

Lecture 4: Probability and Bayesian Classifier

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- A key concept in the field in machine learning is that of uncertainty
 - Through noise on measurements
 - Through the finite size of data sets
- Probability theory provides a consistent framework for the quantification and manipulation of uncertainty
- Forms one of the central foundations for pattern recognition.

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Kolmogorov's Axioms of Probability (1933)

- To each sentence a , a numerical degree of belief between 0 and 1 is assigned

$$0 \leq p(a) \leq 1$$

$$p(\text{true})=1, p(\text{false})=0$$

- The probability of disjunction is given by

$$p(a \vee b) = p(a) + p(b) - p(a \wedge b)$$



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Where do these numerical degrees of belief come from?

- Humans can *believe* in a subjective viewpoint from *experience*. This approach is called **Bayesian**
- For a finite sample we can estimate the true fraction. We count the *frequency* of an event in a *sample*. We do not know the true value because we cannot access the whole population of events. This approach is called **frequentist**
- From the true nature of the universe, for example, for a fair coin, the probability of heads is 0.5. This approach is related to the **Platonic world** of ideas. However, we can never verify whether a fair coin exists

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- From the frequentist approach, one can determine the probability of an event a by counting
- If Ω is the set of all possible events, $p(\Omega) = 1$, then $a \in \Omega$.
- $\text{card}(\Omega)$ is the number of elements of the set Ω , $\text{card}(a)$ is the number of elements of the set a and

$$p(a) = \frac{\text{card}(a)}{\text{card}(\Omega)}$$

$$p(a \wedge b) = \frac{\text{card}(a \wedge b)}{\text{card}(\Omega)}$$

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- Now we can define the posterior probability, the probability of a after the evidence b is obtained

$$p(a|b) = \frac{\text{card}(a \wedge b)}{\text{card}(b)}$$

- using

$$p(a \wedge b) = \frac{\text{card}(a \wedge b)}{\text{card}(\Omega)}$$

- we get

$$p(a|b) = \frac{p(a \wedge b)}{p(b)} \quad p(b|a) = \frac{p(a \wedge b)}{p(a)}$$

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Bayes' Rule

$$p(a|b) = \frac{p(a \wedge b)}{p(b)} \quad p(b|a) = \frac{p(a \wedge b)}{p(a)}$$

- The Bayes' rule follows from both equations

$$p(b|a) = \frac{p(a|b) \cdot p(b)}{p(a)}$$

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Law of Total Probability

- For mutually exclusive events b_1, \dots, b_n with

$$\sum_{i=1}^n p(b_i) = 1$$

- the **law of total probability** is represented by

$$p(a) = \sum_{i=1}^n p(a \wedge b_i) = \sum_{i=1}^n p(a, b_i)$$

$$p(a) = \sum_{i=1}^n p(a|b_i) \cdot p(b_i)$$

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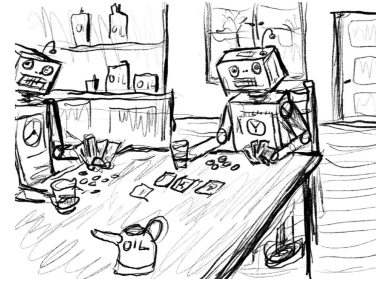
The Rules of Probability

Sum Rule

$$p(X) = \sum_Y p(X, Y)$$

Product Rule

$$p(X, Y) = p(Y|X)p(X)$$



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Bayes' Rule

$$p(a|b) = \frac{p(a \wedge b)}{p(b)} \quad p(b|a) = \frac{p(a \wedge b)}{p(a)}$$

- The Bayes' rule follows from both equations

$$p(b|a) = \frac{p(a|b) \cdot p(b)}{p(a)}$$

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Reverent Thomas Bayes (1702-1761)



- He set down his findings on probability in “Essay Towards Solving a Problem in the Doctrine of Chances” (1763), published posthumously in the Philosophical Transactions of the Royal Society of London.
 - The drawing after a portrait of Bayes used in a 1936 book, it is not known if the portrait is actually representing him.

