Cancer survival analysis using population-based data Granada, 27-28 March 2017

Age-standardised Net Survival

Cohort and Period Net Survival

Estimates









### **Outline**

- 1. Age-standardised Net Survival.
- 2. Cohort and Period Estimation of Age-standardised Net Survival.
- 3. Practical implementation.



### Introduction

- Net Survival deals with Competing Risk.
- Pohar Perme...to deal with censoring (IPCW): because the probability of death is increased for the majority of cancer sites during the first years of follow-up.

OK, great. But ...

The excess risk of death from cancer is itself often dependent from the structure of AGE of the Background POPULATION.



### Why Age-Standardised Net Survival?

So, the Comparison of all-ages survival between two or more populations (regions or countries), or when analysing survival trends over time can be misleading if the age structure of the cancer patient populations is very different.

Imaging a Net Survival comparison between JAPAN and GUATEMALA

We need to control for the different STRUCTURE OF AGE when comparing Net Survival between populations



### What can we do about it?

- Compare age-specific survival
   (not convenient if there are too many strata)
- Calculate a summary survival value which "controls" or "adjusts" for the effect of age
  - Direct age-standardisation is one such approach
    - commonly used in the analysis of routinely-collected data



### **Direct age-standardisation**

Produces an **overall summary measure** for each population whilst **removing the effect of differences in age structure** of cancer patients that may compromise the comparison of un-standardised survival



# Constraints of the direct age standardisation method

Requires estimation of survival for <u>each</u> age group

 Requires at least one patient at risk within each age group for each time interval of follow-up, unless all patients died before the end of the study

May not be feasible when data are sparse



### **Interpretation**

- Age-standardised survival does not reflect the actual experience of the population being examined.
- It is an artificial (hypothetical) survival estimate that allow comparisons across populations while "adjusting" for age.
- It is interpretable as the overall survival that would have occurred if the age distribution of the patient group under study had been the same as that of the standard population.
- Age-standardised survival estimates should be compared to each other only if they have been calculated using the same reference population.

# Examples of (external) standard cancer patient populations

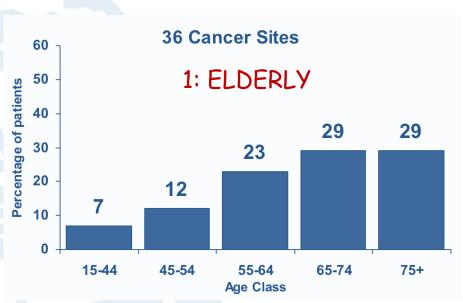
# 3. International cancer survival standard (ICSS) Recommended standard population

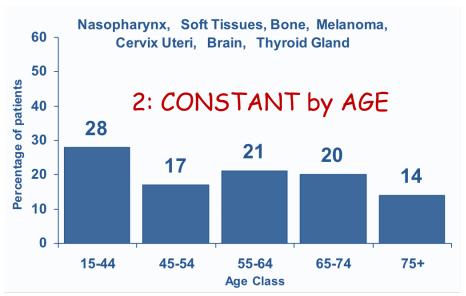
3 standards for 3 types of cancer derived from the EUROCARE-2 dataset

- Type 1: increasing with age (91% of cases)
- Type 2: constant with age (7% of cases)
- Type 3: mainly young adults (<2%)</li>
  - Minimise difference between raw and standardised
  - Smallest number of standards
  - Not arbitrary (specific to one study)
  - Males and females combined



**Type 1: increasing with age** (91% of cases) **Type 2: constant with age** (7% of cases)





Type 3: mainly young adults (<2%)



