

State-Space Stock Assessment Models

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Hidden process models

- Separate stochastic nature of hidden variables of interest from that of observations.
- Generalizes time series models and models for independent observations.
- Includes all types of mixed effects models, time series, and state-space models as special cases.
- Newman et al. (2006) is a good introduction with a focus on population models.

Hidden process models

- The hidden process $\theta_1, \dots, \theta_T$ evolves stochastically over time.
- The probability model for the hidden process θ ,

$$P(\theta_t) = f(\theta_1, \dots, \theta_{t-1})$$

is equivalent to a statistical time series model (e.g., ARIMA), that may or may not be linear and Gaussian.

- The probability model for observations

$$P(x_t) = f(\theta_t)$$

generally treat the observations at time t , conditional on the state of the hidden process, as independent.

State-space models

- For state-space models, probability of current state can be written as a function of the state at the previous time step

$$P(\theta_t) = f(\theta_{t-1})$$

- The Kalman Filter (Kalman 1960) to estimate linear Guassian state-space models was a major breakthrough.
- A number of previous applications of state-space models in fisheries spanning decades.
 - Biomass Dynamics/Production models (Millar and Meyer 2000)
 - Delay-difference models (Meyer and Millar 1999)
 - Length-based models (Sullivan et al. 1990)
 - Age-structured models (Mendelssohn 1988)
- But applications in management are relatively recent (Nielsen and Berg 2014)

Traditional Population models

We model population at time t as a function of population at time $t - 1$.

- For a (difference equation) Shaefer model

$$B_t = B_{t-1}r \left(1 - \frac{B_{t-1}}{K} \right) - C_{t-1}$$

- For an age-structured model

$$N_{t,a} = N_{t-1,a-1}e^{-Z_{t,a}}$$

where $Z_{t,a} = F_{t,a} + M_{t,a}$

Traditional Population models

All stochasticity is in the observation equations

$$P(\log I_t) = N \left(\log \hat{I}_t, \sigma_t^2 \right)$$

- For a Shaefer model

$$\hat{I}_t = qB_t$$

- For an age-structured model

$$\hat{I}_{t,a} = q_a N_{t,a} e^{-Z_{t,a} f}, \quad \hat{I}_t = \sum_{a=1}^A \hat{I}_{t,a}$$

There are other probability models for age composition observations and catch for statistical catch at age models.

State-space versions of population models

Include stochasticity in the evolution of the population over time so population size is the hidden state variable θ

- For a Shaefer model, the expected biomass is the same

$$\hat{B}_t = B_{t-1}r \left(1 - \frac{B_{t-1}}{K}\right) - C_{t-1}$$

The realized biomass is a stochastic outcome. For example,

$$P(\theta_t | \theta_{t-1}) = P(B_t | B_{t-1}) = \mathcal{N}(\log \hat{B}_t, \sigma_B^2)$$

- For an age-structured model, the expected abundance at age is the same

$$\hat{N}_{t,a} = N_{t-1,a-1}e^{-Z_{t-1,a-1}}$$

and $\theta_t = N_{t,1}, \dots, N_{t,A}$ is a vector. For example,

$$P(\theta_{t,a} | \theta_{t-1,a-1}) = P(N_{t,a} | N_{t-1,a-1}) = \mathcal{N}(\log \hat{N}_{t,a}, \sigma_N^2)$$

State-space probability models

Let D_t be the observations at time t , and modeling process from time $t = 1$ to $t = T$

- For a Shaefer model, the joint probability of the process and observations is

$$P(\boldsymbol{\theta}, \mathbf{D}) = P(B_1, \dots, B_T, D_1, \dots, D_T) = \prod_{t=1}^T P(D_t | B_t) P(B_t | B_{t-1})$$

- For an age-structured model

$$\begin{aligned} P(\boldsymbol{\theta}, \mathbf{D}) &= P(N_{1,a}, \dots, N_{T,A}, D_1, \dots, D_T) \\ &= \prod_{t=1}^T P(D_t | N_{t,1}, \dots, N_{t,A}) \prod_{a=1}^A P(N_{t,a} | N_{t-1,a-1}) \end{aligned}$$

- There are various ways of dealing with the initial state.

