

# Economics 144

## Economic Forecasting

### Lecture 6

#### Characterizing Cycles

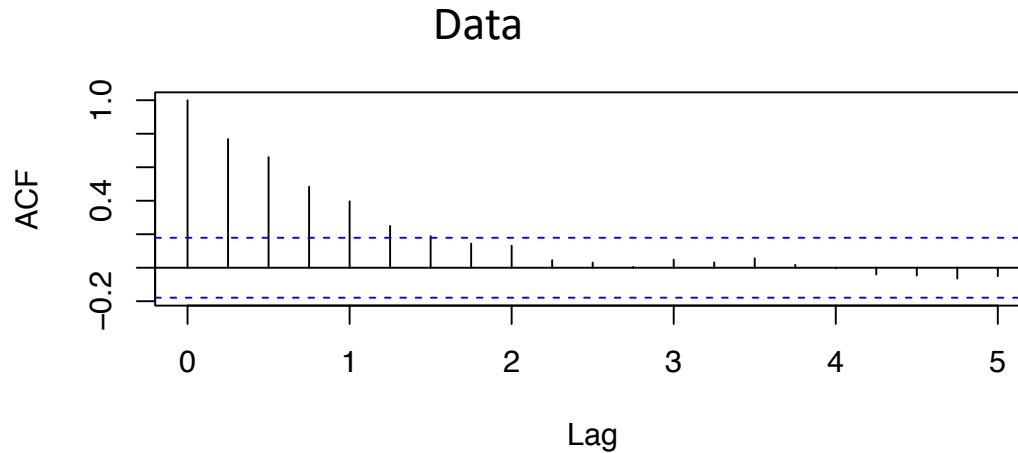
#### Moving Average Models

Dr. Randall R. Rojas

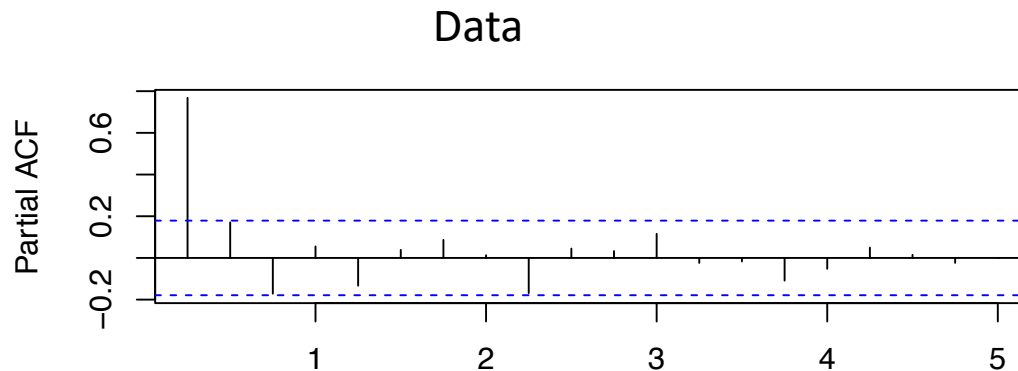
# Today's Class

- Covariance Stationary Time Series
- White Noise
- The Lag Operator
- Wold's Theorem
- Characteristics of the  $MA(q)$  Process
  - Example:  $MA(1)$  Process

# Covariance Stationary Time Series



In R: `acf(data)`



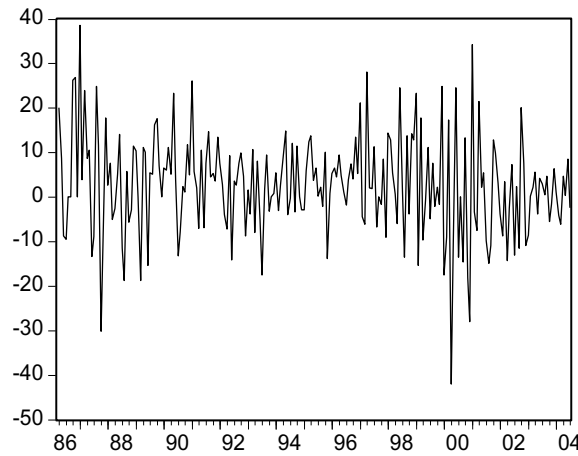
In R: `pacf(data)`

Recall that covariance stationary processes have  $\rho(k)$  and  $p(k)$  that  $\rightarrow 0$  as  $k \rightarrow \text{infinity}$ .

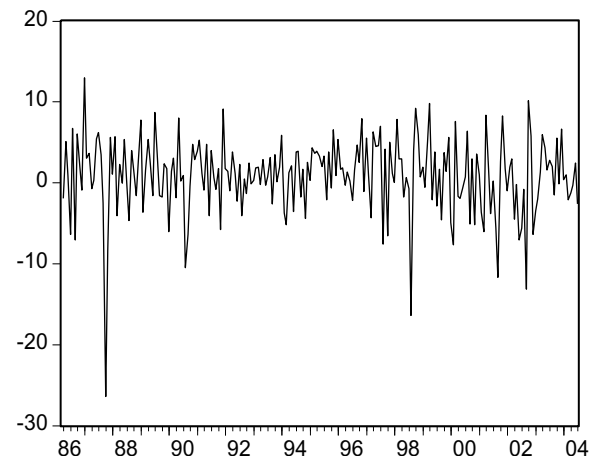
# White Noise 1 of 3

- **White Noise:** Time series process with zero mean, constant variance, and no serial correlation.
- Since  $\rho(k) = 0$  and  $p(k) = 0$  for  $k \geq 1$ , there is no link between past and present observations.
  - Cannot predict the future.
- **Examples:** S&P500 returns, interest rates, etc.

# Example: Autocorrelation Functions of Monthly Returns for Microsoft and the Dow Jones Index



— Microsoft monthly returns



— Dow Jones monthly returns

Sample: 1986:03 2004:07 Included observations: 220				
Autocorrelation	Partial Correlation	AC	PAC	
		1 -0.081	-0.081	
		2 -0.094	-0.101	
		3 0.132	0.117	
		4 -0.017	-0.006	
		5 0.008	0.030	
		6 -0.013	-0.029	
		7 0.106	0.113	
		8 0.015	0.023	
		9 -0.006	0.024	
		10 0.131	0.112	
		11 0.013	0.035	
		12 -0.016	0.005	
		13 -0.020	-0.045	
		14 0.030	0.013	
		15 -0.075	-0.091	
		16 0.064	0.068	
		17 0.085	0.051	
		18 -0.094	-0.064	
		19 -0.049	-0.079	
		20 0.046	0.005	

Sample: 1986:03 2004:07 Included observations: 220				
Autocorrelation	Partial Correlation	AC	PAC	
		1 -0.021	-0.021	
		2 -0.044	-0.044	
		3 -0.056	-0.058	
		4 -0.126	-0.131	
		5 0.048	0.037	
		6 -0.031	-0.045	
		7 0.092	0.081	
		8 -0.044	-0.057	
		9 -0.043	-0.030	
		10 0.043	0.035	
		11 -0.012	0.006	
		12 0.015	-0.006	
		13 -0.003	0.003	
		14 -0.034	-0.034	
		15 -0.059	-0.060	
		16 0.039	0.042	
		17 0.031	0.012	
		18 0.079	0.076	
		19 -0.026	-0.028	
		20 -0.016	0.005	

# Wold's Theorem (Part I)

- **Q:** What's left after filtering the trend and seasonal components?
- **A:** Covariance stationary (short memory) residuals! How should we model them? → **Wold's Theorem!**
- **Wold's Representation Theorem:** Let  $\{y_t\}$  be any zero-mean covariance-stationary process. Then

$$y_t = B(L)\varepsilon_t = \sum_{i=0}^{\infty} b_i \varepsilon_{t-i} \text{ and } \varepsilon_t \sim WN(0, \sigma^2)$$

↑  
(Innovations)

where  $\sum_{i=0}^{\infty} b_i^2 < \infty$  and  $b_0=1$ .

