

Overview of Time-to-Event Analysis Theory and Concepts

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<https://github.com/migariane/SVA-ULB>

March 7, 2018

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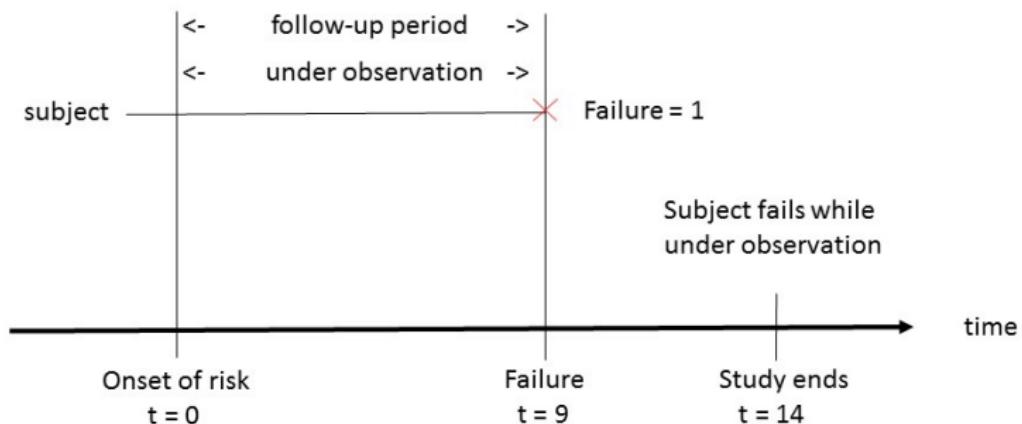
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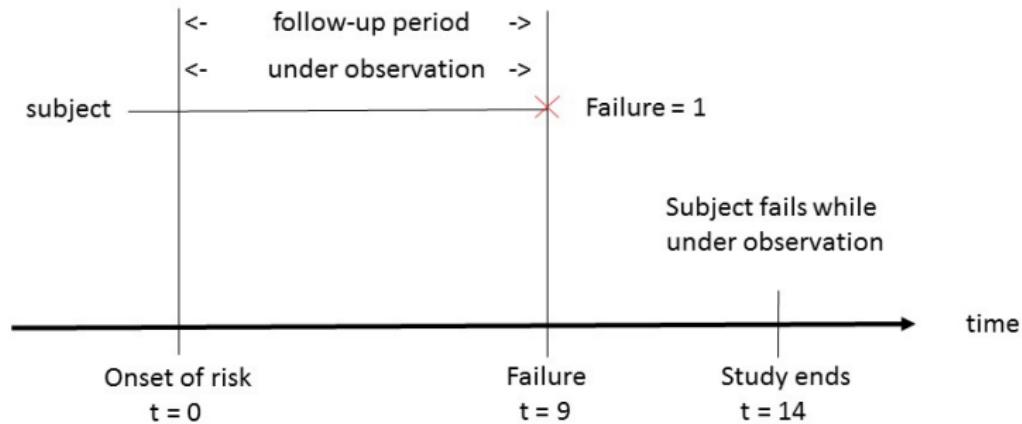
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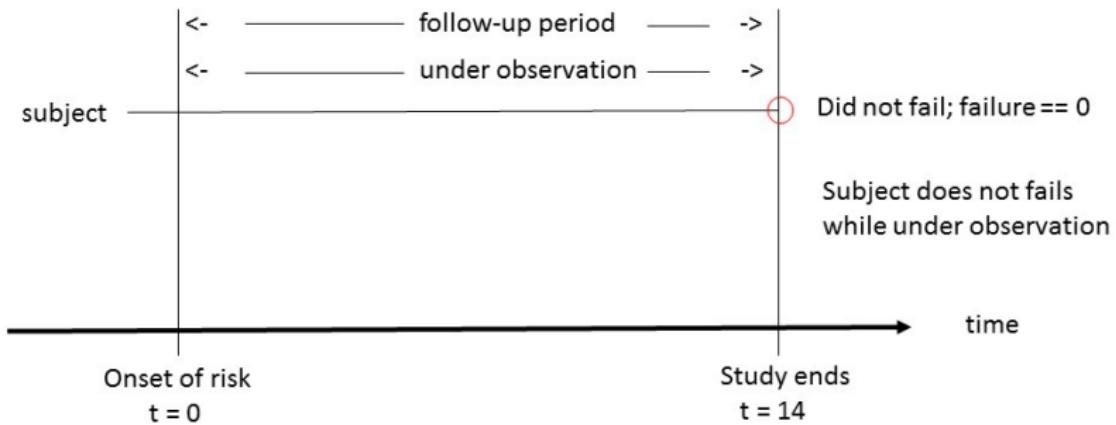
It is censoring?



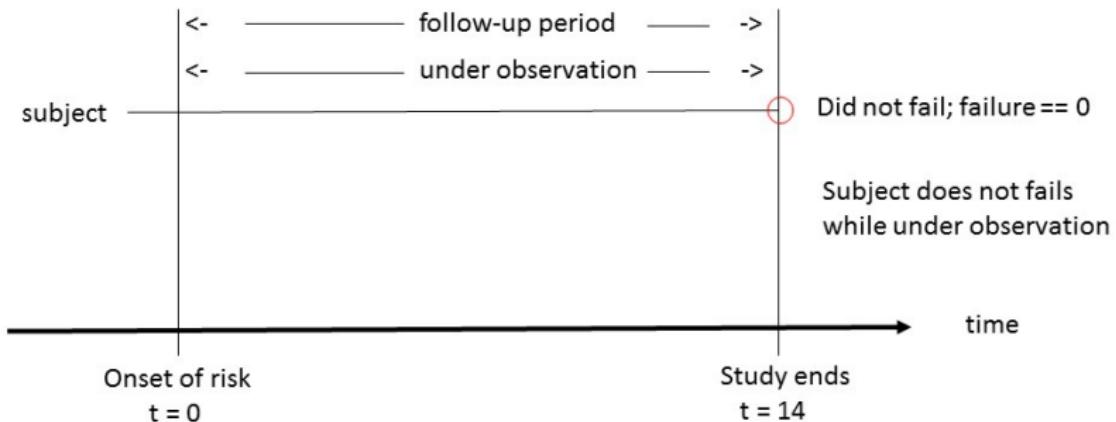
No, it is just a failure (the event of interest)



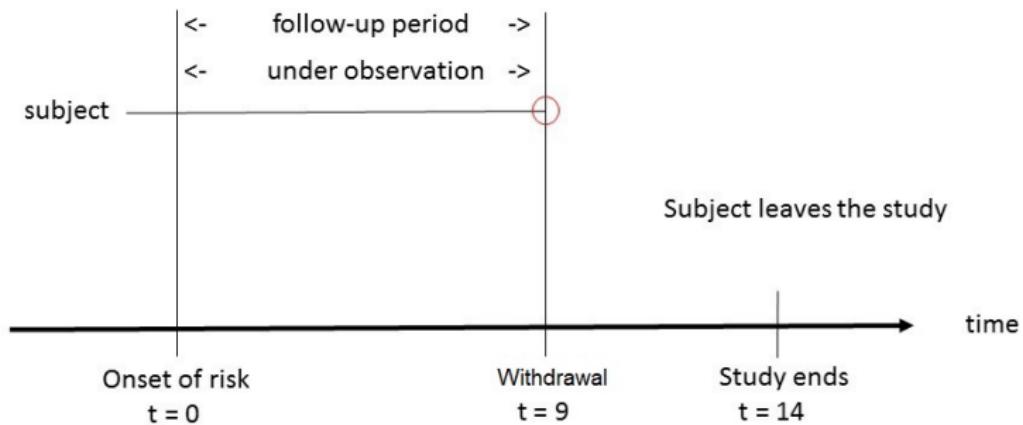
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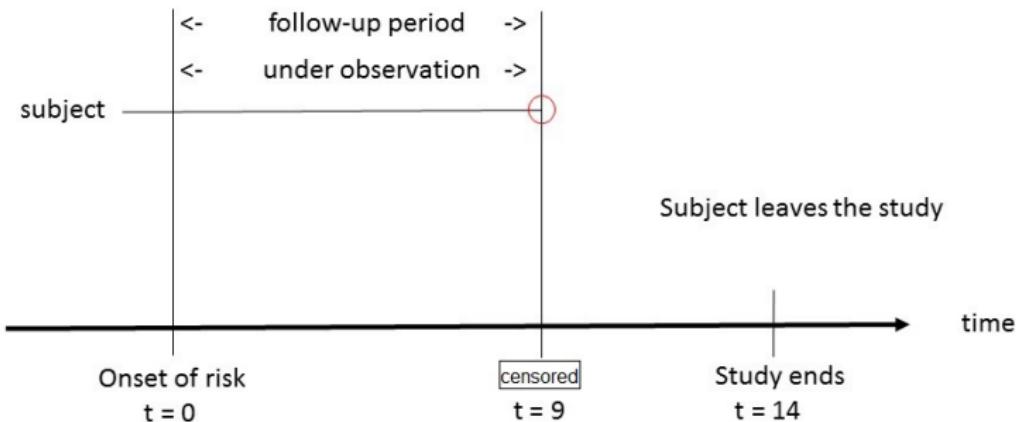
Yes, right censoring at the end of the study



It is censoring?