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Bayes' Rule

$$p(h_k|D) = \frac{p(D|h_k) \cdot p(h_k)}{p(D)} = \frac{p(D, h_k)}{p(D)}$$

- $p(h_k)$ is called the **prior** (before)
 - For example, what is the probability of some illness in Portugal
- $p(D|h_k)$ is called **likelihood** and can be easily estimated
 - For example, what is the probability that some illness generates some symptoms?
- $p(D, h_k)$ is called **joint distribution**
- $p(h_k|D)$ is called **posterior probability**

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Bayes' Rule

$$p(h_k|D) = \frac{p(D|h_k) \cdot p(h_k)}{p(D)} = \frac{p(D, h_k)}{p(D)}$$

- Bayes rule can be used to determine the total posterior probability $p(h_k|D)$ of hypothesis h_k given data D
 - For example, what is the probability that some illness is present?
- The most probable hypothesis h_k out of a set of possible hypothesis h_1, h_2, \dots given some present data is according to the Bayes rule

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Maximum a Posteriori (MAP) Hypothesis

- $p(h_k|D)$ and $p(D, h_k)$ are related in a linear manner

$$p(h_k|D) \propto p(D|h_k) \cdot p(h_k)$$

posterior \propto likelihood \times prior

- to determine the **maximum posteriori hypothesis** h_{MAP} we maximize

$$h_{\text{MAP}} = \arg \max_{h_k} \frac{p(D|h_k) \cdot p(h_k)}{p(D)}$$

- we can see, the maximization is independent of $p(D)$, it follows

$$h_{\text{MAP}} = \arg \max_{h_k} p(D|h_k) \cdot p(h_k)$$

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Maximum Likelihood (ML) hypothesis

- If we assume $p(h_k) = p(h_v)$ for all h_k and h_v , then we can further
- simplify, and choose the maximum likelihood (ML) hypothesis

$$h_{ML} = \arg \max_{h_k} p(D|h_k)$$

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Bayesian Interpretation

- In the Bayesian (or epistemological) interpretation, **probability measures a “degree of belief”** and Bayes’ rule links the degree of belief in a proposition before and **after** accounting for evidence
- with *prior probability* $p(h_k)$
- $p(D|h_k)$ represents the likelihood of the data D if we assume h_k to be true
 - if we, in fact, observe D , we can update our belief about h_k through the rule

$$p(h_k|D) = \frac{p(D|h_k) \cdot p(h_k)}{p(D)}$$

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Bayesian Interpretation and bias

- Objective likelihood is biased by the prior belief

$$\text{posterior} \propto \text{likelihood} \times \text{prior} = \text{likelihood} \times \text{bias}$$

- Bias is a disproportionate weight in favor of or against an idea or thing, usually in a way that is closed-minded, prejudicial, or unfair.
- Biases can be innate or learned. People may develop biases for or against an individual, a group, or a belief.

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Cancer screening

- Cancer screening aims to detect cancer before symptoms appear
- This may involve for example a blood test.
- Suppose that a patient tests positive...
- The test is secure because in **99** percent of the cases the test returns a correct positive result (= positive) in which a rare form of cancer is actually present.
- Should the doctor tell the patient, that he has cancer?

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- The test has correct negative result (= negative) in 99 percent of the cases where the rare form of cancer is not present
- It is also known that 0.001 of the entire population have the rare form of cancer ($h = \text{cancer}$)
- $p(\text{cancer}) = 0.001$, $p(\neg\text{cancer}) = 0.999$
- $p(\text{positive}|\text{cancer}) = 0.99$, $p(\text{positive}|\neg\text{cancer}) = 0.01$,
- $p(\text{negative}|\text{cancer}) = 0.01$, $p(\text{negative}|\neg\text{cancer}) = 0.99$

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- We determine h_{map} according to the linear relation

$$\text{posterior} \propto \text{likelihood} \times \text{prior}$$

$$p(\text{cancer}|\text{positive}) \propto p(\text{positive}|\text{cancer}) \cdot p(\text{cancer}) = 0.99 \cdot 0.001$$

$$p(\neg\text{cancer}|\text{positive}) \propto p(\text{positive}|\neg\text{cancer}) \cdot p(\neg\text{cancer}) = 0.01 \cdot 0.999$$

It follows

$$h_{\text{map}} = \neg\text{cancer}$$

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$$p(\text{cancer}|\text{positive}) \propto p(\text{positive}|\text{cancer}) \cdot p(\text{cancer}) = 0.99 \cdot 0.001$$

$$p(\neg\text{cancer}|\text{positive}) \propto p(\text{positive}|\neg\text{cancer}) \cdot p(\neg\text{cancer}) = 0.01 \cdot 0.999$$

It follows

$$h_{\text{map}} = \neg\text{cancer}$$

- So, despite the positive result, **we are still more confident** that the patient is healthy than otherwise.
- The right thing to do would be to another test to try to accumulate more evidence in favor of the hypothesis that patient has the disease.

