

Assignment 3

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**1****1.a**

$$\begin{aligned}
\ell(\theta) &= x_1 \log(2 + \theta) + (x_2 + x_3) \log(1 - \theta) + x_4 \log \theta + c \\
&= 125 \log(2 + \theta) + (18 + 19) \log(1 - \theta) + 35 \log \theta + c \\
\ell'(\theta) &= \frac{125}{2 + \theta} + \frac{37}{1 - \theta} + \frac{35}{\theta} \\
&= \frac{-197\theta^2 + 16\theta + 70}{(\theta + 2)(1 - \theta)\theta}
\end{aligned}$$

Thus the MLE is found by setting the numerator of the derivative of the log-likelihood to zero.

$$\begin{aligned}
0 &= -197\theta^2 + 16\theta + 70 \\
\theta &= \frac{1}{197} \left( 8 \pm \sqrt{13854} \right) \\
&\approx -0.556868, 0.638086
\end{aligned}$$

Since  $0 \leq \theta \leq 1$ , then the MLE is  $\theta = \frac{1}{197} \left( 8 + \sqrt{13854} \right) \approx 0.638086$ .

**1.b**

E-Step: Calculate

$$\begin{aligned}
Q(\theta | \theta^{(k)}) &= \mathbb{E}_{\theta^{(k)}} \{ \ell_c(\theta) | y \} \\
&= \mathbb{E}_{\theta^{(k)}} \{ (x_{12} + x_4) \log \theta + (x_2 + x_3) \log(1 - \theta) | y \} \\
&= \mathbb{E}_{\theta^{(k)}} \{ x_{12} | y \} \log \theta + x_4 \log \theta + (x_2 + x_3) \log(1 - \theta) \\
&= \frac{\theta^{(k)} (1 + x_2 + x_3 + x_4)!}{(1 - \theta^{(k)})^2 (x_2 + x_3 + x_4)!} \log \theta + x_4 \log \theta + (x_2 + x_3) \log(1 - \theta)
\end{aligned}$$

M-Step: Maximize  $Q(\theta | \theta^{(k)})$  with respect to  $\theta$ .

$$\begin{aligned}
\frac{dQ(\theta^{(k+1)} | \theta^{(k)})}{d\theta} &= 0 \\
&= \frac{1}{\theta^{(k+1)}} \frac{\theta^{(k)}}{(1 - \theta^{(k)})^2} \frac{(1 + x_2 + x_3 + x_4)!}{(x_2 + x_3 + x_4)!} + \frac{x_4}{\theta^{(k+1)}} - \frac{x_2 + x_3}{1 - \theta^{(k+1)}} \\
\theta^{(k+1)} &= \frac{\theta^{(k)} (1 + x_2 + x_3 + x_4)! + x_4 (1 - \theta^{(k)})^2 (x_2 + x_3 + x_4)!}{\theta^{(k)} (1 + x_2 + x_3 + x_4)! + (x_2 + x_3 + x_4) (1 - \theta^{(k)})^2 (x_2 + x_3 + x_4)!}
\end{aligned}$$

### 1.c

I don't know actually how to calculate  $Q \dots$  so  $\theta^{(k)} \rightarrow 1$  as  $k \rightarrow \infty$  which is not right...

## 2

### 2.a

First, find the CDF of the Weibull distribution,

$$\begin{aligned} w(x|\alpha) &= \int_0^x W(t|\alpha) dt \\ &= \int_0^x \frac{2t}{\alpha^2} e^{-\frac{t^2}{\alpha^2}} dt \\ &= -e^{-\frac{t^2}{\alpha^2}} \Big|_0^x \\ &= 1 - e^{-\frac{x^2}{\alpha^2}} \end{aligned}$$

Then, find the complete log-likelihood function

$$\begin{aligned} \ell_c(\alpha) &= \sum_u \log(W(x_i|\alpha)) + \sum_c \log(1 - w(x_i|\alpha)) \\ &= \sum_u \log\left(\frac{2x_i}{\alpha^2} e^{-\frac{x_i^2}{\alpha^2}}\right) + \sum_c \log\left(e^{-\frac{x_i^2}{\alpha^2}}\right) \\ &= \sum_u \log\left(\frac{2x_i}{\alpha^2}\right) - \sum_{i=1}^n \frac{x_i^2}{\alpha^2} \end{aligned}$$

Then, the E-step is

$$Q(\alpha|\alpha^{(k)}) = \mathbb{E}_{\alpha^{(k)}} \{\ell_c(\theta) | y\}$$

and the M-step is

$$\frac{dQ(\alpha^{(k+1)}|\alpha^{(k)})}{d\alpha} = 0$$

### 2.b

## 3

$$f(x) = ce^{-x}, \quad 0 < x < 2$$

$$\begin{aligned} c^{-1} &= \int_0^2 e^{-x} dx \\ &= -e^{-x} \Big|_0^2 \\ &= 1 - e^{-2} \end{aligned}$$

$$\begin{aligned} F(x) &= \frac{\int_0^x e^{-t} dt}{1 - e^{-2}} \\ &= \frac{1 - e^{-x}}{1 - e^{-2}} \end{aligned}$$

First sample  $u$  from  $\text{unif}(0, 1)$ , then find  $F^{-1}(u)$ . Equivalently solve  $F(x) - u = 0$ .

$$\begin{aligned} 0 &= F(x) - u \\ 0 &= \frac{1 - e^{-x}}{1 - e^{-2}} - u \\ 0 &= 1 - e^{-x} - (1 - e^{-2})u \\ x &= -\log(1 - (1 - e^{-2})u) \end{aligned}$$

The results are displayed in Figure 1. The gray bars represent frequency. The white bars represent the cumulative frequency. The lines represent the estimated densities.