Yes, right censoring before the end of the study (LT FU)



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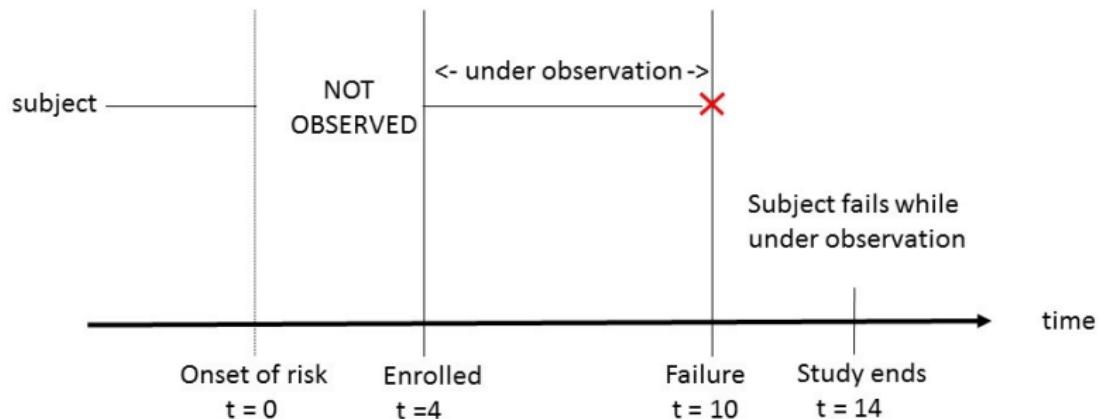
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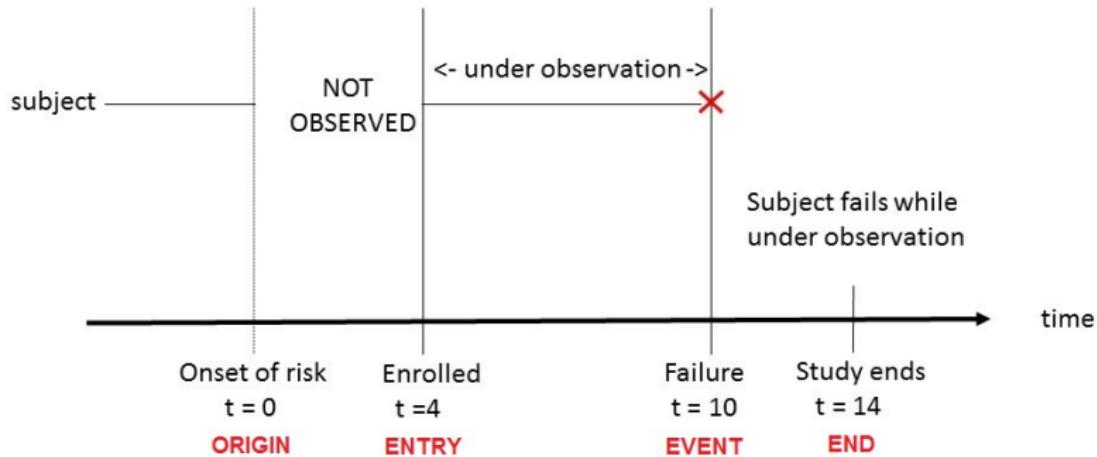
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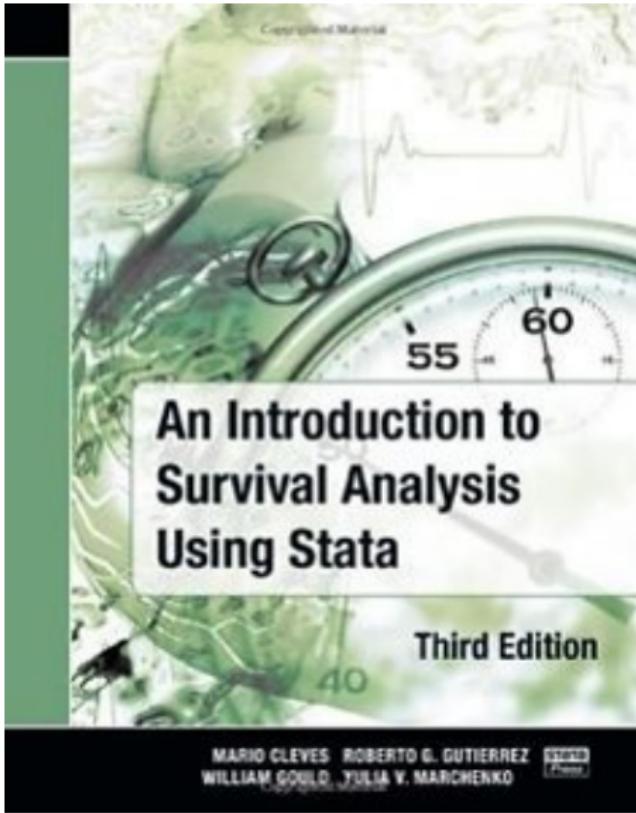


Left Truncation



Setting TIME in Statistical Software

Setting of time: chapters 5,6,7



Stata stset command

stset finmdy, fail(dead==1) origin(time diagmdy) enter(time diagmdy)

```
failure event: dead == 1
obs. time interval: (origin, finmdy]
enter on or after: time diagmdy
exit on or before: failure
t for analysis: (time-origin)
    origin: time diagmdy
```

```
355801 total observations
0 exclusions
```

```
355801 observations remaining, representing
212121 failures in single-record/single-failure data
920269050 total analysis time at risk and under observation
                           at risk from t =
                           earliest observed entry t =
                           last observed exit t =      0
                                              0
                                              12052
```

Listing the output of the Stata stset command

list diagmdy finmdy _t0 _t _d _st in 1/10

	diagmdy	finmdy	_t0	_t	_d	_st
1.	01jan2000	31dec2003	0	3.9972621	1	1
2.	29jul1998	31dec2003	0	5.4236824	0	1
3.	30jan1998	31dec2003	0	5.9164956	0	1
4.	09jul1998	31dec2003	0	5.4784394	0	1
5.	22dec1998	31dec2003	0	5.0239562	0	1
6.	23jan1998	31dec2003	0	5.9356605	0	1
7.	16jul1998	31dec2003	0	5.4592745	0	1
8.	07jul1999	31dec2003	0	4.4845996	0	1
9.	18aug1998	31dec2003	0	5.3689254	0	1
10.	01sep1999	01oct2003	0	4.0821355	1	1

Origin and Entry

Important

It is critical to distinguish between the **time** of the **STUDY ENTRY** and the **ORIGIN** of **time**.

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Time

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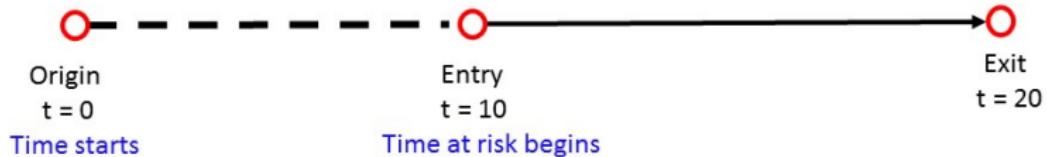
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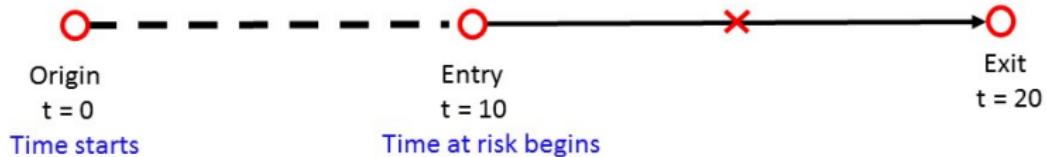
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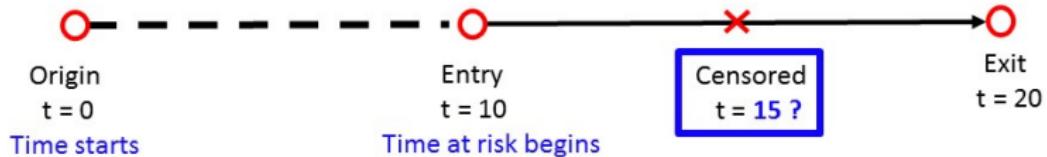
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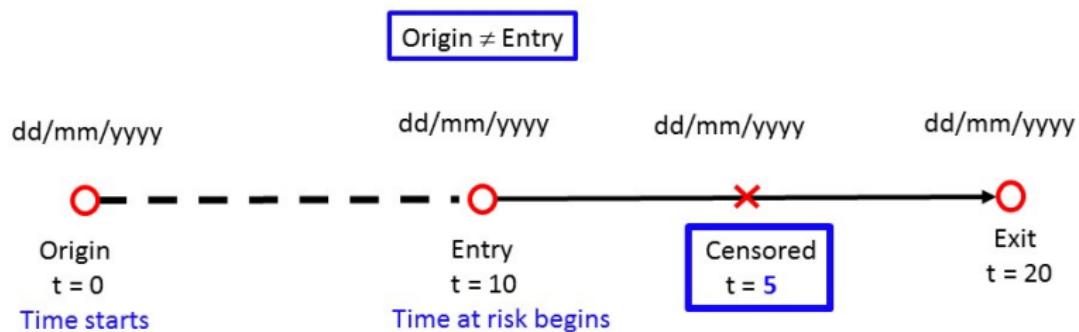


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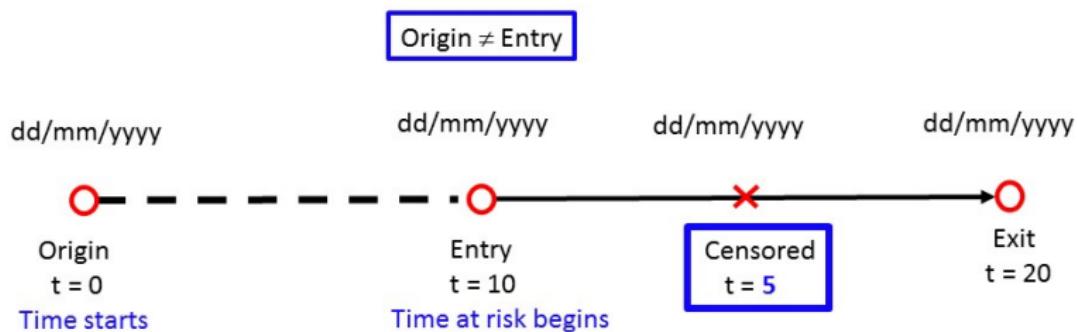


