

# Bayesian Inference: Conjugate Models

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August 15, 2022

# Bayes' Rule

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$$p(\theta|x) = \frac{\underset{\text{likelihood}}{p(x|\theta)} \underset{\text{prior}}{p(\theta)}}{\underset{\text{evidence}}{p(x)}} = \frac{p(x|\theta)p(\theta)}{\int p(x|\theta)p(\theta) d\theta}$$

## Posterior

The posterior distribution is proportional to the prior times the likelihood:

$$p(\theta|x) \propto p(x|\theta)p(\theta)$$

The posterior distribution *is a distribution* over  $\theta$ .

## Evidence

The evidence, or *marginal likelihood*, can be used for model comparison.

## Example: Beta-bernoulli model

Sometimes, we can compute the posterior distribution by hand, given prior and likelihood.

### Setup: flipping a coin

Probability that it lands heads is (unknown)  $\theta$ .

Prior probability over  $\theta$  assumed to follow a  $Beta(3, 3)$  distribution:

$$p(\theta) = \frac{\theta^{3-1}(1-\theta)^{3-1}}{B(3, 3)}$$

Note:  $\theta \sim Beta(a, b)$  means  $p(\theta) \propto \theta^{a-1}(1-\theta)^{b-1}$

Will collect data by flipping coin once. Likelihood of observing heads ( $x = 1$ ) or tails ( $x = 0$ ) is given by a Bernoulli distribution:

$$p(x|\theta) = \theta^x(1-\theta)^{1-x}$$

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Note:  $\theta \sim Beta(a, b)$  means  $p(\theta) = \frac{1}{B(a, b)} \theta^{a-1} (1-\theta)^{b-1}$ , where  $B(a, b) := \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$ .

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### Computing the posterior after observing $x=1$

$$p(\theta|x) \propto p(x|\theta)p(\theta) = \theta^1(1-\theta)^0\theta^2(1-\theta)^2 = \theta^3(1-\theta)^2 \implies \theta|x \sim Beta(4, 3)$$

# Conjugacy

## The idea

We have conjugacy when the prior and the posterior distributions are in the same family (e.g. in the previous example, the prior and posterior are beta distributions).

## Definition

Conjugacy can be defined as follows (gelman2013bayesian). If  $\mathcal{F}$  is a class of sampling distributions and  $\mathcal{P}$  is a class of prior distributions for  $\theta$ , then the class  $\mathcal{P}$  is *conjugate* for  $\mathcal{F}$  if

$$p(\theta \mid y) \in \mathcal{P} \text{ for all } p(\cdot \mid \theta) \in \mathcal{F} \text{ and } p(\cdot) \in \mathcal{P}$$

Technical note: the condition trivially holds if  $\mathcal{P}$  is taken to be the space of all probability distributions!

# Posterior predictive distribution

## Given

$p(\theta|x)$  - posterior

$p(\theta)$  - prior

$p(x|\theta)$  - likelihood

## Posterior predictive distribution

Consider the probability of new data  $x'$ . Posterior predictive distribution is:

$$p(x'|x) = \int p(x', \theta|x) d\theta = \int p(x'|\theta, x)p(\theta|x) d\theta = \int p(x'|\theta)p(\theta|x) d\theta$$

Incorporates the knowledge and uncertainty about  $\theta$  that we still had after seeing data  $x$ .

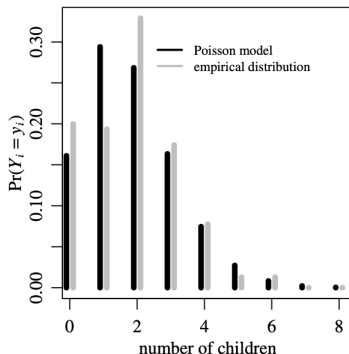
## Another example: The Gamma-Poisson model

Some measurements, such as a person's number of children or number of friends, have values that are whole numbers. Perhaps the simplest probability model on whole numbers is the Poisson model.

- **Sample:**  $X$  the observed number.
- **Sample space:**  
 $\mathcal{X} = \{0, 1, 2, \dots\}$
- **Density:**

$$p(x \mid \lambda) = \frac{\lambda^x e^{-\lambda}}{x!} \quad (1)$$

- **Parameter:**  $\lambda = \mathbb{E}[X]$
- **Parameter space:**  
 $\lambda \in (0, \infty)$



A Poisson distribution with mean 1.83, along with the empirical distribution of the number of children of women of age 40 from the GSS during the 1990's

# The Gamma-Poisson model

The Gamma distribution is a conjugate prior for the Poisson likelihood.  
Can you show this?