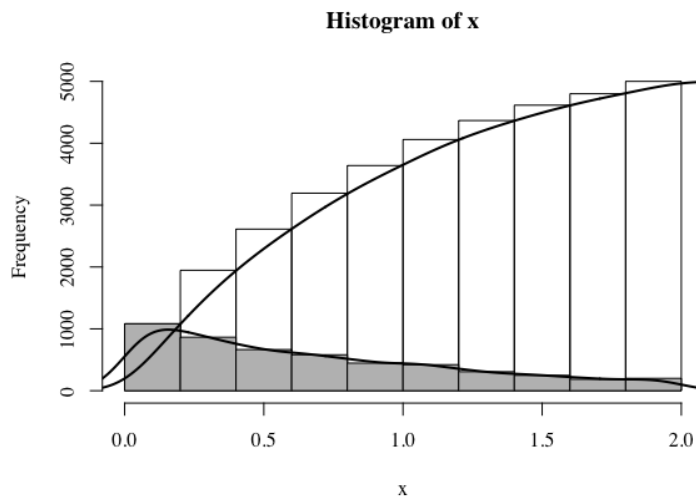


Figure 1: **Problem 3.**

**4**

**4.a**

$$\begin{aligned} \alpha_1 &= \sup_{0 < x} \frac{q(x)}{g_1(x)} = \sup_{0 < x} \frac{\frac{e^{-x}}{1+x^2}}{e^{-x}} = \sup_{x > 0} \frac{1}{1+x^2} = 1 \\ \alpha_2 &= \sup_{0 < x} \frac{q(x)}{g_2(x)} = \sup_{0 < x} \frac{\frac{e^{-x}}{1+x^2}}{\frac{2}{\pi(1+x^2)}} = \sup_{x > 0} \frac{\pi}{2} e^{-x} = \frac{\pi}{2} \end{aligned}$$

Use the inverse transform to sample  $x$  from  $g_1(x)$  and  $g_2(x)$  using  $u$  from  $\text{unif}(0, 1)$ .

$$\begin{aligned} G_1(x) &= \int_0^x e^{-t} dt \\ &= 1 - e^{-x} \\ x &= -\log(1 - u) \end{aligned} \qquad \begin{aligned} G_2(x) &= \frac{2}{\pi} \int_0^x \frac{1}{1+t^2} dt \\ &= \frac{2}{\pi} \tan^{-1} x \\ x &= \tan\left(\frac{\pi}{2} u\right) \end{aligned}$$

The results are shown in Table 1.

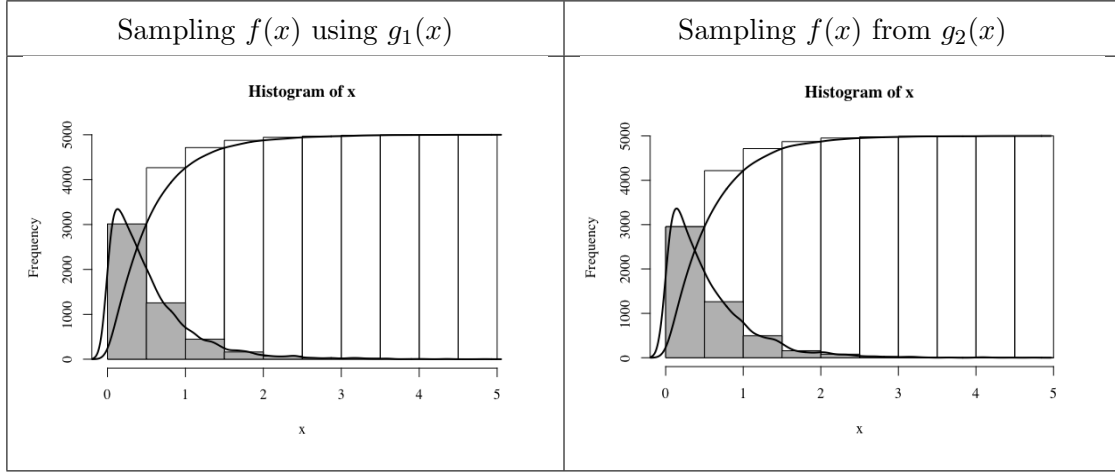


Table 1: **Problem 4.**

#### 4.b

The number of iterations needed to produce 5,000 data points required approximately 8,000 iterations using  $g_1(x)$  and 11,000 iterations using  $g_2(x)$ . This makes sense since the  $\alpha_1 < \alpha_2$  and the acceptance probability, which is inversely proportional to the  $\alpha$  value, is higher for  $g_1(x)$ .