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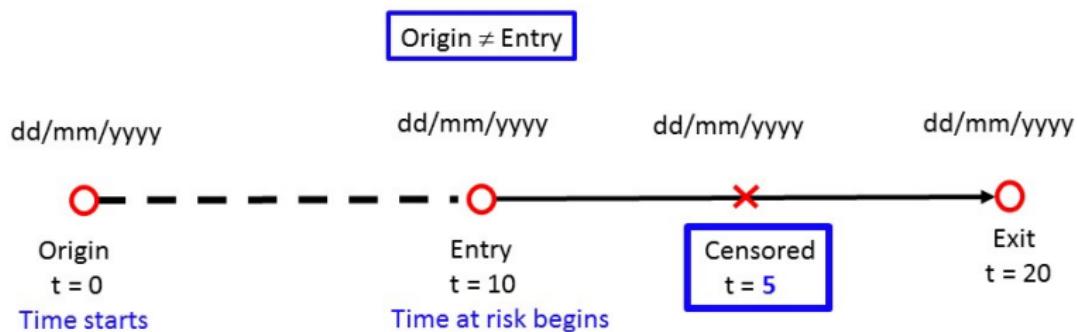
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The Maths behind Survival

The **anatomy** of time-to-event analysis

Anatomische les door Dokter van der Meer, 1617



Epidemiologists

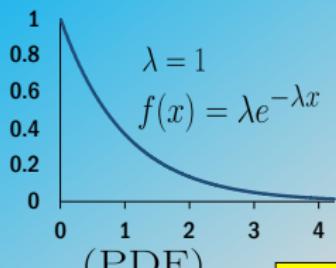
$$x = BD = \sqrt{a^2 + b^2} \quad \alpha^2 + b^2 = (\alpha - b)(\alpha + b) \quad \begin{cases} f(x) = \\ g(x) = \end{cases}$$
$$AB = \sqrt{AB_x^2 + AB_y^2}$$
$$= mx + b$$
$$\frac{1}{\operatorname{ctg} \alpha}$$
$$\cos \alpha = x$$
$$\sin \alpha = y$$
$$\operatorname{tg} \alpha = \frac{y}{x}$$
$$\pi = \int_{-\infty}^{\infty} \frac{dx}{1+x^2}$$
$$y = \sqrt{x}$$
$$Ax + By = C$$

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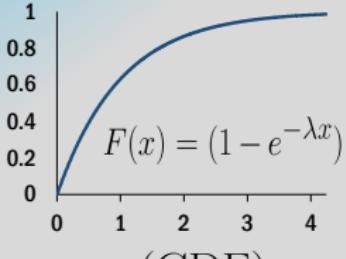
The Probability Density Function **PDF** and Cumulative Density Function **CDF** of a random variable

PDF and CDF

Continuous Probability and Cumulative Density Functions



(PDF)



(CDF)

$$F(x) = \int_{-\infty}^{+\infty} f(x) dx$$

$$f(x) = \lambda e^{-\lambda x}$$

$$F(x) = (1 - e^{-\lambda x})$$

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Derivative and antiderivative

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$$H(t) = \int_0^t \frac{f(u)}{S(u)} du = \int_0^t \frac{1}{S(u)} (\mathcal{F}') =$$

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$$\lambda(t) = \frac{\#\text{events(failures)}}{\#\text{survivors}} = P(\text{event at } t \mid \text{survival up to } t) = \frac{f(t)}{S(t)}$$

$H(t)$: Cumulative hazard or Λ

$$H(t) = \int_0^t \frac{f(u)}{S(u)} du = \int_0^t \frac{1}{S(u)} (F') = \int_0^t \frac{1}{S(u)} (1 - S(u))' = -\log(S(t))$$

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$$\lambda(t) = \frac{f(t)}{S(t)} = \frac{-d\log(S(t))}{dt}$$

The soul of Survival Analysis

From the $h(t)$ we can derive:

$$H(t) =$$

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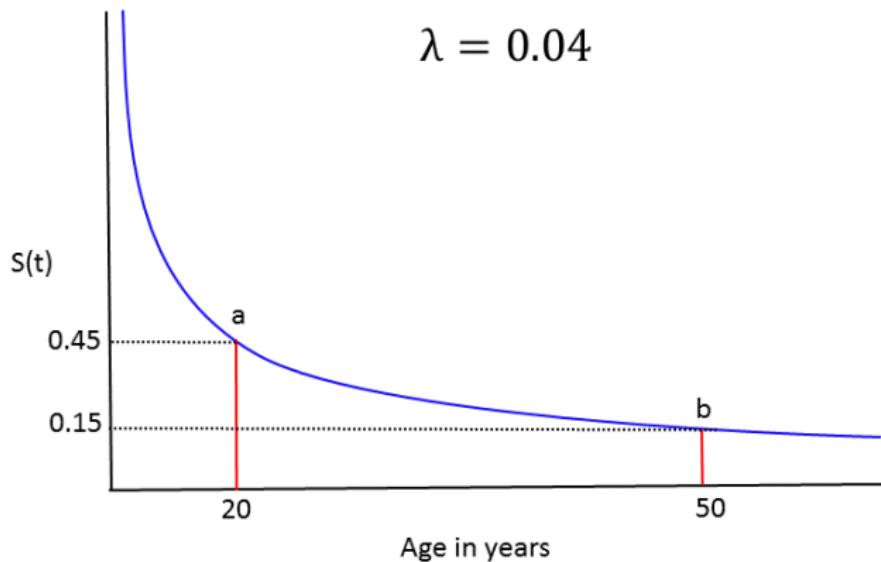
Formulae to derive h(t), H(t) and S(t): Pintilie M.(2007)

Table 2.1 Basic mathematical formulae.

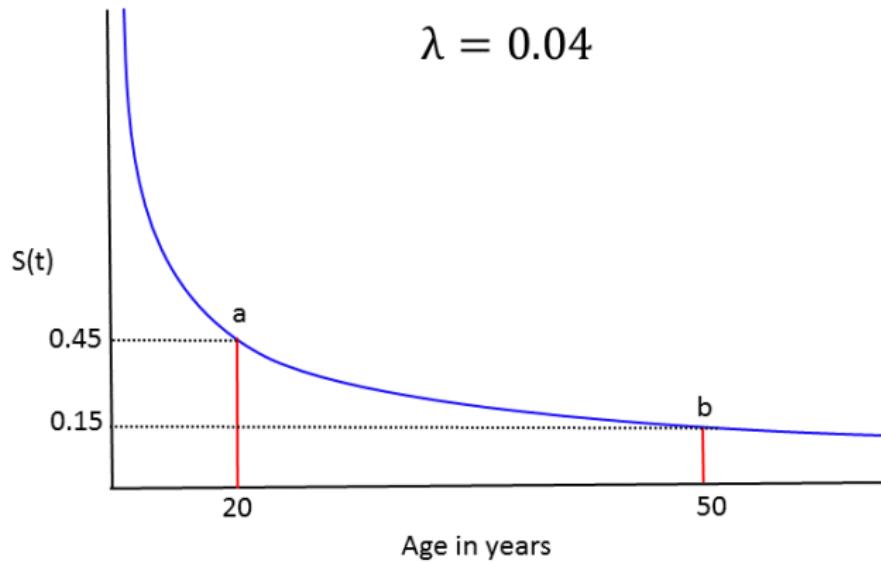
Function	Continuous distributions		Discrete distributions	
	Definition	Relationship to other functions	Definition	Relationship to other functions
Cumulative distribution	$F(t) = P(T \leq t)$	$F(t) = \int_0^t f(x)dx$	$F(t) = P(T \leq t)$	$F(t) = \sum_{t_j \leq t} p(t_j)$
Probability density	$f(t) = \frac{d}{dt} F(t)$	$f(t) = -\frac{d}{dt} S(t)$	$p(t_j) = P(T = t_j)$	$p(t_j) = S(t_{j-1}) - S(t_j)$
Survivor	$S(t) = P(T > t)$	$S(t) = 1 - F(t)$ $= \int_t^\infty f(x)dx$ $= \exp\{-H(t)\}$	$S(t) = P(T > t)^*$	$S(t) = 1 - F(t)$ $= \sum_{t_j > t} p(t_j)$ $= \exp\{-H_2(t)\}$
Hazard	$h(t) = \lim_{\delta t \rightarrow 0} \left\{ \frac{P(t \leq T < t + \delta t T > t)}{\delta t} \right\}$	$h(t) = f(t)/S(t)$ $= -\frac{d}{dt} \log S(t)$	$h(t_j) = P(T = t_j T \geq t_j)$ $H_1(t) = \sum_{t_j \leq t} h(t_j)$	$h(t_j) = p(t_j)/S(t_{j-1})$
Cumulative hazard	$H(t) = \int_0^t h(x)dx$	$H(t) = -\log S(t)$	$H_2(t) = -\sum_{t_j \leq t} \log \{1 - h(t_j)\}$	$H_2(t) = -\log S(t)$