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Can you show this?

Given observations  $\mathbf{x} = (x_1, \dots, x_n)$ , we have

$$p(\lambda) \propto \lambda^{\alpha-1} \exp\{-\beta\lambda\}$$

$$p(\mathbf{x} \mid \lambda) \stackrel{iid}{=} \prod_{i=1}^n p(x_i \mid \lambda)$$

$$\stackrel{(Poisson)}{\propto} \lambda^{\sum_{i=1}^n x_i} \exp\{-n\lambda\}$$

so

$$p(\lambda \mid \mathbf{x}) \propto p(\mathbf{x} \mid \lambda)p(\lambda)$$

$$= \lambda^{(\alpha + \sum_i x_i)-1} \exp\{-(\beta + n)\lambda\}$$

# The Gamma-Poisson model

That is,

$$\lambda \sim \text{Gamma}(\alpha, \beta)$$

$$x_i \mid \lambda \stackrel{\text{iid}}{\sim} \text{Poisson}(\lambda)$$

$$\implies \lambda \mid \mathbf{x} \sim \text{Gamma}\left(\alpha + \sum_{i=1}^n x_i, \beta + n\right)$$

## Gamma-Poisson Model: Posterior Expectation

Can you write the posterior expectation as a compromise between the prior expectation and sample mean (like we did for the Beta-Bernoulli model)?

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Gamma dist.

$$\mathbb{E}[\lambda \mid \mathbf{x}] = \frac{\alpha + \sum_{i=1}^n x_i}{\beta + n}$$

Gamma dist.

$$= \frac{\beta}{\beta + n} \underbrace{\frac{\alpha}{\beta}}_{\text{prior expectation}} + \frac{n}{\beta + n} \underbrace{\frac{\sum_{i=1}^n x_i}{n}}_{\text{sample mean}}$$

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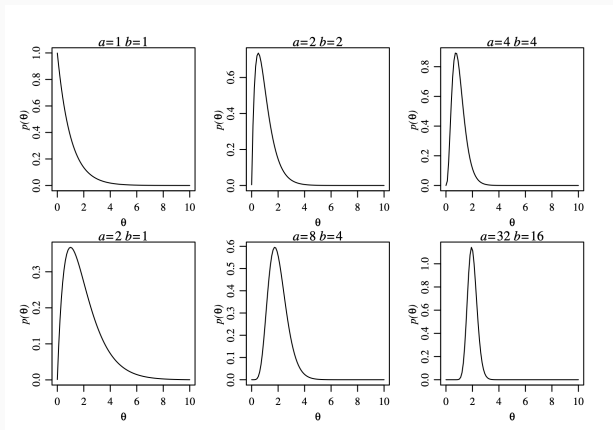
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*Interpretation?*

- $\beta$ : number of prior observations
- $\alpha$ : sum of counts from  $\beta$  prior observations

# Selecting a prior

The previous slide gives us one way to specify the prior distribution:



“My prior belief is that the average count in the population is  $\frac{\alpha}{\beta}$ , and I am as certain about this as if I had observed this in a sample of size  $\beta$ .”

# Gamma-Poisson Model: Posterior Predictive

The posterior predictive for the Gamma Poisson is given by

$$x_{\text{new}} \sim \text{NegativeBinomial}\left(\alpha', \frac{1}{1 + \beta'}\right)$$

where  $(\alpha', \beta')$  are the posterior parameters of the Gamma.

The Negative Binomial is a *two-parameter* alternative to the Poisson model which provides a flexible model of count data (e.g., it can handle overdispersion)

Generally, computing the posterior distribution is much harder than in this example!

Consider the denominator in  $p(\theta|x) = \frac{p(x|\theta)p(\theta)}{\int p(x|\theta)p(\theta) d\theta}$  - integrals are hard

In nonconjugate examples, we need approaches to work with the posterior distribution when we cannot calculate it directly. Stay tuned!



# Exponential families

The Bernoulli and the Poisson distributions are both **exponential families**.

## Exponential family

An *exponential family* is a set of probability distributions whose probability density functions have the following form

$$p(x | \theta) = h(x) \exp\{\eta(\theta)^T t(x) - a(\theta)\}$$

where we refer to  $h$  as the base measure,  $\eta$  as the natural parameter,  $t$  as the sufficient statistics, and  $a$  as the log normalizer.

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- Any probability model in the exponential family has a **conjugate prior**.
- When the conjugate prior is used, you get the posterior and posterior predictive in **closed form**!