From Corazziari et al., 2004



### Example: Age-standardised cancer survival (salivary gland cancer)

#### **Unstandardised Survival Estimates at 5 years:**

	1971-75	1976-80	1981-85	1986-90
All ages	67.3	63.1	60.7	58.6

#### Age Specific Survival Estimates at 5 years:

Age Group	1971-75	1976-80	1981-85	1986-90
15-44	91.3	86.8	84.4	84.7
45-54	72.1	76.8	65.1	67.3
55-64	61.9	55.6	57.9	51.2
65-74	45.8	53.4	48.7	51.4
75+	29.2	33.9	37.7	32.3

#### Standardised Survival Estimates at 5 years

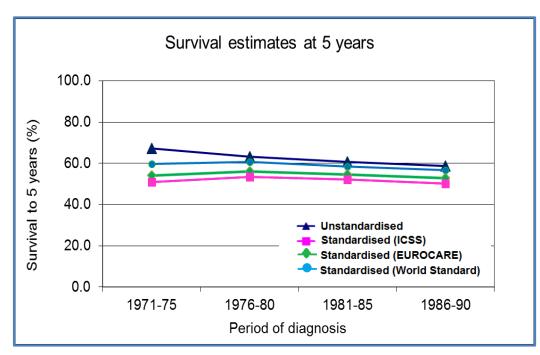
Otanuar diseu c	ui vivai LSt	illiates at 5	years	
Age Group	1971-75	1976-80	1981-85	1986-90
15-44	6.4	6.1	5.9	5.9
45-54	8.6	9.2	7.8	8.1
55-64	14.2	12.8	13.3	11.8
65-74	13.3	15.5	14.1	14.9
75+	8.5	9.8	10.9	9.4
			-	
Age standardised	51.0	53.4	52.1	50.1
	<b>↓</b> 51.0=6.4+	8.6+14.2+	13.3+8.5	
	5.9 = 84 8.1 = 67	3.7*0.07 <b>←</b> 7.3*0.12		
	11.8 = 5	1.2*0.23		
	14.9 = 5	1.4*0.29		
	9.4 = 32	2.3*0.29		

#### **ICSS Population Standard**

Age Group	weight	proportion
15-44	7	0.07
45-54	12	0.12
55-64	23	0.23
65-74	29	0.29
75+	29	0.29
All ages	100	1.00

Using other standards

	EUR	OCARE	World	Standard
Age Group	weight	proportion	weight	proportion
15-44	515	0.15	18.7	0.19
45-54	452	0.13	19.1	0.19
55-64	674	0.19	23.1	0.23
65-74	905	0.26	18.9	0.19
75+	992	0.28	20.2	0.20
All ages	3539	1	100	1.0



### Calculation: age-standardised survival

- 1. Estimate age-specific survival for each age-group
- 2. Combine these estimates in a weighted average

Age-standardised survival (AS) 
$$= S_1.w_1 + S_2.w_2 + ... + S_n.w_n =$$
 
$$= \sum_{i=1}^n S_{i.}w_i \longrightarrow \text{Sum weights: } w_1 + w_2 + ... + w_n = 1$$

where

S<sub>i</sub> = survival estimate for age group i

w<sub>i</sub> = proportion of patients in age group i in the standard population

# Calculation: standard error for age-standardised survival

$$SE(AS) = \sqrt{\sum_{i=1}^{n} w^{2}_{i} \cdot SE(S_{i})^{2}}$$

where,

SE(S<sub>i</sub>)= standard error for the survival estimate in age group i

 $w_i$  = proportion of patients in age group i in the standard population

# Calculation: 95% confidence interval for age standardised survival

Lower Limit =  $AS / \exp[1.96 * SE(AS) / AS]$ 

Upper limit = 
$$AS * \exp[1.96 * SE(AS) / AS]$$



Based on a logarithmic transformation of the survival function

### **Conclusions**

- When comparing net survival estimates between groups there is still a residual impact of patient age on the estimates.
- This can be remedied by age-standardisation of the net survival estimates. Confidence intervals can be generated for age-standardised estimates, allowing comparisons between them.
- The reference population can be chosen to maximise:
  - Comparability with other studies
  - Maximum agreement between un-standardised and standardised estimates
  - Computational ease
- ICSS reference population is a very good choice for all these... though you may be required to use another one!



