Supplement to: On Model Averaging Regression Coefficients

setup

```
library(ggplot2)
library(data.table)
library(mvtnorm)
```

"there are, literally, an infinite number of generating models with different combinations of effector ("predictor") variables that could generate precisely the same probabilistic effect (b_1). Here are two"

This header comes from the larger argument that there is not a specific generating model of the data when using probabilistic conditioning because this conditioning makes no assumptions what variables generate the response but only about how the response is related to a specific set of X variables once both are generated and observed.

Here I create fake data generated from different sets of variables but the conditional expectation of the probabilistic effect of ParentSex on the response NCalls conditioned on no other variables is precisely the same.

Simplest example – all variables have unit variance and the conditional response is normally distributed.

```
seed_i <- 1
n <- 10<sup>6</sup>
rho <- 0.5 # the true correlation between X1 and X2
X <- rmvnorm(n, sigma=matrix(c(1,rho,rho,1), nrow=2))</pre>
# Model A - only X1 and U generate the data
# U is the component due to uncorrelated factors that creates "noise"
set.seed(seed_i)
n <- 10<sup>5</sup>
beta_1.A <- 0.4
U <- sqrt(1-beta_1.A^2)*rnorm(n)</pre>
v1 \leftarrow beta 1.A*X[,1] + U
coef(lm(y1 \sim X[,1]))
## (Intercept)
                        X[, 1]
## -0.002056911 0.400308073
# Model B - X1, X2, and U generate the data
set.seed(seed_i)
n <- 10<sup>5</sup>
beta_2.B <- -0.3
beta_1.B <- beta_1.A - rho*beta_2.B
```

```
sigma_exp <- beta_1.8^2 + beta_2.8^2 + 2*beta_1.A*beta_1.B*rho
U <- sqrt(1-(sigma_exp))*rnorm(n)
y2 <- (X%*%c(beta_1.B, beta_2.B))[,1] + U
coef(lm(y2 ~ X[,1]))

## (Intercept) X[, 1]
## -0.001119979 0.400302096</pre>
```

More general example – No variables have unit variance and the conditional response is normally distributed.

 $E(b_1)$ for both models is 0.4. The generating effect β_1 in Model B is 0.55.

```
n <- 10<sup>6</sup>
seed_i <- 1
# create two columns of X that have a correlation of rho
rho <- 0.5
sigma_1 <- 0.5
sigma_2 <- 1.2
sigma.sq_12 <- rho*sigma_1*sigma_2 # cov
Sigma <- matrix(c(sigma_1^2, rho*sigma_1*sigma_2, rho*sigma_1*sigma_2, sigma_2^2), nrow=2)
X <- rmvnorm(n, sigma=Sigma)</pre>
beta 0 <- 10 # add an intercept
# model A without X2, again U is the component due to uncorrelated factors, or "noise"
set.seed(seed_i)
beta 1.A <- 2.2
sigma.sq.explained <- (beta_1.A*sigma_1)^2</pre>
R2 < -0.5
sigma.sq.total <- sigma.sq.explained/R2</pre>
U <- sqrt(sigma.sq.total - sigma.sq.explained)*rnorm(n)</pre>
y1 \leftarrow beta_0 + beta_1.A*X[,1] + U
coef(lm(y1 ~ X[,1]))
## (Intercept)
                     X[, 1]
     10.000052
                   2.200432
# model B - with X2
set.seed(seed_i)
beta_2.B <- -0.7
beta_1.B <- beta_1.A - sigma.sq_12/sigma_1^2*beta_2.B
sigma.sq.explained <- (beta_1.B*sigma_1)^2 + (beta_2.B*sigma_2)^2 +
  2*beta 1.B*beta 2.B*rho*sigma 1*sigma 2
U <- sqrt(sigma.sq.total - sigma.sq.explained)*rnorm(n)</pre>
y2 \leftarrow beta_0 + (X_* c(beta_1.B, beta_2.B))[,1] + U
# summary(lm(y2 ~ X))$r.squared
coef(lm(y2 \sim X[,1]))
## (Intercept)
                     X[, 1]
     10.000143
                   2.200426
\hat{b}_1 for Model A is 2.2004322
```

Generalized example – fake owl data

Here the conditional response of generating model is Poisson. The second generating model is constructed to have the same P(NCalls|ParentSex) as the first generating model by

