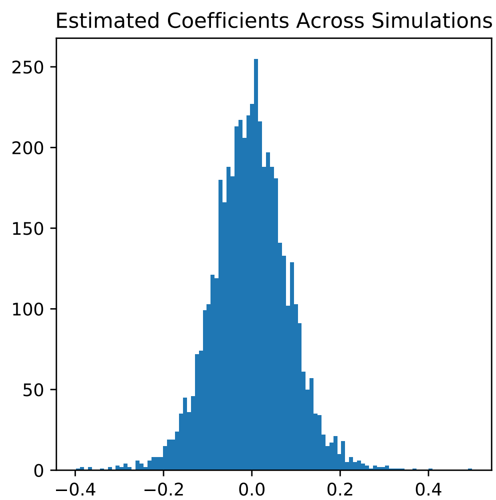
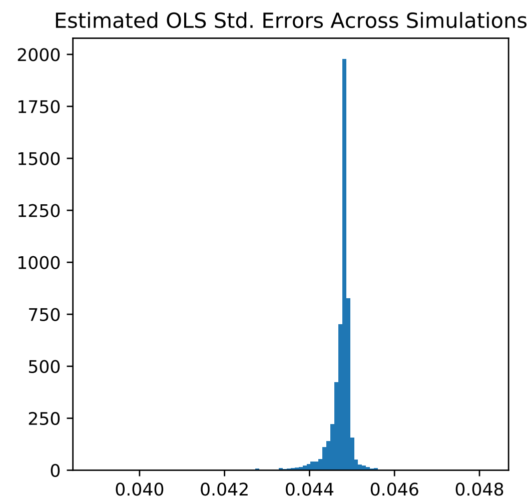
Problem Set 2
Econ 6413–Ng
Miguel Acosta
September 22, 2017

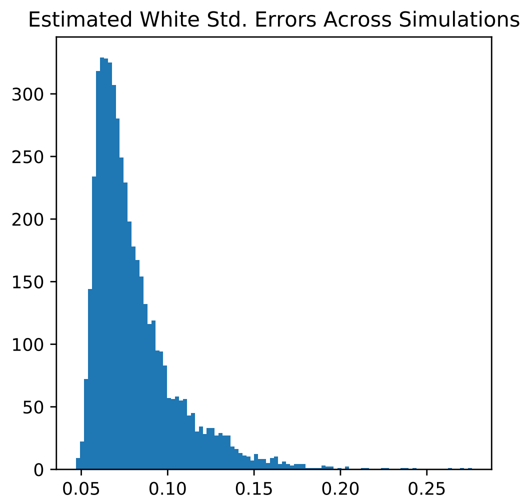
Q1. In the results below, I've assumed that $v_t \sim N(0,1)$, as in Hamilton (2008). I am able to replicate table 2 of Hamilton (2008). Here are the results for $T = 500$ and 5000 replications:

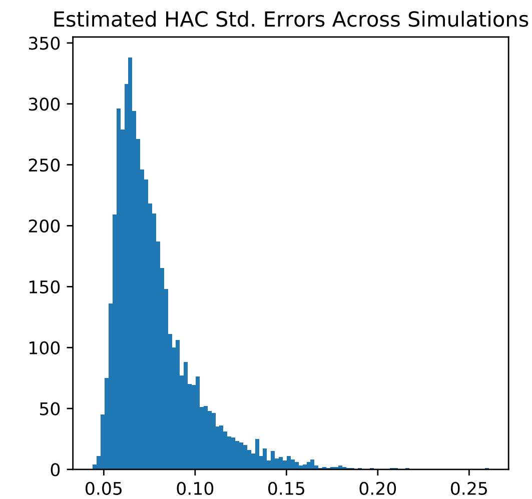
Standard Error	Size
OLS	0.280
White	0.062
HAC, $q = 5$	0.062

For fun, here are the distributions of the $\widehat{\beta}$ and standard errors across the simulations.





