

Chapter 13: Non parametric Least Squares:

13.1

$$\begin{aligned} a) G(t) &= E[(Z-t)^2] \\ &= E[(Z-E[Z])^2] \\ &= E[Z^2] - 2E[Z]E[E[Z]] + E[E[Z]]^2 \\ &= \sigma^2 \end{aligned}$$

(Tower Property, Law of Total Expectation)

$$E[X] = E[E(X|Y)]$$

$$b) \text{ (Equation 13.1) } \tilde{L}_F := E_{X,Y}[(Y - F^*(X))^2]$$

$$\begin{aligned} \tilde{L}_F &:= E_{X,Y}[(Y - F^*(X))^2] \\ &= E_{X,Y}[Y^2] - 2E_{X,Y}[Y]E[F^*(X)] + E[F^*(X)]^2 \\ &= E_{X,Y}[Y^2] - 2E_{X,Y}[Y]E[E[Y|X=X]] + E[E[Y|X=X]]^2 \\ &= E_{X,Y}[Y^2] - 2E_{X,Y}[Y] \cdot E_{X,Y}[Y] + E[Y]^2 \\ &= \sigma_{XY}^2 \end{aligned}$$

$$\begin{aligned} c) \tilde{L}_F - \tilde{L}_{F^*} &= E_{X,Y}[(Y - F(X))^2] - E[(Y - F^*(X))^2] \\ &= E_{X,Y}[(F(X) - F^*(X))^2] \\ &= \|F - F^*\|_2^2 \end{aligned}$$

13.2

$$\begin{aligned} E[\|F_\theta - F_\theta^*\|_n^2] &= E[\|F_\theta\|_n^2] - 2E[\|F_\theta\|] \cdot E[\|F_\theta^*\|] + E[\|F_\theta^*\|]^2 \\ &= \frac{\|X(\theta - \theta^*)\|_n^2}{n} \\ &= \frac{\|X\|_n^2}{n} \sigma^2 \\ &= \text{rank}(X) \cdot \sigma^2 / n \end{aligned}$$

Relationship / Test:

$$\|X\|_n^2 = \left(\sum_{i=1}^N \sum_{j=1}^m X_{ij}^2 \right)^{2/n} = \text{rank}(X)$$

when $X_{ij} \geq 1$

13.3 (Equation 13.10) "Cubic Spline"

$$\hat{f} \in \arg\min_f \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda_n \int_0^1 f''(x)^2 dx \right\}$$

Book's method: $\hat{f} \in \arg\min_f \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda_n \int_0^1 f''(x)^2 dx \right\}$
 $= \arg\min_f \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda_n \|f\|^2 \right\}$

Where $f(x) = \langle f, K(\cdot, x) \rangle$

$$= \sum_{j=1}^n f \cdot \phi_j(x) \phi_j(x)$$

$$= \sqrt{n} K \alpha$$

$$\|f\|^2 = \frac{1}{n} \langle \sum K(\cdot, x), \sum K(\cdot, x) \rangle$$

$$= \alpha^T K \alpha$$

$$\hat{f} \in \arg\min_f \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda_n \|f\|^2 \right\}$$

$$\in \arg\min_{\theta} \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - \sqrt{n} K \alpha)_2^2 + \lambda_n \alpha^T K \alpha \right\}$$

$$= 0$$

$$\hat{\alpha} = \frac{1}{(K + \lambda_n I)} \left(\frac{y}{\sqrt{n}} \right)$$

$$\hat{f}(x) = \theta_0 + \theta_1 x + \sum_{i=1}^n \hat{\alpha}_i K(x, x_i)$$

Alternative

Method #1

A minimal function sums terms

$$\hat{f}(x) = \hat{F}(x) + F(x)$$

Minimal
Function

Original
"Linear"

Correction
terms

$$= \theta_0 + \theta_1 x + \sum_{i=1}^N \alpha K(x, x_i)$$

Alternative
Method #2

$$\begin{aligned}\hat{f} &= \underset{\hat{f}}{\operatorname{argmin}} \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x))^2 + \lambda_n \int_0^1 f''(x)^2 dx \right\} \\ &= \underset{\hat{f}, \hat{\alpha}}{\operatorname{argmin}} \left\{ \frac{1}{n} \sum_{i=1}^n (y_i - f(x))^2 + \lambda_n \left(\sum_{i=1}^n \hat{\alpha}^T K \right)^2 \right\} \\ &\in \underset{\hat{f}, \hat{\alpha}}{\operatorname{argmin}} \left\{ \sum_{i=1}^N \frac{y_i^2}{n} - \frac{2}{n} \sum_{i=1}^N y_i f(x) + \sum_{i=1}^N \frac{f(x)^2}{n} + \lambda \left(\sum_{i=1}^n \hat{\alpha}^T K \right)^2 \right\}\end{aligned}$$