* Probability function

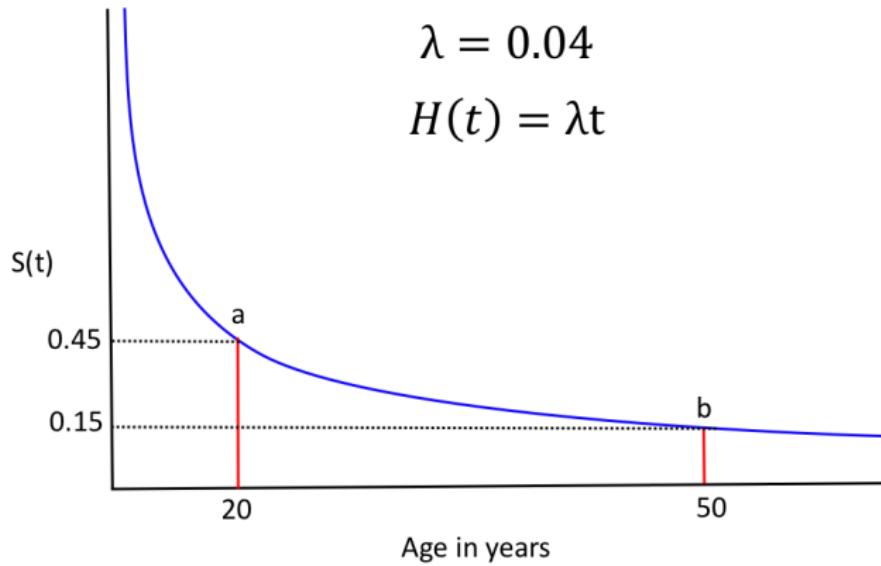
Deriving $h(t)$, $H(t)$, and $S(t)$ from the Exponential distribution



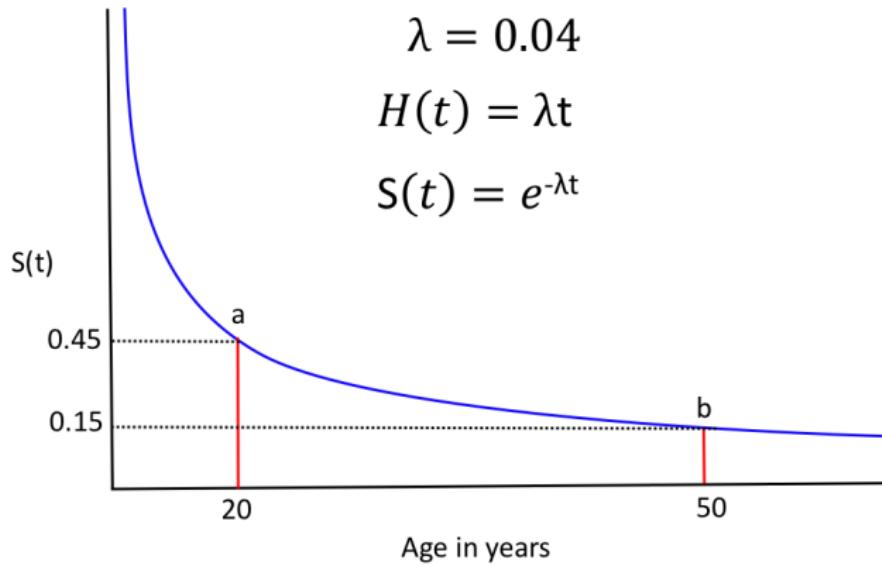
Calculus and Geometry



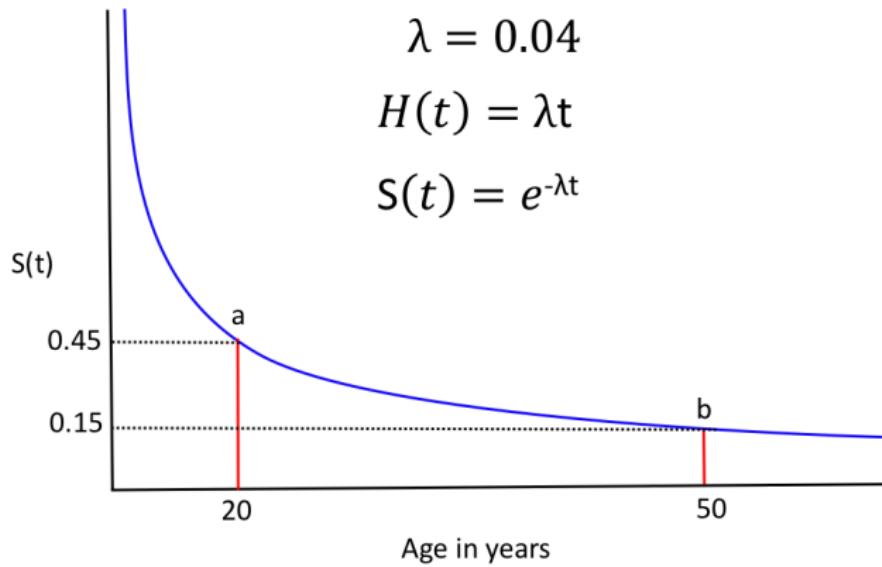
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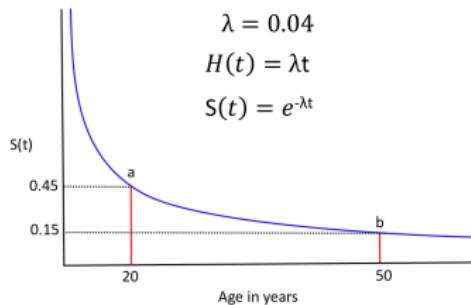
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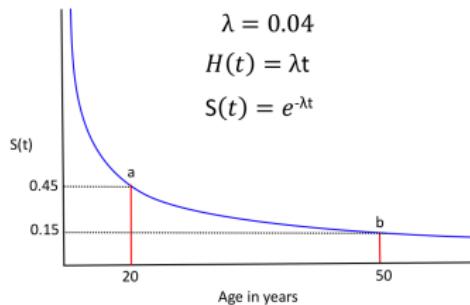
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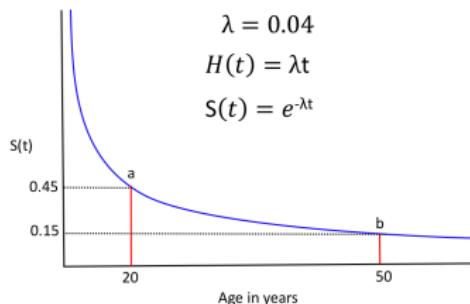


Calculus and Geometry



Link between hazard and velocity

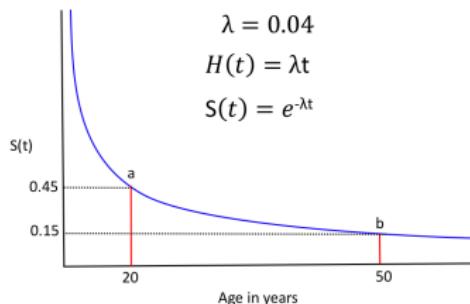
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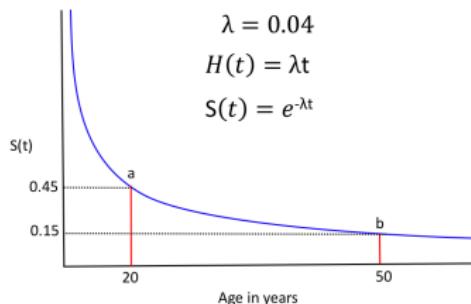
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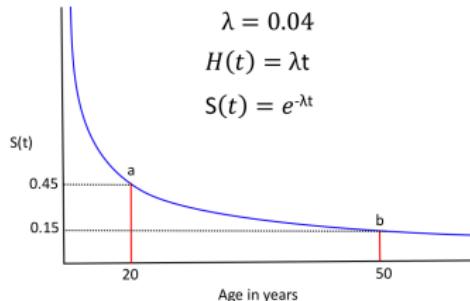
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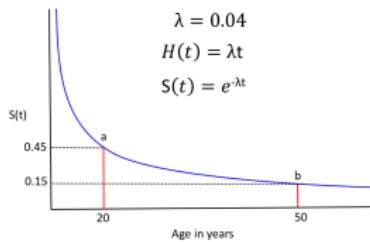
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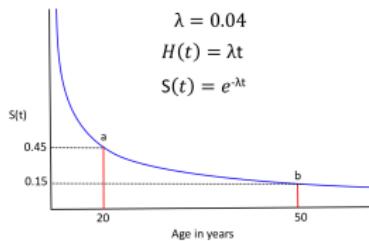
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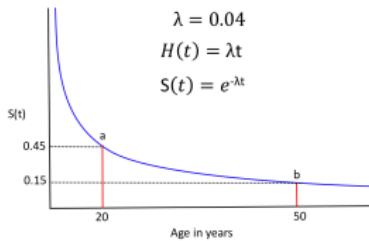


Calculus and Geometry



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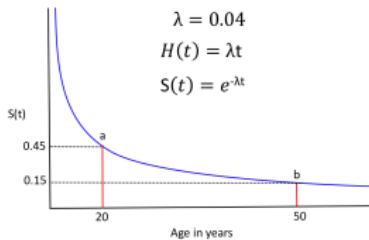
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Calculus and Geometry

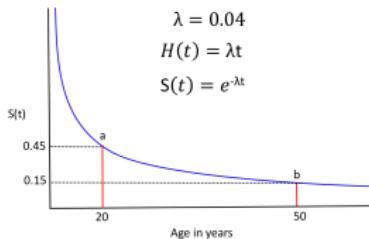


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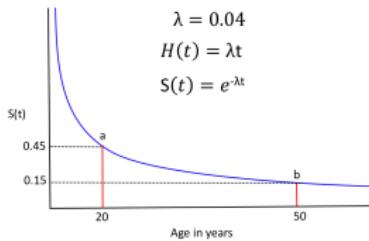


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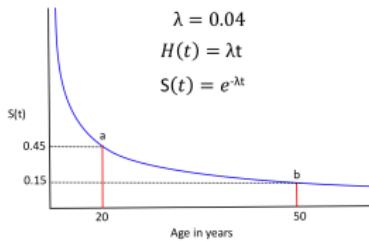


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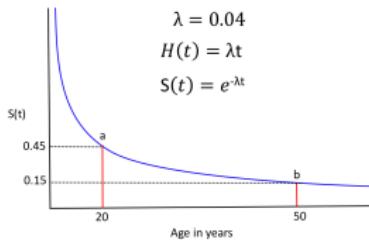