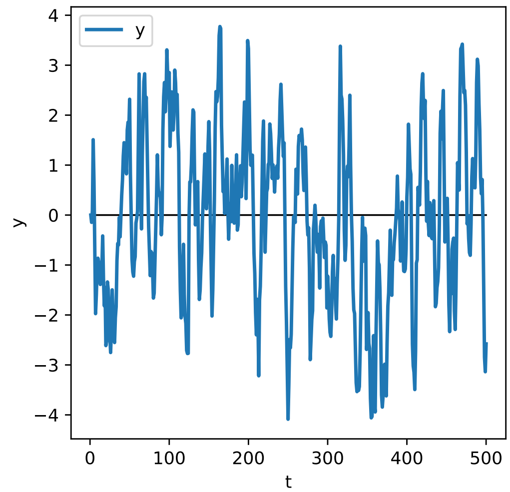




Q2. We are considering an ARMA$(1,1)$:

$$y_t = \alpha y_{t-1} + \beta e_{t-1} + e_t \qquad\qquad e_t \sim N(0, \sigma^2)$$

Here is a simulated series with 500 time periods and $\beta \equiv (\alpha, \beta, \sigma^2) = (0.8, 0.5, 0.5)$.



i. In problem set 1, I calculated γ_0 and γ_1 as:

$$\gamma_0 = \frac{\sigma^2(2\alpha\beta + \beta^2 + 1)}{1-\alpha^2} \qquad\qquad \gamma_1 = \sigma^2 \cdot \frac{\alpha^2\beta + \beta^2\alpha + \beta + \alpha}{1-\alpha^2}. \tag{1}$$

Remember that the series is mean-0. So, the second-order autocovariance is:

$$\gamma_2 = \mathbb{E}[y_t y_{t-2}] = \alpha y_{t-1}y_{t-2} + \beta e_{t-1}y_{t-2} + e_t y_{t-2}] = \alpha\gamma_1 \tag{2}$$

ii. The GMM objective function, therefore, is:

$$\overline{g}(\theta) \equiv \frac{1}{T-3}\sum_{t=3}^{T} \widehat{\psi} - \gamma(\theta) = \frac{1}{T-3}\sum_{t=3}^{T} \begin{bmatrix} y_t^2 - \frac{\sigma^2(2\alpha\beta+\beta^2+1)}{1-\alpha^2} \\ y_t y_{t-1} - \sigma^2 \cdot \frac{\alpha^2\beta+\beta^2\alpha+\beta+\alpha}{1-\alpha^2} \\ y_t y_{t-2} - \alpha\sigma^2 \cdot \frac{\alpha^2\beta+\beta^2\alpha+\beta+\alpha}{1-\alpha^2} \end{bmatrix}$$

Results follow:

α	0.8315
β	0.4478
σ^2	0.4546

iii. The third and fourth centered moments of an ARMA$(1,1)$ will be difficult to calculate, as you can see here (thanks to SymPy for these expressions):

$$\begin{aligned}
\mathbb{E}[y_t^3] &= \mathbb{E}[\alpha^3 y_{t-1}^3 + 3\alpha^2\beta e_{t-1} y_{t-1}^2 + 3\alpha^2 e_t y_{t-1}^2 + 3\alpha\beta^2 e_{t-1}^2 y_{t-1} \\
&\quad + 6\alpha\beta e_{t-1} e_t y_{t-1} + 3\alpha e_t^2 y_{t-1} + \beta^3 e_{t-1}^3 + 3\beta^2 e_{t-1}^2 e_t + 3\beta e_{t-1} e_t^2 + e_t^3] \\
\mathbb{E}[y_t^4] &= \mathbb{E}[\alpha^4 y_{t-1}^4 + 4\alpha^3\beta e_{t-1} y_{t-1}^3 + 4\alpha^3 e_t y_{t-1}^3 + 6\alpha^2\beta^2 e_{t-1}^2 y_{t-1}^2 \\
&\quad + 12\alpha^2\beta e_{t-1} e_t y_{t-1}^2 + 6\alpha^2 e_t^2 y_{t-1}^2 + 4\alpha\beta^3 e_{t-1}^3 y_{t-1} + 12\alpha\beta^2 e_{t-1}^2 e_t y_{t-1} \\
&\quad + 12\alpha\beta e_{t-1} e_t^2 y_{t-1} + 4\alpha e_t^3 y_{t-1} + \beta^4 e_{t-1}^4 + 4\beta^3 e_{t-1}^3 e_t + 6\beta^2 e_{t-1}^2 e_t^2 \\
&\quad + 4\beta e_{t-1} e_t^3 + e_t^4]
\end{aligned}$$