= 0

$$\begin{aligned}f^*(x) &= \frac{-b \pm \sqrt{b^2 - 4ac}}{2a} \\ &= \frac{\frac{2}{n} \sum_{i=1}^N y_i \pm \sqrt{\left(\frac{2}{n} \sum_{i=1}^N y_i\right)^2 - 4(1/n) \left(\sum_{i=1}^N \frac{y_i^2}{n} + \lambda \left(\sum_{i=1}^n \hat{\alpha}^T K\right)^2\right)}}{2 \cdot (1/n)}\end{aligned}$$

$$= y_i \pm \sqrt{\lambda \left(\sum_{i=1}^n \hat{\alpha}^T K\right)^2 / n}$$

$$= \theta_0 + \theta_1 x \pm \frac{1}{\sqrt{n}} \sum \hat{\alpha}^T K(x, x)$$

"Again, an assumption about a linear term."

$$b) (\hat{\theta}, \alpha) = \underset{(\theta, \alpha) \in \mathbb{R}^2 \times \mathbb{R}^n}{\operatorname{argmin}} \left\{ \frac{1}{2n} \|y - X\theta - \sqrt{n} K \alpha\|_2^2 + \lambda_n \alpha^T K \alpha \right\}$$

Where $K \in \mathbb{R}^{n \times n}$ and $X \in \mathbb{R}^{n \times 2} = \begin{bmatrix} 1 & x_0 \\ 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix}$

$$f^*(\hat{\theta}, \alpha) = \theta_0 x_0 + \theta_1 x_1 + \sqrt{n} K \alpha$$

$$= X \theta - \sqrt{n} \sum_{i=1}^n \hat{\alpha}^T K(x, x) \quad \text{from part a.}$$

(Star Domain)

$$x, y \in \mathbb{R}^n$$

$$[x, y] = \{tx + (1-t)y : 0 \leq t \leq 1\}$$

$$= x - (y-x)[0, 1]$$

(Star Set) A subset C , $x \in C$ and $x^* \in C$ for every

$$[x^*, x] \subseteq C, \text{ has a } x^* + (x - x^*)\alpha \in C.$$

when $\alpha \in \mathbb{R}$

13.4.

$$a) [x^*, x] = \{ \alpha x + (1-\alpha)x^* : 0 \leq \alpha \leq 1 \}$$

"The star domain is an affine relationship or weighted-function"

b) The definition is above in star set.

$$13.5. \text{ (Equation 13.71)} \quad \frac{G_n(\delta; F)}{\delta} \leq \frac{\delta}{2\sigma}$$

$$\text{Where } G(\delta; F) = E_w \left[\sup_{\substack{g \in F \\ \|g\| \leq \sigma}} \left| \frac{1}{n} \sum_{i=1}^n w_i g(x_i) \right| \right]$$

$$a) \text{ If } \delta^2 = 4\sigma^2, \text{ then } \frac{G_n(\delta; F)}{\delta} \leq \frac{\delta^2}{2\sigma}$$

$$\leq \frac{2\sigma}{2\sigma}$$

$$\leq 1$$

$$b) \text{ If } \delta^2 \geq \min\left[1, \frac{9}{\pi} \frac{\sigma^2}{n}\right], \text{ then } \frac{G_n(\delta; F)}{\delta} \leq \delta/2\sigma$$

$$\leq \frac{\min \left[1, \sqrt{\frac{9}{n\pi}} \sigma \right]}{2\sigma}$$

$$\leq \min \left[\frac{1}{2}\sigma, \sqrt{\frac{2}{n\pi}} \right]$$

13.6.

