# Control charts methodology

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#### **Definitions**

- $n_i$  Number of observations of the ith sample batch (i = 1...m)
- m Number of batches of samples
- p Fraction of unfit products/units/other
- $D_i$  Number of unfit products/units/other in the ith batch (i = 1...m)
- $\hat{x}_i$  Sample mean of the ith batch of observations (i = 1...m)
- $\hat{x}^* = \sum_{i=1}^m \frac{\hat{x}_i}{m}$  Average of sample means
- $\hat{s}_i$  Sample standard deviation of the ith batch of observations (5.29 Montgomery)
- $\hat{r}_i = \max x_i \min x_i$  Range of the *i*-th sample batch
- $\hat{R} = \sum_{i=1}^{m} \frac{\hat{r}_i}{m}$  Average range
- $W = \frac{\hat{R}}{\sigma}$  Relative range

## P chart

Standards given: p is known p = P

Central line (given p = P)

$$CL = P$$

Upper control limit

$$UCL = P + 3\sqrt{P(1-P)/n_i}$$

Lower control limit

$$LCL = P - 3\sqrt{P(1-P)/n_i}$$

Measures of quantity of interest

$$\hat{P}_i = \frac{D_i}{n_i}$$

No standards given: p is unknown

Average fraction of unfit for each batch

$$\hat{p}_i = \frac{D_i}{n_i}$$

P estimator

$$P^* = \frac{\sum_{i=1}^{m} D_i}{\sum_{i=1}^{m} n_i}$$

Central line

$$CL = P^*$$

Upper control limit

$$UCL_i = P^* + 3\sqrt{P^*(1 - P^*)/n_i}$$

Lower control limit

$$LCL_i = P^* - 3\sqrt{P^*(1 - P^*)/n_i}$$

## s chart

Standards given:  $\sigma$  and  $\mu$  are known

Central line

$$CL = c_4 \sigma$$

Upper control limit

$$UCL = c_4\sigma + 3\sigma\sqrt{1 - c_4^2}$$

Lower control limit

$$LCL = c_4\sigma - 3\sigma\sqrt{1 - c_4^2}$$

No standards given:  $\sigma$  is unknown

Central line

$$CL = \hat{s}^* = \sqrt{\frac{\sum_{i=1}^{m} (n_i - 1)\hat{s}_i^2}{\sum_{i=1}^{m} n_i - m}}$$

Upper control limit

$$UCL = \hat{s}^* + 3\frac{\hat{s}^*\sqrt{1 - c_4^2}}{c_4}$$

Lower control limit

$$LCL = \hat{s}^* - 3\frac{\hat{s}^*\sqrt{1 - c_4^2}}{c_4}$$

## R chart

Sample size is assumed constant for the R chart

Central line

$$CL = \hat{R} = \sum_{i=1}^{m} \frac{\hat{r}_i}{m}$$

Upper control limit

$$UCL = \hat{R} + 3d_3 \frac{\hat{R}}{d_2}$$

Lower control limit

$$LCL = \hat{R} - 3d_3\frac{\hat{R}}{d_2}$$

## xbar chart

## r method

Sample size is assumed constant for the r method xbar chart. Should the sample size be different then it is strongly suggested to use the s method xbar chart

Central line

$$CL = x^* = \sum_{i=1}^{m} \frac{\hat{x}_i}{m}$$

Upper control limit

$$UCL = x^* + \frac{3\hat{R}}{d_2\sqrt{n}}$$

Lower control limit

$$LCL = x^* - \frac{3\hat{R}}{d_2\sqrt{n}}$$

## s method

Standards given:  $\sigma$  and  $\mu$  are known

Central line

$$CL = \mu$$

Upper control limit

$$UCL = \mu + 3\frac{\sigma}{\sqrt{n}}$$

Lower control limit

$$LCL = \mu - 3\frac{\sigma}{\sqrt{n}}$$

No standards given:  $\sigma$  and  $\mu$  are unknown

Central line

$$CL = x^* = \frac{\sum_{i=1}^{m} n_i \hat{x}_i}{\sum_{i=1}^{m} n_i}$$

Upper control limit

$$UCL = \hat{x}^* + 3\frac{\hat{\sigma}}{\sqrt{n}} = \hat{x}^* + 3\frac{\hat{s}^*}{c_4\sqrt{n}}$$