Luckily, most of these terms are 0, or simplify substantially. In particular, note that

$$\mathbb{E}[f(e_t)y_{t-1}] = \mathbb{E}[e_t e_{t-1} y_{t-1}] = \mathbb{E}[e_t f(y_{t-1})] = 0 \qquad \mathbb{E}[e_t^3] = 0 \qquad \mathbb{E}[e_t^4] = 3\sigma^4$$

The first expression is therefore:

$$\mathbb{E}[y_t^3] = \alpha^3\,\mathbb{E}[y_{t-1}^3] \Longrightarrow \ \mathbb{E}[y_{t-1}^3] =$$

(assuming $\mathbb{E}[y_0^3] = 0$, which is consistent with the assumption that $y_0 = 0$). The second expression reduces to

$$\mathbb{E}[y_t^4] = \alpha^4\,\mathbb{E}[y_{t-1}^4] + 3\sigma^4(\beta^4 + 4\alpha\beta^3 + 6\alpha^2\beta^2 + 4\alpha^3\beta + 1).$$

Since we only have covariance stationarity, the 4th moment depends on time, and can be calculated recursively assuming $\mathbb{E}[y_t^4] = 0$.

With all this, I present results that didn't work.

α	-0.8128
β	0.9982
σ^2	2.5513

Q3. **i.** Remember what OLS does:

$$\begin{bmatrix}\alpha_1\\ \alpha_2\end{bmatrix} = \begin{bmatrix}\gamma_0 & \gamma_1\\ \gamma_1 & \gamma_0\end{bmatrix}^{-1}\begin{bmatrix}\gamma_1\\ \gamma_2\end{bmatrix}$$

We can take advantage of the expressions in 1 and 2 to back out the estimates for $(\alpha, \beta, \sigma^2)$. Multiply both sides by the projection matrix to get

$$\begin{bmatrix}\gamma_0 & \gamma_1\\ \gamma_1 & \gamma_0\end{bmatrix}\begin{bmatrix}\alpha_1\\ \alpha_2\end{bmatrix} = \begin{bmatrix}\gamma_1\\ \gamma_2\end{bmatrix}$$

We can estimate the RHS easily. So, this identifies α since $\frac{\gamma_2}{\gamma_1} = \alpha$. Next, taking the ratio of γ_0 to γ_1, we have one equation to solve for $\widehat{\beta}$. I asked SymPy to do this. Finally, we can use the γ_0 equation to get σ^2.

ii. The OLS estimates are:

	Estimate	Newey-West S.E.
ϕ_1	1.2144	0.0394
ϕ_2	-0.3424	0.0393

The estimate of $\phi(1)$ is 0.87 [0.017]. The number in brackets is the Newey-West standard error with $q = 5$. Inverting this to get the parametets of the ARMA$(1,1)$ model as discussed in the previous subquestion results in:

α	0.8315
β	0.4168
σ^2	0.4717

iii. I don't know what this means...

Q4. I went with qustion (i). Here are the results for SMM, using autocovariances of order 0, 1, and 2 as the moment conditions:

α	0.8383
β	0.4365
σ^2	0.4573

Here are the results using indirect inference, with an AR(2) as the auxiliary model:

α	0.8405
β	0.4379
σ^2	0.5000

For each of the results above, I used 501 simulations, and kept $T = 500$. I used the same sample of iid shocks in both simulation estimations.