$$a) G_n(\delta; F^*(1)) := \mathbb{P} \left[\left| \frac{1}{n} \sum_{i=1}^n w_i g(x_i) \right| \geq \delta \right]$$

$$= \mathbb{P} \left[\left| \sum_{i=1}^n w_i g(x_i) \right| \geq n\delta \right]$$

$$= \mathbb{P} \left[\lambda \left| \sum_{i=1}^n w_i g(x_i) \right| \geq n\lambda\delta \right]$$

$$= \mathbb{P} \left[e^{1 + \lambda \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x_i) \right| \right] + \frac{\lambda^2 \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x_i) \right|^2 \right]}{2}} \geq e^{n\lambda\delta} \right]$$

$$\leq \mathbb{P} \left[e^{\frac{\lambda^2 \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x) \right|^2 \right]}{2}} \geq e^{n\lambda\delta} \right]$$

$$\arg \min_{\lambda} \left\{ \frac{\lambda^2 \cdot \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x) \right|^2 \right]}{2} - n\lambda\delta \right\} = 0$$

$$\lambda^* = n\delta / \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x) \right|^2 \right]$$

$$= n\delta / \sigma^2$$

$$G_n(\delta; F^*(1)) \leq \mathbb{P} \left[e^{\frac{\lambda^2 \mathbb{E} \left[\left| \sum_{i=1}^n w_i g(x) \right|^2 \right]}{2}} \geq e^{n\lambda\delta} \right]$$

$$\leq \mathbb{P} \left[e^{\frac{n^2\delta^2}{2\sigma^2}} \geq e^{n^2\delta^2/\sigma^2} \right]$$

$$\leq \text{tr} \left(e^{\frac{n^2\delta^2}{2\sigma^2}} \right) \cdot e^{-n^2\delta^2/\sigma^2}$$

$$\leq \text{rank}(\sigma^2) \cdot e^{\frac{n^2\delta^2}{2\sigma^2} - n^2\delta^2/\sigma^2}$$

$$\leq n \cdot e^{-\frac{n^2\delta^2}{2\sigma^2}}$$

Note



- The trace and rank coefficient derive from Theorem 6.12/6.15.
- The derivation fits a maximal limit to the trace of the largest eigenvalue.

$$G(\sigma, F(1)) = (\sigma)$$

$$= \sqrt{\frac{2\sigma^2 \log n}{n}} \quad \dots \text{ for a } \mathbb{P}[X] = \text{Normal} \text{ where } = 1$$

$$= C_1 \sqrt{\frac{\log n}{n}}$$

$$\|\hat{\theta} - \theta^*\|_2^2 = \sigma$$

$$= C_1' \sigma \sqrt{\frac{\log n}{n}} \quad \text{where } C_1 = \sqrt{2}\sigma, C_1' = \sqrt{2}$$

$$b) \text{ From part a, } G(\sigma, F(1)) = \sqrt{\frac{2\sigma^2 \log n}{n}}$$

$$= \frac{C_2 \sigma \sqrt{\log n}}{n} \quad \text{when } C_2 = \sqrt{\frac{2}{n}} \sigma$$

$$\|\hat{\theta} - \theta^*\|_2^2 = \sqrt{\frac{2\sigma^2 \log n}{n}}$$

$$= \frac{C_2' \sigma^2 \log n}{n} \quad \text{when } C_2' = \sigma \sqrt{\frac{\log n}{n}}$$

$$13.7. \mathcal{P}_m = \{f_\theta: \mathbb{R}_n \rightarrow \mathbb{R}_n \mid \theta \in \mathbb{R}^m\} \quad \text{where } f_\theta(x) = \sum_{j=0}^{m-1} \theta_j x^j$$

$$\begin{aligned} \mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \frac{\sigma^2 m \log n}{n}] &= \mathbb{P}\left[\lambda \|\hat{F} - F^*\|_n^2 \geq c_0 \frac{\lambda \sigma^2 m \log n}{n}\right] \\ &:= \mathbb{P}\left[e^{1 + \lambda \mathbb{E}[\|\hat{F} - F^*\|_n^2] + \frac{\lambda^2 \mathbb{E}[\|\hat{F} - F^*\|_n^4]^2}{2} + \dots} \geq e^{c_0 \frac{\lambda \sigma^2 m \log n}{n}}\right] \\ &\leq \mathbb{P}\left[e^{\frac{\lambda^2 \mathbb{E}[\|\hat{F} - F^*\|_n^4]^2}{2}} \geq e^{c_0 \frac{\lambda \sigma^2 m \log n}{n}}\right] \end{aligned}$$

$$\arg \min_{\lambda} \left\{ \frac{\lambda^2 \cdot \mathbb{E}[(\hat{F} - F^*)^2]^2}{2} - c_0 \frac{\lambda \sigma^2 m \log n}{n} \right\} = 0$$

$$\lambda^* = \frac{c_0 \sigma^2 m \log n}{n \cdot \mathbb{E}[(\hat{F} - F^*)^2]^2}$$