### References

- Corazziari I, Quinn M, Capocaccia R. Standard cancer patient population for age standardising survival ratios. Eur J Cancer 2004;40:2307-2316.
- Capocaccia R, Gatta G, Roazzi P, Carrani E, Santaquilani M, De Angelis R, Tavilla R, EUROCARE Working Group. The EUROCARE-3 database: methodology of data collection, standardisation, quality control and statistical analysis. Ann Oncol 2003; 14: v14-v27.
- Sankaranarayanan, R.; Black, R. J. & Parkin, D. M. Cancer survival in developing countries (IARC Scientific Publications No. 145); 1998.
- Ito Y, Ohno Y, Rachet B, Coleman MP, Tsukuma H, Oshima A. Cancer survival trends in Osaka, Japan: the influence of age and stage at diagnosis. Jpn J Clin Oncol 2007; 37: 452-458.
- Quaresma, M, Coleman M, Rachet B. 40-year trends in an index of survival for all cancers combined and survival adjusted for age and sex for each cancer in England and Wales, 1971-2011: a population-based study. The Lancet 2015; 385: 1206-1218.

# Cancer Survival Analysis using Population-Based Data

Granada, 27-28 March - 2017



# Day 2 - Practical Study designs

Improving health worldwide http://csg.lshtm.ac.uk









### Learning objectives

1.To estimate net survival using cohort and period approach, and to compare the results.



### Stata: stset

#### Setting the structure of your survival data

stset(end of follow-up: date), fail(dead==1) origin(time date) enter(time date) exit(time date) scale(value)

Scale: scale(30) = monthly data; scale(365.25) = yearly data

$$t = \frac{time - origin()}{scale()}$$
; by default:  $t = \frac{time - origin(0)}{scale(1)}$ 

enter: When the subject enter the study (date)

origin: When the subject becomes at risk and needed if you have only calendar dates in your dataset

exit: When the subject leaves the study (date)

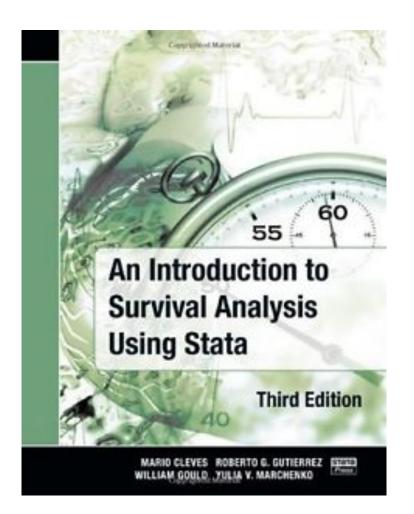
\_t0 and \_t: Record the time span, in analysis time t. \_t0= start and \_t= end

**d:** Record the outcome at the end of the span.

\_st: Record whether the observation is going to be used in the current analysis







Chapters 4, 5, 6



### Breast\_stns.dta

_															Cale	endar	year of	f follov	/-up															_
	1971	1972	1973	1974	1975	1976	1977	1978	1979	1980	1981	1982	1983	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	1996	1997	1998	1999	2000	2001	2002	2003	<u>;                                    </u>
971	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32	
972		0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	
973			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	
974				0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	
975					0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	
976						0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	
977							0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	
978								0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	
979									0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	
980										0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	
981											0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	
982												0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	
983													0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	
984														0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	
985															0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	
986																0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	
987																	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	
988															1			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	
989																			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	
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### Breast\_stns.dta