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The Statistical Methods

The Statistical Methods used in time-to-event analysis

Methods

Statistical Methods

Non-parametric

Statistical Methods

Non-parametric

- Life Tables

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	Total a_i	Deaths d_i	Lost c_i	N.A.R $n_i = (a_i - (c_i/2))$	Rate $r_i = d_i/n_i$	$s(t) = 1 - r_i$	$S(t) = S(i-1) \times s_i$
00-01	100	35	0	100	0.350	0.650	0.650
01-02	65	6	10	60	0.100	0.900	0.585
02-03	49	5	18	40	0.125	0.875	0.512
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In Stata use sts list

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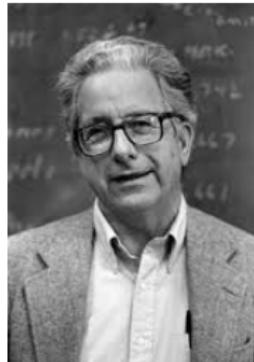
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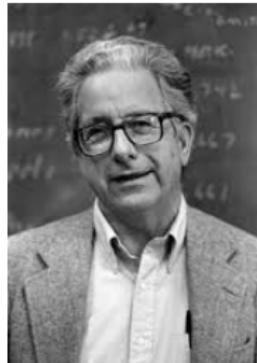
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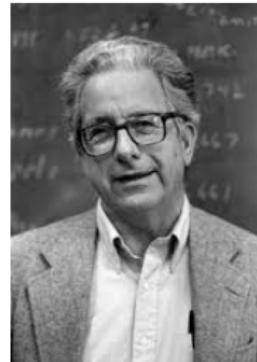
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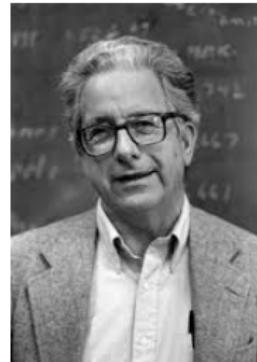


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Kaplan-Meier by group

Table 2. Kaplan-Meier survival function by treatment group

Group	time	n	d	censor	(n-d)/n	S	d/(n(n-d))	sum	In(-In(S))	InIn(SE)	InIn(L95%)	InIn(U95%)	L95%CI	U95%CI
Placebo	0	5	0	0	1.0000	1.0000	0.0000	0.0000						
Placebo	1	5	1	0	0.8000	0.8000	0.0500	0.0500	-1.4999	1.0021	-3.4640	0.4641	0.2038	0.9692
Placebo	2	4	1	0	0.7500	0.6000	0.0833	0.1333	-0.6717	0.7148	-2.0727	0.7293	0.1257	0.8818
Placebo	7	3	0	1	1.0000	0.6000	0.0000	0.1333	-0.6717	0.7148	-2.0727	0.7293	0.1257	0.8818
Placebo	8	2	1	0	0.5000	0.3000	0.5000	0.6333	0.1856	0.6610	-1.1099	1.4812	0.0123	0.7192
Placebo	12	1	1	0	0.0000	0.0000								
	time	n	d	censor	(n-d)/n	S	d/(n(n-d))	sum	In(-In(S))	InIn(SE)	InIn(L95%)	InIn(U95%)	L95%CI	U95%CI
Drug	0	5	0	0	1.0000	1.0000	0.0000	0.0000						
Drug	4	5	1	0	0.8000	0.8000	0.0500	0.0500	-1.4999	1.0021	-3.4640	0.4641	0.2038	0.9692
Drug	6	4	0	1	1.0000	0.8000	0.0000	0.0500	-1.4999	1.0021	-3.4640	0.4641	0.2038	0.9692
Drug	9	3	1	0	0.6667	0.5333	0.1667	0.2167	-0.4642	0.7405	-1.9156	0.9871	0.0683	0.8631
Drug	10	2	1	0	0.5000	0.2667	0.5000	0.7167	0.2790	0.6405	-0.9764	1.5343	0.0097	0.6861
Drug	13	1	1	0	0.0000	0.0000								

Kaplan-Meier by group

Table 2. Kaplan-Meier survival function by treatment group

Group	time	n	d	censor	(n-d)/n	S	d/(n(n-d))	sum	ln(-ln(S))	lnln(SE)	lnln(L95%)	lnln(U95%)	L95%CI	U95%CI
Placebo	0	5	0	0	1.0000	1.0000	0.0000	0.0000	-1.4999	1.0021	-3.4640	0.4641	0.2038	0.9692
Placebo	1	5	1	0	0.8000	0.8000	0.0500	0.0500	-0.6717	0.7148	-2.0727	0.7293	0.1257	0.8818
Placebo	2	4	1	0	0.7500	0.6000	0.0833	0.1333	-0.6717	0.7148	-2.0727	0.7293	0.1257	0.8818
Placebo	7	3	0	1	1.0000	0.6000	0.0000	0.1333	-0.6717	0.7148	-2.0727	0.7293	0.1257	0.8818
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Inference

Table 3. Long-rank test statistic by treatment group.

time	Placebo			Drug			Log-rank estimation				
	n_{1i}	d_{1i}	c_{1i}	n_{0i}	d_{0i}	c_{0i}	n_i	d_i	e_{1i}	v_{1i}	$d_{1i} - e_{1i}$
0	5	0	0	5	0	0	10	0	0	0	0
1	5	1	0	5	0	0	10	1	0.500	0.250	0.500
2	4	1	0	5	0	0	9	1	0.444	0.247	0.556
4	3	0	0	5	1	0	8	1	0.375	0.234	-0.375
6	3	0	0	4	0	1	7	0	0.000	0.000	0.000
7	3	0	1	3	0	0	6	0	0.000	0.000	0.000
8	2	1	0	3	0	0	5	1	0.400	0.240	0.600
9	1	0	0	3	1	0	4	1	0.250	0.188	-0.250
10	1	0	0	2	1	0	3	1	0.333	0.222	-0.333
12	1	1	0	1	0	0	2	1	0.500	0.250	0.500
13	0	0	0	1	1	0	1	1	0.000	0.000	0.000
									1.63	1.20	

Log-rank test

$$\widehat{e}_{1i} = \frac{n_{1i}d_i}{n_i}; U = \sum_{i=1}^r \left(d_{1i} - \frac{n_{1i}d_i}{n_i} \right) = 1.20$$

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	n_{1i}	d_{1i}	c_{1i}	n_{0i}	d_{0i}	c_{0i}			e_{1i}	v_{1i}	$d_{1i} - e_{1i}$
0	5	0	0	5	0	0	10	0	0	0	0
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$$\widehat{\text{Var}}(\mathbf{U}) = \sum_{i=1}^r \frac{n_{1i}n_{0i}d_i(n_i - d_i)}{n_i^2(n_i - 1)} = 1.63$$

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$$\text{Log-rank test (Q)} = \frac{U^2}{\widehat{\text{Var}}(U)} = \frac{1.43}{1.63} = 0.88 \text{ Chi-square (1df), p-value=0.348}$$