$$\begin{aligned} \mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \frac{\sigma^2 m \log n}{n}] &\leq \mathbb{P}\left[e^{\frac{\lambda^2 \mathbb{E}[(\hat{F} - F^*)^2]^2}{2}} \geq e^{c_0 \frac{\lambda \sigma^2 m \log n}{n}}\right] \\ &\leq \mathbb{P}\left[e^{\frac{c_0^2 \sigma^4 m^2 \log^2 n \cdot \mathbb{E}[(\hat{F} - F^*)^2]^2}{n^2 \mathbb{E}[(\hat{F} - F^*)^4]}} \geq e^{\frac{c_0^2 \sigma^4 m^2 \log^3 n}{n^2 \mathbb{E}[(\hat{F} - F^*)^2]}}\right] \\ &\leq \text{tr} \left(e^{\frac{2c_2 m \log n}{n}} \right) \cdot e^{-c_2 m \log n} \\ &\quad - c_2 m \log n \\ &\leq m \cdot e^{-c_2 m \log n} \\ &\leq c_1 e^{-c_2 m \log n} \end{aligned}$$

Notes: Two approximations appear in the book's derivation, 1) Taylor Expansion
2) Trace/Eigenvalue upper limit.

Where $C_1 = m$, $C_2 = c_0 \left(\frac{\sigma}{n} \right)^2 m \log n$

$$13.8. \quad \mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \left(\frac{\sigma^2}{n} \right)^{4/5}]$$

The book offers other representations with standard deviation $[\sigma^2]$, Lipschitz constant $[L]$ and a random constant $[\delta]$. Past experiences involved a random constant or scalable-Lipschitz constant, rather than standard deviation, with manual and most probable fits. An example is an arbitrary intensity adjustment in an image where standard deviation is unknown without a reference in the image. A best guess becomes a decrease from maximum intensity toward a random intensity constant or slope-dependence constant, such as Lipschitz.

$$\mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \left(\frac{\sigma^2}{n} \right)^{4/5}] = \mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \left(\frac{\sigma_n}{n^{4/5}} \right)^{4/5}]$$

$$= \mathbb{P} \left[e^{\frac{1 + \lambda \mathbb{E}[\|F - F^*\|_n^2] + \frac{\lambda^2 \mathbb{E}[\|F - F^*\|_n^2]}{2} + \dots}{c_0 \lambda \left(\frac{\sigma^2}{n}\right)^{4/5}} \geq e \right]$$

$$\leq \mathbb{P} \left[e^{\frac{\lambda^2 \mathbb{E}[\|F - F^*\|_n^2]}{2}} \geq e^{c_0 \lambda \left(\frac{\sigma^2}{n}\right)^{4/5}} \right]$$

$$\underset{\lambda}{\operatorname{argmin}} \left\{ \frac{\lambda^2 \mathbb{E}[\|F - F^*\|_n^2]}{2} - c_0 \lambda \left(\frac{\sigma^2}{n}\right)^{4/5} \right\} = 0$$

$$\lambda^* = \frac{c_0 \left(\frac{\sigma^2}{n}\right)^{4/5}}{\mathbb{E}[\|F - F^*\|_n^2]}$$

$$\mathbb{P}[\|\hat{F} - F^*\|_n^2 \geq c_0 \left(\frac{\sigma^2}{n}\right)^{4/5}] \leq \mathbb{P} \left[e^{\frac{\lambda^2 \mathbb{E}[\|\hat{F} - F^*\|_n^2]}{2}} \geq e^{c_0 \lambda \left(\frac{\sigma^2}{n}\right)^{4/5}} \right]$$

$$\leq \operatorname{tr} \left(e^{\frac{c_0^2 \left(\frac{\sigma^2}{n}\right)^{2/5} \mathbb{E}[\|\hat{F} - F^*\|_n^4]}{2 \mathbb{E}[\|\hat{F} - F^*\|_n^2]^4}} \right) \cdot e^{\frac{-c_0^2 \left(\frac{\sigma^2}{n}\right)^{8/5}}{\mathbb{E}[\|F - F^*\|_n^2]^2}}$$

$$\leq n \cdot e^{-c_0^2 (n/\sigma^2)^{1/5}}$$

$$\leq n \cdot e^{-C_2 (n/\sigma^2)^{1/5}}$$

$$\leq C_1 e^{-C_2 (n/\sigma^2)^{1/5}}$$

$$\text{where } C_1 = n, C_2 = c_0^2$$

13.9:

$$a) \underbrace{\frac{\sigma}{n} \left| \sum_{i=1}^n w_i \hat{\Delta}(x_i) \right|}_{\text{Optimal Mean}} \leq \underbrace{\|\Delta\|_2^2}_{\text{Feasible Mean}} + \underbrace{\|\Delta\|_2^2}_{\text{Error}}$$

