

Elements of Applied Biostatistics using Statistical Modeling

Jeffrey A. Walker

2018-08-28

Contents

1	Statistical Modeling	5
1.1	Statistical modeling with linear models	5
1.2	Model fitting	8
1.3	Statistical modeling vs. Null hypothesis testing	13
1.4	Multilevel models	13
1.5	Linear models versus non-linear models	13
2	Variability and Uncertainty (Standard Deviations and Standard Errors)	15
2.1	The sample standard deviation vs. the standard error of the mean	15
2.2	Using Google Sheets to generate fake data to explore uncertainty	17
2.3	Using R to generate fake data to explore uncertainty	18
	Appendix 1: Getting Started with R	23
2.4	Get your computer ready	23
2.5	Start learning	23
2.6	Getting Data into R	24
2.7	Additional R learning resources	24
2.8	Packages used extensively in this text	24
	Appendix 2: Online Resources for Getting Started with Linear Modeling in R	25

Chapter 1

Statistical Modeling

More cynically, one could also well ask “Why has medicine not adopted frequentist inference, even though everyone presents P -values and hypothesis tests?” My answer is: Because frequentist inference, like Bayesian inference, is not taught. Instead everyone gets taught a misleading pseudo-frequentism: a set of rituals and misinterpretations caricaturing frequentist inference, leading to all kinds of misunderstandings. – Sander Greenland

We use statistics to learn from data with uncertainty. Traditional introductory textbooks in biostatistics implicitly or explicitly train students and researchers to “discover by p -value” using hypothesis tests (appendix xxx). Over the course of many chapters, the student is trained to use something like a dichotomous key to choose the correct “test” for the data at hand, compute a test statistic for their data, compute a p -value based on the test statistic, and compares the p -value to 0.05. Textbooks typically give very little guidance about what can be concluded if $p < 0.05$ or if $p > 0.05$, but many researchers conclude they have “discovered” something if $p < 0.05$ but found “no effect” if $p > 0.05$.

Researchers learn almost nothing useful from a hypothesis test. If we are investigating the effects of an increasingly acidified ocean on coral growth, $p = 0.002$ may be evidence that pH affects growth, but, from everything we know about pH and cell biology, it would be absurd to conclude from any data that ocean acidification does not affect growth. Instead, we want to know the magnitude of the effect and our uncertainty in estimating this magnitude. We can use this magnitude and uncertainty to make predictions about the future of coral reefs, under different scenarios of ocean acidification. We can use the estimated effects and uncertainty to model the consequences of the effects of acidification on coral growth on fish production or carbon cycling.

The “discovery by p -value” strategy, or Null-Hypothesis Significance Testing (NHST), has been criticized by statisticians for many, many decades. Nevertheless, introductory biostatistics textbooks written by both biologists and statisticians continue to organize textbooks around a collection of hypothesis tests, with little emphasis on estimation and uncertainty.

1.1 Statistical modeling with linear models

This textbook is an introduction to the analysis of biological data using a statistical modeling approach. As an introduction, the focus will be linear models and extensions of the linear models including linear mixed models and generalized linear models. Here, I refer to all of these as “linear models” because all are a function of a linear predictor. Linear models are the engine behind many hypothesis tests but the emphasis in statistical modeling is estimation and uncertainty instead of test statistics and p -values. A modeling view of statistics is also more coherent than a dichotomous key strategy.

Box

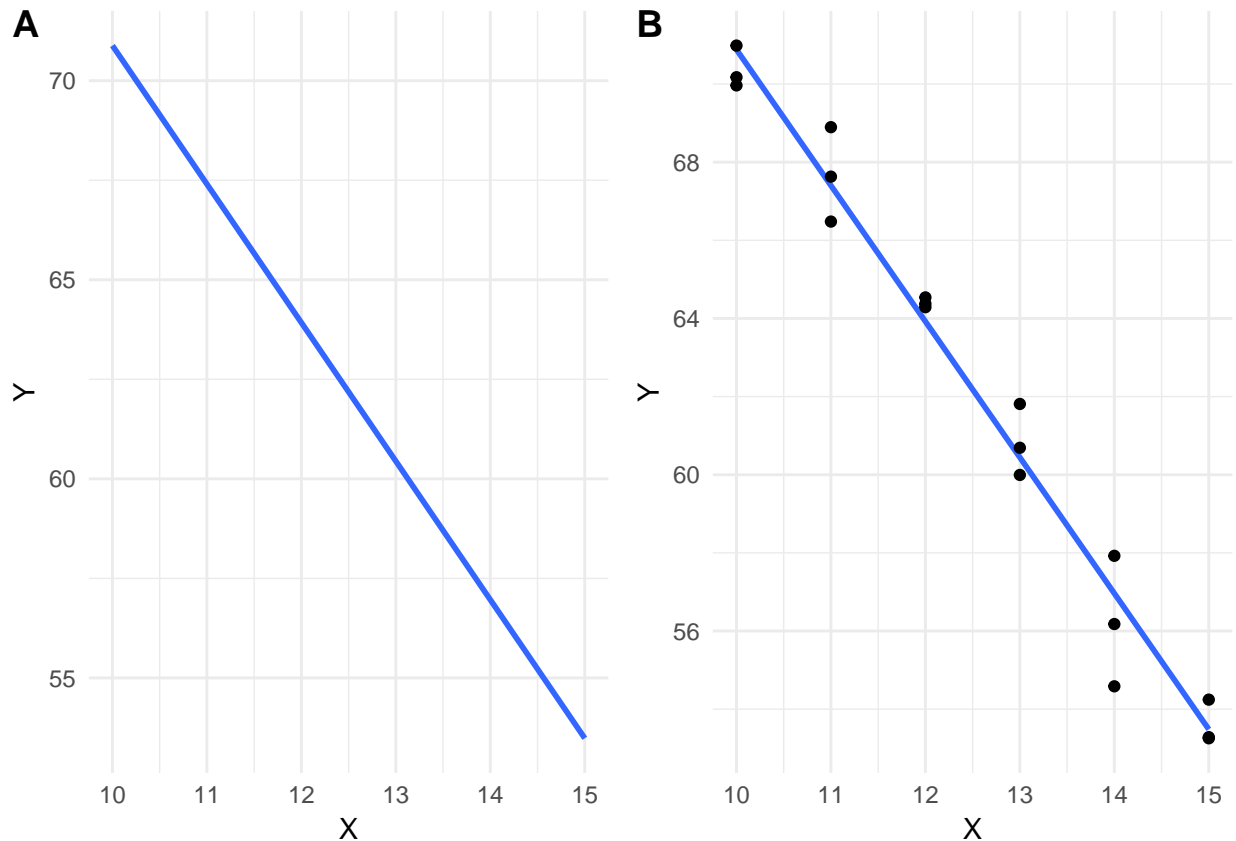


Figure 1.1: A line vs. a linear model. (A) the line $y = -3.48X + 105.7$ is drawn. (B) A linear model fit to the data. The model coefficients are numerically equal to the slope and intercept of the line in A.

linear mixed models are also known as multilevel models and hierarchical models. Generalized linear models (GLMs) are frequently called non-linear models. While it is true that the response (Y) is usually a non-linear function of the X in a GLM, the expected values of Y are a non-linear transformation of a linear predictor function like that in equation (1.2). A common phrase is that GLMs are “linear in the parameters.”