_															Cal	endar	year of	follov	v-up															_
	1971 1	972	1973	1974	1975	1976	1977	1978	1979	1980	1981	1982	1983	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	1996	1997	1998	1999	2000	2001	2002	2003	3
971	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32	1
972		0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	ı
973			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	ı
974				0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	ı
975					0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	ı
976						0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	ı
977							0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	ı
978								0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	
79									0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	
980										0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	
81											0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	
82												0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	
83													0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	1
84														0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	1
85															0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	1
86																0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	1
987																	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	1
88															1			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	1
989																			0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	1
990																				0	1	2	3	4	5	6	7	8	9	10	11	12	13	1
991																					0	1	2	3	4	5	6	7	8	9	10	11	12	1
992																						0	1	2	3	4	5	6	7	8	9	10	11	1
993																							0	1	2	3	4	5	6	7	8	9	10	1
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#### . stset finmdy, fail(dead==1) origin(time diagmdy) enter(time diagmdy)

failure event: dead == 1

obs. time interval: (origin, finmdy]

enter on or after: time diagmdy

exit on or before: failure

t for analysis: (time-origin)

origin: time diagmdy


355801 total observations

0 exclusions

355801 observations remaining, representing

212121 failures in single-record/single-failure data

920269050 total analysis time at risk and under observation

at risk from t = 0

earliest observed entry t = 0

last observed exit t = 12052



### Stata: stset

stset finmdy, fail(dead==1) origin(time diagmdy) enter(time diagmdy) scale(365.25)

list diagmdy finmdy \_t0 \_t \_d \_st in 1/10

-	<b></b>					
	diagmdy	finmdy	_ <sup>t0</sup>	_t	_d	_st
1.	01jan2000	31dec2003	0	3.9972621	1	1
2.	29jul1998	31dec2003	0	5.4236824	0	1
3.	30jan1998	31dec2003	0	5.9164956	0	1
4.	09jul1998	31dec2003	0	5.4784394	0	1
5.	22dec1998	31dec2003	0	5.0239562	0	1
6.	23jan1998	31dec2003	0	5.9356605	0	1
7.	16jul1998	31dec2003	0	5.4592745	0	1
8.	07jul1999	31dec2003	0	4.4845996	0	1
9.	18aug1998	31dec2003	0	5.3689254	0	1
10.	01sep1999	01oct2003	0	4.0821355	1	1
+						



### **Net Survival Estimates**

#### **Hypothetical Scenario:**

9 Cohorts from 1971 to 1979

5 incident cases per year

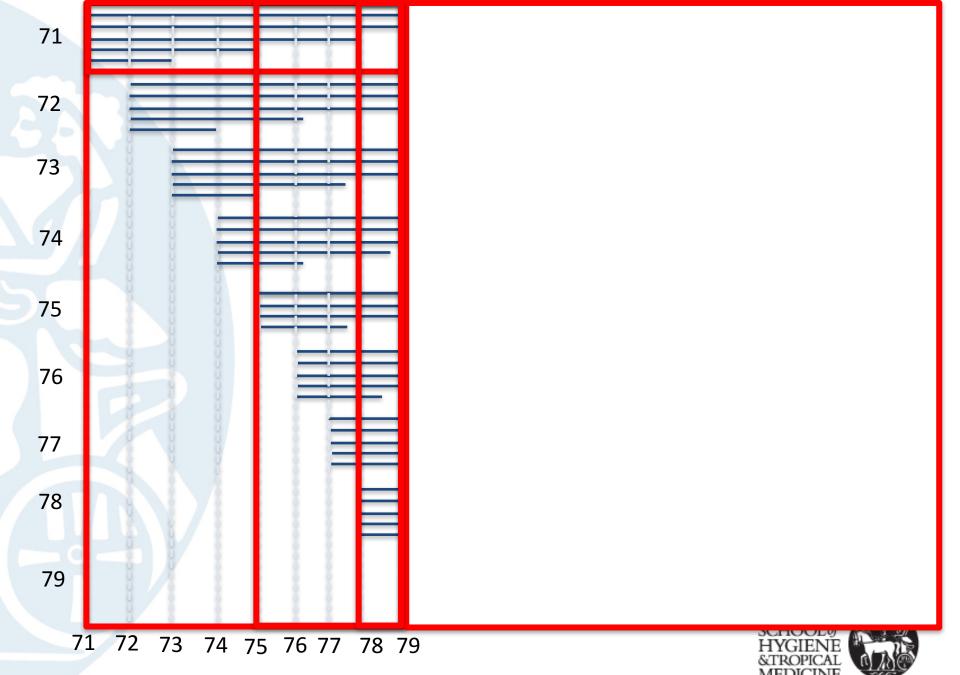
Recursive survival pattern:

- 1 individual survives 15 years
- 1 individual survives 10 years
- 1 individual survives 7 years
- 1 individual survives 4 years
- 1 individual survives 2 years

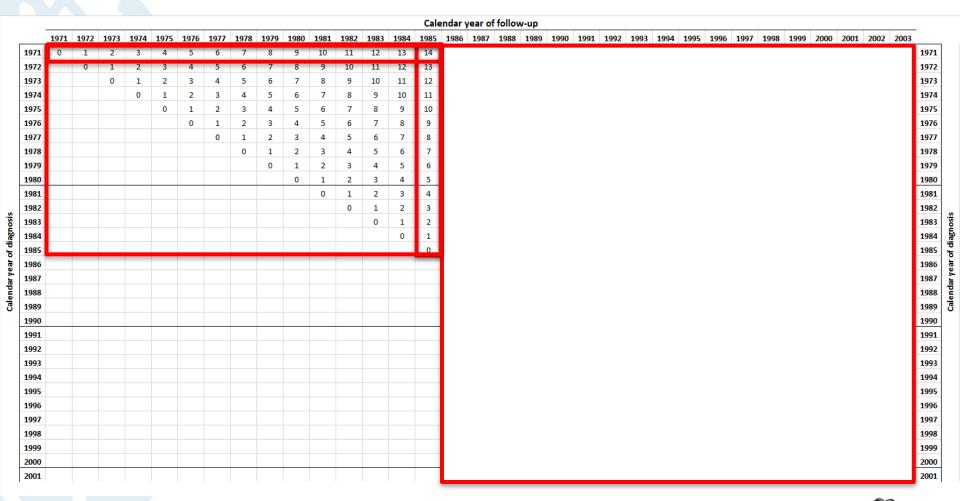
### Several options to estimate Net Survival answering different questions:

- 1. The complete data (considering the last update date)
- 2. From one particular cohort
- 3. From one or several aggregated calendar periods





### Q1: Cohort and Period





# Cohort approach, 1971

### \*cohort approach 1971

```
stset finmdy, fail(dead==1) origin(time diagmdy) enter(time diagmdy) scale(365.25)
```

```
stns list using lifetable_stns.dta if year(diagmdy)==1971, //
age(agediagindays=_age) period(diagmdy=yearindays) ///
strata(sex dep) rate(rate_day) ///
at(.5 1 5 10 15, scalefactor(365.25) unit(year)) ///
saving(cohort_1971, replace)
```

type of estimate: kaplan-meier

failure \_d: dead == 1

analysis time \_t: (finmdy-origin)

origin: time diagmdy enter on or after: time diagmdy

	Time	Event	Beg.		Net	Net	Surv.		
	(year)	Time	Total	Fail	Lost	Fun	ction	[95% Con	f. Int.]
-									
	.5	182	6061	838	0	0	. 8722	0.8634	0.8810
	1	365	5223	480	1	0	.8013	0.7906	0.8120
	5	1826	4742	2056	7	0	. 4992	0.4844	0.5139
	10	3646	2679	1024	1	0	. 3485	0.3306	0.3664
	15	5467	1654	474	3	0	.3126	0.2907	0.3345



# **Cohort, 1971**

The survival of women diagnosed in 1971, as computed with the cohort approach, was:

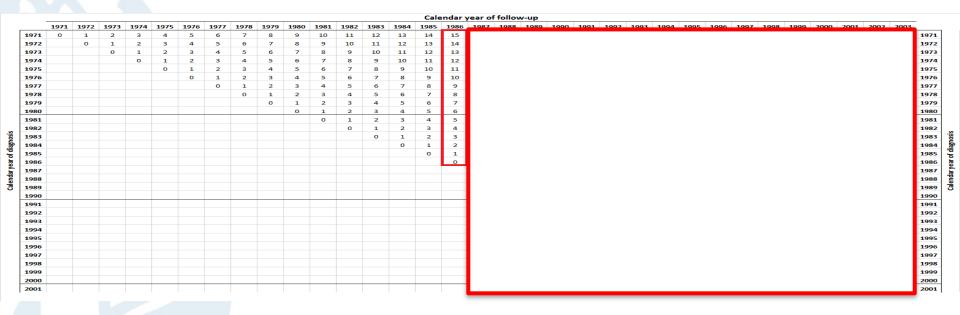
**one-year** survival: 0.80 (95% CI 0.79-0.81)

**five-year** survival: 0.50 (95% CI 0.48-0.51)

**ten-year** survival: 0.35 (95% CI 0.33-0.37)

**fifteen-year** survival: 0.31 (95% CI 0.31-0.33)



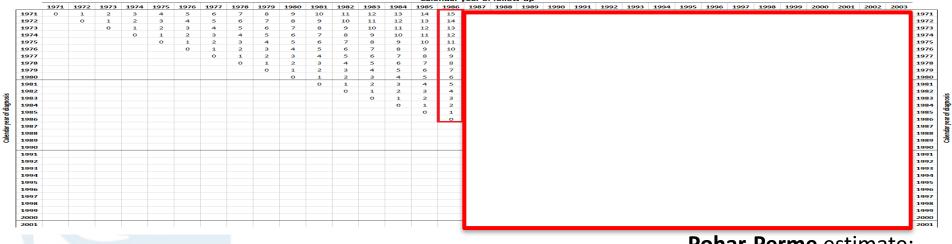


#### **JUSTIFICATION:**

Need of **UPDATED SURVIVAL estimates** that account for recent **improvements in treatment** 

- All observations included in the analysis are LEFT TRUNCATED and RIGHT CENSORED at 1986
- Then from the survivors of the cohort of 1971 we will have 15 years of follow-up (5480 days)
- STNS account for the informative censoring (cancer deaths that occurs during the first years following diagnosis)



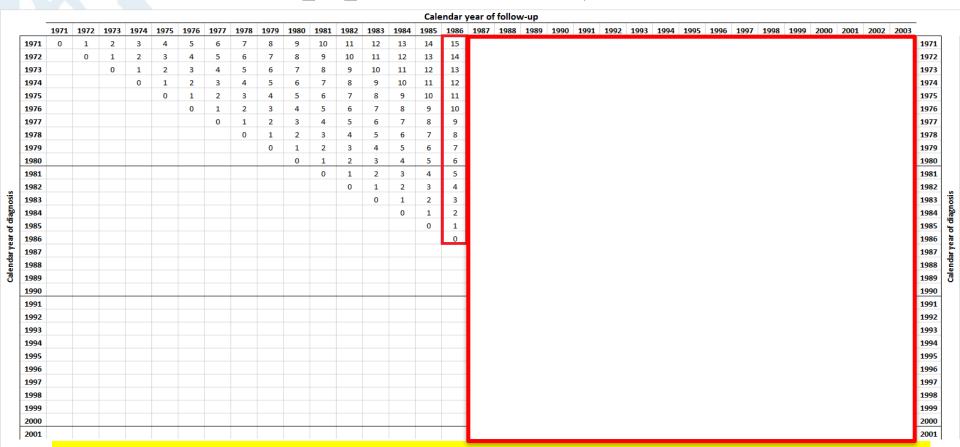


```
*Period approach 1986

stset finmdv, fail(dead==1) origin(time diagmdv) ///
enter(time mdy(1, 1, 1986)) exit(time mdy(12, 31, 1986))
```