# Moving Average Models

- Def: **MA(q)** = Moving Average process of order  $q \geq 0$ :  $Y_t = \mu + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \dots + \theta_q \varepsilon_{t-q}$ , where  $\varepsilon_t \sim (0, \sigma^2)$
- Examples:
  - MA(**1**):  $Y_t = \mu + \theta_1 \varepsilon_{t-1} + \varepsilon_t$
  - MA(**5**):  $Y_t = \mu + \theta_1 \varepsilon_{t-1} + \theta_5 \varepsilon_{t-5} + \varepsilon_t$
  - MA(**10**):  $Y_t = \mu + \theta_{10} \varepsilon_{t-10} + \varepsilon_t$

# Moving Average Models

For every MA( $q$ ) process, we need to address the following 3 questions:

1. What does a time series of an MA process look like?
2. What do the corresponding ACFs and PACFs look like?
3. What is the optimal forecast?



# Moving Average Models

## Example: MA(1) Process

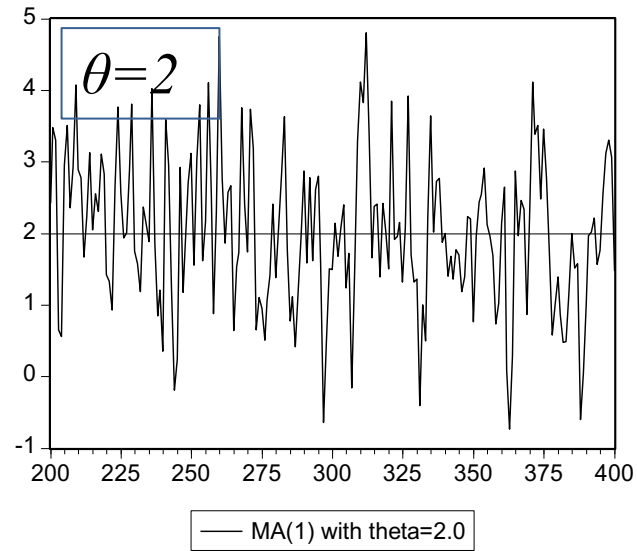
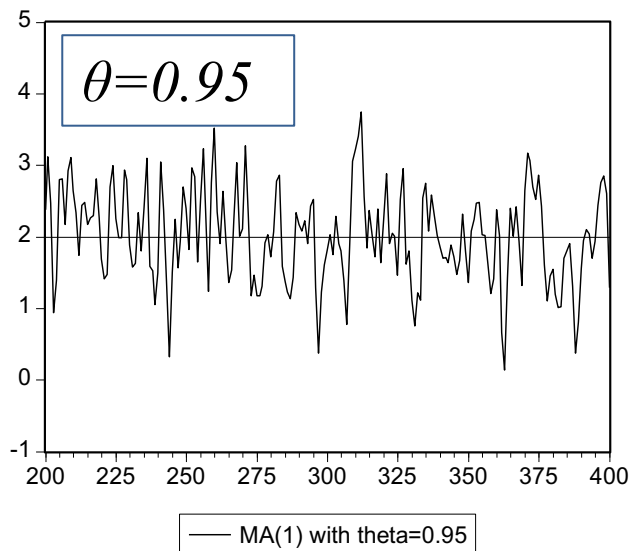
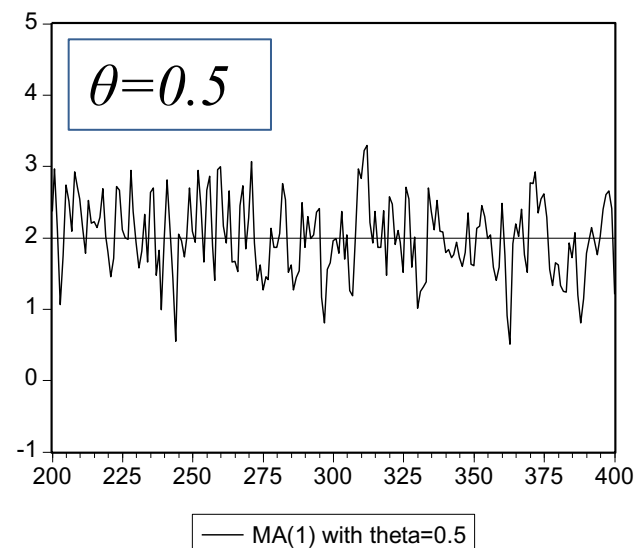
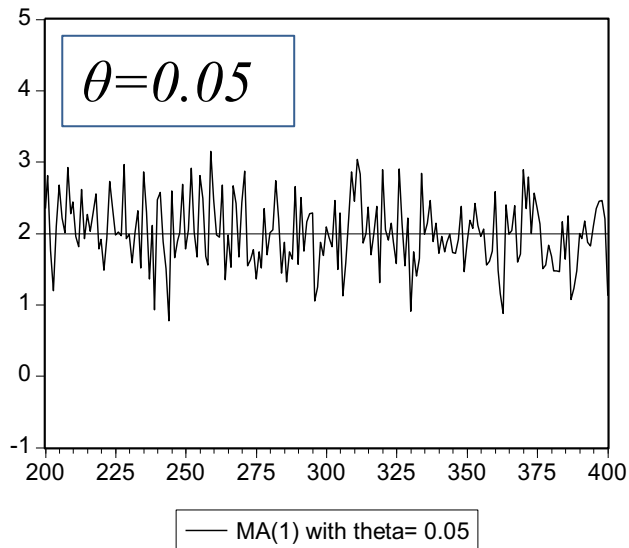
- (1) What does a time series of an MA process look like?

- Consider the following MA(1) process:

$$Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t$$

- We can plot this process for  $\mu = 2$ , and different values of  $\theta$ , e.g., for  $\theta = 0.05$ ,  $\theta = 0.5$ ,  $\theta = 0.95$ , and  $\theta = 2$ .
- We can show that  $E(Y_t) = \mu$ , and  $\sigma^2(Y_t) = (1 + \theta^2)\sigma^2_\varepsilon$

$$\text{MA}(1): Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t$$



# Moving Average Models

## Example: MA(1) Process

- (2) What do the corresponding ACFs and PACFs look like?
- ACF:
  - (i) We would expect to see only 1 spike different from zero, i.e.,  $\rho_1 \neq 0$ , and all others equal to zero ( $\rho_k = 0$ ,  $k > 1$ ).
  - (ii) The magnitude of the spike should be proportional to  $\theta$  for  $|\theta| < 1$ .
  - Note: You can show (please fill in the steps) that  $\rho_1 = \theta / (1 + \theta^2)$
  - (iii) Given the expression above, the sign of the ACF is the same as the sign of  $\theta$ .