All students are familiar with the idea of a linear model from learning the equation of a line, which is

$$Y = mX + b \quad (1.1)$$

where m is the slope of the line and b is the Y -intercept. It is useful to think of equation (1.1) as a function that maps values of X to values of Y . Using this function, if we input some value of X , we always get the same value of Y as the output.

A linear model is a function, like that in equation (1.1), that is fit to a set of data, often to model a process that generated the data or something like the data. The line in Figure 1.1A is just that, a line, but the line in Figure 1.1B is a model of the data in Figure 1.1B. The basic structure of a linear model is

$$Y = \beta_0 + \beta_1 X + \varepsilon \quad (1.2)$$

A linear model has two parts: the “model” ($Y = \beta_0 + \beta_1 X$) and the “error” (ε). The model part looks like the equation for a line except that I’ve used β_0 for the intercept and β_1 for the slope and I’ve put the intercept term first. This re-labeling and re-arrangement make the notation for a linear model more flexible

for more complicated linear models. For example $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$ is a model where Y is a function of two X variables.

As with the equation for a line, the model part of a linear model is a function that maps a value of X to a specific value of Y . This mapped value is the **expected value** given a specific input value of X . The error part of a linear model is a random variable that adds some random value to this expected value. Nothing about the model part of a linear model can predict its value.

The inputs to a linear model (the X variables) have many names including “independent variables,” “predictor variables,” “explanatory variables,” “treatment variables,” and “covariates”. The output of a linear model (the Y variable or variables if the model is multivariate) is the “dependent variable,” “response,” or “outcome.” The β in the linear model are model **parameters**. There can be additional parameters in more sophisticated models. The coefficients of the X in a linear model (β_1 in model (1.2)) are often called “the effects” (so β_1 is the effect of X_1).

Although a linear model is a model of a data-generating process, linear models are not typically used to actually generate any data. Instead, when we use a linear model to understand something about a real dataset, we think of our data as one realization of a process that generates data like ours. A linear model is a model of that process. That said, it is incredibly useful to use linear models to create fake datasets for at least two reasons: to probe our understanding of statistical modeling generally and, more specifically, to check that a model actually creates data like that in the real dataset that we are analyzing.

1.1.1 Linear models are used for prediction, explanation, and description

Researchers typically use linear models to understand relationships between one or more Y variables and one or more X variables. These relationships include

1. Descriptive modeling. Sometimes a researcher merely wants to describe the relationship between Y and a set of X variables, perhaps to discover patterns. For example, the arrival of a spring migrant bird (Y) as a function of sex (X_1) and age (X_2) might show that males and younger individuals arrive earlier. Importantly, if another X variable is added to the model (or one dropped), the coefficients, and therefore, the precise description, will change. That is, the interpretation of a coefficient as a descriptor is *conditional* on the other covariates (X variables) in the model. In a descriptive model, there is no implication of causal effects and the goal is not prediction. Nevertheless, it is very hard for humans to discuss a descriptive model without using causal language, which probably means that it is hard for us to think of these models as *mere description*. Like natural history, descriptive models are useful as patterns in want of an explanation, using more explicit causal models including experiments.
2. Predictive modeling. Predictive modeling is very common in applied research. For example, fisheries researchers might model the relationship between population density and habitat variables to predict which subset of ponds in a region are most suitable for brook trout (*Salvelinus fontinalis*) reintroduction. The goal is to build a model with minimal prediction error, which is the error between predicted and actual values for a future sample. In predictive modeling, the X (“predictor”) variables are largely instrumental – how these are related to Y is not a goal of the modeling, although sometimes an investigator may be interested in the relative importance among the X for predicting Y (for example, collecting the data may be time consuming, or expensive, or environmentally destructive, so know which subset of X are most important for predicting Y is a useful strategy).
3. Explanatory (causal) modeling. Very often, researchers are explicitly interested in *how* the X variables are causally related to Y . The fisheries researchers that want to reintroduce trout may want to develop and manage a set of ponds to maintain healthy trout populations. This active management requires intervention to change habitat traits in a direction, and with a magnitude, to cause the desired response. This model is predictive – a specific change in X predicts a specific response in Y – because the coefficients of the model provide knowledge on how the system functions – how changes in the inputs *cause* change in the output. Causal interpretation of model coefficients requires a set of strong assumptions about the X variables in the model.

Biologists are often not very explicit about which of these is the goal of the modeling and use a combination of descriptive, predictive, and causal language to describe and discuss results. By contrast, researchers in economics and other social sciences, as well as epidemiology and medicine more generally, are usually very explicit if their model is descriptive, predictive, or causal.

1.2 Model fitting

In order to use a linear model to describe, predict, or explain, we need to fit a model to data. Instead of using an abstract model like that in model (1.2), I will introduce model fitting using data from Dryad Data Repository.

1.2.1 A linear model with a single, continuous X

The data are from ?, who showed that North American red squirrel (*Tamiasciurus hudsonicus*) mothers from Yukon, Alaska produce faster growing pups in years with increased squirrel density. Remarkably, they even showed that perceived (but not actual) density results in faster growing pups. To begin to investigate how pregnant mothers control the future growth rate of pups, the researchers measured the relationship between local squirrel density and the amount of fecal cortisol metabolites from pregnant mothers. Cortisol is a hormone that is secreted as part of stress response. The researchers were interested in cortisol because it had previously been shown that, in mammals, blood cortisol levels in pregnant mothers have numerous effects on offspring long past birth. If increased squirrel density causes increased blood cortisol levels then we would expect to find a positive relationship between *Density* and

Figure 1.2 is a **scatterplot** of the data with the amount of cortisol metabolites in the feces on the Y axis and local squirrel density on the X axis. The line through the data is a graphical representation of a linear model fit to the data and the gray cloud around the line is a graphical representation of the uncertainty in the model. The researchers wanted to model the “effect” of squirrel density on the amount of cortisol metabolites in the feces of the pregnant mothers. Graphically, this effect is the slope of the line in Figure 1.2.

The model is

$$E[FCM|Density] = \beta_0 + \beta_1 Density \quad (1.3)$$

In words, model (1.3) reads “the expected value of FCM conditional on density is beta-knot plus beta-one times density”. An **expected value** is a long run average – if we were to sample lots and lots of red squirrel populations with $Density = x$ (where x is a specific value), we’d expect the average FCM across these samples to be $\beta_0 + \beta_1 x$.

In model (1.3), there is a single X variable (FCM). While the X variables are often called the “dependent” variables, in this model FCM does not “depend” on the independent variable $Density$ in any causal sense – meaning if I were to intervene and set $Density$ to some value x , I would expect FCM to equal $\beta_0 + \beta_1 x$. Rather, FCM only “depends” on $Density$ in a probabilistic sense – if $Density = x$ then the most probable value of FCM is $\beta_0 + \beta_1 x$. With some strong assumptions model (1.3) can be turned into a model of causal dependency, which is the focus of chapter xxx.