Pohar-Perme estimate:
Needs to account for
Informative censoring
from the date of diagnosis

```
stns list using lifetable_stns.dta ///
age(agediagindays=_age) period(diagmdy=yearindays) ///
begintime(origin) strata(sex dep) rate(rate_day) ///
at(.5 1 5 10 15, scalefactor(365.25) unit(year) end_followup(5480) ///
saving(period_1986, replace)
```



The **PERIOD** approach corresponds to the **survival which would be observed** amongst **NEWLY DIAGNOSED patients** if the currently observed conditional survival probabilities were to remain constant.

In 1986, the predicted survival for women with breast cancer, obtained through **period analysis**:

**one-year** survival: 0.88 (95% CI 0.87-0.88)

**five-year** survival: 0.63 (95% CI 0.62-0.64)

ten-year survival: 0.48 (95% CI 0.46-0.50)

**fifteen-year** survival: 0.41 (95% CI 0.39-0.44)

In situations where **SURVIVAL** is improving over time the PERIOD approach will always be lower than the corresponding cohort estimate.

#### **ADVANTAGE:**

It can estimate the NET SURIVAL for the cohort 1986 SEVERAL YEARS EARLIER.



The survival of **women diagnosed in 1971**, as computed with the **cohort approach**, was:

one-year survival: 0.80 (95% CI 0.79-0.81) five-year survival: 0.50 (95% CI 0.48-0.51) ten-year survival: 0.35 (95% CI 0.33-0.37) fifteen-year survival: 0.31 (95% CI 0.31-0.33)

In 1986, the predicted survival for women with breast cancer, obtained through period analysis:

one-year survival: 0.88 (95% CI 0.87-0.88) five-year survival: 0.63 (95% CI 0.62-0.64) ten-year survival: 0.48 (95% CI 0.46-0.50) fifteen-year survival: 0.41 (95% CI 0.39-0.44)

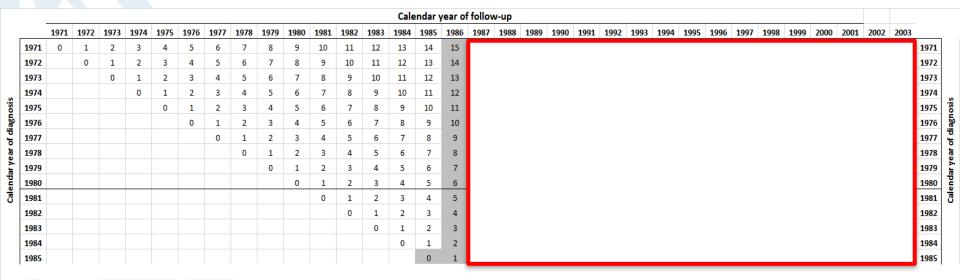


															Cale	ndar	ear of	follo	w-up																
	1971	1972	1973	1974	1975	1976	1977	1978	1979	1980	1981	1982	1983	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	1996	1997	1998	1999	2000	2001	2002	2003		
197	L 0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15																		1971	
1972	2	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14																		1972	
197	3		0	1	2	3	4	5	6	7	8	9	10	11	12	13																		1973	
1974	ı			0	1	2	3	4	5	6	7	8	9	10	11	12																		1974	<b>(A</b>
197	;				0	1	2	3	4	5	6	7	8	9	10	11																		1975	osis
Calendar year of diagnosis 1975 1976 1977 1986 1985 1985 1985 1985 1985 1985 1985 1985	j					0	1	2	3	4	5	6	7	8	9	10																		1976	diagnosis
197	'						0	1	2	3	4	5	6	7	8	9																	- 1	1977	of d
1978	3							0	1	2	3	4	5	6	7	8																		1978	ear
1979	)								0	1	2	3	4	5	6	7																		1979	ar y
1980	)									0	1	2	3	4	5	6																		1980	end
1982	L										0	1	2	3	4	5																		1981	Ö
1982	2											0	1	2	3	4																	- 1	1982	
198	3												0	1	2	3																	- 1	1983	
1984	ı													0	1	2																		1984	
198	;									Col	ho	rt 8	35		0	1																	╝	1985	