# Moving Average Models



## Example: MA(1) Process

- (2) What do the corresponding ACFs and PACFs look like?
- PACF:
  - Note: You can show (please fill in the steps) that the autocovariance of order 1 is given by:  $\gamma_1 = \theta\sigma_\varepsilon^2$ , and all other orders are equal to 0.
  - (i) The PACF decreases to zero in an alternating fashion, according to  $p_k > 0$  ( $k=\text{odd}$ ), and  $p_k < 0$  ( $k=\text{even}$ ).
  - (ii) We can also show that  $p_1 = \rho_1$ .

# Autocorrelation Functions of Simulated MA(1) Processes



Sample: 1 2000  
Included observations: 1999

$\theta = 0.05$

Autocorrelation	Partial Correlation	AC	PAC
		1 0.072 0.072	
		2 -0.027 -0.032	
		3 -0.061 -0.057	
		4 -0.024 -0.017	
		5 0.016 0.016	
		6 -0.004 -0.011	
		7 0.022 0.022	
		8 0.003 0.001	
		9 0.000 0.000	
		10 -0.011 -0.009	


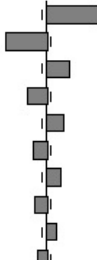
Sample: 1 2000  
Included observations: 1999

$\theta = 0.5$

Autocorrelation	Partial Correlation	AC	PAC
		1 0.406 0.406	
		2 -0.038 -0.243	
		3 -0.076 0.053	
		4 -0.038 -0.039	
		5 0.006 0.029	
		6 0.009 -0.015	
		7 0.020 0.030	
		8 0.011 -0.012	
		9 -0.003 0.002	
		10 -0.010 -0.009	



Sample: 1 2000  
Included observations: 1999

$\theta = 0.95$

Autocorrelation	Partial Correlation	AC	PAC
		1 0.498 0.498	
		2 -0.042 -0.386	
		3 -0.081 0.220	
		4 -0.042 -0.183	
		5 0.003 0.157	
		6 0.013 -0.127	
		7 0.019 0.131	
		8 0.013 -0.112	
		9 -0.004 0.092	
		10 -0.010 -0.088	

Sample: 1 2000  
Included observations: 1999

$\theta = 2$

Autocorrelation	Partial Correlation	AC	PAC
		1 0.406 0.406	
		2 -0.039 -0.243	
		3 -0.076 0.053	
		4 -0.038 -0.039	
		5 0.006 0.030	
		6 0.009 -0.016	
		7 0.020 0.030	
		8 0.010 -0.013	
		9 -0.004 0.002	
		10 -0.010 -0.008	

Can you distinguish them?

# Moving Average Models

## Example: MA(1) Process

- From our previous example, based on the ACFs alone, we could not distinguish the MA(1) process with  $\theta=0.5$  from the MA(1) process with  $\theta=2$ . This property is known as invertibility.

- **Def: Invertibility**

An MA(1) process is invertible if  $|\theta| < 1$ . Otherwise, if  $|\theta| \geq 1$ , the process is noninvertible.

# Moving Average Models

## Example: MA(1) Process

- **Meaning of Invertibility:** You can transform the MA(1) process to an 'autoregressive' function of its own past (lagged values), such that the recent past has more weight than the distant past.
- **Example:**  $Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t = \mu + (1 - (-\theta)L) \varepsilon_t$

- Solve for  $\varepsilon_t \rightarrow \boxed{\varepsilon_t = \frac{1}{1 - (-\theta)L} (Y_t - \mu)}$

# Moving Average Models

## Example: MA(1) Process

- Since  $|\theta| < 1$ , we can perform a Taylor series expansion of the denominator:

$$\begin{aligned}\frac{1}{1 - (-\theta)L} &= \lim_{i \rightarrow \infty} (1 + (-\theta)L + (-\theta)^2 L^2 + \dots + (-\theta)^i L^i + \dots) \\ &= (1 - \theta L + \theta^2 L^2 + \dots)\end{aligned}$$

- Therefore, we can re-express  $\varepsilon_t = \frac{1}{1 - (-\theta)L} (Y_t - \mu)$  as:  $\varepsilon_t = (1 - \theta L + \theta^2 L^2 + \dots) (Y_t - \mu)$

$$\Rightarrow Y_t - \mu = \theta(Y_{t-1} - \mu) - \theta^2(Y_{t-2} - \mu) + \dots + \varepsilon_t$$

= Autoregressive Process



# Moving Average Models

## Example: MA(1) Process

- If an MA process is invertible, you can always find an Autoregressive representation.
- To predict the future, we need the information contained in the past.
- As noted earlier, the recent past has more weight than the distant past since:  $\lim_{t \rightarrow \infty} \theta^t \rightarrow 0$

# Moving Average Models

## Example: MA(1) Process

- **Q:** What can we conclude for the case when  $|\theta| \geq 1$ ?
- **A:** We cannot perform the Taylor series expansion on  $\theta$  but we can on  $1/\theta$ .
- Since  $\theta > 1$ , consider expanding  $1/(1-\theta L)$  as follows:

$$\frac{1}{1 - \theta L} = -\frac{\frac{1}{\theta L}}{1 - \frac{1}{\theta L}} = -\frac{1}{\theta L} \left( 1 + \frac{1}{\theta L} + \left( \frac{1}{\theta L} \right)^2 + \dots \right)$$

Infinite lag polynomial 18

# Moving Average Models

## Example: MA(1) Process

- Def: **Forward Operator (F)** =  $1/L$ , where  $FY_t = Y_{t+1} \rightarrow F(L(Y_t)) = F(Y_{t-1}) = Y_t$ . The forward operator is inverse of the lag operator. It delivers the process at a future date.

$$\frac{1}{1 - \theta L} = -\frac{1}{\theta} F \left( 1 + \frac{1}{\theta} F + \frac{1}{\theta^2} F^2 + \dots \right)$$

$$\varepsilon_t = \frac{1}{1 - \theta L} Y_t = -\frac{1}{\theta} F \left( 1 + \frac{1}{\theta} F + \frac{1}{\theta^2} F^2 + \dots \right) Y_t$$

$$\boxed{-\frac{1}{\theta} Y_{t+1} = +\frac{1}{\theta^2} Y_{t+2} + \frac{1}{\theta^3} Y_{t+3} \dots + \varepsilon_t}$$