β_0 and β_1 are the **parameters** of model (1.3). Specifically β_0 is the model **intercept** and β_1 is the modeled **effect** of $Density$. Again, the effect (β_1) has a probabilistic, and not causal, interpretation. This interpretation is

$$\beta_1 = E[FCM|Density = x + 1] - E[FCM|Density = x] \quad (1.4)$$

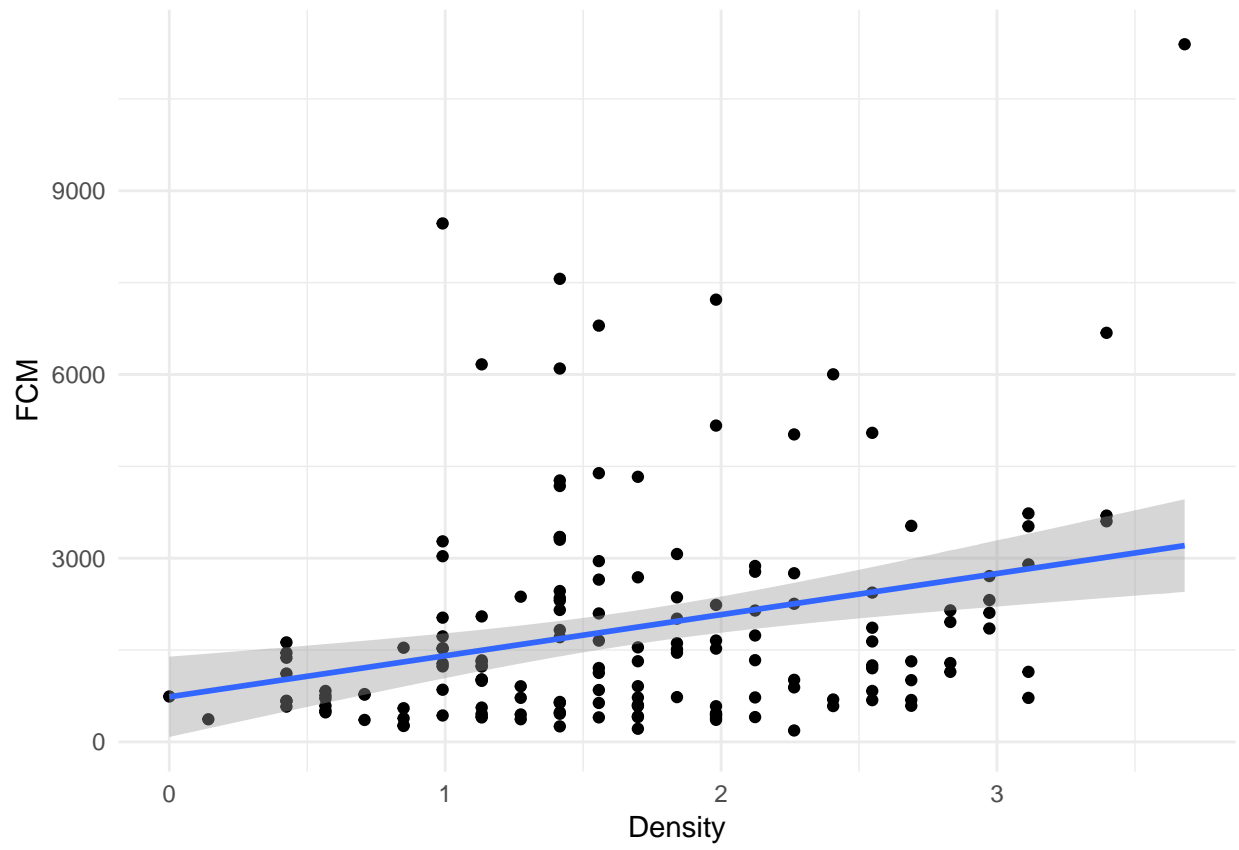


Figure 1.2: A scatterplot of Fecal cortisol metabolites and squirrel density.

Or, in words, “beta-1 is the expected value of FCM when density equals $x + 1$ minus the expected value of FCM when the density equals x .” β_1 is simply the difference in expected values given a one unit difference in *Density*.

1.2.1.1 Using a linear model to estimate effects

The goal of the statistical model here is to estimate β_1 – the probabilistic effect of *Density* on *FCM*. This estimate, and a measure of the uncertainty of this estimate, are in the table of coefficients of the fit model

##	Estimate	Std. Error	t value	Pr(> t)
## (Intercept)	735.9604	331.9395	2.217152	0.0280776547
## Density	671.1380	178.8963	3.751549	0.0002483802

where the entries in the column “Estimate” are estimates of the parameters β_0 and β_1 in model (1.3). The entries in the column “Std. Error” are the standard errors (SE) of the estimates, which are measures of the uncertainty of the estimates.

The parameter estimates in the table above are the coefficients of the fitted model

$$FCM_i = b_0 + b_1 \text{Density}_i + e_i \quad (1.5)$$

where the subscript i refers to the i th individual. The coefficients b_0 and b_1 are the y-intercept and the slope of the line in Figure 1.2. The coefficient for *Density* (b_1) is 671.1, and (given the definition of the parameter β_1 in equation (1.4)) we expect squirrel mothers with a local density of 2 squirrels within a 150 m radius of her midden to average 671.1 more units of FCM (ng of fecal cortisol metabolites per gram dry food) than mother squirrels with a local density of only 1 squirrel within a 150 m radius of her midden. Remember that this coefficient is estimating a probabilistic parameter. Consequently, the coefficient b_1 is simply a descriptor of a pattern of relationship between local density and fecal cortisol metabolites - no causal effect is implied. With the strong assumptions explained in chapter xxx, however, b_1 can estimate a causal effect.

1.2.1.2 Using a linear model for prediction

Model (1.5) gives the measured value of *FCM* for each squirrel. The equation includes the modeled part ($b_0 + b_1 \text{Density}_i$) and the **residual** from the model (e_i). The modeled part is the modeled or **predicted value**,

$$\widehat{FCM} = b_0 + b_1 \text{Density} \quad (1.6)$$

where \widehat{FCM} is read as “FCM hat”. Very often, we use the model part (equation (1.6)) to predict unknown or future values given different modeled inputs (the X).

1.2.2 Linear models with categorical X are the same as linear models with continuous X

Singh et al. (xxx) studied the effect of parasite infection on the production of recombinant offspring in several lines of fruit fly *Drosophila melanogaster*. Recombinant offspring are those with allele combinations that do not occur in either parent.