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$$p(\text{positive}, \text{cancer}) = p(\text{positive}|\text{cancer}) \cdot p(\text{cancer}) = 0.99 \cdot 0.001$$

$$p(\text{positive}, \neg\text{cancer}) = p(\text{positive}|\neg\text{cancer}) \cdot p(\neg\text{cancer}) = 0.01 \cdot 0.999$$

$$p(\text{positive}|\text{cancer}) = \frac{p(\text{positive}, \text{cancer})}{p(\text{positive}, \text{cancer}) + p(\text{positive}, \neg\text{cancer})}$$

$$\text{law of total probability: } p(\text{positive}) = p(\text{positive}, \text{cancer}) + p(\text{positive}, \neg\text{cancer})$$

$$p(\text{positive}|\text{cancer}) = \frac{p(\text{positive}|\text{cancer}) \cdot p(\text{cancer})}{p(\text{positive})}$$

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Estimating $p(h)$

- Let us draw some principles to estimate

$$p(h_k|D) = \frac{p(D|h_k) \cdot p(h_k)}{p(D)}$$

- Let us first start with $p(h)$
 - given no prior knowledge that *one hypothesis is more likely* than another
 - $p(h)$ can be uniformly distributed

$$\forall_{h \in H} p(h) = \frac{1}{|H|}$$

- otherwise, estimate the prior base on the observed frequency

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Estimating $p(D|h)$

- If data is **discrete**:
 - probability of each possible occurrence using class-conditional probability **mass function** (probability that a discrete random variable gives exact values)
 - we use the frequentist approach
 - e.g. I observe 2 out of 10 individuals with blue eyes and brown in shift A and 1 out of 8 in shift B, then $p(\mathbf{x} = [\text{blue eyes}, \text{brown}]|A) = 0.2$ and $p(\mathbf{x} = [\text{blue eyes}, \text{brown}]|B) = 0.125$
- If data is **real-valued**:
 - probability based on class-conditional probability density function
 - we can use distributions

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Bayesian optimal classifier

- What is the most probable classification of the new instance given the training data?

$$h_{MAP} = \arg \max_h p(h|\mathbf{x}_{new}) = \arg \max_h \frac{p(\mathbf{x}_{new}|h)p(h)}{p(\mathbf{x}_{new})} = \arg \max_h p(\mathbf{x}_{new}|h)p(h)$$

... where the hypotheses correspond to our classes

- we ignore the denominator as it does not alter decision
- The Bayesian classifier has as many parameter as:
 - the number of priors minus 1
 - we can deduce one prior from the remaining ones
 - e.g. given h_1, h_2 and h_3 , $p(h_3) = 1 - p(h_2) - p(h_1)$
 - the number of parameters associated with the class-conditional distributions, $p(\mathbf{x}|h)$

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Bayesian optimal classifier: example

- Priors

$$p(c=0) = \frac{\text{card}(c=0)}{\text{card}(\Omega)} = \frac{3}{7}, \quad p(c=1) = 1 - p(c=0) = \frac{4}{7}$$

- Joint Probability

$$p(v_1=0, v_2=A, v_3=0, c=0) = \frac{\text{card}(v_1=0, v_2=A, v_3=0, c=0)}{\text{card}(\Omega)} = \frac{1}{7}$$

- Likelihood

$$p(v_1=0, v_2=A, v_3=0|c=0) = \frac{1}{3} = \frac{\text{card}(v_1=0, v_2=A, v_3=0, c=0)}{\text{card}(c=0)} = \frac{1}{3} \quad p(v_1=0, v_2=A, v_3=0, c=0) = \frac{1}{7} \quad \frac{1}{3} = \frac{1}{7} \cdot \frac{1}{3}$$

- Data Joint

$$p(v_1=0, v_2=A, v_3=0) = \frac{\text{card}(v_1=0, v_2=A, v_3=0, c=0)}{\text{card}(\Omega)} = \frac{2}{7}$$

- Data: law of total probability

$$p(v_1=0, v_2=A, v_3=0) = p(v_1=0, v_2=A, v_3=0, c=0) + p(v_1=0, v_2=A, v_3=0, c=1) = \frac{1}{7} + \frac{1}{7}$$