Lower control limit

$$LCL = \hat{x}^* - 3\frac{\hat{\sigma}}{\sqrt{n}} = \hat{x}^* - 3\frac{\hat{s}^*}{c_4\sqrt{n}}$$

# I and MR chart (Chart for individual observations)

the moving range MR

$$MR_i = |x_i - x_{i-1}|$$

and the average moving range

$$\hat{MR} = \sum_{i=1}^{n-1} \frac{MR_i}{n-1}$$

then:

#### I chart

Central line

$$CL = \hat{x}$$

Upper control limit

$$UCL = \hat{x} + \frac{3\hat{MR}}{d_2}$$

Lower control limit

$$LCL = \hat{x} - \frac{3\hat{MR}}{d_2}$$

Note:  $d_2$  depends on the moving range size (e.g. if a moving range of n=2 observations is used, then  $d_2=d_2(2)$ )

#### MR chart

Central line

$$CL = \hat{MR}$$

Upper control limit

$$UCL = D_4 \hat{MR}$$

Lower control limit

$$LCL = D_3 \hat{MR}$$

Note: if a moving range of n=2 observations is used, then  $d_2=d_2(2)$ 

# c chart (Control Chart for Nonconformities)

Constant sample size is assumed for a c chart.

Standards given: c is known (c is the parameter of the Poisson distribution)

Central line

$$CL = c$$

Upper control limit

$$UCL = c + 3\sqrt{c}$$

Lower control limit

$$LCL = max(c - 3\sqrt{c}, 0)$$

#### Standards not given: c is unknown

Central line

$$CL = \hat{c} = \frac{\sum_{i=1}^{m} D_i}{m}$$

Upper control limit

$$UCL = \hat{c} + 3\sqrt{\hat{c}}$$

Lower control limit

$$LCL = \hat{c} - 3\sqrt{\hat{c}}$$

# u chart (Control Chart for Average Number of Nonconformities per Unit)

•  $u_i = \frac{D_i}{n_i}$  is the total number of nonconformities in a sample of  $n_i$  inspection units.

Then,

Center line

$$CL = \hat{u} = \frac{\sum_{i=1}^{m} D_i}{\sum_{i=1}^{m} n_i}$$

Upper control limit

$$UCL = \hat{u} + 3\sqrt{\frac{\hat{u}}{n_i}}$$

Lower control limit

$$LCL = \hat{u} - 3\sqrt{\frac{\hat{u}}{n_i}}$$

#### Standardized u chart

#### Preferred choice with variable sample size

Plotted statistic

$$Z_i = \frac{u_i - \hat{u}}{\sqrt{\frac{\hat{u}}{n_i}}}$$

Central line

$$CL = 0$$

Upper control limit

$$UCL = 3$$

Lower control limit

$$LCL = -3$$

# $C_4$ Bias correction

The c4 constant depends only on the number of samples n and is defined as follows:

$$c_4 = \sqrt{\frac{2}{n-1} \frac{\Gamma(\frac{n}{2})}{\Gamma(\frac{n-1}{2})}}$$

where  $\Gamma$  is Euler's Gamma function.

```
#' c4
#' @description Returns bias correction for estimating sigma.
#' @param n number of samples in the current batch
#' @return numeric bias correction
#' @examples
#' biasc2 <-c4(2)</pre>
```

```
#' @export

c4 <- function(n)
{
    return(sqrt(2/(n-1))*gamma(n/2)/gamma((n-1)/2))
}

# Plot of correction for n from 2 to 20
#plot(2:20, sapply(2:20, c4), type='l')</pre>
```

# $d_2$ constant for estimating $\sigma$ given the estimated average range $\hat{R}$

$$d_2(n) = \int_{-\infty}^{\infty} [1 - [1 - F_z(z)]^n - F_z^n(z)] dz$$

where  $F(z) = \Phi(z)$ 

Note: samples are assumed to come from a normal distribution. The formula used can be found at the following  ${\it link}$ 

```
#' d2
#' @description Calculates expected value of W for estimating sigma given the average range
#' @param n number of samples in the current batch
#' @return Expected value of W
#' @examples
#' biasd2 <-d2(2)
#' @export
d2 <- function(n)
    f <- function(z)
        return(1-(1-pnorm(z))^n-pnorm(z)^n)
    }
    d2_value <- integrate(f,lower=-Inf,upper=Inf)$value</pre>
    return(d2_value)
}
# Plot of correction for n from 2 to 20
#plot(2:20, sapply(2:20, d2), type='l')
```