Biased: we do not have INCIDENT cases for the YEAR 1986



### Hybrid approach: 1986



```
Hybrid 1986
gen entermdy = cond(diagmdy>=mdy(1, 1, 1985), diagmdy, mdy(1, 1, 1986))
```

```
stset finmdy, fail(dead==1) origin(time diagmdy) enter(time entermdy) ///
exit(time mdy(12, 31, 1986)) Cohort 85
```

```
stns list using lifetable_stns.dta, ///
age(agediagindays=_age) period(diagmdy=yearindays) ///
begintime(origin) strata(sex dep) rate(rate_day) end_followup(5480) ///
at(.5 1 5 10 15, scalefactor(365.25) unit(year)) ///
saving(hybrid_1986, replace)
```



# Hybrid approach: 1986

														Cal	lendar '	year c	of follo	w-u	p																	,
	1971	1972	1973	1974	1975	1976	1977	1978	1979	1980	1981	1982	1983	1984	1985	5 1986	1987	19	988 19	989	1990	1991	1992	1993	1994	↓ 199	5 <b>199</b> 6	<del>199</del> ز	7 1998	8 1999	2000	2001	2002	200	3	
1971	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15																			1971	.] '
1972		0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	1																		1972	_ '
1973			0	1	2	3	4	5	6	7	8	9	10	11	12	13	1																		1973	,
1974				0	1	2	3	4	5	6	7	8	9	10	11	12	1																		1974	4 5
1975					0	1	2	3	4	5	6	7	8	9	10	11	1																		1975	
1976						0	1	2	3	4	5	6	7	8	9	10	1																		1976	
1977							0	1	2	3	4	5	6	7	8	9	1																		1977	7 5
1978								0	1	2	3	4	5	6	7	8	1																		1978	_ ear
1979									0	1	2	3	4	5	6	7	1																		1979	ar y
1975 1976 1977 1978 1979 1980 1981									'	0	1	2	3	4	5	6	4																		1980	_ <u>_</u>
1981											0	1	2	3	4	5	1																		1981	ı s
1982												0	1	2	3	4	1																		1982	_
1983									<u> </u>	'			0	1	2	3	1																		1983	,
1984														0	1	2	1																		1984	2
1985															0	1	4																		1985	,
																																			_	

#### Hybrid 1986

If date of diagnosis >=1985 (then) → 1985 else date of diagnosis <1985 (then) → 1,1,1986

#### **Combined information:**

Cohort 1985 + most updated information for other cohorts (the most recent year).

stset finmdy, fail(dead==1) origin(time diagmdy) ///
enter(time entermdy) exit(time mdy(12, 31, 1986))



### Hybrid approach: 1986

The 'biased' period estimates using the modified dataset were:

```
one-year survival: 0.84 (95% CI 0.80-0.87)
five-year survival: 0.60 (95% CI 0.57-0.63)
ten-year survival: 0.46 (95% CI 0.43-0.48)
fifteen-year survival: 0.39 (95% CI 0.36-0.42)
```

The estimates obtained with **hybrid analysis** were:

•	one-year survival:	0.88	(95% CI 0.88-0.89)
•	five-year survival:	0.63	(95% CI 0.62-0.64)
•	ten-year survival:	0.48	(95% CI 0.47-0.50)
•	fifteen-year survival:	0.41	(95% CI 0.39-0.44)



# References and further reading

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