## Problem Set 1 GR6493 [Dean] Miguel Acosta & Sara Shahanaghi November 15, 2017

Question 1 For those of us that may have forgotten what NIAS and NIAC are, define:

$$\left[ P_A(a \mid \omega) \equiv \sum_{\gamma \in \Gamma(A)} \pi_A(\gamma \mid \omega) C_A(a \mid \gamma) \right] \quad \left[ g(\gamma, A) \equiv \max_{a \in A} \sum_{\omega \in \Omega} \gamma(\omega) u(a, \omega) \right] \quad \left[ G(\pi, A) \equiv \sum_{\gamma \in \Gamma(\pi)} P(\gamma) g(\gamma, A) \right]$$

**NIAS** For every chosen action a,

$$\sum \mu(\omega) P_A(a \mid \omega) [u(a(\omega)) - u(b(\omega))] \ge 0, \ \forall b \in A$$
 (1)

**NIAC** For an observed sequence of decision problems  $A_1, \ldots, A_K$ , and associated revealed information stuctures  $\overline{\pi}^1, \ldots, \overline{\pi}^K$ ,

$$G(A^1, \overline{\pi}^1) - G(A^1, \overline{\pi}^2) + \dots + G(A^K, \overline{\pi}^K) - G(A^K, \overline{\pi}^1) \ge 0.$$

We claim that the data generated by Gabaix's model does not necessarily satisfy the NIAS condition. We will demonstrate this using a counterexample. The counterexample we present is a one-dimensional version of the quadratic utility example presented in section II.A of Gabaix's paper.

Let a be the agent's choice of action, x the true unobserved state, and m the agent's attention strategy. Fix any choice of  $m \in (0,1)$  (i.e., the agent is not paying full attention). Let the agent's prior belief regarding x given by  $\mu(x)$ . By assumption, the mean and variance of this distribution exist and are given by 0 and  $\sigma^2$ , respectively. Define quadratic utilities as

$$u(a,x) = -\frac{1}{2}(a - cx)^2$$

where c > 0 is some given constant. Define  $x^s \equiv mx$ According to Gabaix's model, the agent chooses

$$a = arg \max_{a} u(a, x^s) = cx^s = cmx.$$

Define  $b \equiv cx$ . Referring back to our utility function, we know u(b, x) = 0 for all x. Now, fix any given realization of the state x. We know that

$$u(a,\omega) = -\frac{1}{2}(cmx - cx)^2 < 0 = u(b,\omega)$$

The above inequality holds strictly since we assumed  $m \in (0,1)$ . So, for any prior  $\mu$ , we have

$$\int_{-\infty}^{\infty} \mu(x) [u(a(\omega) - u(b(\omega))] dx = \int_{-\infty}^{\infty} \mu(x) - \frac{1}{2} (cmx - cx)^2 dx < 0$$

Thus, we conclude that NIAS fails.

## Question 2

1. First, let's establish notation. Each decision problem  $D \in \{8, 9, 10, 11\}$  in this experiment has two possible actions—denote these by  $a_D$  and  $b_D$  (so, for D=9, we have  $a_D=12$  and  $b_D=13$ ). Let the typical state be denoted by  $\omega \in \Omega \equiv \{1, 2, 5, 6\}$ .

The NIAS condition (1) for this experiment requires the following conditions to hold for all D

$$\sum_{n \in \mathcal{D}} P(a_D \mid \omega) [u(a_D(\omega)) - u(b_D(\omega))] \ge 0$$
(2a)

$$\sum_{\omega \in \Omega} P(b_D \mid \omega) [u(b_D(\omega)) - u(a_D(\omega))] \ge 0$$
(2b)

Since there are only two possible actions in each decision problem, we have

$$P(a_D \mid \omega) + P(b_D \mid \omega) = 1, \forall \omega \in \Omega.$$

Using this, and the definition  $\Delta_{\omega} \equiv u(a_D(\omega)) - u(b_D(\omega))$  allows (2) to be written as

$$\sum_{\omega \in \Omega} P(a_D \mid \omega) \Delta_{\omega} \ge 0$$
$$\sum_{\omega \in \Omega} (1 - P(a_D \mid \omega))(-\Delta_{\omega}) \ge 0,$$

i.e.,

$$\sum_{\omega \in \Omega} P(a_D \mid \omega) \Delta_{\omega} \ge 0$$
$$\sum_{\omega \in \Omega} P(a_D \mid \omega) \Delta_{\omega} \ge \sum_{\omega \in \Omega} \Delta_{\omega},$$

which combine to form the (still necessary and sufficient) condition for NIAS:

$$\sum_{\omega \in \Omega} P(a_D \mid \omega) \Delta_{\omega} \ge \max \left\{ \sum_{\omega \in \Omega} \Delta_{\omega}, 0 \right\} = 0$$

where the last equality follows from the symmetry of the specific payoffs of this experiment. I have two ideas for testing.