- Posterior

$$p(c=0|v_1=0, v_2=A, v_3=0) = \frac{p(v_1=0, v_2=A, v_3=0|c=0)p(c=0)}{p(v_1=0, v_2=A, v_3=0)} = \frac{\frac{1}{3} \cdot \frac{1}{3}}{\frac{2}{7}} = \frac{1}{2}$$

	v_1	v_2	v_3	class
x_1	1	C	1	1
x_2	1	C	1	0
x_3	0	B	1	0
x_4	0	A	0	0
x_5	1	C	1	1
x_6	0	B	1	1
x_7	0	A	0	1

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Bayesian optimal classifier: example

	v_1	v_2	v_3	class
x_1	1	C	1	1
x_2	1	C	1	0
x_3	0	B	1	0
x_4	0	A	0	0
x_5	1	C	1	1
x_6	0	B	1	1
x_7	0	A	0	1

- We can "classify" new observations in the same way, e.g. $\mathbf{x}_{\text{new}} = [1, C, 1]$, what is the class, $c = 0$ or $c = 1$?

- $p(c = 1, v_1 = 1, v_2 = C, v_3 = 1) = p(v_1 = 1, v_2 = C, v_3 = 1 | c = 1) p(c = 1) = \frac{2}{7}$
- $p(c = 0, v_1 = 1, v_2 = C, v_3 = 1) = p(v_1 = 1, v_2 = C, v_3 = 1 | c = 0) p(c = 0) = \frac{1}{7}$
- $p(c = 1, v_1 = 1, v_2 = C, v_3 = 1) > p(c = 0, v_1 = 1, v_2 = C, v_3 = 1)$

\mathbf{x}_{new} is classified with class 1

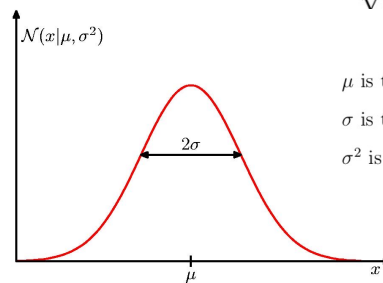
- $p(c = 1 | v_1 = 1, v_2 = C, v_3 = 1) = \frac{p(c = 1, v_1 = 1, v_2 = C, v_3 = 1)}{p(v_1 = 1, v_2 = C, v_3 = 1)} = \frac{\frac{2}{7}}{\frac{2}{7} + \frac{1}{7}} = \frac{2}{3}$
- $p(c = 0 | v_1 = 1, v_2 = C, v_3 = 1) = \frac{p(c = 0, v_1 = 1, v_2 = C, v_3 = 1)}{p(v_1 = 1, v_2 = C, v_3 = 1)} = \frac{\frac{1}{7}}{\frac{2}{7} + \frac{1}{7}} = \frac{1}{3}$
- $p(c = 1 | v_1 = 1, v_2 = C, v_3 = 1) + p(c = 0 | v_1 = 1, v_2 = C, v_3 = 1) = 1$

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Gaussian Distribution

- Gaussian distribution or normal is defined by the probability

$$p(x|\mu, \sigma^2) = \mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2 \cdot \pi \cdot \sigma}} \cdot \exp\left(-\frac{1}{2 \cdot \sigma^2} \cdot (x - \mu)^2\right)$$



μ is the mean

σ is the standard deviation

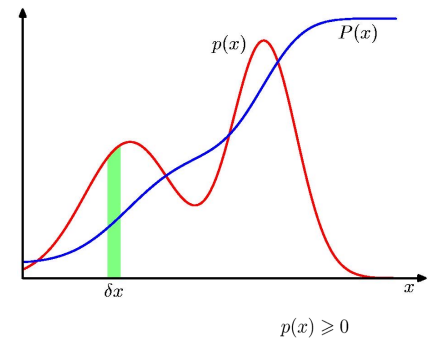
σ^2 is the variance

$$\int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) dx = 1$$

$$\mathcal{N}(x|\mu, \sigma^2) > 0$$

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Probability Density Function (PDF)



$$p(x \in (a, b)) = \int_a^b p(x) dx$$

$$P(z) = \int_{-\infty}^z p(x) dx$$

Cumulative distribution function (CDF)

$$\int_{-\infty}^{\infty} p(x) dx = 1$$

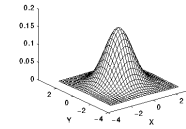
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Relative Probability

- Gaussian distribution is a type of continuous probability distribution for a real-valued random variable.
- The Gaussian distribution or normal distribution is defined as PDF (Probability Density Function) that reflects the **relative** probability.
- The **PDF may give a value greater than one** (small standard deviation).
- It is the area under the curve that represents the probability. However, the PDF reflects the relative probability.
 - Does a continuous probability distribution exist in the real world?