Figure 1.3, shows the results of one of the experiments, specifically, the recombinant frequencies for each replicate of the **treatment levels** “Smarc” (flies who were parasitized by the bacteria *Serratia marcescens*) and “Wounded” (flies who were given a sterile wound as a control). The mean of each treatment level (or group) is shown with the large black dot within the group’s scatter of individual values, and the difference

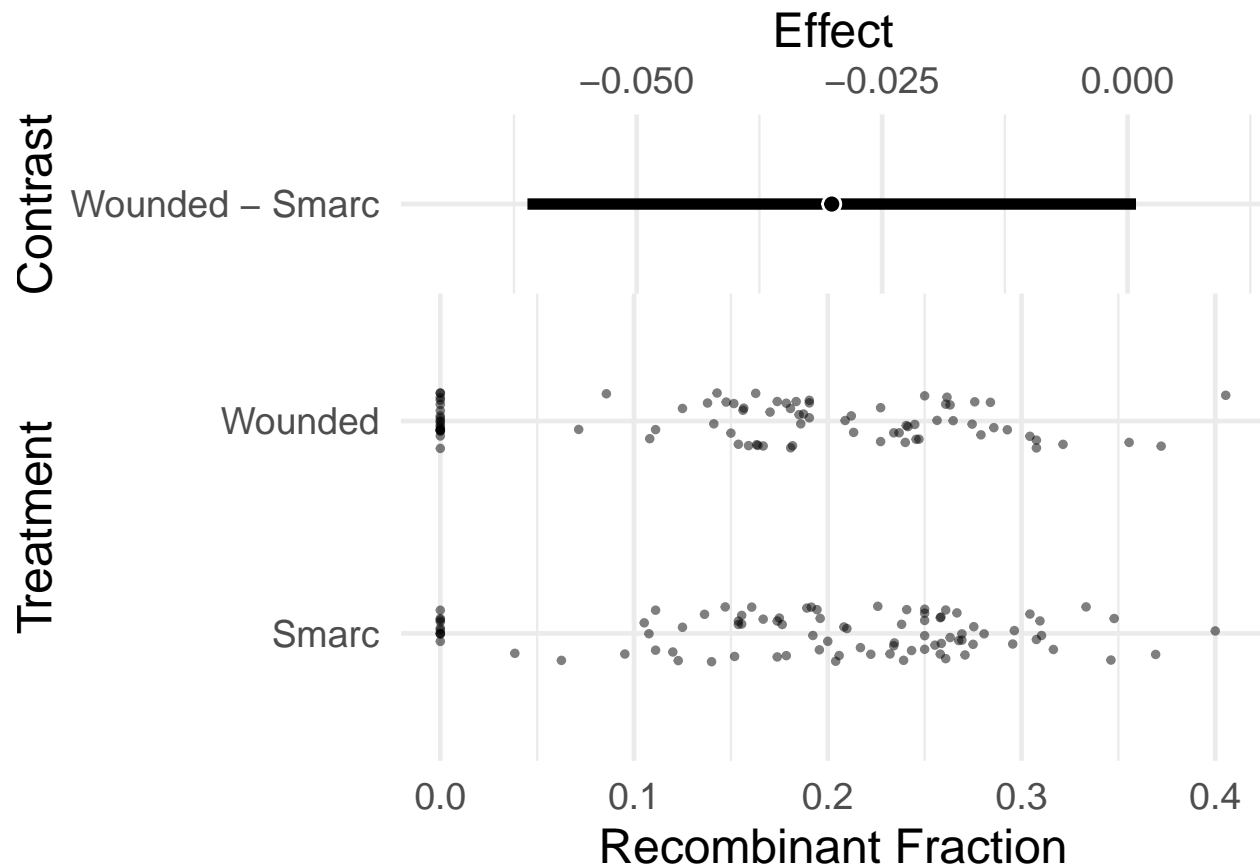


Figure 1.3: Harrell plot of fly data. The bottom part of the graph shows the data while the top part shows the effect estimate and a measure of uncertainty. The specifics of the plot will be explained in Chapter xxx. Briefly, the large black dots within the boxes in the bottom part are the group mean recombination frequencies. The black dot in the top part of the plot is the difference in these group means (the effect).

in the means is shown in the top part with the black dot in the top part. The top plot also shows a measure of the uncertainty in the estimate of this difference (the thick black line).

The means of the two treatment levels (groups) are

```
##      Treatment      mean
## 1:      Smarc 0.1949128
## 2:      Wounded 0.1647802
```

The difference between the means is 0.0301326. This is the estimate of the effect of *S. marcescens* parasitism on recombinant frequency. In general, we wouldn't report these means or this difference in means to this precision because the raw measures are not this precise but I do it here because in order to compare this result to that of analyzing the data with a linear model.

The effect of *Treatment* can be modeled with the linear model

$$E[Recombinant_fraction|Treatment] = \beta_0 + \beta_1 Treatment \quad (1.7)$$

The left side of this equation is read as “the expected recombinant fraction conditional on Treatment” and can be thought of as “the expected recombinant fraction for a specific treatment level is equal to...”. Perhaps surprisingly, this is the same model as that used for the squirrel fecal cortical metabolites, which is, more generally $E[Y|X] = \beta_0 + \beta_1 X$.

What are the estimates of β_0 and β_1 ?

```
##              Estimate Std. Error   t value    Pr(>|t|)
## (Intercept)   0.16478024 0.01129404 14.590016 8.719394e-32
## TreatmentSmarc 0.03013261 0.01570068  1.919191 5.663526e-02
```

Compare the column “Estimate” with the group means and difference in means computed above. The estimate of the intercept is the mean of the Wounded group. The estimate of the *Treatment* coefficient is the difference in means. That is,

In a linear model with a categorical X , the coefficients (b_0 and b_1) are a mean and a difference in means.

The coefficients in the model with categorical X are also an intercept and slope, and this is explored in the problems at the end of the chapter, but it is not especially useful to think of them in this way. In both kinds of models (categorical and continuous X), we are generally less interested in b_0 and more interested in b_1 – the effect of X .

For the recombinant fly experiment, the effect of bacteria infection is 0.03, or 3 additional recombinants per 100 offspring. The SE of this effect (0.016) is our measure of uncertainty. The SE is used to compute the 95% confidence interval in the top part of figure 1.3. The interval contains the range of effects that are consistent with the data. This range includes values up to about 6% and down to about 0%. Importantly, negative values (other than very small ones) are not consistent with the data.

1.2.3 “Statistical model” not “regression model”

Statistical modeling terminology can be confusing. The X variables in a statistical model may be quantitative (continuous or integers) or categorical (names or qualitative amounts) or some mix of the two. Linear models with all quantitative independent variables are often called “regression models.” Linear models with all categorical independent variables are often called “ANOVA models.” Linear models with a mix of quantitative and categorical variables are often called “ANCOVA models” if the focus is on one of the categorical X or “regression models” if there tend to be many independent variables. These names reflect the history of the development of the different kinds of linear models. I advocate using the term “statistical model” for general usage and “linear model” for more specific use, regardless of the combination of variable types.