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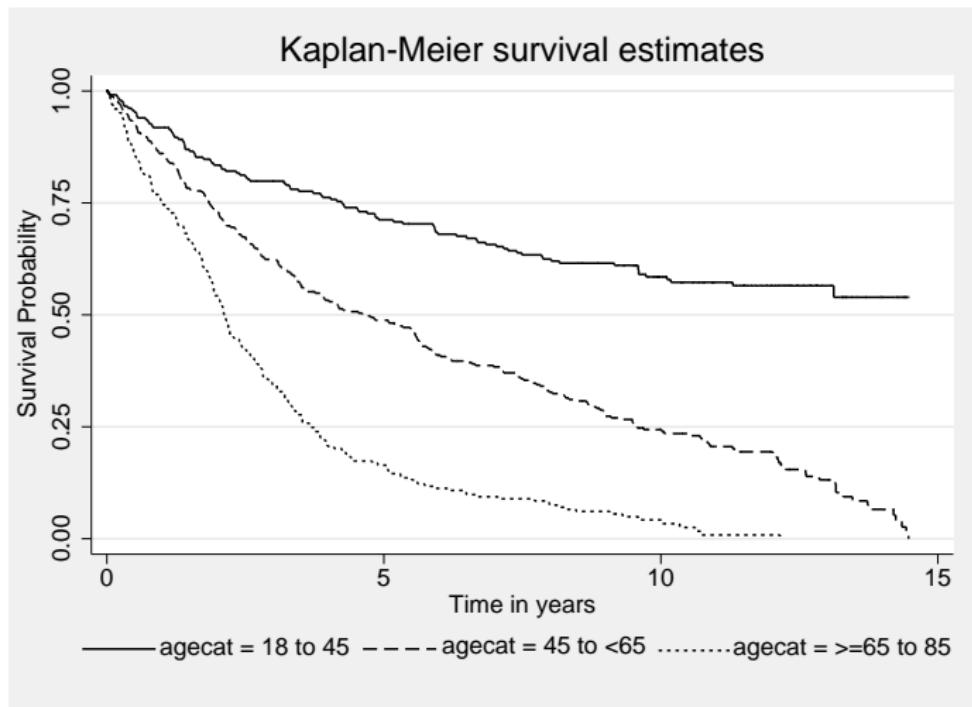
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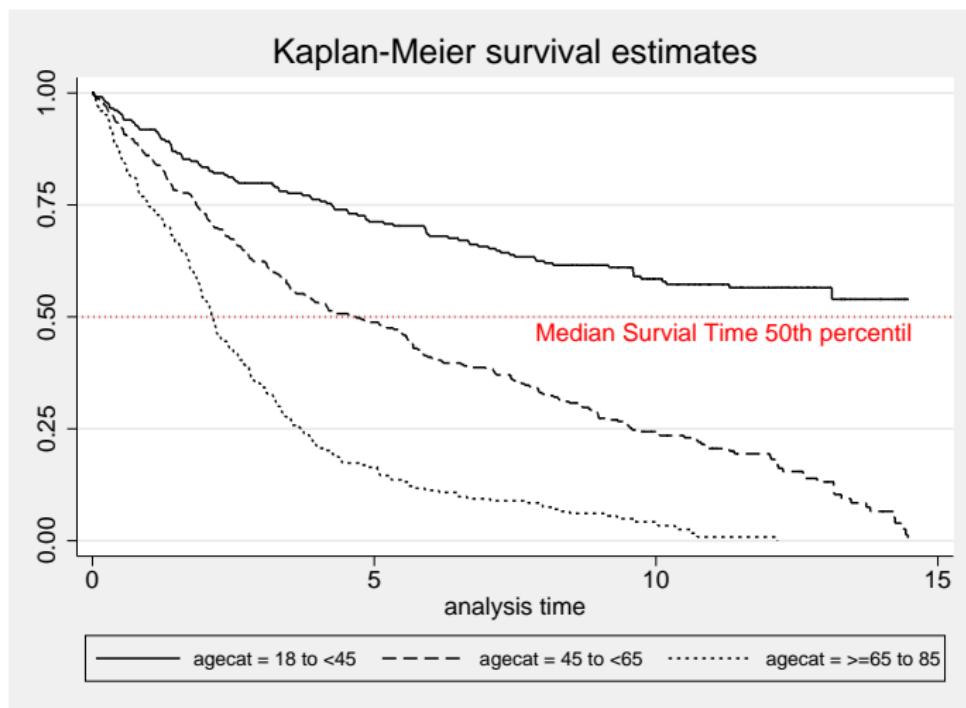
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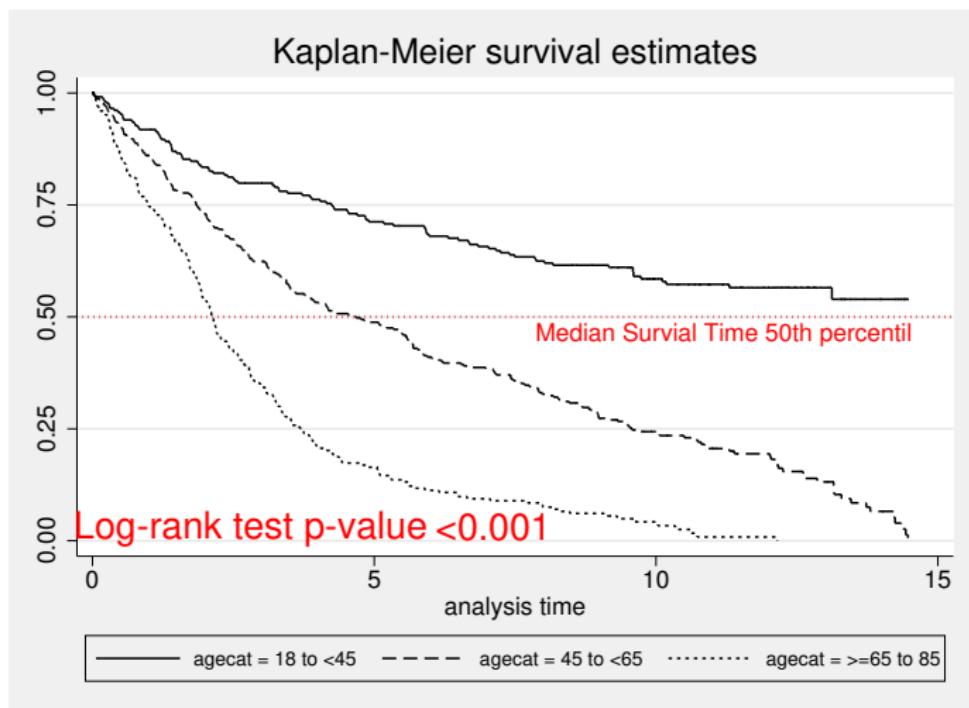
Kaplan-Meier



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Kaplan-Meier



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Hazard Models

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Hazard Models

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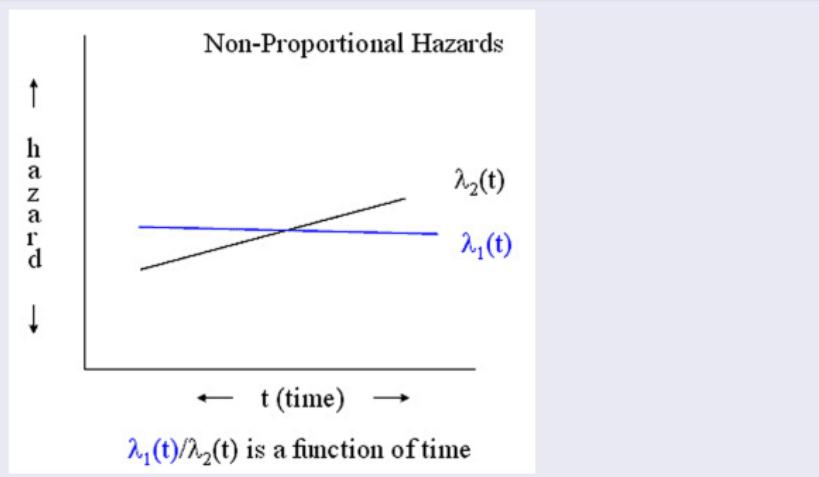
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Hazard Models

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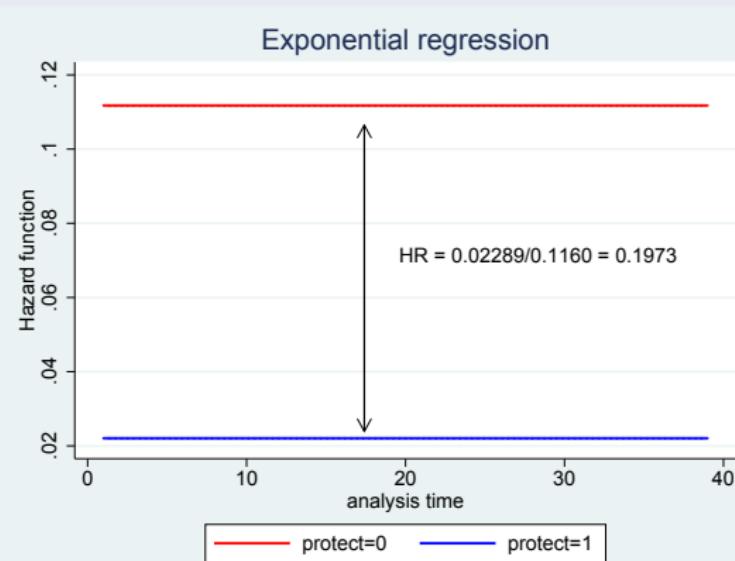
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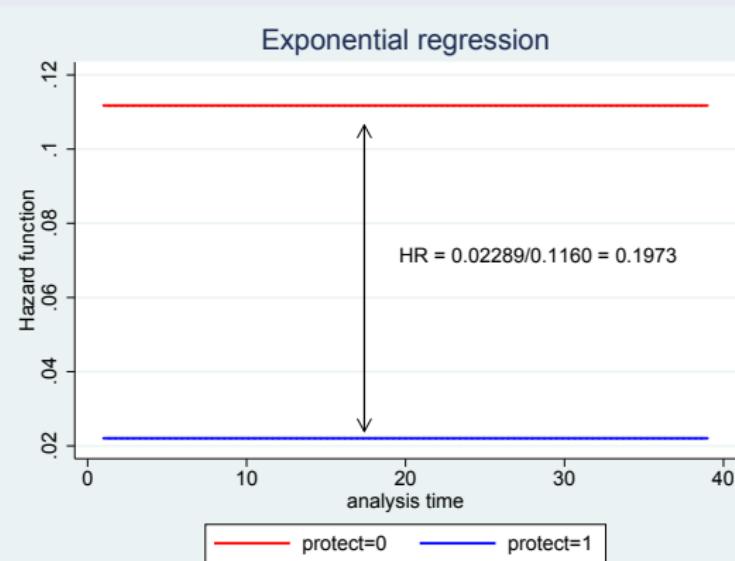
Proportionality of Hazards

Exponential model with constant Hazard



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Exponential Models

```
.use http://www.stata.com/courses/nc631-13/lec4/hip2, clear
. streg protect, dist(exp) hr
    failure _d: fracture
    analysis time _t: time1
    id: id
Exponential regression -- log relative-hazard form
No. of subjects =          48                      Number of obs     =      106
No. of failures =         31
Time at risk     =      714
Log likelihood   = -50.227658
                                         LR chi2(1)      =      19.68
                                         Prob > chi2     =     0.0000
-----+
          _t | Haz. Ratio    Std. Err.      z     P>|z|    [95% Conf. Interval]
-----+
  protect |   .1973684    .0727765    -4.40    0.000    .0958099    .4065789
  _cons |   .1117647    .0256406    -9.55    0.000    .0712895    .17522
-----+
```

Parametric Survival Distributions: Pintilie M. (2007)

Table 2.2 Parametric models for time-to-event data.