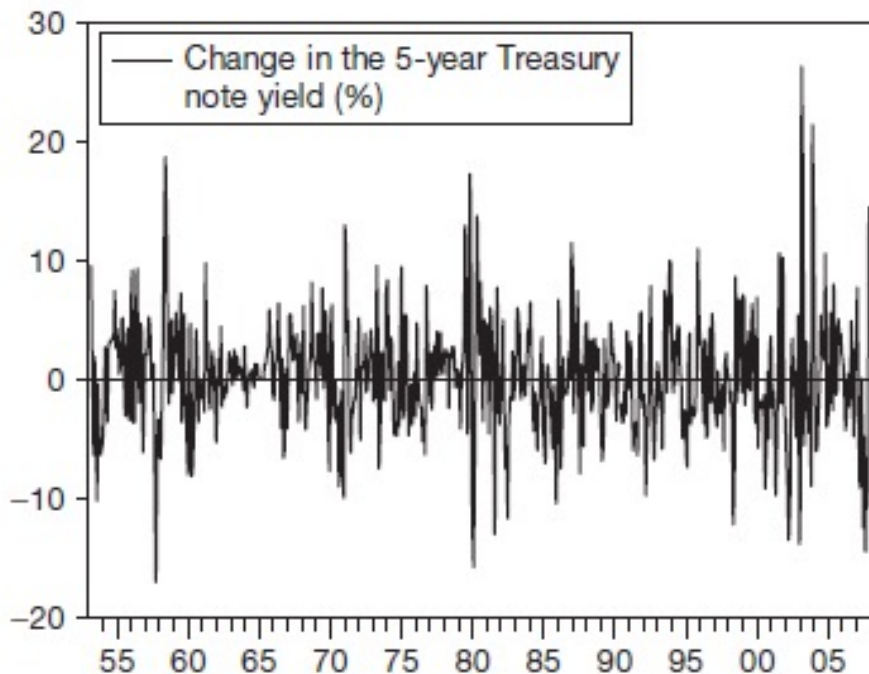
The present is a function of the future.

Autoregressive representation of a noninvertible MA(1) process.

# Moving Average Models

## Example: MA(1) Process

- What is the practical use of all this?



Sample: 1953M04 2008M04  
Included observations: 660

Autocorrelation	Partial Correlation	AC	PAC
		1 0.339	0.339
		2 -0.073	-0.213
		3 0.007	0.129
		4 0.014	-0.063
		5 -0.043	-0.017
		6 -0.073	-0.060
		7 -0.069	-0.035

What process would you suggest?

$$\hat{\rho}_1 = 0.339 \rightarrow \hat{\theta}_1 \approx 0.34$$

→ MA(1)

# Moving Average Models

## Forecasting in an MA(1) Process

- Consider first the 1-step-ahead forecast,  $h = 1$ :  
MA(1):  $Y_t = \mu + \theta\varepsilon_{t-1} + \varepsilon_t \rightarrow Y_{t+1} = \mu + \theta\varepsilon_t + \varepsilon_{t+1}$
- **Optimal Point Forecast:**  $f_{t,1} = E(Y_{t+1}|I_t) = \mu + \theta\varepsilon_t$ 
  - Note: We can write  $\varepsilon_t$  in terms of lags of  $Y_t$
- **One-period-ahead Forecast Error:**  $e_{t,1} = Y_{t+1} - f_{t,1} = \varepsilon_{t+1}$
- **Uncertainty of the Forecast:**  $\sigma^2_{t+1|t} = \text{var}(Y_{t+1}|I_t) = \sigma^2_\varepsilon$
- **Density Forecast:**  $f(Y_{t+1}|I_t) \sim N(\mu + \theta\varepsilon_t, \sigma^2_\varepsilon)$ 
  - Note: We can compute the confidence intervals from the density forecast.

Please go over the steps for  $h=1$  and  $h=2$  (Section 6.2<sup>a</sup>)

# Moving Average Models

## Forecasting in an MA(1) Process

- Consider first the k-step-ahead forecast,  $h = k$ :

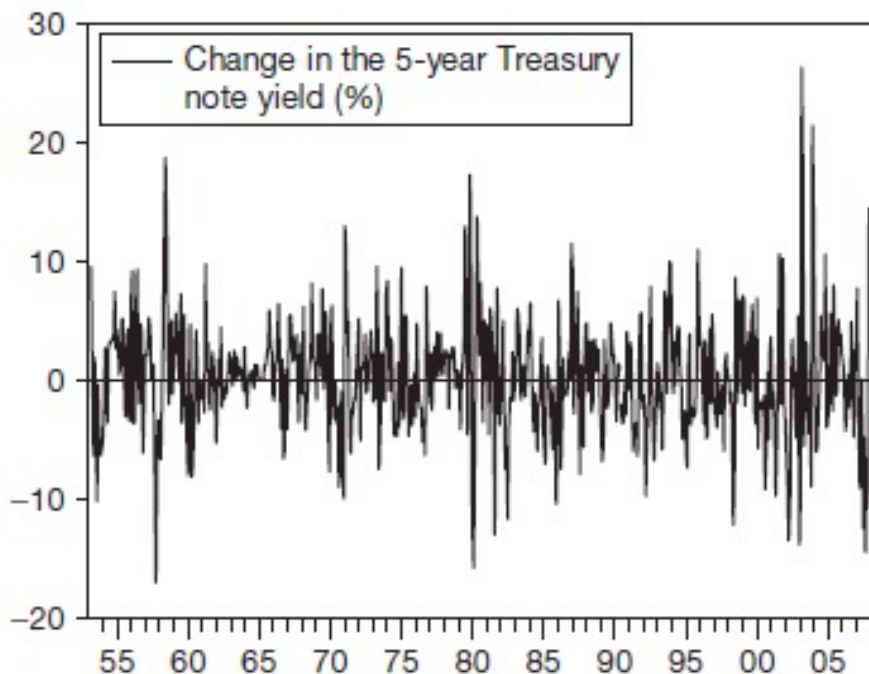
$$\text{MA}(\mathbf{1}): Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t \rightarrow Y_{t+1} = \mu + \theta \varepsilon_t + \varepsilon_{t+1}$$

- Optimal Point Forecast:  $f_{t,k} = E(Y_{t+k}|I_t) = \mu$
- k-period-ahead Forecast Error:  $e_{t,k} = Y_{t+k} - f_{t,k} = \varepsilon_{t+k} + \theta \varepsilon_{t+k-1}$
- Uncertainty of the Forecast:  $\sigma^2_{t+k|t} = \sigma^2_{\varepsilon}(1 + \theta^2) = \sigma^2_Y$
- Density Forecast:  $f(Y_{t+k}|I_t) \sim N(\mu, \sigma^2_Y)$

# Moving Average Models

## Example: MA(1) Process

- Forecasting the 5-year Constant Maturity Yield



Sample: 1953M04 2008M04  
Included observations: 660

Autocorrelation	Partial Correlation	AC	PAC
1	0.339	0.339	0.339
2	-0.073	-0.213	-0.073
3	0.007	0.129	0.007
4	0.014	-0.063	0.014
5	-0.043	-0.017	-0.043
6	-0.073	-0.060	-0.073
7	-0.069	-0.035	-0.069