- 1. setting β_2 to some value. 2. setting $\beta_1^{(B)} = \beta_1^{(A)} - \frac{\sigma_{12}^2}{\sigma_2^2}\beta_2^{(B)}$.
- 3. setting $\beta_0^{(B)} = \beta_0^{(B)} \sigma_{added}$. The subtracted term is the added variance due to the second factor and shifts the intercept on the response scale back to that in model A.

```
seed_i <- sample(1:10^4,1)
n <- 10<sup>6</sup>
z_{code} \leftarrow rep(c(0,1), each=n) # the sex "gene" 1=male
# set up mean for females
# exp_beta_0 <- 175 # Ncalls=175, mean calls in females on response scale
# beta_0 <- log(exp_beta_0) # mean calls in log space
beta_0 \leftarrow 5.1 \# exp(b_0) = 164.02
# set up sex
sex_code <- z_code # i.e. no error mapping gene to sex</pre>
sigma.sex\_code \leftarrow 0.5*sqrt(n*2/(n*2-1)) # sigma(sex\_code)
ParentSex <- ifelse(sex code==0,'F','M')</pre>
# set up correlation between arrival time and sex
beta_z <- 4.5 # 4.5 original effect of sex on arrival time
sigma.arrival.unexplained <- 3
ArrivalTime <- beta_z*z_code + rnorm(n*2, sd=sigma.arrival.unexplained) # females centered
# at zero
sigma.arrival_time <- sqrt(beta_z^2*sigma.sex_code^2 + sigma.arrival.unexplained^2)
# expected correlation between sex and arrival time in log space
rho_12 <- beta_z*sigma.sex_code/(sqrt((beta_z*sigma.sex_code)^2 +
                                          sigma.arrival.unexplained^2))
sigma.sq_12 <- rho_12*sigma.sex_code*sigma.arrival_time # covariance
# this generates expected correlations of .6
# cor(ArrivalTime, sex_code)
# cov(ArrivalTime, sex_code)
# these should generate the same coef of sex
# model A - no arrival time
set.seed(seed_i)
beta_1.A <- -0.04 #-0.04
mu.A <- beta_0 + beta_1.A*sex_code # expected log count
count.A <- rpois(n*2, lambda=exp(mu.A))</pre>
m1 <- glm(count.A ~ ParentSex,
           family=poisson(link = "log"),
           na.action=na.fail)
coef(summary(m1))
```

Estimate Std. Error z value Pr(>|z|)

```
## (Intercept) 5.09982920 7.808815e-05 65308.6156
## ParentSexM -0.03995618 1.115531e-04 -358.1808
exp(predict(m1, newdata=data.frame(ParentSex=c("F", "M"))))
          1
## 163.9939 157.5705
# model B - with arrival time
set.seed(seed i)
beta_2.B <- 0.05 #.05
beta_1.B <- beta_1.A - sigma.sq_12/sigma.sex_code^2*beta_2.B
# expected added variance
sigma.sq.mu.B <- (beta_1.B*sigma.sex_code)^2 + (beta_2.B*sigma.arrival_time)^2 +
 2*beta_1.B*beta_2.B*sigma.sq_12
sigma.sq.mu.A <- (beta_1.A*sigma.sex_code)^2</pre>
sigma.sq.add <- (sigma.sq.mu.B - sigma.sq.mu.A)/2</pre>
#beta_0.B <- beta_0 - 0.011 # shifts to same P(NCalls/ParentSex)</pre>
beta_0.B <- beta_0 - sigma.sq.add # in response space this is dividing by variance
mu.B <- beta_0.B + beta_1.B*sex_code + beta_2.B*ArrivalTime # expected log count
count.B <- rpois(n*2, lambda=exp(mu.B))</pre>
m2 <- glm(count.B ~ ParentSex,
           family=poisson(link = "log"),
           na.action=na.fail)
coef(summary(m2))
                             Std. Error
                  Estimate
                                           z value Pr(>|z|)
## (Intercept) 5.09974609 7.809158e-05 65304.6864
## ParentSexM -0.03981013 1.115538e-04 -356.8692
exp(predict(m2, newdata=data.frame(ParentSex=c("F", "M"))))
          1
## 163.9803 157.5804
# expected
beta_1.B + sigma.sq_12/sigma.sex_code^2*beta_2.B
## [1] -0.04
```