Optimal Mean = Feasible Mean + Error
Squared error

The book suggests: $(\|\Delta\|_2^2 \leq \epsilon \delta) + (\|\Delta\|_2^2 \geq \epsilon \delta)$

I scoured the book for a "t" variable.
 With no hints, t represents a threshold.
 The value of t is unknown and arbitrary. I guess t-table and Student's t-distribution for a threshold-cutoff. Cases involve size (d) and degrees of freedom Knowledge prior to variance.

Proof by Exhaustion:

$$\text{Case \#1: } \|\Delta\|_2^2 \leq t \cdot \sigma \quad , \quad \frac{\sigma}{n} \left\| \sum_{i=1}^n w_i g(x) \right\| \leq d \cdot t \cdot \sigma$$

$$\text{Case \#2: } \|\Delta\|_2^2 \geq t \cdot \sigma \quad , \quad \frac{\sigma}{n} \left\| \sum_{i=1}^n w_i g(x) \right\| \leq 2 \cdot \sqrt{t \cdot \sigma} \cdot \|\Delta\| \leq 2 \sqrt{t \cdot \sigma} \|\Delta\|$$

$$\frac{\sigma}{n} \left\| \sum_{i=1}^n w_i g(x) \right\| \leq d \cdot t \cdot \sigma + 2 \sqrt{t \cdot \sigma} \|\Delta\|$$

$$b) \text{ Suppose } \sqrt{\sum_{j=1}^d \|g_j\|_n^2} \leq \sqrt{K} \left\| \sum_{j=1}^d g_j \right\|_n$$

$$\begin{aligned} \|\hat{f} - f^*\|_n^2 &= \left(\sqrt{K} \left\| \sum_{j=1}^d g_j \right\|_n \right)^2 \\ &\leq \text{tr} \left(K \cdot \left\| \sum_{j=1}^d g_j \right\|_n^2 \right) \\ &\leq d \cdot K \cdot \sigma_{n, \max}^2 \end{aligned}$$

$$13.10 \quad a) \hat{f} = \min_{\theta \in \mathbb{R}^T} \left\{ \frac{1}{n} \|y - \Phi \theta\|_2^2 + \lambda_n \|\theta\|_2^2 \right\}$$

$$= \frac{2}{n} \Phi(y - \Phi \theta) + 2\lambda \theta$$

$$= 0$$

$$\theta^* = \frac{y \Phi}{n(\Phi^2 + \lambda I)}$$

$$b) \inf_{f \in F(1:T)} \|f - f^*\|_2^2 = \inf \left\| \sum_{i=1}^T (\theta_i - \theta_i^*) \phi_i \right\|_2^2 + \inf \left\| \sum_{i=T+1}^{\infty} \theta_i^* \phi_i \right\|_2^2$$

$$= \|\theta_{1:T}^* - \theta_{1:T}\|_2^2 + \|\theta_m^*\|_2^2$$

$$= \|\theta_m^*\|_2^2$$

13.11.

a) Function class:

$$F(1:T) = \left\{ \beta_0 + \sum_{i=1}^K (\beta_m \phi_m + \bar{\beta}_m \phi_m) \mid \beta_0 + \sum_{i=1}^K (\beta_m \phi_m + \bar{\beta}_m \phi_m) \leq 1 \right\}$$

From the problem,

$$\int_0^1 f^{(x)}(mx)^2 dx \leq R$$

$$m^{2x} \int_0^1 f^{(x)}(x)^2 dx \leq R$$

$$\int_0^1 f^{(x)}(x)^2 dx \leq \frac{R}{m^{2x}}$$

$$\sum (\beta_m^2 \phi_m + \bar{\beta}_m^2 \phi_m) \leq \frac{R}{m^{2x}}$$

"Second term in the kernel problem from Chapter 12"

$$\beta_m^2 + \bar{\beta}_m^2 \leq \frac{c R}{m^{2\kappa}}$$

Where $c = \text{constant}$

b) (Equation 13.47) $\|F - F^*\| = \beta_m^2 + \bar{\beta}_m^2$

$$\leq \frac{c R}{m^{2\kappa}}$$

Where $c = \text{arbitrary constant within a bound.}$