$statComp_hw4$

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1

For $x \in [a_{i-1}, a_i], i \in 1, \cdots, j$,

$$\begin{split} P(X^* \leq x | J = i) &= \int_{-\infty}^x f_i(x) dx \\ &= k \int_{-a_{i-1}}^x f(x) dx \\ &= k (F(x) - \frac{i-1}{k}) \\ &= k F(x) - i + 1 \end{split}$$

Hence,

$$P(X^* \leq x | J=i) = \left\{ \begin{array}{ll} kF(x) - i + 1 & \quad x \in [a_{i-1}, a_i] \\ 0 & \quad x \in (-\infty, a_{i-1}) \\ 1 & \quad x \in (a_i, \infty) \end{array} \right.$$

Hence,

$$\begin{split} P(X^* \leq x) &= \sum_{j=1}^n P(X^* \leq x | J = j) P(J = j) \\ &= \frac{1}{k} \sum_{j=1}^n P(X^* \leq x | J = j) \\ &= \frac{1}{k} \left(\sum_{j=1}^{i-1} 1 + P(X^* \leq x | J = i) + \sum_{j=i+1}^n 0 \right) \\ &= \frac{1}{k} \left(i - 1 + [kF(x) - i + 1] \right) \\ &= F(x) \end{split}$$

Hence, $X^* \sim X$.

For Y^* ,

$$Y^* = \frac{g_j(X^*)}{f_j(X^*)} = \frac{g_j(X^*)}{kf(X^*)} \sim \frac{g(X)}{kf(X)} = \frac{Y}{k}$$

with $X^* \sim X$.

 $\mathbf{2}$

exercise 6.1

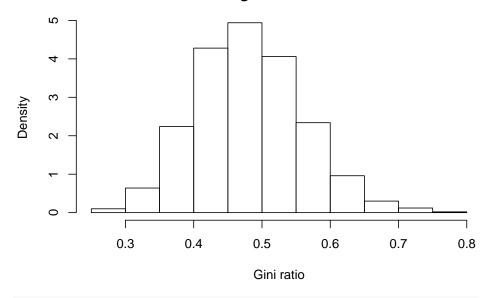
```
n <- 20
m <- 1000
mse <- numeric(9)
for (k in 1:9) {
   tmean <- numeric(m)
   for (i in 1:m) {
      x <- sort((rcauchy(n)))
      tmean[i] <- sum(x[(k+1):(n-k)])/(n-2*k)
   }
   target <- median(tmean)
   mse[k] <- mean((tmean - target)^2)
}</pre>
```

```
k < -c(1:9)
rbind(k, mse)
##
            [,1]
                       [,2]
                                  [,3]
                                             [,4]
                                                        [,5]
                                                                   [,6]
                                                                              [,7]
       1.000000 2.0000000 3.0000000 4.0000000 5.0000000 6.0000000 7.0000000
## mse 1.735897 0.4097611 0.2216231 0.1745251 0.1683093 0.1426654 0.1330567
                        [,9]
##
             [,8]
## k
       8.0000000 9.0000000
## mse 0.1290898 0.1357178
exercise 6.9
n <- 20
m <- 1000
gini.ratio <- numeric(m)</pre>
for (j in 1:m) {
  x <- sort(rlnorm(n))</pre>
  mu <- mean(x)
  G <- 0
  for (i in 1:n) {
    G \leftarrow G + (2*i - n - 1) * x[i]
  gini.ratio[j] \leftarrow G / (n<sup>2</sup> * mu)
}
mean <- mean(gini.ratio)</pre>
median <- median(gini.ratio)</pre>
c(mean, median)
## [1] 0.4811732 0.4759992
quantile(gini.ratio, probs = seq(.1, .9, by = .1))
                                                                            70%
                     20%
                                                                 60%
##
          10%
                                30%
                                           40%
                                                      50%
                                                                                       80%
## 0.3851137 0.4159509 0.4385938 0.4559212 0.4759992 0.4978336 0.5189110 0.5455270
```

```
## 90%
## 0.5847519
```

```
hist(gini.ratio, probability = T,
    main = 'Histogram of Gini ratio generated
    from log normal distribution',
    xlab = 'Gini ratio')
```

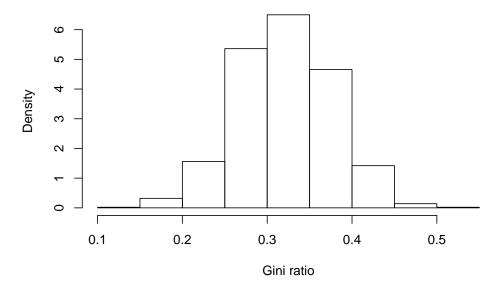
Histogram of Gini ratio generated from log normal distribution



```
n <- 20
m <- 1000
gini.ratio <- numeric(m)
for (j in 1:m) {
    x <- sort(runif(n))
    mu <- mean(x)
    G <- 0
    for (i in 1:n) {
        G <- G + (2*i - n - 1) * x[i]
    }
    gini.ratio[j] <- G / (n^2 * mu)</pre>
```