Estimating state-space models

- Maximum likelihood
 - The joint likelihood is integrated over (random) state variables yielding the marginal likelihood.
 - Maximize the marginal loglikelihood

$$\mathcal{L}(\alpha|D) = P(D|\alpha) = \int P(\theta, D|\alpha) d\theta$$

with respect to parameters α (e.g., K , r , fishing mortality, catchability, variance parameters)

- Different methods can be used to maximize the log-likelihood. E.g., Laplace approximation, importance sampling.
- Various filters (e.g., Kalman) can also be used likelihood calculations, depending on linearity and probability distributions.
- States estimates are still Bayesian because they were random variables to begin with.

$$P(\theta|D, \hat{\alpha}) = \frac{P(\theta, D|\hat{\alpha})}{P(D|\hat{\alpha})}$$

Estimating state-space models

- Full Bayesian

- Specify appropriate priors for all parameters α

$$P(\theta, \alpha | D) = \frac{P(\theta, \alpha, D)}{P(D)} \propto P(\theta, D | \alpha) P(\alpha)$$

- Use your favorite Markov Chain Monte Carlo method to obtain pseudo-independent samples from the posterior distribution.
 - Pretty simple in principle, but results can be sensitive to choice of prior distributions for parameters.
 - Ensuring estimates are appropriate for a given model requires attention by the analyst.
 - Care must be taken in evaluating independence of samples from the posterior.
 - Can be more computationally intensive and take longer to obtain results.

Recent work on incorporating environmental data

State-space age-structured model

More specifics on state-space implementation of age-structured models

- Recruitment is a stochastic outcome

$$\log N_{t,1} | \text{SSB}_{t-1} \sim \text{N} (g(\theta), \sigma_{N_1}^2)$$

where expected log-recruitment might be just an average over the time series

$$g(\theta) = \log(\bar{R})$$

or related to SSB (e.g., Beverton-Holt)

$$g(\theta) = \log \left[\frac{a \text{SSB}_{t-1}}{1 + b \text{SSB}_{t-1}} \right].$$

State-space age-structured model

- Log-numbers surviving is normally distributed with expectations defined by Baranov equations.

$$\log N_{t+1,a} | N_{t,a-1} \sim N(\log N_{t,a-1} - Z_{t,a-1}, \sigma_{N_a})$$

because $\hat{N}_{t+1,a} = N_{t,a-1} e^{-Z_{t,a-1}}$.

- For the plus group $a = A$, The expected abundance is the sum of those surviving in the plus group and those of the younger age class surviving

$$\hat{N}_{t,A} = N_{t,a-1} e^{-Z_{t,a-1}} + N_{t,A} e^{-Z_{t,A}}$$

so,

$$\log N_{t+1,A} | N_{t,a-1}, N_{t,A} \sim N(\log(N_{t,a-1} e^{-Z_{t,a-1}} + N_{t,A} e^{-Z_{t,A}}), \sigma_{N_A}).$$

- Initial numbers-at-age $N_{1,1}, \dots, N_{1,A}$ are estimated parameters.

State-space assessment model with environmental effects

The environmental covariate(s) can be viewed as a component of the state-space

$$x_{t+1}|x_t \sim N(f(x_t), \sigma_x^2)$$

(a simple random walk in this case) and demographic parameters can be reparameterized as functions of the environment state. For recruitment we considered two types of effects,

$$g(\theta, x_{t-1}) = \log(\bar{R}) + \beta_R x_{t-1}$$

and

$$g(\theta, x_{t-1}) = \log \left[\frac{aSSB_{t-1}}{1 + (e^{b_0 + \beta_R x_{t-1}}) SSB_{t-1}} \right].$$

State-space assessment model with environmental effects

- Observations of the environmental covariates are normally distributed

$$y_t | X_t \sim N \left(X_t, \sigma_{y,t}^2 \right)$$

- aggregate log-catch observations are normally distributed

$$\log C_t | \mathbf{N}_t \sim N \left[f(\mathbf{N}_t, \mathbf{F}_t, \mathbf{M}_t), \sigma_{C,t}^2 \right]$$

where

$$f(\mathbf{N}_t, \mathbf{F}_t, \mathbf{M}_t) = \log \left[\sum_{a=1}^A N_{a,t} W_{a,t} \frac{F_{a,t}}{F_{a,t} + M_{a,t}} \left(1 - e^{-F_{a,t} - M_{a,t}} \right) \right]$$