1.3 Statistical modeling vs. Null hypothesis testing

Most biostatistics textbooks for biologists guide a student/researcher toward the “correct” statistical test for experimental data. The concept of a statistical test of inference is explored more in Appendix xxx but for now, a typical textbook would probably steer a researcher into analyzing the recombinant fly data with a t -test of the difference between means.

```
fly_t <- t.test(Recombinant_fraction ~ Treatment, data=sub_fly, var.equal=TRUE)
fly_t.p <- fly_t$p.value
fly_t.t <- abs(fly_t$statistic)
```

The output of a t -test is a test statistic (t) and a p -value, which, roughly, is the probability of finding a test statistic as large or larger than the observed test statistic if we were to repeat the experiment many, many times using hypothetical data in which there is no effect. This hypothetical data with no effect is the **null hypothesis**.¹ A very small p (say 0.01 or 0.0001) would be unlikely if the null hypothesis were true, and, consequently, a very small p is evidence “against the null” (or evidence that a low-probability event occurred, but we don’t have additional evidence for this so this conclusion is generally dismissed). Again, researchers typically conclude that a difference is “statistically significant” if p is less than 0.05 and, consequently, the treatment has “an effect”.

The t -statistic for the fly recombinant data is 1.9191906 and the p -value is 0.0566353, which means there is a 5.6% probability of finding $t \geq 1.9191906$ if “the null were true”. The p -value of 0.056 is a pretty small probability, and is close to but not smaller than 0.05, and so is not “statistically significant.” How is this reported? If the researchers have an *a priori* hypothesis of an effect, they will often report a p -value of 0.056 as “marginally significant” or worse “trending toward significant” (why not trending “away”?). But if the researchers have an *a priori* hypothesis of no effect, then they will often report a p -value of 0.056 as simply “not significant” or worse “no effect” (see xxx why a p -value is not evidence of no effect).

Most importantly, no part of null hypothesis testing is concerned with estimating the effect size and our uncertainty in this estimate. Test statistics and p -values are not measures of effect size even though each is a function of effect size. This is because each is also a function of sample size and variation. A large t and small p could result from a large effect, or a large sample, or small variability. Null hypothesis testing encourages a focus on the trivia (presence or absence of an effect) instead of the information that we need to model a system. If we want to model the biological consequences of an intervention, such as a drug, or of changing conditions, such as ocean acidification, or if we just want to model relationships within a system, then we need measures of effect size and uncertainty from statistical models.

A p -value can be a useful, but limited tool, and a researcher can use the statistical model to test specific hypotheses if desired. The coefficients of the model are the simplest of these tests. Look at the column “Pr(>|t|)” in the table of coefficients from the recombinant fly experiment above. This column contains the probability of a t -test for each coefficient. The p value for b_1 (the Smarc treatment) is precisely the p -value for the t -test of the means. This is because the math behind a t -test is a special case of the linear model in model (1.7). And the math behind ANOVA is a special case of the linear model. And the math behind regression is a special case of the linear model. In other words, there is little reason to learn these special cases as unrelated tests. There is no reason to teach (or learn) the dichotomous key to the tests of inference.

1.4 Multilevel models

1.5 Linear models versus non-linear models

In this text, I use “linear model” for any model that is linear in the parameters, which means that the different components of the model are added together. Or, using the language of matrix algebra, the predictor is a

¹The effect in the null hypotheses can be any pre-specified value. The nil null (zero effect) is the most common

simple dot product of the model matrix and the coefficients. For example, a cubic polynomial model

$$Y = \beta_0 + \beta_1 X + \beta_2 X^2 + \beta_3 X^3 + \varepsilon \quad (1.8)$$

is a linear model, even though the function is non-linear, because the different components are added (or, using matrix algebra, the predictor is $\mathbf{X}\boldsymbol{\beta}$).

A generalized linear model (GLM) has the form $g(\mu_i) = \eta_i$ where η (the greek letter eta) is the linear predictor, which is linear in the parameters.

$$\eta = \mathbf{X}\boldsymbol{\beta} \quad (1.9)$$

Many sources do not consider a GLM to be a “linear model” but an “extension” of a linear model. Regardless, a GLM is linear in the parameters and in this textbook, I include GLMs under the “linear model” umbrella.

Non-linear models, in contrast to a GLM or classical linear model, are not linear in the parameters (the predictor is not a simple dot product of the model matrix and a vector of parameters). For example, the Michaelis-Menten model is a nonlinear model

$$Y = \frac{\beta_1 X}{\beta_2 + X} + \varepsilon \quad (1.10)$$

Chapter 2

Variability and Uncertainty (Standard Deviations and Standard Errors)

Uncertainty is the stuff of science. A major goal of statistics is measuring uncertainty. What do we mean by uncertainty? Uncertainty is the error in estimating a parameter, such as the mean of a sample, or the difference in means between two experimental treatments, or the predicted response given a certain change in conditions. Uncertainty is measured with a **variance** or its square root, which is a **standard deviation**. The standard deviation of a statistic is also (and more commonly) called a **standard error**.

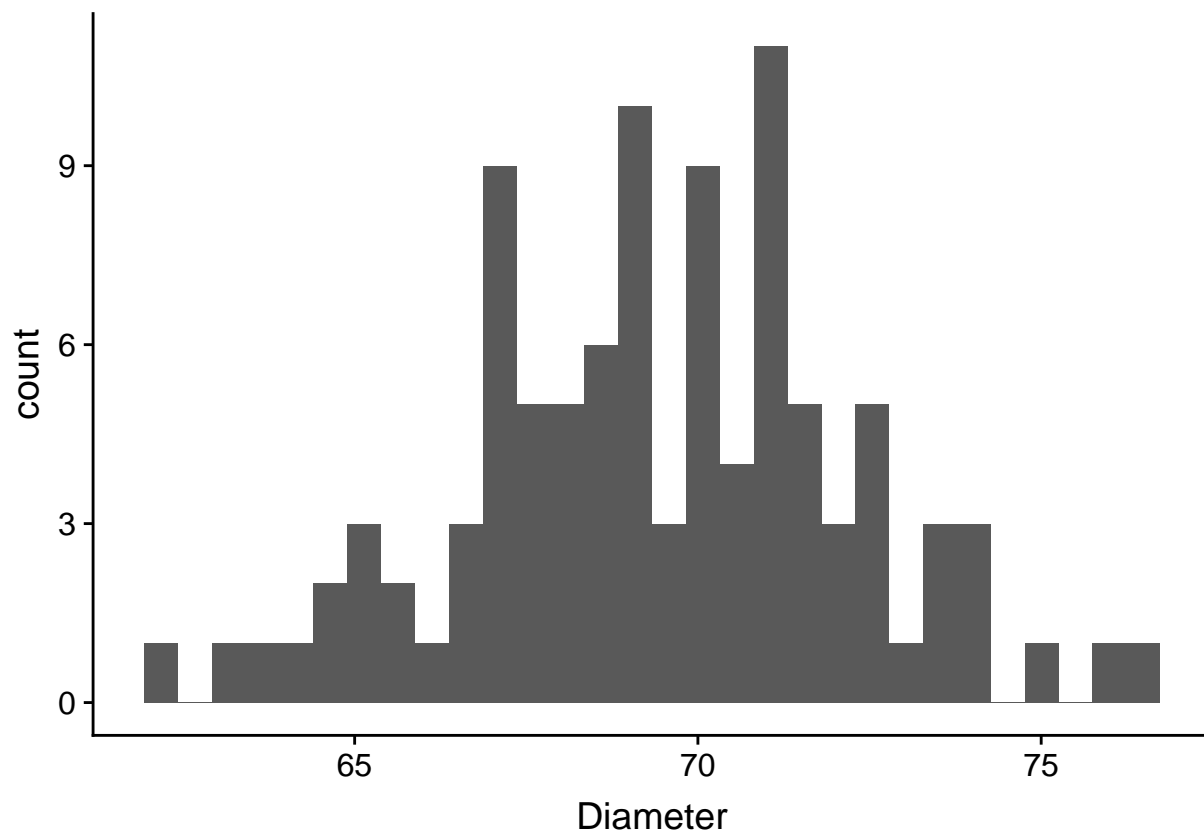
Uncertainty emerges because of variability. In any introductory statistics class, students are introduced to two measures of variability, the “standard deviation” and the “standard error.” These terms are absolutely fundamental to statistics – they are the start of everything else. Yet, many biology professors confuse these terms and certainly, introductory students do too.