Distribution	$f(t)$	$F(t)$	$S(t)$	$h(t)$
Exponential	$\lambda e^{-\lambda t}$	$1 - e^{-\lambda t}$	$e^{-\lambda t}$	λ
Weibull	$\lambda \theta t^{\theta-1} \exp\{-\lambda t^\theta\}$	$1 - \exp\{-\lambda t^\theta\}$	$\exp\{-\lambda t^\theta\}$	$\lambda \theta t^{\theta-1}$
Gamma	$\lambda^\theta t^{\theta-1} e^{-\lambda t}$	$\text{II}(\theta, \lambda t)^\dagger$	$1 - \text{II}(\theta, \lambda t)$	$f(t)/S(t)$
Generalized gamma	$\frac{\alpha \lambda^\theta t^{\alpha\theta-1} \exp\{-\lambda t^\alpha\}}{\Gamma(\theta)}$	$\text{II}(\theta, \lambda t^\alpha)^\dagger$	$1 - \text{II}(\theta, \lambda t^\alpha)$	$f(t)/S(t)$
Lognormal	$\frac{\exp\left\{-\frac{1}{2}\left(\frac{\log t - \mu}{\sigma}\right)^2\right\}}{\sigma t \sqrt{2\pi}}$	$\Phi\left(\frac{\log t - \mu}{\sigma}\right)^\ddagger$	$1 - \Phi\left(\frac{\log t - \mu}{\sigma}\right)$	$f(t)/S(t)$
Log-logistic	$\frac{\lambda \theta t^{\theta-1}}{(1 + \lambda t^\theta)^2}$	$\frac{\lambda t^\theta}{1 + \lambda t^\theta}$	$\frac{1}{1 + \lambda t^\theta}$	$\frac{\lambda \theta t^{\theta-1}}{1 + \lambda t^\theta}$

† II is the incomplete gamma function: $\text{II}(\theta, \lambda t) = \int_0^t \lambda^\theta t^{\theta-1} e^{-\lambda t} dt / \Gamma(\theta)$.

‡ Φ is the cdf of the standard normal distribution.

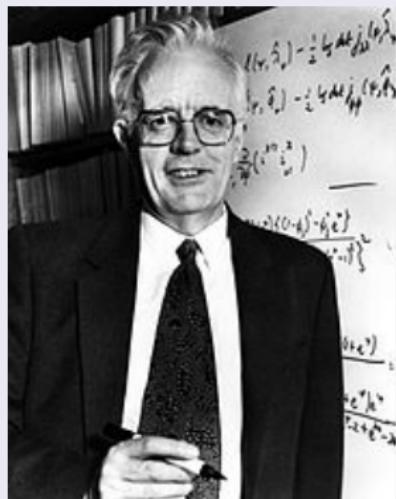
Parametric Survival Distributions

Shiny application

<http://watzilei.com/shiny>

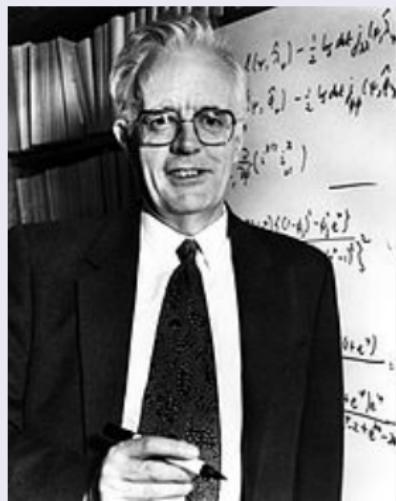
Semi-parametric Cox Proportional Hazard Model

Sir David Cox



Semi-parametric Cox Proportional Hazard Model

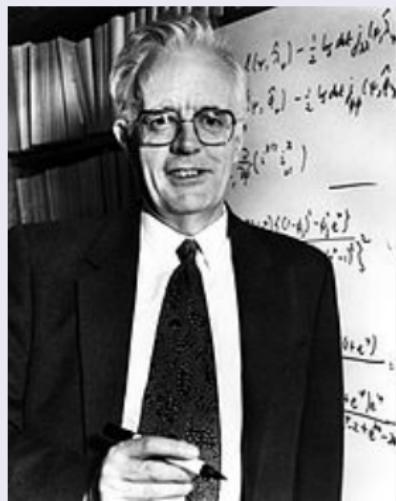
Sir David Cox



Sir David Cox has made pioneering and important contributions to numerous areas of statistics and applied probability, of which the best known is perhaps the semi-parametric proportional hazards model

Semi-parametric Cox Proportional Hazard Model

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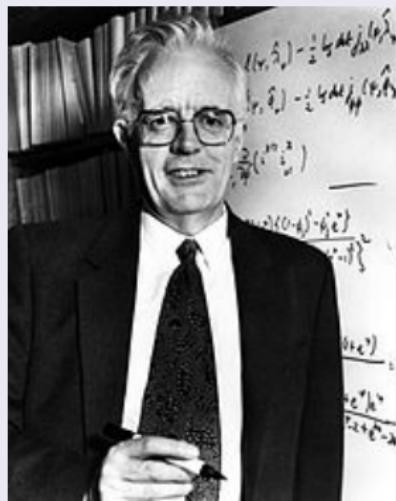


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$$h(t|x_j) =$$

Semi-parametric Cox Proportional Hazard Model

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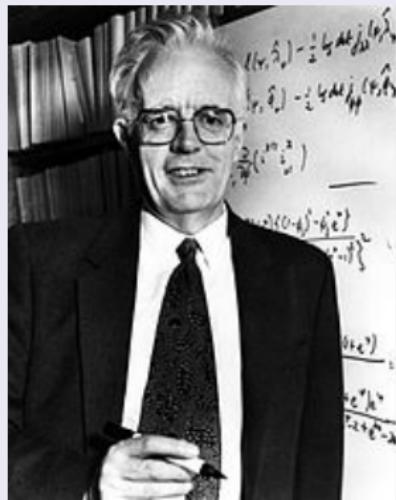


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$$h(t|x_j) = [h_0(t)$$

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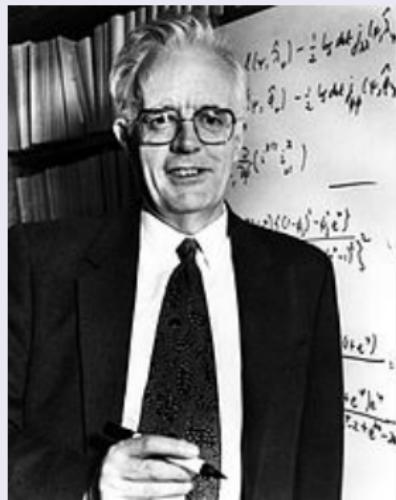


Sir David Cox has made pioneering and important contributions to numerous areas of statistics and applied probability, of which the best known is perhaps the semi-parametric proportional hazards model

$$h(t|x_i) = [h_0(t) \times \exp(\beta_0)]$$

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Sir David Cox

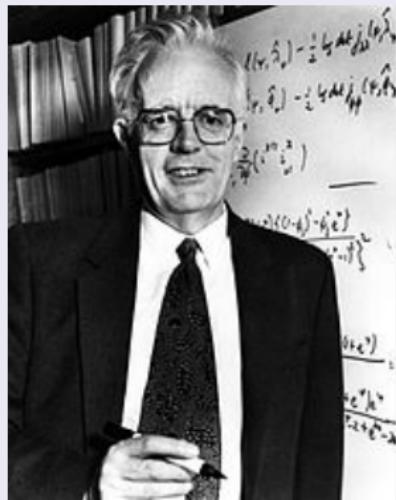


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$$h(t|x_j) = [h_0(t) \times \exp(\beta_0)] \times$$

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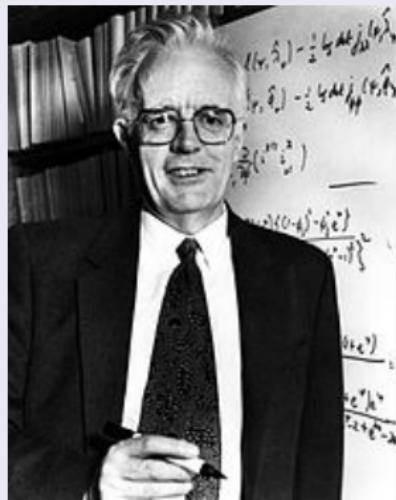


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$$h(t|x_j) = [h_0(t) \times \exp(\beta_0)] \times \exp(\beta x_j)$$

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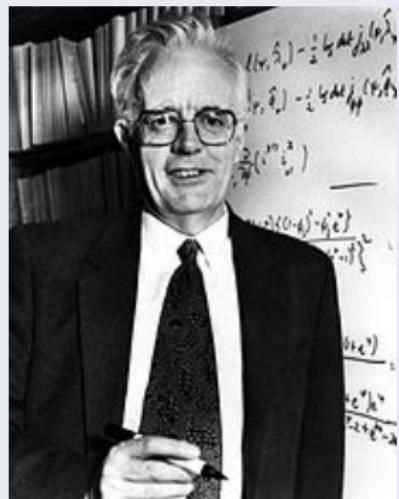


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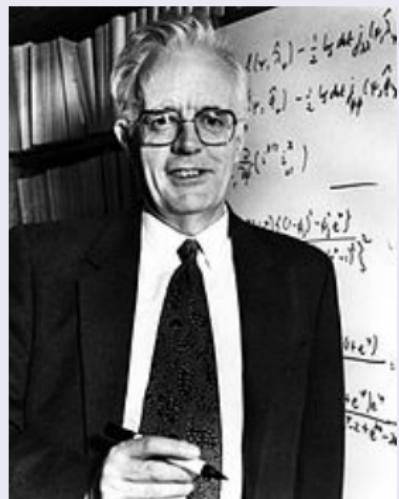
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Semiparametric Cox Model

- Assumptions about the functional form of the $h_0(t)$ is unspecified

Semi-parametric Cox Proportional Hazard Model

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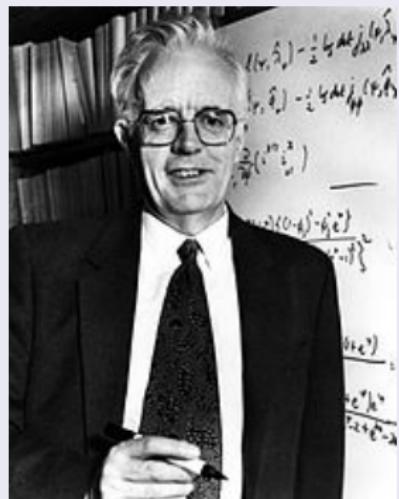
$$h(t|x_j) = [h_0(t) \times \exp(\beta_0)] \times \exp(\beta x_j)$$