- i. Test this condition for each person. Then, take the mean and standard deviation of those probabilities to do a standard t-test (or whatever... mean/sd)
- ii. Boostrapping. Drop observations that aren't state omega. Then sample (with replacement) from this pool and calculate the condition. Then use the s.d. of those bootstrapped probabilities to do a t-test.

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## Question 3

1. For concreteness, the payoffs are

$$\begin{array}{c|cccc} & I & H \\ \hline d & -c_I & -c_H \\ n & -s & 0 \end{array} & \leq & \begin{array}{c|ccccc} & I & H \\ \hline d & -\overline{c}_I & -\overline{c}_H \\ n & -s & 0 \end{array}$$

with  $\bar{c}_I < c_I < s$  and  $c_H > \max\{0, \bar{c}_H\}$ .

**NIAS** In this example, the following conditions constitute NIAS:

$$\sum_{\omega \in \{I, H\}} \mu_{\omega} P(d \mid \omega) [u(d(\omega)) - u(n(\omega))] \ge 0$$
$$\sum_{\omega \in \{I, H\}} \mu_{\omega} P(n \mid \omega) [u(n(\omega)) - u(d(\omega))] \ge 0$$

i.e.,

$$\mu_I P(d \mid I)[-c_I + s] + (1 - \mu_I)P(d \mid H)[-c_H] \ge 0$$
 (3a)

$$\mu_I P(n \mid I)[-s + c_I] + (1 - \mu_I)P(n \mid H)[c_H] \ge 0.$$
 (3b)

Replacing  $P(n \mid \omega)$  with  $1 - P(d \mid \omega)$  in (3b)

$$\mu_I(1 - P(d \mid I))[-s + c_I] + (1 - \mu_I)(1 - P(d \mid H))[c_H] \ge 0$$

i.e.,

$$\mu_I P(d \mid I)[-c_I + s] + (1 - \mu_I)P(d \mid H)[-c_H] \ge \mu_I(s - c_I) - (1 - \mu_I)c_H$$

Which combines with (3a) to form a single NIAS condition

$$\mu_I P(d \mid I)[-c_I + s] + (1 - \mu_I) P(d \mid H)[-c_H] \ge \max\{\mu_I(s - c_I) - (1 - \mu_I)c_H, 0\}$$
 (4)

Let's examine the RHS:

$$\mu_I(s - c_I) - (1 - \mu_I)c_H > 0 \iff \frac{\mu_I}{1 - \mu_I} > \frac{c_H}{s - c_I}$$
 (5)

When this is true, then condition (4) becomes

$$\mu_{I}P(d \mid I)[-c_{I} + s] + (1 - \mu_{I})P(d \mid H)[-c_{H}] \ge \mu_{I}(s - c_{I}) - (1 - \mu_{I})c_{H}$$

$$\iff \frac{\mu_{I}}{1 - \mu_{I}}P(d \mid I) - P(d \mid H)\frac{c_{H}}{s - c_{I}} \ge \frac{\mu_{I}}{1 - \mu_{I}} - \frac{c_{H}}{s - c_{I}} > 0$$

$$\iff \frac{P(d \mid I)}{P(d \mid H)} > \frac{c_{H}}{s - c_{I}} \left[\frac{\mu_{I}}{1 - \mu_{I}}\right]^{-1}$$
(6)

Otherwise, when (5) is reversed, then (4) becomes

$$\mu_{I}P(d \mid I)[-c_{I} + s] + (1 - \mu_{I})P(d \mid H)[-c_{H}] \ge 0$$

$$\iff \frac{\mu_{I}}{1 - \mu_{I}}P(d \mid I) - P(d \mid H)\frac{c_{H}}{s - c_{I}} \ge 0$$

$$\iff \frac{P(d \mid I)}{P(d \mid H)} \ge \frac{c_{H}}{s - c_{I}} \left[\frac{\mu_{I}}{1 - \mu_{I}}\right]^{-1} > 1$$

$$\implies P(d \mid I) \ge P(d \mid H)$$

Which seems sensible. Not sure what to make of (6). These conditions are the same under the subsidy, just replacing the non-barred variables by their barred-counterparts.

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