*Model:*  $Y_t = \mu + \theta \varepsilon_{t-1} + \varepsilon_t$

*Estimated Model:*  $Y_t = 0.160 + 0.485\varepsilon_t + \varepsilon_t$

where  $\hat{\sigma}_Y^2 = 23.683$  and  $\hat{\sigma}_\varepsilon = 4.449$

# Estimation Output: 5-Year Treasury Yield (Monthly Percentage Changes)

Dependent Variable: DY Method: Least Squares Sample (adjusted): 1953M05 2007M11 Included observations: 655 after adjustments Convergence achieved after 7 iterations Backcast: 1953M04				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.160159	0.258095	0.620544	0.5351
MA(1)	0.485011	0.034468	14.07130	0.0000
R-squared	0.165370	Mean dependent var	0.168613	
Adjusted R-squared	0.164092	S.D. dependent var	4.866609	
S.E. of regression	4.449443	Akaike info criterion	5.826484	
Sum squared resid	12927.79	Schwarz criterion	5.840177	
Log likelihood	-1906.173	F-statistic	129.3829	
Durbin-Watson stat	2.055799	Prob(F-statistic)	0.000000	
Inverted MA Roots	-.49			

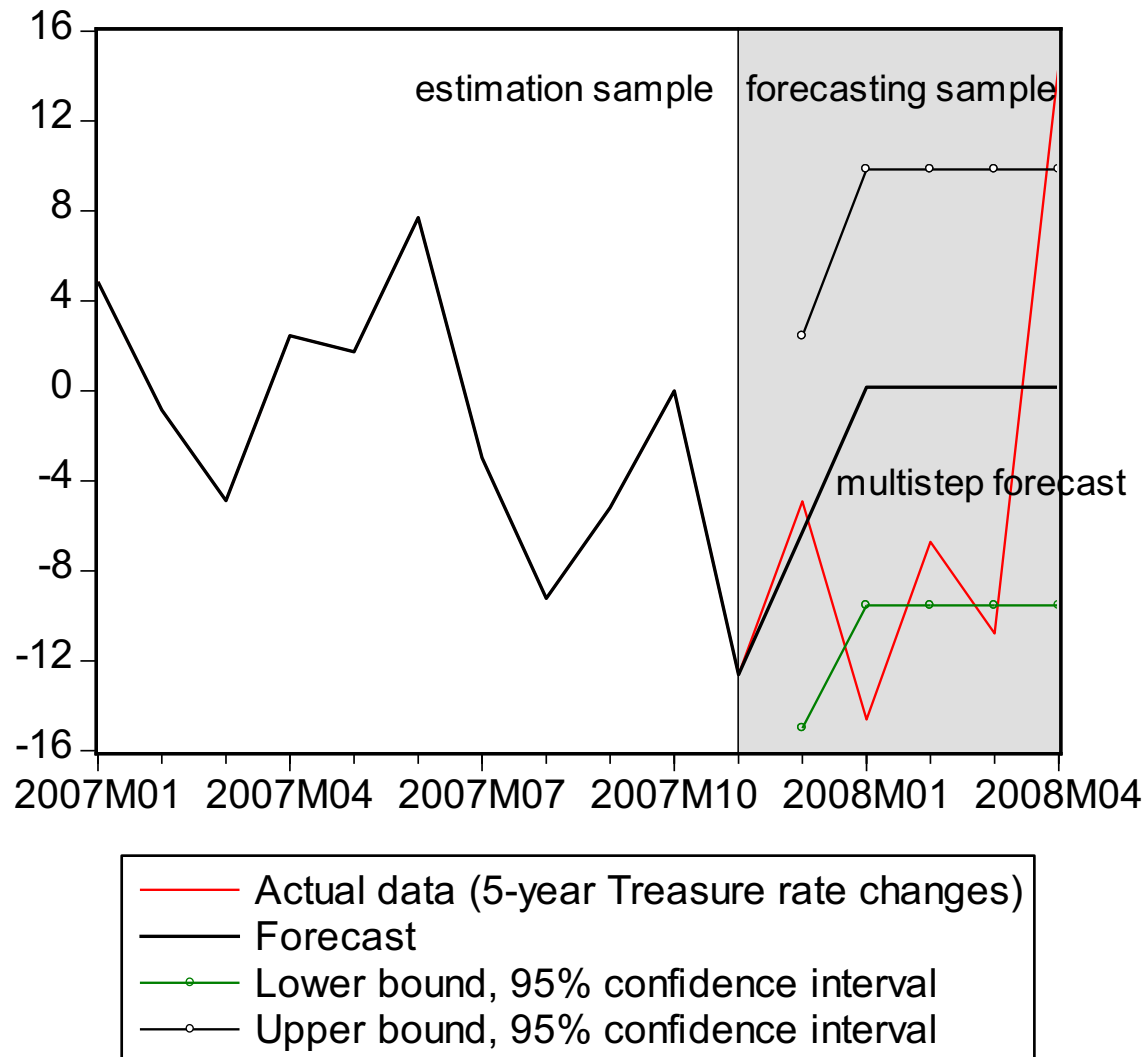


# December 2007-April 2008 Forecasts of 5-year Treasury Yield Changes

$h = 1$ 12/2007	$f_{t,1} = \hat{\mu} + \hat{\theta}\varepsilon_t =$ $= 0.160 + 0.485\hat{\varepsilon}_t$ $= -6.276\%$	$\sigma_{t+1 t}^2 = \hat{\sigma}_\varepsilon^2 = 4.449^2$	$f(Y_{t+1} I_t) \rightarrow N(\mu_{t+1 t}, \sigma_{t+1 t}^2)$ $= N(-6.276, 4.449^2)$
$h = 2$ 1/2008	$f_{t,2} = \hat{\mu} = 0.160\%$	$\sigma_{t+2 t}^2 = \hat{\sigma}_\varepsilon^2(1 + \hat{\theta}^2)$ $= 4.449^2(1 + 0.485^2)$ $= 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+2} I_t) \rightarrow N(0.16, 23.683)$
$h = 3$ 2/2008	$f_{t,3} = \hat{\mu} = 0.160\%$	$\sigma_{t+3 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+3} I_t) \rightarrow N(0.16, 23.683)$
$h = 4$ 3/2008	$f_{t,4} = \hat{\mu} = 0.160\%$	$\sigma_{t+4 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+4} I_t) \rightarrow N(0.16, 23.683)$
$h = 5$ 4/2008	$f_{t,5} = \hat{\mu} = 0.160\%$	$\sigma_{t+5 t}^2 = 23.683 = \hat{\sigma}_Y^2$	$f(Y_{t+5} I_t) \rightarrow N(0.16, 23.683)$