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Normal Distribution in D dim

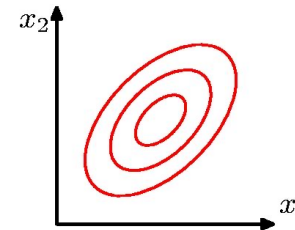


Over D dimensional space

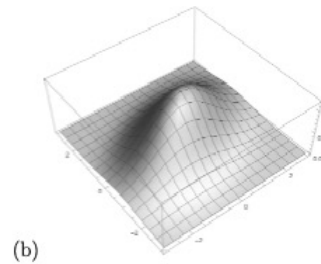
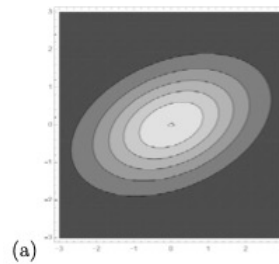
$$p(\mathbf{x}|\boldsymbol{\mu}, \Sigma) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \Sigma) = \frac{1}{(2 \cdot \pi)^{D/2}} \cdot \frac{1}{|\Sigma|^{1/2}} \cdot \exp\left(-\frac{1}{2} \cdot (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} \cdot (\mathbf{x} - \boldsymbol{\mu})\right)$$

where

- $\boldsymbol{\mu}$ is the D dimensional mean vector
- Σ is a $D \times D$ covariance matrix
- $|\Sigma|$ is the determinant of Σ



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- (a) The Gaussian distribution over 2 dimensional space with $\boldsymbol{\mu} = (0, 0)^T$ and the covariance matrix Σ

$$\Sigma = \begin{pmatrix} 2 & 0.5 \\ 0.5 & 1 \end{pmatrix}.$$

- (b) Three dimensional plot of the Gaussian.

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Covariance Matrix

- A position $c_{ij} = \Sigma_{ij}$ of this matrix measures the tendency of two features, x_i and x_j , to vary in the same direction, for N features indexed by k

$$c_{ij} = \frac{\sum_{k=1}^N (x_{k,i} - \bar{x}_i) \cdot (x_{k,j} - \bar{x}_j)}{N - 1}$$

- with \bar{x}_i and \bar{x}_j being the arithmetic mean of the two variables of the sample
- Covariances are symmetric; $c_{ij} = c_{ji}$ and, so, the resulting covariance matrix Σ is symmetric and positive-definite

$$\Sigma = \begin{pmatrix} c_{11} & c_{12} & \cdots & c_{1m} \\ c_{21} & c_{22} & \cdots & c_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ c_{m1} & c_{m2} & \cdots & c_{mm} \end{pmatrix}$$

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Multivariate Gaussian: example

Approximate a multivariate Gaussian distribution using the following points: $\{(-2,2)^T, (-1,3)^T, (0,1)^T, (-2,1)^T\}$

$$\mu = \frac{1}{4} \left(\begin{bmatrix} -2 \\ 2 \end{bmatrix} + \begin{bmatrix} -1 \\ 3 \end{bmatrix} + \begin{bmatrix} 0 \\ 1 \end{bmatrix} + \begin{bmatrix} -2 \\ 1 \end{bmatrix} \right) = \begin{bmatrix} -1.25 \\ 1.75 \end{bmatrix}$$

$$c_{12} = c_{21} = \frac{(-2+1.25)(2-1.75) + (-1+1.25)(3-1.75) + (0+1.25)(1-1.75) + (-2+1.25)(1-1.75)}{3} = -0.83$$

$$c_{11} = \frac{(-2+1.25)^2 + (-1+1.25)^2 + (0+1.25)^2 + (-2+1.25)^2}{3} = 0.92$$

$$c_{22} = \frac{(2-1.75)^2 + (3-1.75)^2 + (1-1.75)^2 + (1-1.75)^2}{3} = 0.92$$

$$\Sigma = \begin{pmatrix} c_{11} & c_{12} \\ c_{12} & c_{22} \end{pmatrix} = \begin{pmatrix} 0.92 & -0.83 \\ -0.83 & 0.92 \end{pmatrix}$$

$$\Sigma^{-1} = \begin{pmatrix} 1.1 & 0.1 \\ 0.1 & 1.1 \end{pmatrix}, \quad \text{Det}(\Sigma) = |\Sigma| = 0.833 \text{ Type equation here.}$$