State-space assessment model with environmental effects

- aggregate log-survey observations are normally distributed

$$\log I_t | N_t \sim N \left[\log \left(\sum_{a=1}^A Q_a N_{a,t} e^{-(F_{a,t} + M_{a,t}) \delta_{I,t}} \right), \sigma_{I,t}^2 \right]$$

- catch and survey age composition are logistic normal distributed

Maximum Likelihood Estimation

The parameters θ estimated by maximum likelihood include

- variance parameters for state transitions and age composition observations,
- annual fishing mortalities and selectivities
- survey catchability and selectivities, and
- mean recruitment parameters.

We used AD Model Builder to maximize the Laplace approximation of the marginal likelihood.

- MSY-based reference points also estimated as functions of environmental covariates with uncertainty.
- Projection years also included in likelihood to estimate future state of environment and stock and propagate uncertainties.

Southern New England yellowtail flounder and Mid-Atlantic Cold Pool

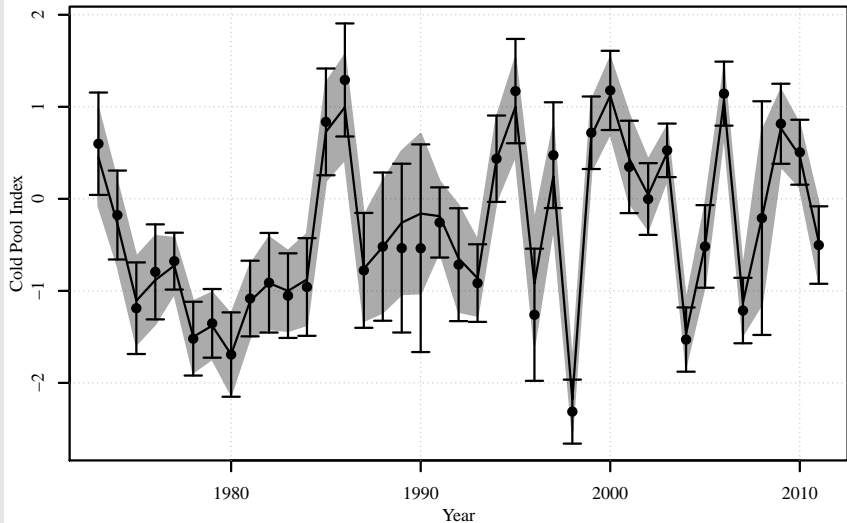
- Use the same data inputs from most recent benchmark assessment.
- Some simplifying assumptions about fishery and survey selectivity to reduce parameters.
- Mid-Atlantic Bight cold pool is thought to affect recruitment.
- Fit 4 models with the different recruitment assumptions with and without environmental effects and compared performance using AIC and Mohn's ρ .

Model Performance

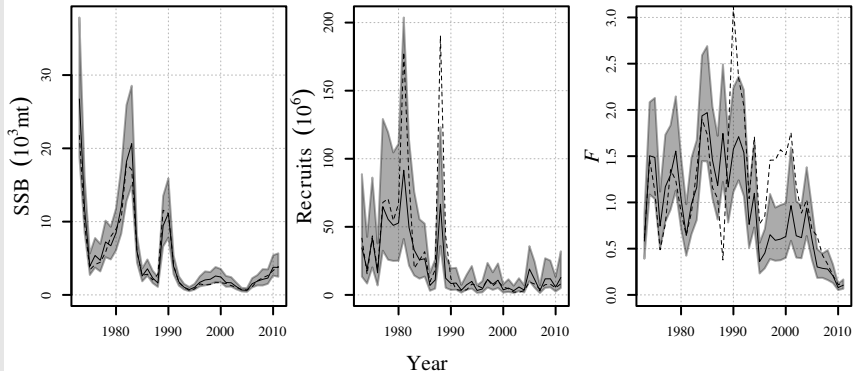
Model	$-\log(L)$	n_p	AIC	$\rho(SSB)$	$\rho(F)$	$\rho(R)$
R	-798.748	65	-1467.496	0.112	-0.145	0.266
$R(\text{CP})$	-805.714	66	-1479.428	0.115	-0.146	0.236
$R(\text{SSB})$	-804.080	66	-1476.160	0.106	-0.136	0.238
$R(\text{SSB}, \text{CP})$	-809.327	67	-1484.654	0.109	-0.139	0.219