# Moving Average Models

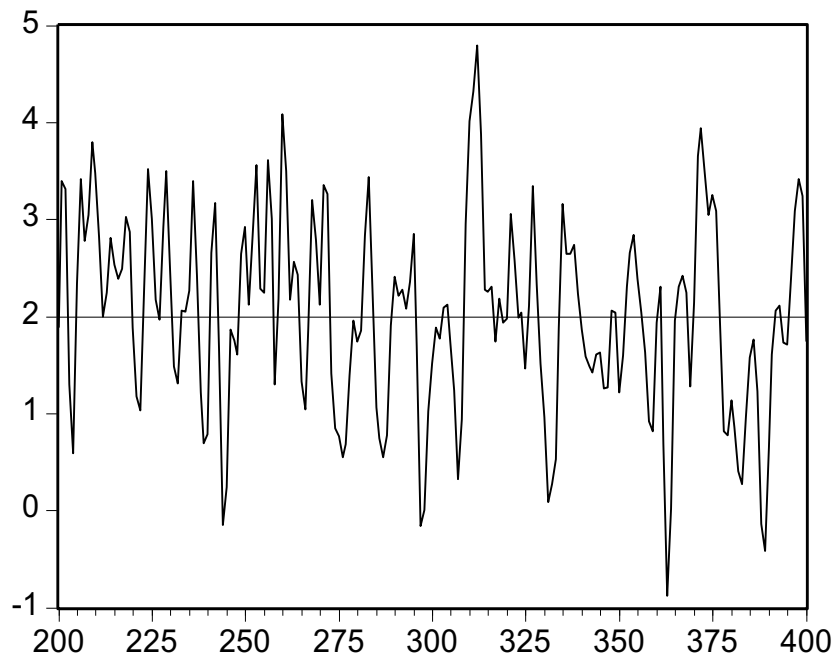
## Example: MA(1) Process



# Moving Average Models

## Example: MA(2) Process

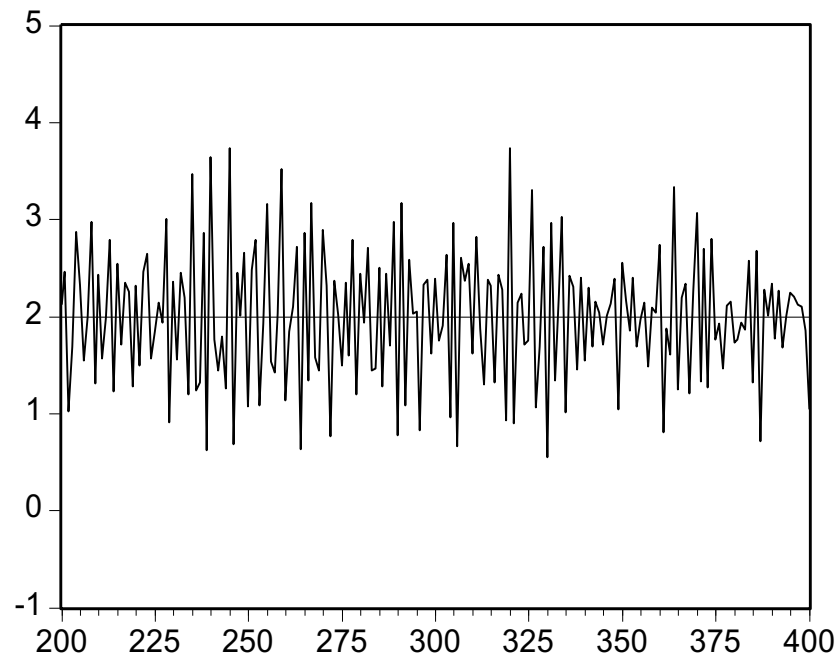
$$Y_t = 2 + 1.7e_{t-1} + 0.72e_{t-2} + e_t$$



— MA(2) with  $\theta_1 = 1.70$  and  $\theta_2 = 0.72$

(a)

$$Y_t = 2 - \varepsilon_{t-1} + 0.25\varepsilon_{t-2} + \varepsilon_t$$

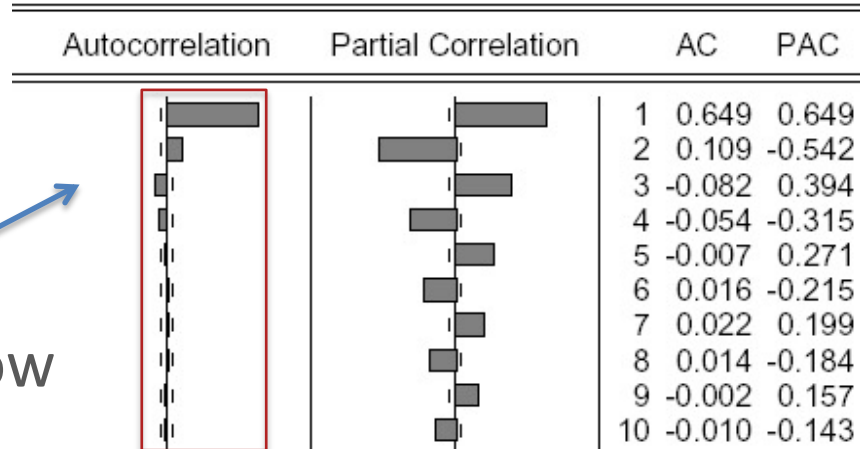


— MA(2) with  $\theta_1 = -1$  and  $\theta_2 = 0.25$

(b)

Sample: 1 2000  
Included observations: 1998

$$Y_t = 2 + 1.7\varepsilon_{t-1} + 0.72\varepsilon_{t-2} + \varepsilon_t$$

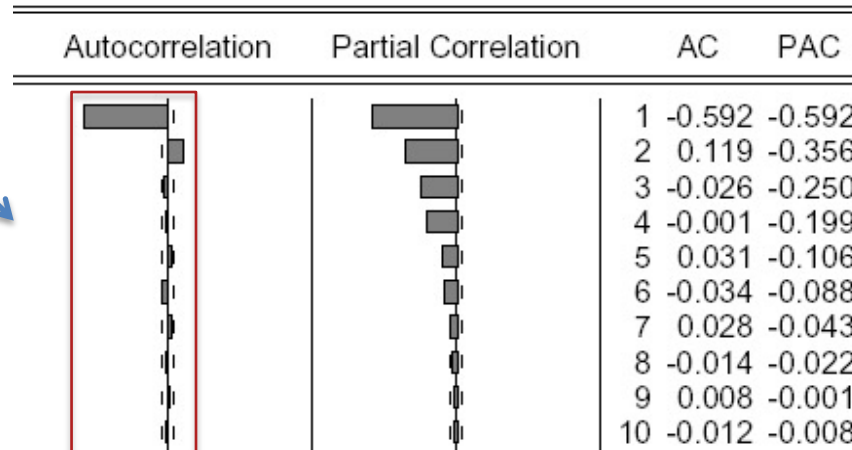


Both ACFs show  
2 spikes only

→ MA(2)