If $A = \begin{bmatrix} a & b \\ c & d \end{bmatrix}$ then

$$A^{-1} = \frac{1}{ad-bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$$

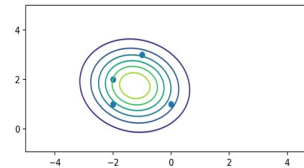
Inverse of A Determinant of A Adjoint of A

$$N(\mathbf{x}|\mu, \Sigma) = \frac{1}{(2\pi)^{2/2} \sqrt{0.833}} \exp \left(-\frac{1}{2} \left(\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} - \begin{bmatrix} -1.25 \\ 1.75 \end{bmatrix} \right)^T \begin{bmatrix} 1.1 & 0.1 \\ 0.1 & 1.1 \end{bmatrix} \left(\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} - \begin{bmatrix} -1.25 \\ 1.75 \end{bmatrix} \right) \right)$$

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Multivariate Gaussian: example

- What is the shape of the previous 2-dimensional Gaussian?
 - fixing μ and Σ inspection...



- What is the probability of observing (0,0)?

$$\bullet N\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix} \middle| \mu, \Sigma\right) = \frac{1}{2\pi\sqrt{0.083}} \exp\left(-\frac{1}{2}\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix} - \begin{bmatrix} -1.25 \\ 1.75 \end{bmatrix}\right)^T \begin{bmatrix} 1.1 & 0.1 \\ 0.1 & 1.1 \end{bmatrix} \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix} - \begin{bmatrix} -1.25 \\ 1.75 \end{bmatrix}\right)\right) = 0.0145$$

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Bayesian optimal classifier: example

- Consider a population of 100 individuals
 - 30 individuals have phenotype A, 30 have B, and remaining ones have C
 - the expression of three genes (variables) are characterized by the following 3-dimensional Gaussians

$$N_A\left(\mu_A = \begin{bmatrix} 0.375 \\ 0.875 \\ 0.25 \end{bmatrix}, \Sigma_A = \begin{bmatrix} 3.41 & 1.34 & 2.6 \\ 1.34 & 2.125 & 1.18 \\ 2.6 & 1.18 & 2.8 \end{bmatrix}\right), N_B\left(\mu_B = \begin{bmatrix} 0.5 \\ 0.125 \\ 0.875 \end{bmatrix}, \Sigma_B = \begin{bmatrix} 0.286 & 0.07 & -0.07 \\ 0.07 & 0.125 & 0.018 \\ -0.07 & 0.018 & 0.125 \end{bmatrix}\right), N_C\left(\mu_C = \begin{bmatrix} 0 \\ -0.125 \\ 0.125 \end{bmatrix}, \Sigma_C = \begin{bmatrix} 1.7 & 1.14 & 1 \\ 1.14 & 1.55 & 0.73 \\ 1 & 0.73 & 0.98 \end{bmatrix}\right)$$

$$p(A) = \frac{30}{100}, p(B) = \frac{30}{100}, p(C) = \frac{40}{100} \text{ prior, called mixture parameters}$$

classify observations $\mathbf{x}_1 = [0, 1.1, -0.8]$

$$p(\mathbf{x}_1|N_A)=0.019, \quad p(\mathbf{x}_1|N_B)=5.4E-14, \quad p(\mathbf{x}_1|N_C)=0.0088$$

$$p(\mathbf{x}_1, N_A) = p(A)p(\mathbf{x}_1|N_A), \quad p(\mathbf{x}_1, N_B) = p(B)p(\mathbf{x}_1|N_B), \quad p(\mathbf{x}_1, N_C) = p(C)p(\mathbf{x}_1|N_C)$$

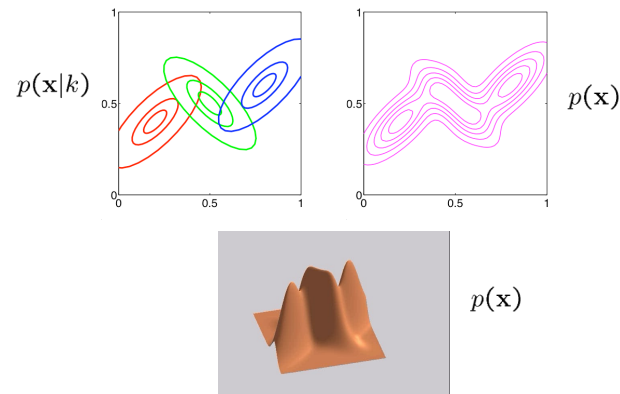
$$p(\mathbf{x}_1) = p(\mathbf{x}_1, N_A) + p(\mathbf{x}_1, N_B) + p(\mathbf{x}_1, N_C)$$

$$\bullet p(A|\mathbf{x}_1) = \frac{p(A)p(\mathbf{x}_1|N_A)}{p(\mathbf{x}_1)} = 0.0057, \quad p(B|\mathbf{x}_1) = \frac{p(B)p(\mathbf{x}_1|N_B)}{p(\mathbf{x}_1)} = 0, \quad p(C|\mathbf{x}_1) = \frac{p(C)p(\mathbf{x}_1|N_C)}{p(\mathbf{x}_1)} = 0.0035$$