When a research biologist uses the term “standard deviation,” they are probably referring to the sample standard deviation which is a measure of the variability of a sample. When a research biologist uses the term “standard error,” they are probably referring to the standard error of a mean, but it could be the standard error of another statistics, such as a regression slope. An important point to remember and understand is that all standard errors *are* standard deviations. This will make more sense soon.

2.1 The sample standard deviation vs. the standard error of the mean

2.1.1 Sample standard deviation

The sample standard deviation is a measure of the variability of a sample. For example, were we to look at a histological section of skeletal muscle we would see that the diameter of the fibers (the muscle cells) is variable. We could use imaging software to measure the diameter of a sample of 100 cells and get a **distribution** like this



The mean of this sample is 69.4 and the standard deviation is 2.8. The standard deviation is the square root of the variance, and so computed by

$$s_y = \sqrt{\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n - 1}} \quad (2.1)$$

Memorize this equation. To understand the logic of this measure of variability, note that $y_i - \bar{y}$ is the deviation of the i th value from the sample mean, so the numerator is the sum of squared deviations. The numerator is a sum over n items and the denominator is $n - 1$ so the variance is (almost!) an averaged squared deviation. More variable samples will have bigger deviations and, therefore, bigger average squared deviations. Since the standard deviation is the square root of the variance, a standard deviation is the square root of the average squared deviation, it can be thought of as a (not “the”) measure of an average deviation.

Notes on the variance and standard deviation 1. Variances are additive but standard deviations are not. This means that the variance of the sum of two independent (uncorrelated) random variables is simply the sum of the variances of each of the variables. This is important for many statistical analyses. 2. The units of variance are the square of the original units, which is awkward for interpretation. The units of a standard deviation is the same as that of the original variable, and so is much easier to interpret. 3. For variables that are approximately normally distributed, we can map the standard deviation to the quantiles of the distribution. For example, 68% of the values are within one standard deviation of the mean, 95% of the values are within two standard deviations, and 99% of the values are within three standard deviations.

2.1.2 Standard error of the mean

A standard error of a statistic is a measure of the precision of the statistic. The standard error of the mean is a measure of the precision of the mean. The smaller the value the more precise the estimate. The standard error of the mean (SEM) is computed as

$$SEM = \frac{s_y}{\sqrt{n}} \quad (2.2)$$

The SEM is often denoted $s_{\bar{y}}$ to indicate that it is a standard deviation of the mean (\bar{y}). In what sense is a standard error a measure of variability? This is kinda weird. If we sample 100 cells in the slide one time and compute the mean diameter, how can the mean have a standard deviation? There is only one value! To understand how the SEM is a standard deviation, imagine resampling 100 cells and recomputing a mean an infinite number of times and each time, you write down the newly computed mean. The standard error of the mean is the standard deviation of this infinitely long column of means.

Notes on standard errors

1. The SEM is only one kind of standard error. A standard deviation can be computed for any statistics – these are all standard errors. For some statistics, such as the mean, the standard error can be computed directly using an equation, such as that for the SEM (equation @ref{eq:se}). For other statistics, a computer intensive method such as the **bootstrap** is necessary to compute a standard error. We will return to the bootstrap shortly.
2. The units of a standard error are the units of the measured variable.
3. A standard error is proportional to sample variability (the sample standard deviation, s_y) and inversely proportional to sample size (n). Sample variability is a function of both natural variation (there really is variation in diameter among fibers in the quadriceps muscle) and measurement error (imaging software with higher resolution can measure a diameter with less error). Since the SEM is a measure of the precision of estimating a mean, this means this precision will increase (or the SEM will decrease) if 1) an investigator uses methods that reduce measurement error and 2) an investigator computes the mean from a larger sample.
4. This last point (the SEM decreases with sample size) seems obvious when looking at equation @ref{eq:se}, since n is in the denominator. Of course n is also in the denominator of equation @ref{eq:variance} for the sample standard deviation (eq) but the standard deviation does not decrease as sample size increases. First this would make any sense – variability is variability. A sample of 10,000 cell diameters should be no more variable than a sample of 100 cell diameters. Second, this should also be obvious from equation @ref{eq:variance}. The standard deviation is the square root of an average and averages don't increase with the number of things measured since both the the numerator (a sum) and denominator increase with n .

2.2 Using Google Sheets to generate fake data to explore uncertainty

In statistics we are interested in estimated parameters of a **population** using measures from a **sample**. The goal in this section is to use Google Sheets (or Microsoft Excel) to use fake data to discover the behavior of sampling and to gain some intuition about uncertainty using standard errors.

2.2.1 Steps

1. Open Google Sheets
2. In cell A1 type “Mu”. Mu is the greek letter μ and is very common notation for the poplation value (the TRUE value!) of the mean of some hypothetical measure. In cell B1, insert some number as the value of μ . Any number! It can be negative or positive.
3. In cell A2 type “Sigma”. Sigma is the greek letter σ . σ^2 is very common (universall!) notation for the population (TRUE) variance of some measure or parameter. Notice that the true (population) values of the mean and variance are greek letters. This is pretty standard in statistics. In cell B2, insert some positive number (standard deviations are the positive square roots of the variance).
4. In cell A8 type the number 1
5. In cell A9 insert the equation “=A8 + 1”. What is this equation doing? It is adding the number 1 to the value in the cell above, so the resulting value should be 2.

6. In Cell B8, insert the equation “=normsinv(rand())*\$B\$2 + \$B\$1“. The first part of the equation creates a random normal variable with mean 0 and standard deviation 1. multiplication and addition transform this to a random normal variable with mean μ and standard deviation σ (the values you set in cells B1 and B2).
7. copy cell B8 and paste into cell B9. Now Highlight cells A9:B9 and copy the equations down to row 107. You now have 100 random variables sampled from a infinite population with mean μ and standard deviation σ .
8. In cell A4 write “mean 10”. In cell B4 insert the equation “=average(B8:B17)”. The resulting value is the **sample mean** of the first 10 random variables you created. Is the mean close to μ ?
9. In cell A5 write “sd 10”. In cell B5 insert the equation “=stdev(B8:B17)”. The result is the **sample standard deviation** of the first 10 random variables. Is this close to σ ?
10. In cell A6 write “mean 100”. In cell B6 insert the equation “=average(B8:B107)”. The resulting value is the **sample mean** of the all 100 random variables you created. Is this mean closer to μ than mean 10?
11. In cell A7 write “sd 100”. In cell B7 insert the equation “=stdev(B8:B107)”. The resulting value is the **sample standard deviation** of the all 100 random variables you created. Is this SD closer to σ than sd 10?