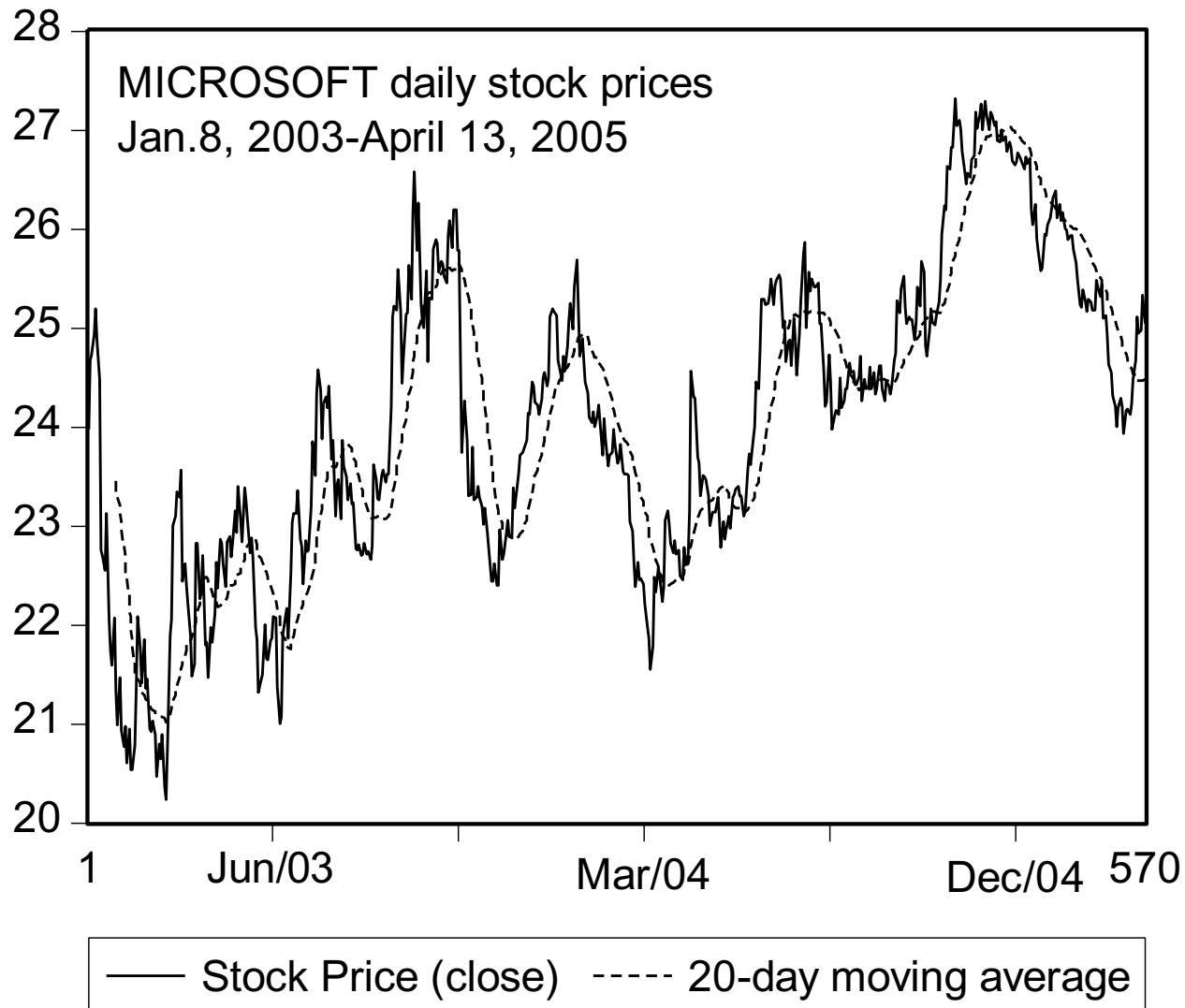
Sample: 1 2000  
Included observations: 1998

$$Y_t = 2 - \varepsilon_{t-1} + 0.25\varepsilon_{t-2} + \varepsilon_t$$



# Moving Average Models

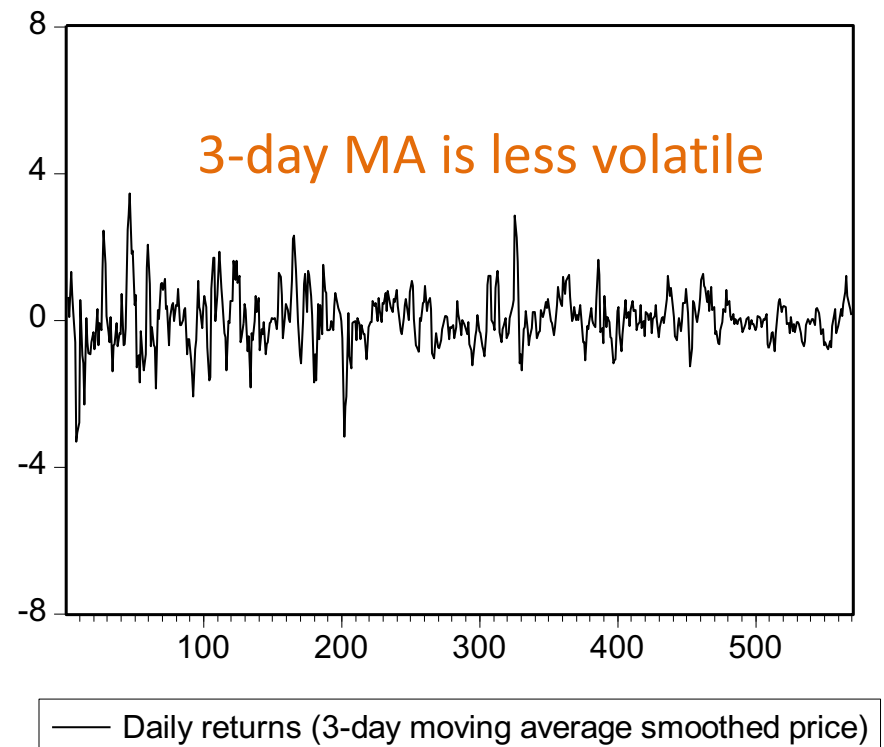
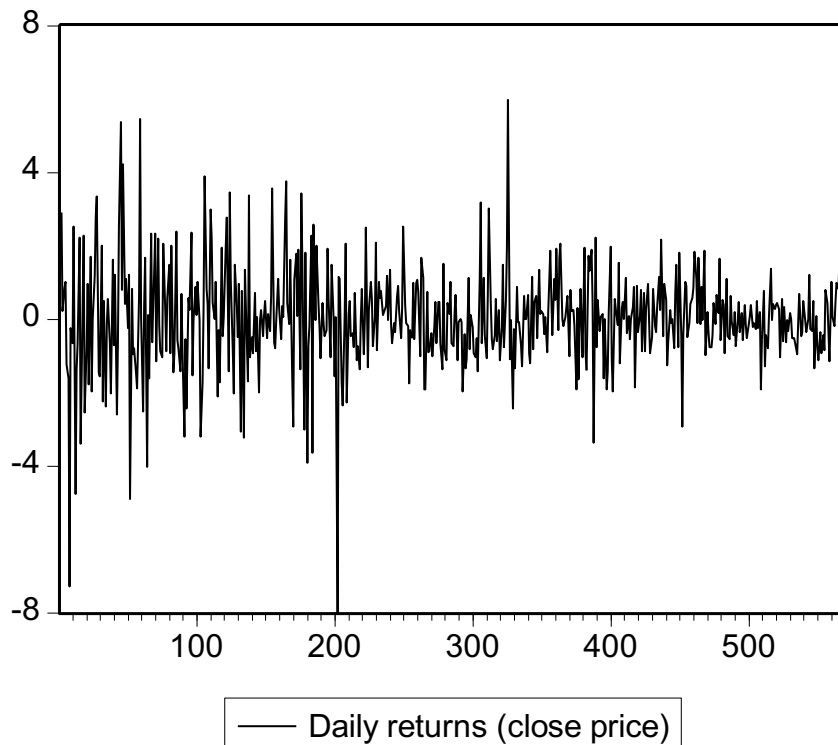
## Example: MA(?) Process