\mathbf{x}_1 is classified with phenotype A

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Example: Mixture of 3 Gaussians k



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Bayes optimal classifier

• Advantages

- when data distributions are well-approximated, provides highly **accurate** results
- priors can be easily neglected to not bias posteriors

• Disadvantages

- requires a good amount of data to estimate joint distributions
 - impracticable in the presence of **high-dimensional data**
- can be computationally **expensive**
 - discrete data: need to compute the posterior probability for every hypothesis
 - numeric data: need to approximate distributions
 - e.g. fitting multivariate Gaussians can be expensive due covariance matrix inversion

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Joint distribution

- A joint distribution for toothache, cavity, catch, *dentist's probe catches in my tooth* ☹

- we need to know the conditional probabilities of the conjunction of toothache and cavity
- what can a dentist conclude if the probe catches in the aching tooth?

$$P(\text{cavity} \mid \text{toothache} \wedge \text{catch}) = \frac{P(\text{toothache} \wedge \text{catch} \mid \text{cavity})P(\text{cavity})}{P(\text{toothache} \wedge \text{cavity})}$$

- **Problem?**

- For n possible variables there are 2^n possible combinations

	toothache		no toothache	
	catch	no catch	catch	no catch
cavity	0.108	0.012	0.072	0.008
no cavity	0.016	0.064	0.144	0.576

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Conditional independence

- Once we know that the patient has cavity we do not expect the probability of the probe catching to depend on the presence of toothache

- **independence** $P(\text{catch} \mid \text{cavity} \wedge \text{toothache}) = P(\text{catch} \mid \text{cavity})$
 $P(\text{toothache} \mid \text{cavity} \wedge \text{catch}) = P(\text{toothache} \mid \text{cavity})$

- The decomposition of large probabilistic domains into weakly connected subsets via conditional independence is one of the most important developments in the recent history of AI

$$P(a \wedge b) = P(a)P(b)$$

$$P(\text{toothache}, \text{catch}, \text{cavity}, \text{Weather} = \text{cloudy}) = \\ = P(\text{Weather} = \text{cloudy})P(\text{toothache}, \text{catch}, \text{cavity})$$

$$P(a \mid b) = P(a)$$

$$P(b \mid a) = P(b)$$

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Naive Bayes Classifier

- Along with *decision trees*, neural networks, *nearest neighbor*, one of the **most practical learning methods**
- When to use:
 - Moderate or large training set available
 - Attributes that describe instances are conditionally independent given classification
- Successful applications:
 - Diagnosis
 - Classifying text documents

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Naive Bayes Classifier

- Assume target function $f: X \rightarrow V$, where each instance x described by attributes $a_1, a_2 \dots a_n$
- Most probable value of $f(x)$ is:

$$\begin{aligned}
 v_{MAP} &= \arg \max_{v_j \in V} P(v_j | a_1, a_2 \dots a_n) \\
 v_{MAP} &= \arg \max_{v_j \in V} \frac{P(a_1, a_2 \dots a_n | v_j) P(v_j)}{P(a_1, a_2 \dots a_n)} \\
 &= \arg \max_{v_j \in V} P(a_1, a_2 \dots a_n | v_j) P(v_j)
 \end{aligned}$$

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V_{NB}

- Naive Bayes assumption:

$$P(a_1, a_2 \dots a_n | v_j) = \prod_i P(a_i | v_j)$$

- which gives

$$\text{Naive Bayes classifier: } v_{NB} = \arg \max_{v_j \in V} P(v_j) \prod_i P(a_i | v_j)$$

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Naive Bayes Algorithm

- For each target value v_j
- $\hat{P}(v_j) \leftarrow$ estimate $P(v_j)$
- For each attribute value a_i of each attribute a
- $\hat{P}(a_i | v_j) \leftarrow$ estimate $P(a_i | v_j)$

$$v_{NB} = \arg \max_{v_j \in V} \hat{P}(v_j) \prod_{a_i \in x} \hat{P}(a_i | v_j)$$

