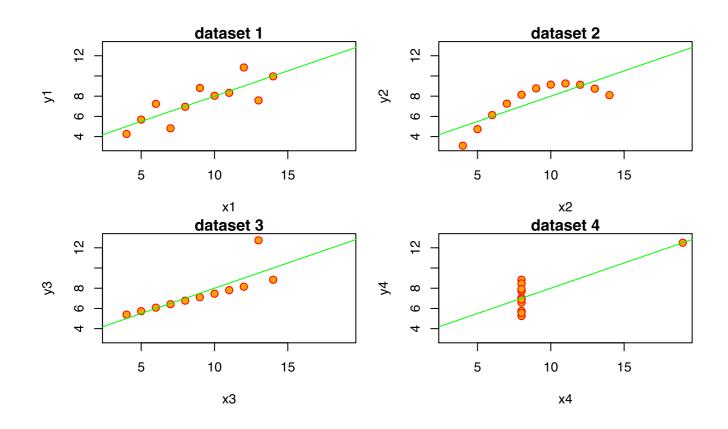


• Is that even possible?



Linearity Assumption: Anscombe

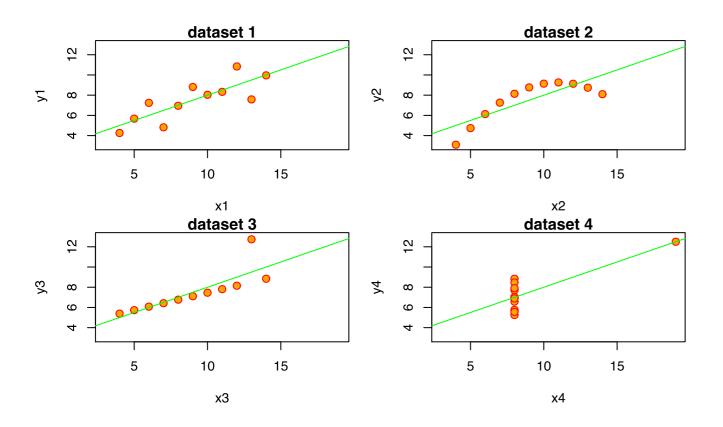
• Francis Anscombe (1973) came up with 4 datasets with identical stats. But look!





Linearity Assumption: Anscombe

• Francis Anscombe (1973) came up with 4 datasets with identical stats. But look!



dataset	cov	var(y)	var(x)
1	5.501	4.127	11
2	5.500	4.128	11
3	5.497	4.123	11
4	5.499	4.123	11



Nonlinear Relationships in Data?

- We can accomodate non-linear relationships in regressions.
- Just add a *higher order* term like this:

$$y_i = b_0 + b_1 x_i + b_2 x_i^2 + e_i$$

• This is **multiple regression** (in 2 weeks!)



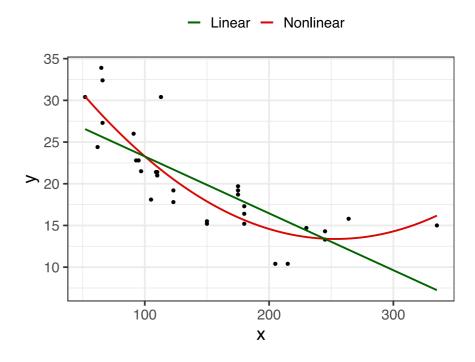
Nonlinear Relationships in Data?

- We can accomodate non-linear relationships in regressions.
- Just add a *higher order* term like this:

$$y_i = b_0 + b_1 x_i + b_2 x_i^2 + e_i$$

• This is **multiple regression** (in 2 weeks!)

 For example, suppose we had this data and fit the previous regression model:
 Nonlinear relationship between x and y





Analysis of Variance (ANOVA)

- ullet Remember that $y_i = \hat{y}_i + e_i.$
- We have the following decomposition:

$$egin{aligned} Var(y) &= Var(\hat{y} + e) \ &= Var(\hat{y}) + Var(e) + 2Cov(\hat{y}, e) \ &= Var(\hat{y}) + Var(e) \end{aligned}$$

• Because:

$$egin{array}{ll} \circ \ Var(x+y) = Var(x) + Var(y) + 2Cov(x,y) \ \circ \ Cov(\hat{y},e) = 0 \end{array}$$

• Total variation (SST) = Model explained (SSE) + Unexplained (SSR)

- SST = sum of squares total
- ∘ SSE = sum of squares error
- SSE = sum of squares regresion





$$R^2 = rac{ ext{variance explained}}{ ext{total variance}} = rac{SSE}{SST} = 1 - rac{SSR}{SST} \in [0,1]$$



$$R^2 = rac{ ext{variance explained}}{ ext{total variance}} = rac{SSE}{SST} = 1 - rac{SSR}{SST} \in [0,1]$$

- R^2 close to 1 indicates a **very** *high* **explanatory power** of the model.
- R^2 close to 0 indicates a **very** *low* **explanatory power** of the model.



$$R^2 = rac{ ext{variance explained}}{ ext{total variance}} = rac{SSE}{SST} = 1 - rac{SSR}{SST} \in [0,1]$$

- R^2 close to 1 indicates a very high explanatory power of the model.
- R^2 close to 0 indicates a **very** low **explanatory** power of the model.
- Interpretation: an \mathbb{R}^2 of 0.5, for example, means that the variation in x "explains" 50% of the variation in y.



$$R^2 = rac{ ext{variance explained}}{ ext{total variance}} = rac{SSE}{SST} = 1 - rac{SSR}{SST} \in [0,1]$$

- R^2 close to 1 indicates a very high explanatory power of the model.
- R^2 close to 0 indicates a **very** *low* **explanatory power** of the model.
- Interpretation: an \mathbb{R}^2 of 0.5, for example, means that the variation in x "explains" 50% of the variation in y.
- \triangle Low R^2 does **NOT** mean it's a useless model! Remember that econometrics is interested in causal mechanisms, not prediction!



Task 4: \mathbb{R}^2 and goodness of fit (10 minutes)

- 1. Regress classize on avgmath_cs. Assign to an object math_reg.
- 2. Pass $math_{reg}$ in the summary() function. What is the (multiple) \mathbb{R}^2 for this regression? How can you interpret it?
- 3. Compute the squared correlation between classize and avgmath_cs. What does this tell you of the relationship between \mathbb{R}^2 and the correlation in a regression with only one regressor?
- 4. Install and load the broom package. Pass math_reg in the broom::augment() function and assign it to a new object. Use the variance in avgmath_cs (SST) and the variance in .fitted (predicted values; SSE) to find the R^2 using the formula in the previous slide.
- 5. Repeat steps 1 and 2 for avgverb_cs. For which exam does the variance in class size explain more of the variance in students' scores?





SEE YOU NEXT WEEK!

- michele.fioretti@sciencespo.fr
- **%** Slides
- % Book
- @ScPoEcon
- @ScPoEcon