# Moving Average Models

## Example: MA(?) Process

### Daily Returns to Microsoft

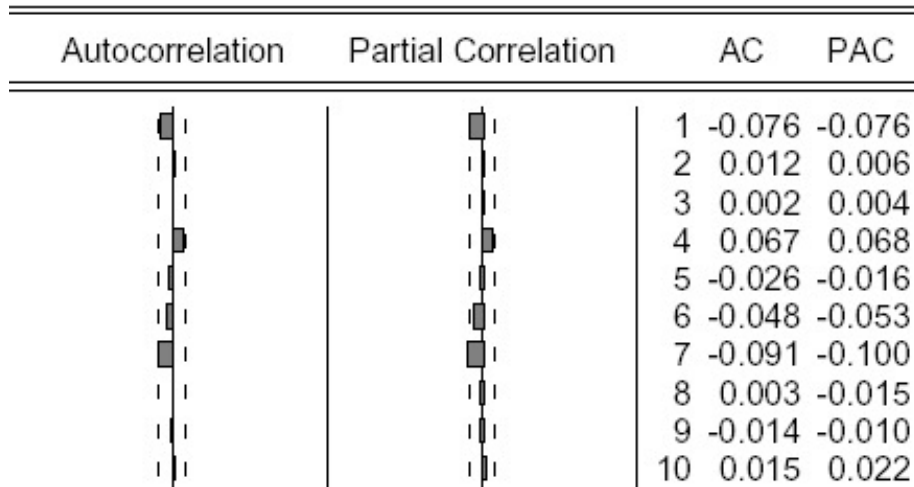


# Moving Average Models

## Example: MA(2) Process

Sample: 1 570  
Included observations: 569

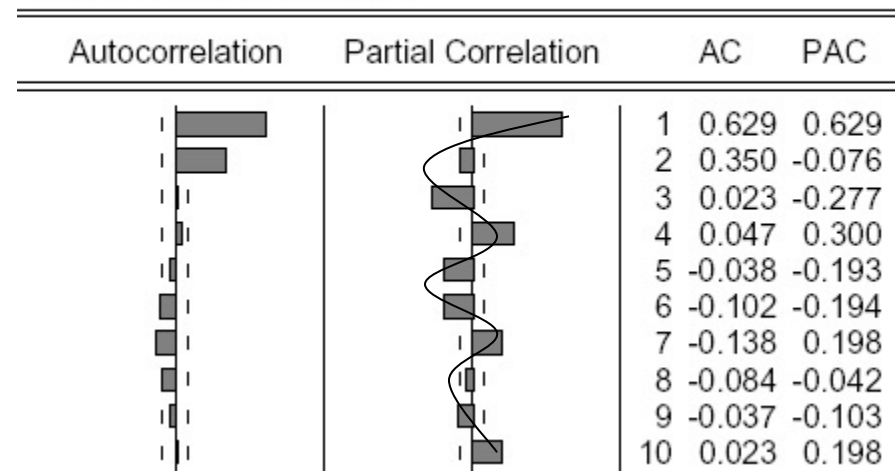
Returns



Looks like white-noise

Sample: 1 570  
Included observations: 569

3-day MA



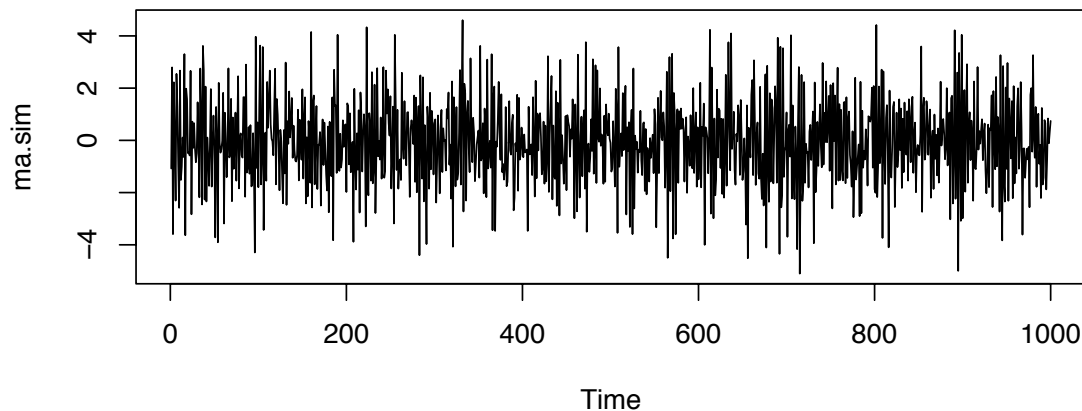
Looks like an MA(2) process

$$Y_t = 0.95\varepsilon_{t-1} + 0.94\varepsilon_{t-2} + \varepsilon_t$$

# Moving Average Models

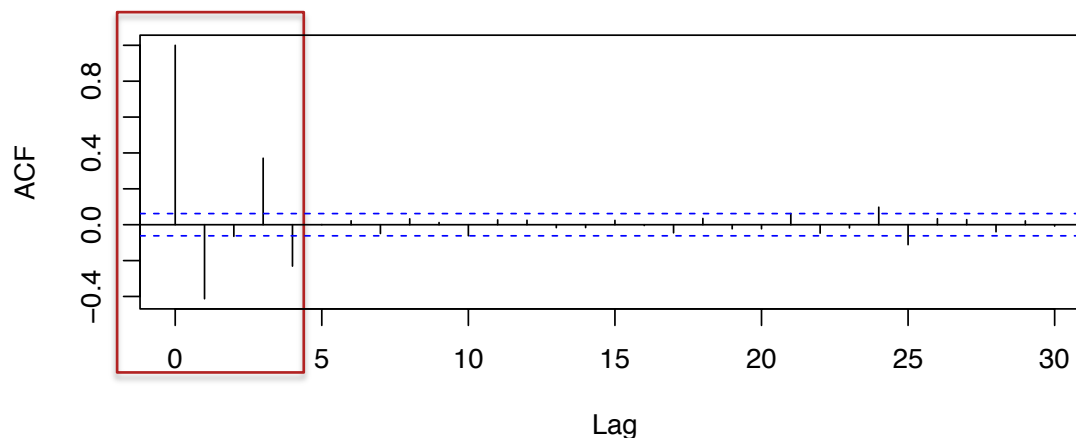
## Example: MA(q) Process

What type of a process is this?



Series ma.sim

MA(4) Process



In general, for any MA(q) process,  $\rho_k = 0$  for any  $k > q$ .



# For Next Class

- Readings about today's class:  
Chapter 6<sup>a</sup>, 8<sup>b</sup>
- Review Exercises / Problems:  
Chapter 7<sup>b</sup>: 1, 2, 3
- Readings for next class:  
6<sup>a</sup>, 7<sup>a</sup>, 8<sup>b</sup>, 9<sup>b</sup>