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Training dataset

Class:
C1:buys_computer='yes'
C2:buys_computer='no'

Data sample:

X =
(age<=30,
Income=medium,
Student=yes
Credit_rating=fair)

age	income	student	credit_rating	buys_computer
<=30	high	no	fair	no
<=30	high	no	excellent	no
30...40	high	no	fair	yes
>40	medium	no	fair	yes
>40	low	yes	fair	yes
>40	low	yes	excellent	no
31...40	low	yes	excellent	yes
<=30	medium	no	fair	no
<=30	low	yes	fair	yes
>40	medium	yes	fair	yes
<=30	medium	yes	excellent	yes
31...40	medium	no	excellent	yes
31...40	high	yes	fair	yes
>40	medium	no	excellent	no

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Naïve Bayesian Classifier: Example

- Compute $P(X|C_i)$ for each class

$$P(\text{age}="<30" \mid \text{buys_computer}="yes") = 2/9 = 0.222$$

$$P(\text{age}="<30" \mid \text{buys_computer}="no") = 3/5 = 0.6$$

$$P(\text{income}="medium" \mid \text{buys_computer}="yes") = 4/9 = 0.444$$

$$P(\text{income}="medium" \mid \text{buys_computer}="no") = 2/5 = 0.4$$

$$P(\text{student}="yes" \mid \text{buys_computer}="yes") = 6/9 = 0.667$$

$$P(\text{student}="yes" \mid \text{buys_computer}="no") = 1/5 = 0.2$$

$$P(\text{credit_rating}="fair" \mid \text{buys_computer}="yes") = 6/9 = 0.667$$

$$P(\text{credit_rating}="fair" \mid \text{buys_computer}="no") = 2/5 = 0.4$$

$$P(\text{buys_computer}="yes") = 9/14$$

$$P(\text{buys_computer}="no") = 5/14$$

- X = (age<=30, income =medium, student=yes, credit_rating=fair)

$$P(X|C_1) : P(X \mid \text{buys_computer}="yes") = 0.222 \times 0.444 \times 0.667 \times 0.667 = 0.044$$

$$P(X|C_2) : P(X \mid \text{buys_computer}="no") = 0.6 \times 0.4 \times 0.2 \times 0.4 = 0.019$$

$$P(X|C_1) \cdot P(C_1) : P(X \mid \text{buys_computer}="yes") \cdot P(\text{buys_computer}="yes") = 0.028$$

$$P(X|C_2) \cdot P(C_2) : P(X \mid \text{buys_computer}="no") \cdot P(\text{buys_computer}="no") = 0.007$$

$$X \text{ belongs to class "buys_computer=yes"} \quad P(C_1|X) = 0.028 / (0.028 + 0.007)$$

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Estimating probabilities in small samples

- We have estimated probabilities by the times the event is observed, n_c , over total opportunities, n
 - poor estimates when n_c is very small
 - **problem**: what if none of the training instances with target value v_j have attribute value a_i ? $\rightarrow n_c$ is 0!
- when n_c is very small: $\hat{P}(a_i|v_j) = \frac{n_c + mp}{n + m}$ $v_{NB} = v_j \in V$ $P(v_j) \prod_i \hat{P}(a_i|v_j)$
 - n is number of training examples for which $v = v_j$
 - n_c number of examples for which $v = v_j$ and $a = a_i$
 - p is the prior estimate
 - m is the weight given to prior (i.e. number of “virtual” examples)

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Naïve Bayes: comments

- **Advantages**
 - easy to implement, good results obtained in most of the cases
 - The decomposition of large probabilistic domains into weakly connected subsets via conditional independence is one of the most important developments in the recent history of AI
- Conditional independence assumption is often violated
- ...but it works surprisingly well anyway

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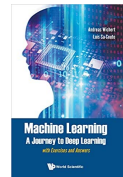
Naïve Bayes: comments

- **Disadvantages**

- Assumption: class conditional independence , therefore loss of accuracy
- Practically, dependencies exist among variables
- E.g., hospitals: patients: Profile: age, family history etc
Symptoms: fever, cough etc., Disease: lung cancer, diabetes etc
- Dependencies among these cannot be modeled by Naïve Bayesian Classifier
- How to deal with these dependencies?
 - Bayesian - Belief Networks

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Literature



- Machine Learning - A Journey to Deep Learning, A. Wichert, Luis Sa-Couto, World Scientific, 2021
 - Chapter 2

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