Semiparametric Cox Model

- Assumptions about the functional form of the $h_0(t)$ is unspecified
- The constant of the model cannot be estimated and is absorbed into the baseline hazard $h_0(t)$

Semi-parametric Cox Proportional Hazard Model

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$$h(t|x_j) = [h_0(t) \times \exp(\beta_0)] \times \exp(\beta x_j)$$

Semiparametric Cox Model

- Assumptions about the functional form of the $h_0(t)$ is unspecified
- The constant of the model cannot be estimated and is absorbed into the baseline hazard $h_0(t)$

Cox Model in Stata

```
. use http://www.stata.com/courses/nc631-13/lec4/hip2, clear  
(hip fracture study)  
. stcox protect, nohr  
failure _d: fracture  
analysis time _t: time1  
id: id
```

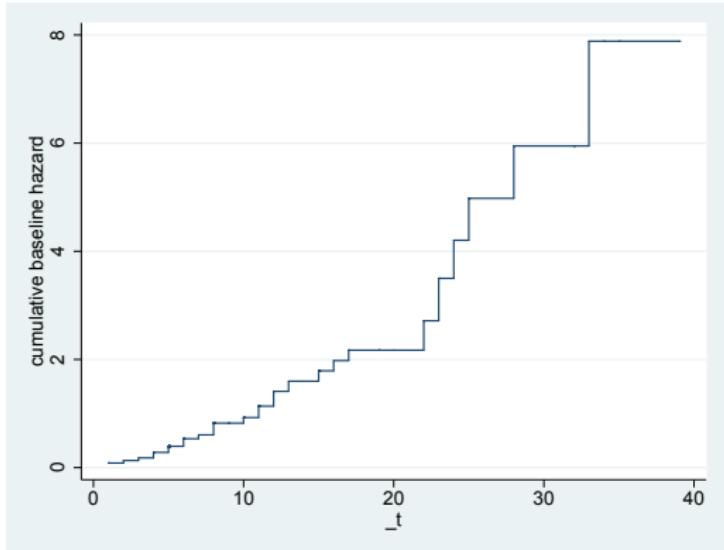
_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
protect	-2.047599	.4404029	-4.65	0.000	-2.910773 -1.184426

Hazard Ratio for wearing the protect device = ? ==> $\exp(-2.05) = 0.129$

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
protect	.1290443	.0568315	-4.65	0.000	.0544336 .3059218

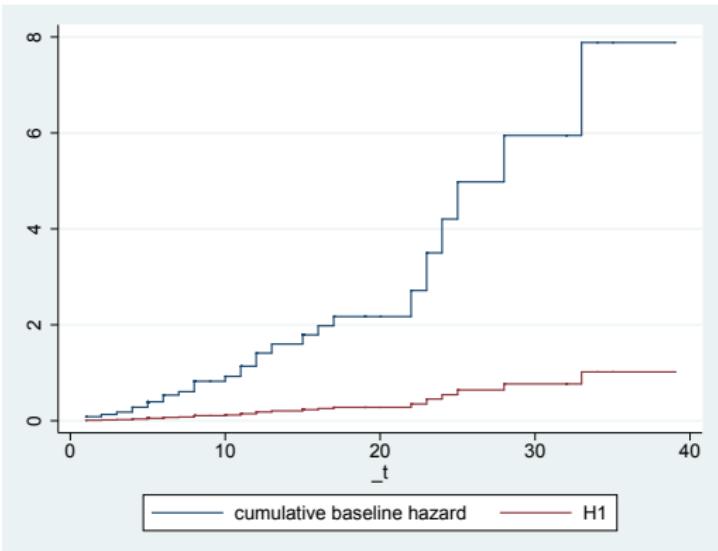
Cumulative Hazard in Stata

- predict H0, basehazard
- twoway line H0 _t,
connect(J) sort



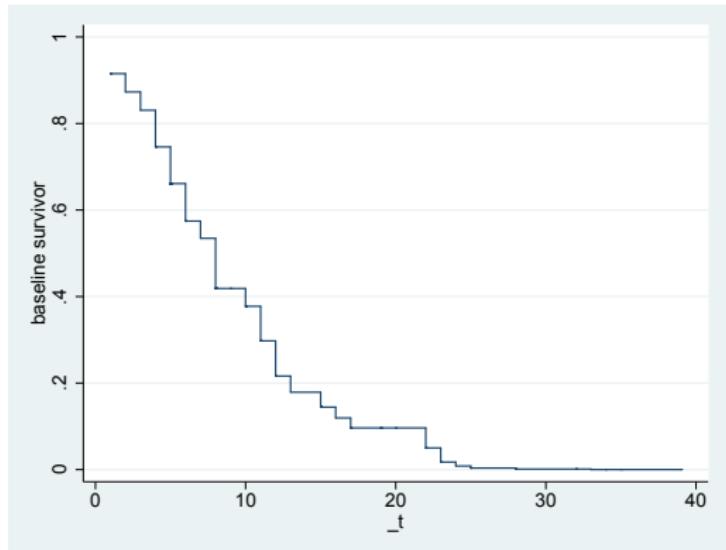
Cumulative Hazard in Stata

- generate $H1 = H0 * \exp(-_b[protect])$
- twoway line H0 H1 _t, sort connect(J J)
- stcurve, cumhaz
at1(protect = 0)
at2(protect=1)



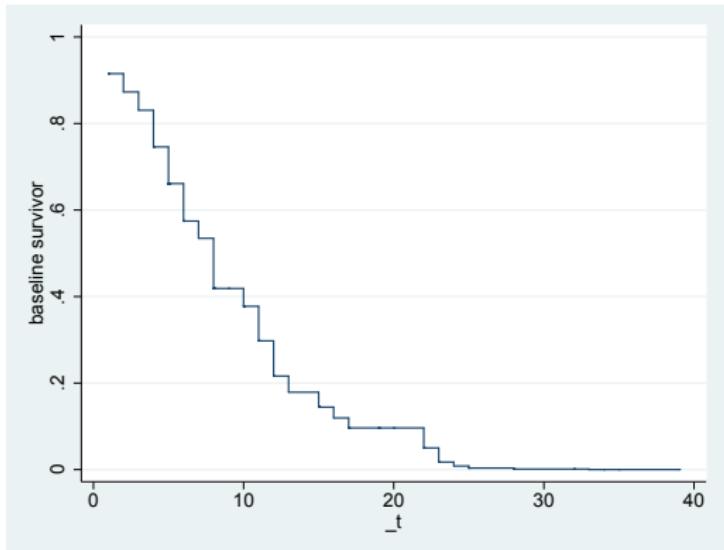
Baseline Survival in Stata

- predict S0, basesurv
- twoway line S0 _t,
connect(J) sort



Baseline Survival in Stata

- predict S0, basesurv
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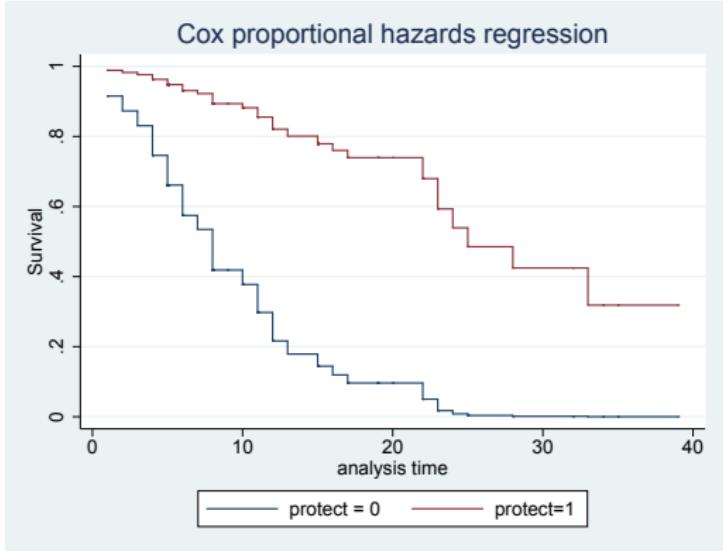


Survival by groups in Stata

- generate S1 =
 $S_0 \cdot \exp(-b[\text{protect}])$
- twoway line S0 S1 _t,
sort connect(J J)
- stcurve, surv at1(protect
= 0) at2(protect=1)

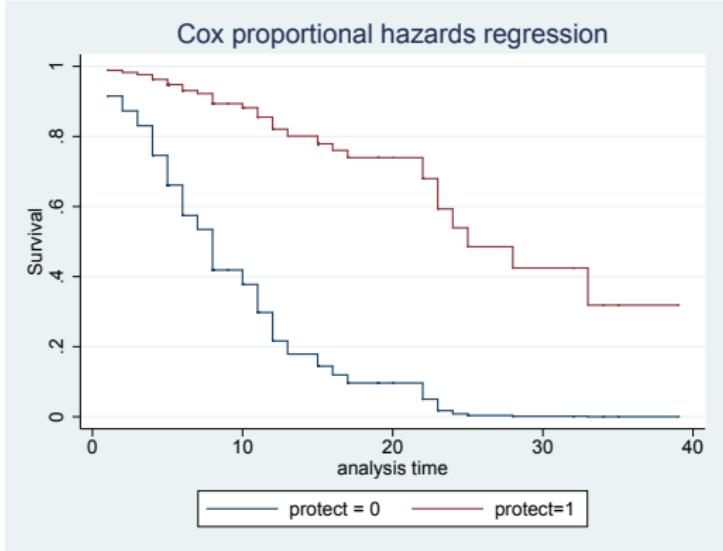
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Survival by groups in Stata

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Final overview

<https://www.mailman.columbia.edu/research/population-health-methods/time-event-data-analysis>

Thank you!

Muchas Gracias!