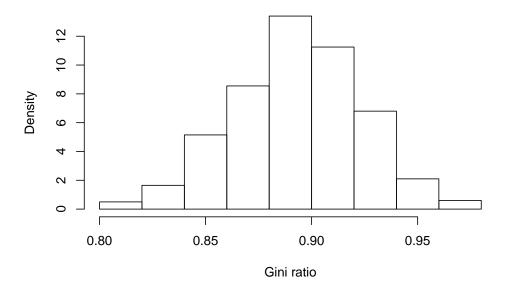
```
}
mean <- mean(gini.ratio)</pre>
median <- median(gini.ratio)</pre>
c(mean, median)
## [1] 0.3214615 0.3202069
quantile(gini.ratio, probs = seq(.1, .9, by = .1))
         10%
                    20%
                               30%
                                         40%
                                                    50%
                                                              60%
                                                                         70%
                                                                                   80%
##
## 0.2531085 0.2730138 0.2914422 0.3059615 0.3202069 0.3356224 0.3515961 0.3691293
         90%
##
## 0.3937447
hist(gini.ratio, probability = T,
     main = 'Histogram of Gini ratio generated
     from uniform distribution',
     xlab = 'Gini ratio')
```

Histogram of Gini ratio generated from uniform distribution



```
n <- 100
m <- 1000
gini.ratio <- numeric(m)</pre>
for (j in 1:m) {
  x <- sort(rbinom(n, size = 1, prob = .1))
 mu <- mean(x)</pre>
  G <- 0
 for (i in 1:n) {
    G \leftarrow G + (2*i - n - 1) * x[i]
  gini.ratio[j] \leftarrow G / (n^2 * mu)
}
mean <- mean(gini.ratio)</pre>
median <- median(gini.ratio)</pre>
c(mean, median)
## [1] 0.89808 0.90000
quantile(gini.ratio, probs = seq(.1, .9, by = .1))
## 10% 20% 30% 40% 50% 60% 70% 80% 90%
## 0.86 0.87 0.88 0.89 0.90 0.91 0.91 0.92 0.94
hist(gini.ratio, probability = T,
     main = 'Histogram of Gini ratio generated
     from Bernoulli distribution',
     xlab = 'Gini ratio')
```

Histogram of Gini ratio generated from Bernoulli distribution



3

As $X \sim N(0,1)$, denote the pdf and cdf of standard normal as f(x) and F(x) respectively, then

$$\begin{split} f_{X_{(i)}}(x) &= \frac{n!}{(i-1)!(n-i)!} F^{i-1}(x) (1-F(x))^{n-i} f(x) \\ f_{X_{(n-i+1)}}(x) &= \frac{n!}{(i-1)!(n-i)!} F^{n-i}(x) (1-F(x))^{i-1} f(x) \end{split}$$

Known F(-x) = 1 - F(x), f(x) = f(-x), we know that

$$\begin{split} E\left[X_{(i)} + X_{(n-i+1)}\right] &= \int_{-\infty}^{\infty} \frac{n!}{(i-1)!(n-i)!} F^{i-1}(x) (1-F(x))^{n-i} f(x) dx + \\ &\int_{-\infty}^{\infty} \frac{n!}{(i-1)!(n-i)!} F^{n-i}(x) (1-F(x))^{i-1} f(x) dx \\ &= \frac{n!}{(i-1)!(n-i)!} \left(\int_{-\infty}^{\infty} F^{i-1}(x) (1-F(x))^{n-i} f(x) dx + \\ &\int_{-\infty}^{\infty} F^{n-i}(-x) (1-F(-x))^{i-1} f(-x) dx\right) \\ &= \frac{n!}{(i-1)!(n-i)!} \left(\int_{-\infty}^{\infty} F^{i-1}(x) (1-F(x))^{n-i} f(x) dx + \int_{-\infty}^{\infty} (1-F(x))^{n-i} F^{i-1}(x) f(x) dx\right) \\ &= 0 \end{split}$$

Hence,

$$\begin{split} E\left[\bar{X}_{[-k]}\right] &= E\left[\frac{1}{n-2k} \sum_{i=k+1}^{n-k} X_{(i)}\right] \\ &= \frac{1}{n-2k} \sum_{i=k+1}^{n-k} E\left[X_{(i)}\right] \\ &= \frac{1}{2(n-2k)} \sum_{i=k+1}^{n-k} E\left[X_{(i)} + X_{(n-i+1)}\right] \\ &= 0 \end{split}$$

4

Denote $\frac{X_i - \mu}{\sigma}$ as Z_i , Y = AZ, where A is an orthogonal matrix, and the first row of A is $(\frac{1}{\sqrt{n}}, \cdots, \frac{1}{\sqrt{n}})$, then $Y \sim N(0, I)$,

$$||Y||^2 = \sum_i Y_i^2 \sim \chi^2(n)$$

and

$$Y_1 = \frac{\sum_i (X_i - \mu)}{\sqrt{n}\sigma} = \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1)$$

where Y_1 is the first value of vector Y.

Hence

$$\begin{split} \frac{\sum_i (X_i - \bar{X})^2}{\sigma^2} &= \frac{\sum_i (X_i - \mu)^2}{\sigma^2} - \frac{n(\bar{X} - \mu)^2}{\sigma^2} \\ &= \|Z\|^2 - Y_1^2 \\ &= \|Y\|^2 - Y_1^2 \\ &\sim \chi^2 (n-1) \end{split}$$

and $\frac{\sum_i (X_i - \bar{X})^2}{\sigma^2}$ is independent with $Y_1 = \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma}$, i.e., $\frac{\sum_i (X_i - \bar{X})^2}{\sigma^2}$ is independent with \bar{X}