The sample standard deviation is a measure of the variability of the sample. The more spread out the sample (the further each value is from the mean), the bigger the sample standard deviation. The sample standard deviation is most often simply known as “The” standard deviation, which is a bit misleading since there are many kinds of standard deviations!

Remember that your computed mean and standard deviations are estimates computed from a sample. They are estimates of the true values μ and σ . Explore the behavior of the sample mean and standard deviation by re-calculating the spreadsheet. In Excel, a spreadsheet is re-calculated by simultaneously pressing the command and equal key. In Google, command-R recalculates but is painfully slow. Instead, if using Google Sheets, just type the number 1 into a blank cell, and the sheet recalculates quickly. Do it again. And again.

Each time you re-calculate, a new set of random numbers are generated and the new means and standard deviations are computed. Compare mean 10 and mean 100 each re-calculation. Notice that these estimates are variable. They change with each re-calculation. How variable is mean 10 compared to mean 100? The variability of the estimate of the mean is a measure of **uncertainty** in the estimate. Are we more uncertain with mean 10 or with mean 100? This variability is measured by a standard deviation. This **standard deviation of the mean** is also called the **standard error of the mean**. Many researchers are loose with terms and use “The” standard error to mean the standard error of the mean, even though there are many kinds of standard errors. In general, “standard error” is abbreviated as “SE.” Sometimes “standard error of the mean” is specifically abbreviated to “SEM.”

The standard error of the mean is a measure of the precision in estimating the mean. The smaller the value the more precise the estimate. The standard error of the mean *is* a standard deviation of the mean. This is kinda weird. If we sample a population one time and compute a mean, how can the mean have a standard deviation? There is only one value! And we compute this value using the sample standard deviation: $SEM = \frac{SD}{\sqrt{N}}$. To understand how the SEM is a standard deviation, Imagine recalculating the spread sheet an infinite number of times and each time, you write down the newly computed mean. The standard error of the mean is the standard deviation of this infinitely long column of means.

2.3 Using R to generate fake data to explore uncertainty

due by the beginning of our next class

note that I use “standard deviation” to refer to the sample standard deviation and “standard error” to refer to the standard error of the mean (again, we can compute standard errors as a standard deviation of any kind of estimate)

2.3.1 part I

In the exercise above, you used Google Sheets to generate p columns of fake data. Each column had n elements, so the matrix of fake data was $n \times m$ (it is standard in most fields to specify a matrix as rows by columns). This is *much* easier to do in R and how much grows exponentially as the size of the matrix grows.

To start, we just generate a $n \times m$ matrix of normal random numbers.

```
# R script to gain some intuition about standard deviation (sd) and standard error (se)
# you will probably need to install ggplot2 using library(ggplot2)
n <- 6 # sample size
p <- 100 # number of columns of fake data to generate
fake_data <- matrix(rnorm(n*p, mean=0, sd=1), nrow=n, ncol=p) # create a matrix
```

the 3rd line is the cool thing about R. In one line I'm creating a dataset with n rows and p columns. Each column is a sample of the standard normal distribution which by definition has mean zero and standard deviation of 1. But, and this is important, any sample from this distribution will not have exactly mean zero and standard deviation of 1, because it's a sample, the mean and standard deviation will have some small error from the truth. The line has two parts to it: first I'm using the function "rnorm" (for random normal) to create a vector of $n*m$ random, normal deviates (draws from the random normal distribution) and then I'm organizing these into a matrix (using the function "matrix")

To compute the vector of means, standard deviations, and standard errors for each column of `fake_data`, use the `apply()` function.

```
means <- apply(fake_data,2,mean) # the apply function is super useful
sds <- apply(fake_data,2,sd)
sems <- sds/sqrt(n)
```

`apply()` is a workhorse in many R scripts. Learn it. Know it. Live it.

The SEM is the standard deviation of the mean, so let's see if the standard deviation of the means is close to the true standard error. We sampled from a normal distribution with $SD=1$ so the true standard is

```
1/sqrt(n)
```

```
## [1] 0.4082483
```

and the standard deviation of the p means is

```
sd(means)
```

```
## [1] 0.3731974
```

Questions

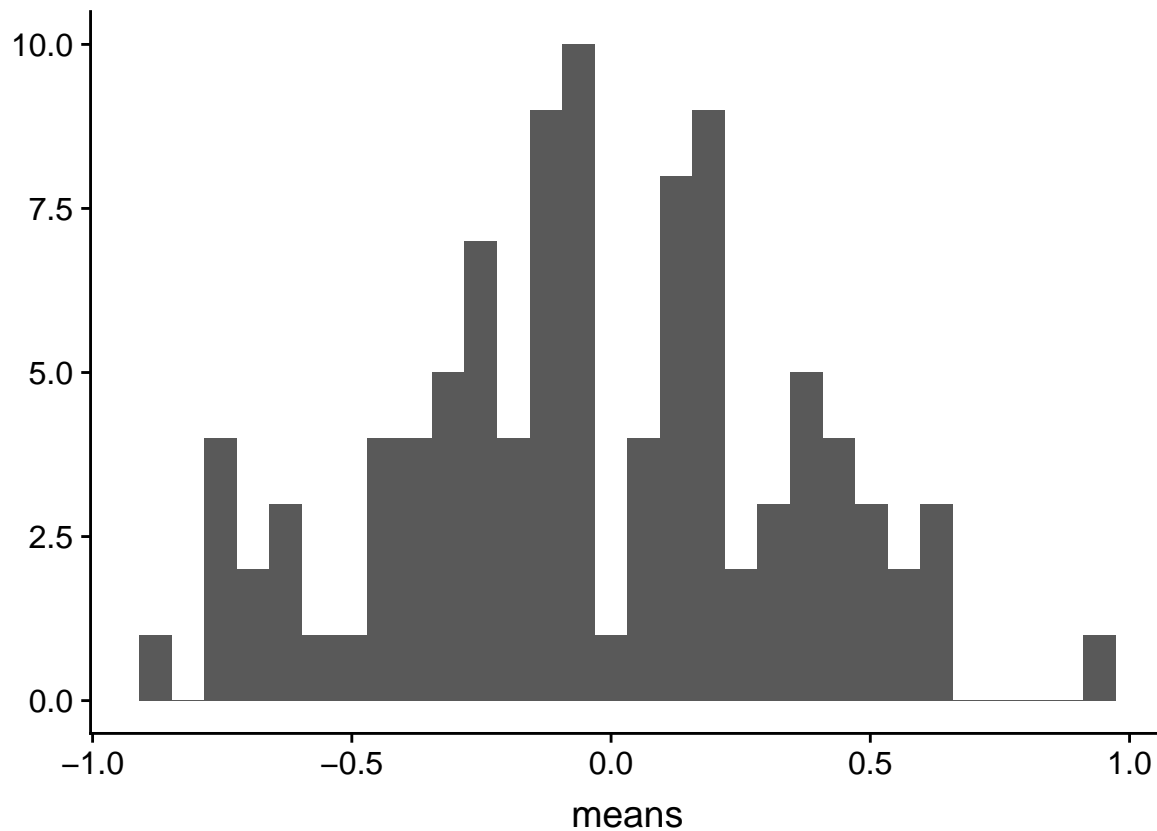
1. how close is `sd(means)` to the true SE?
2. change p above to 1000. Now how close is `sd(means)` to the true SE?
3. change p above to 10,000. Now how close is `sd(means)` to the true SE?