### 5

First, the problem can be converted into polar coordinates with the form

$$f(r, \theta) \propto r^\alpha \cos^\alpha \theta r \sin \theta = r^{\alpha+1} \cos^\alpha \theta \sin \theta = q(r, \theta), \quad 0 < r \leq 1, \quad 0 < \theta < \frac{\pi}{2}$$

where  $\alpha > -1$  since otherwise  $f(r, \theta)$  would not be a probability distribution since the integral would not converge over the domain for  $\alpha \leq -1$ .  $f(r, \theta)$  is now easily integrable where  $f(r, \theta) = c q(r, \theta)$  and

$$\begin{aligned} c^{-1} &= \int_0^{\frac{\pi}{2}} \int_0^1 r^{\alpha+1} \cos^\alpha \theta \sin \theta r dr d\theta \\ &= \frac{1}{\alpha+3} r^{\alpha+3} \Big|_0^1 \frac{1}{\alpha+1} \cos^{\alpha+1} \theta \Big|_0^{\frac{\pi}{2}} \\ &= \frac{1}{\alpha+3} \frac{1}{\alpha+1} \\ c &= (\alpha+3)(\alpha+1) > 1, \quad \alpha > -1 \end{aligned}$$

Then sample  $(x, y)$  from  $[0, 1] \times [0, 1]$ . If  $x^2 + y^2 \leq 1$ , then accept  $(x, y)$ . Now sample  $u$  uniformly from  $[0, 1]$  and accept  $(x, y)$  if  $u \leq f(x, y)/(\beta g(x, y))$  where  $\beta = \sup_{(x, y) \in \mathbb{R}^2} f(x, y)/g(x, y)$  and  $f(x, y) \leq c g(x, y)$  for all  $(x, y)$  in the upper right quarter of the unit disc and  $c > 0$

```
par(family = 'serif')
setwd("/Users/mikhailgaerlan/Box Sync/Education/UC Davis/2016-2017
Spring/STA 243 Computational Statistics/Assignments/Assignment 2")
rm(list=ls())

n = 5000
u = runif(n)
x = -log(1-(1-exp(-2))*u)

h = hist(x,plot=F)
hcum = h
hcum$counts=cumsum(hcum$counts)
plot(hcum)
plot(h,add=T,col='grey')

d = density(x)
lines(d$x,d$y*length(x)*diff(h$breaks)[1],lwd=2)
lines(d$x,cumsum(d$y)/max(cumsum(d$y))*length(x),lwd=2)
```

```

par(family = 'serif')
setwd("/Users/mikhailgaerlan/Box Sync/Education/UC Davis/2016-2017
Spring/STA 243 Computational Statistics/Assignments/Assignment 2")
rm(list=ls())

alpha1 = 1
alpha2 = pi/2

q = function(x){
  return(exp(-x)/(1+x^2))
}
g1 = function(x){
  return(exp(-x))
}
g2 = function(x){
  return(2/(pi*(1+x^2)))
}
sampleg1 = function(n){
  x = 0*(1:n)
  for(i in 1:n){
    u = runif(1)
    x[i] = -log(1-u)
  }
  return(x)
}
sampleg2 = function(n){
  x = 0*(1:n)
  for(i in 1:n){
    u = runif(1)
    x[i] = tan(pi*u/2)
  }
  return(x)
}
samplef = function(n,o){
  x2 = 0*(1:n)
  j = 0
  for (i in 1:n){
    test = T
    while (test){
      if (o == 1){
        test2 = T
        while (test2){
          x = sampleg1(1)
          if (x < 5){
            test2 = F
          }
        }
      }
      j = j + 1
      if (runif(1) < q(x)/(alpha1*g1(x))){
        x2[i] = x
      }
    }
  }
}

```

```

        test = F
    }
} else {
    test2 = T
    while (test2){
        x = sampleg2(1)
        if (x < 5){
            test2 = F
        }
    }
    j = j + 1
    if (runif(1) < q(x)/(alpha2*g2(x))){
        x2[i] = x
        test = F
    }
}
}
}
print(j)
return(x2)
}

x = samplef(5000,1)
h = hist(x,plot=F)
hcum = h
hcum$counts=cumsum(hcum$counts)
plot(hcum)
plot(h,add=T,col='grey')
d = density(x)
lines(d$x,d$y*length(x)*diff(h$breaks)[1],lwd=2)
lines(d$x,cumsum(d$y)/max(cumsum(d$y))*length(x),lwd=2)

x = samplef(5000,2)
h = hist(x,plot=F)
hcum = h
hcum$counts=cumsum(hcum$counts)
plot(hcum)
plot(h,add=T,col='grey')
d = density(x)
lines(d$x,d$y*length(x)*diff(h$breaks)[1],lwd=2)
lines(d$x,cumsum(d$y)/max(cumsum(d$y))*length(x),lwd=2)

```