Peel	R	$R(\text{CP})$	$R(\text{SSB})$	$R(\text{SSB}, \text{CP})$
1	16.83	4.80	8.69	0
2	15.94	4.65	7.96	0
3	15.79	4.22	8.29	0
4	14.83	3.74	7.73	0
5	12.42	2.84	6.79	0
6	13.34	4.04	5.45	0
7	15.16	6.61	3.51	0

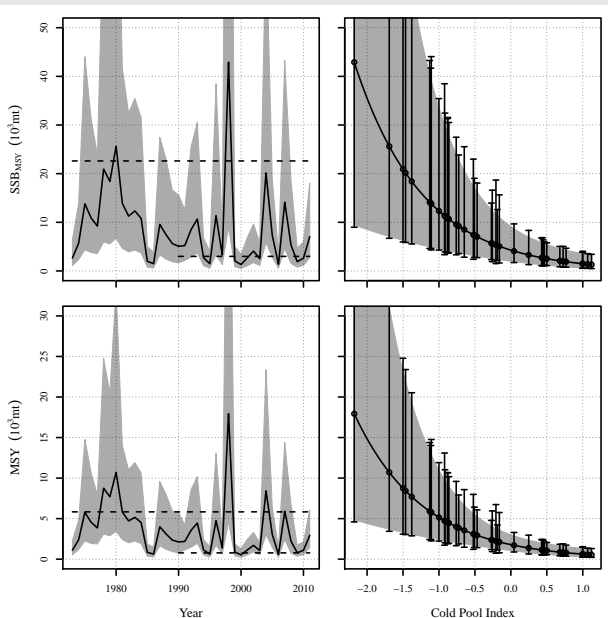
Estimated Cold Pool



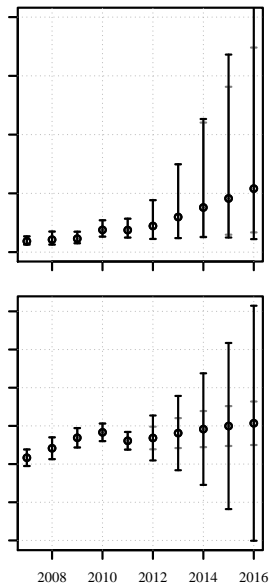
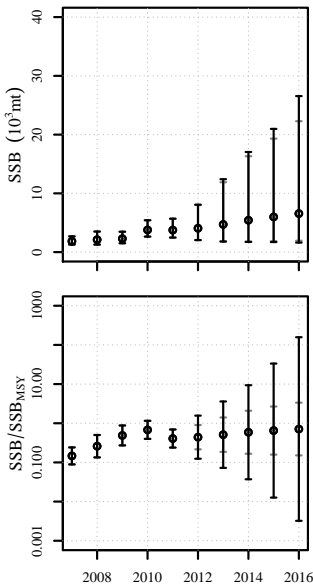
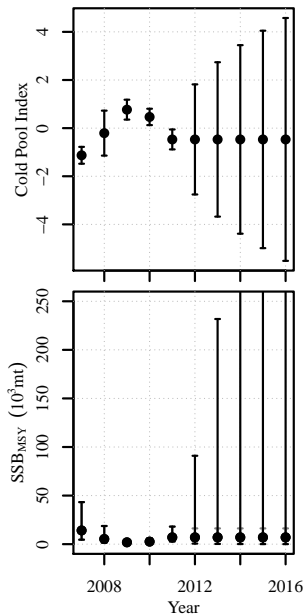
Spawning Stock Biomass, Recruitment, and F



Cold Pool effects on Biological Reference Points



Projections



Year

Last Slide

- Ability to statistically compare models with different recruitment assumptions is an important feature of state space approach.
- Treating environmental covariate as a component of the state space
 - accounts for sampling error of environmental observations,
 - draws attention to the need to better understand error in these observations,
 - propagates uncertainty in environment into population effects, and
 - allows predictions of missing observations based on time-series model.
- The state-space approach also provides a convenient tool for error propagation in forecasting population.

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