2.3.2 part II - means

This is a visualization of the spread, or variability, of the sampled means

```
qplot(means)
```

```
## `stat_bin()` using `bins = 30`. Pick better value with `binwidth`.
```



Compute the mean of the means

```
mean(means)
```

```
## [1] -0.039961
```

Questions

1. Remember that the true mean is zero. How close, in general, are the sampled means to the true mean. How variable are the means? How is this quantified?
2. change n to 100, then replot. Are the means, in general, closer to the true mean? How variable are the means now?
3. Is the mean estimated with $n = 100$ closer to the truth, in general, then the mean estimated with $n = 6$?
4. Redo with $n = 10000$

2.3.3 part III - how do SD and SE change as sample size (n) increases?

```
mean(sds)
```

```
## [1] 1.017144
```

Questions

1. what is the mean of the standard deviations when $n=6$ (set $p=1000$)
2. what is the mean of the standard deviations when $n=100$ (set $p=1000$)
3. when $n = 1000$? (set $p=1000$)
4. when $n = 10000$? (set $p=1000$)

5. how does the mean of the standard deviations change as n increases (does it get smaller? or stay about the same size)
6. repeat the above with SEM

```
mean(sems)
```

```
## [1] 0.4152472
```

Congratulations, you have just done a Monte Carlo simulation!

2.3.4 Part IV – Generating fake data with “for loops”

There are many other strategies for generating fake data – two that will be used extensively here are the function `rmvnorm()` and **for loops**

2.3.4.1 `rmvnorm` and a Covariance Matrix

We used the base R function `rnorm` above to generate `fake_data`, a matrix of random normal values. The columns are generated independently of each other so the expected correlation between any two columns is zero. The `rmvnorm()` (“random multivariate normal”) function from the package `mvtnorm` (“multivariate normal”) returns a matrix of random values drawn from a multivariate normal distribution with a specified **covariance matrix**. A covariance matrix is matrix of the variances and covariances of the p columns of a data matrix. The **diagonal** of the covariance matrix contains the variances of the p columns of the data matrix and the off-diagonal elements contain the $p(p - 1)$ pairwise covariances. The upper right set of covariances is the mirror of the lower left set of covariates, so there are $p(p - 1)/2$ unique covariances.

For our fake data, we want columns that are independent ($E(\text{COV}) = 0$) and have expected variance of 1. This covariance matrix has a special name – the **identity matrix** (or sometimes “unit” matrix). Thus we could use this script to generate fake data

```
n <- 6 # sample size
p <- 10^2 # number of columns of fake data to generate
fake_data <- mvtnorm::rmvnorm(n, mean=rep(0, p), sigma=diag(p))

# compute the vectors of means, sds, and sems and the sd of the means
means <- apply(fake_data, 2, mean) # the apply function is super useful
sds <- apply(fake_data, 2, sd)
sems <- sds/sqrt(n)
sd(means)
```

```
## [1] 0.4539342
```

```
mean(sems)
```

```
## [1] 0.4083748
```

The vector of column means is specified using `mean=` and the multivariate covariance matrix is specified with `sigma=` (the lower case greek letter sigma (σ) is often used to denote a population standard deviation. The upper case greek letter sigma (Σ) is often used to denote a population covariance matrix). This raises two questions

Questions.

1. without using the console, what is returned with `rep(0, p)`?
2. without using the console, what is returned with `diag(p)`?
3. What are the `sd(means)` and `mean(sems)` comparing? What is the pedagogical purpose for adding this?

Now use the console to check your answers.

2.3.4.2 A for loop

```
n <- 6 # sample size
n_iter <- 10^5 # number of iterations of loop (equivalent to p)
means <- numeric(n_iter)
sds <- numeric(n_iter)
sems <- numeric(n_iter)
for(i in 1:n_iter){
  y <- rnorm(n) # mean=0 and sd=1 are default so not necessary to specify
  means[i] <- mean(y)
  sds[i] <- sd(y)
  sems[i] <- sd(y)/sqrt(n)
}
sd(means)
```

```
## [1] 0.4090381
```

```
mean(sems)
```

```
## [1] 0.3883677
```

Questions

1. What do `sd(means)` and `mean(sems)` converge to as `n_iter` is increased from 100 to 1000 to 10,000?
2. Do they converge to the same number?
3. Should they?
4. What is the correct number?

Question number 4 is asking what is $E(\text{SEM})$, the “expected standard error of the mean”. There is a very easy formula to compute this. What is it?

Appendix 1: Getting Started with R

2.4 Get your computer ready

2.4.1 Install R

R is the core software

Download R for your OS

2.4.2 Install R Studio

R Studio is a slick (very slick) GUI interface for developing R projects

Download R Studio Desktop

2.4.3 Resources for installing R and R Studio

On Windows

On a Mac

2.4.4 Install LaTeX

LaTeX (“la-tek”) is necessary to use the pdf output of R Markdown.

On Windows

On a Mac

2.5 Start learning

2.5.1 Start with Data Camp Introduction to R

Data Camp: Introduction to R (free online course)

2.5.2 Then Move to Introduction to R Studio

R Studio Essentials, Programming Part 1 (Writing code in RStudio)

2.5.3 Develop your project with an R Studio Notebook

Getting Started with R Markdown

Introducing Notebooks with R Markdown

2.6 Getting Data into R

Getting your data into R

2.7 Additional R learning resources

Getting used to R, RStudio, and R Markdown

Link to list of R Studio webinars

Link to set of R package cheat sheets (amazing!)

Bookdown online books

2.8 Packages used extensively in this text

1. ggplot2
2. data.table
3. mvtnorm
4. lme4
5. nlme
6. emmeans
7. readxl
8. reshape2

Data Visualisation chapter from *R for Data Science*

Graphics for communication chapter from *R for Data Science*

Youtube: An Introduction to The data.table Package

Coursera: The data.table Package

Appendix 2: Online Resources for Getting Started with Linear Modeling in R

Regression Models for Data Science in R by Brian Caffo

Broadening Your Statistical Horizons: Generalized Linear Models and Multilevel Models by J. Legler and P. Roback

The Art of Data Science by Roger D. Peng and Elizabeth Matsui

Bibliography