# $d_3$ constant

$$E(W^{2}) = 2 \int_{u}^{\infty} \int_{-\infty}^{v} [1 - [1 - F(u)]^{n} - F^{n}(v) + [F(v) - F(u)]^{n}] du dv$$
$$d_{3} = \sqrt{E(W^{2} - d_{2}^{2})}$$

where  $F(z) = \Phi(z), u = \frac{x}{\sigma}$  and  $v = \frac{y}{\sigma}$ 

```
#' @description Calculates correction constant for estimating sigma from the range
#' @param n number of samples in the current batch
#' @return Correction constant d3
#' @examples
#' d3_correction <-d3(2)
#' @export
d3 <- function (n)
    d2 < - d2(n)
    e <- vector()
    for (i in 1:length(n)) {
        int <- integrate(function(w) {</pre>
            w * (1 - ptukey(w, n[i], Inf))
        }, 0, Inf)
        e <- append(e, sqrt(2 * int[[1]] - (d2[1])^2))
    return(e)
}
```

## Group definition function

Assumption about the data structure. The data in input is assumed to be a data.frame object. Additionally, the data.frame should have the following structure

Note that subsequent measurement of the x and y variable have been stacked on top of each other.

```
r <- samples %% n
   # Total subgroups. r > 0. So if r == 0 m is the number of subgroups
   sg \leftarrow m + sign(r)
   # Debug information. Uncomment to print.
   #v \leftarrow c(samples, n, sg, r)
   \#names(v) \leftarrow c("Total samples", "Group numerosity", "Number of subgroups", "Remaining samples")
   #print(v)
   # We can assign a unique id to each sample and then group the sample by id.
   data <- data %>% mutate(sample_group_id = c(rep(1:m,each=n),rep(sg,r)))
   # List to return
   sample_groups <- list()</pre>
   # Populate the list
   for(i in 1:sg)
   {
      sample_groups[[i]] <- filter(data,sample_group_id==i)</pre>
   }
   return(sample_groups)
}
# Example. Divide 20 samples into subgroups of numerosity 3
data <- data.frame(x = 1:20, y = 21:40)
define_subgroups(data,3)
## [[1]]
## x y sample group id
## 1 1 21
                    1
## 2 2 22
                    1
## 3 3 23
                    1
##
## [[2]]
## x y sample_group_id
## 1 4 24
## 2 5 25
                    2
## 3 6 26
                    2
##
## [[3]]
## x y sample_group_id
## 1 7 27
                    3
## 2 8 28
                    3
## 3 9 29
                    3
## [[4]]
## x y sample_group_id
## 1 10 30
## 2 11 31
```

```
## 3 12 32
##
## [[5]]
## x y sample_group_id
## 1 13 33
## 2 14 34
                        5
## 3 15 35
##
## [[6]]
## x y sample_group_id
## 1 16 36
## 2 17 37
                        6
## 3 18 38
##
## [[7]]
## x y sample_group_id
## 1 19 39
                        7
                        7
## 2 20 40
```

Should the samples be arranged in the following format

```
# x1 x2 x3 x4 x5

#Sample1 1 2 3 NA NA

#Sample2 4 5 6 7 8

#Sample3 NA 9 10 11 12
```

Then the following function could be used:

```
define_subgroups <- function(data)</pre>
   sample_groups <- list()</pre>
   for(i in 1:nrow(data))
   {
       sample_groups[[i]] <- as.numeric(data[i,])</pre>
   }
   return(sample_groups)
}
## Example
d <- data.frame()</pre>
d \leftarrow rbind(d,c(1,2,3,NA,NA))
d <- rbind(d,c(4:8))</pre>
d \leftarrow rbind(d,c(NA,9,10,11,12))
colnames(d) <- c("x1","x2","x3","x4","x5")</pre>
define_subgroups(d)
```

```
## [[1]]
```

```
## [1] 1 2 3 NA NA
##

## [[2]]
## [1] 4 5 6 7 8

##

## [[3]]
## [1] NA 9 10 11 12
```