

## Chapter 4: Uniform laws of large numbers:

### (Glivenko - Cantelli Theorem)

The empirical cumulative distribution function  $\hat{F}$  estimates a population's cumulative distribution function in a uniform norm:

$$\|\hat{F}_n - F\|_{\infty} \xrightarrow{a.s.} 0$$

4.1.

a)  $\hat{F}_n = \text{C.D.F.}$  By the "Fundamental Theorem of Calculus" also known as the Glivenko - Cantelli

theorem:  $\|\hat{F}_n - F\|_{\infty} \xrightarrow{a.s.} 0$

$$\delta(\|\hat{F}_n - F\|_{\infty}) \xrightarrow{a.s.} 0$$

$$\delta(\hat{F}_n) \rightarrow \delta(F)$$

b) i) Mean Functional:  $F \mapsto \int x dF(x)$

$$\hat{F}_n = \frac{1}{n} \sum_{i=1}^n x_i = \frac{(n+1)}{2} ; F = \frac{1}{n} \int_1^n x dx = \frac{1}{n} \left[ \frac{n^2-1}{2} \right]$$

A continuous function  $\hat{F}_n - F \xrightarrow{a.s.} 0$ , so the mean functional implies discontinuity. (Plot A)

ii) Cramer-von Mises Functional:  $F \mapsto \int [F(x) - F_0(x)]^2 dF_0(x)$

$$\hat{F}_n = \frac{1}{n} \sum_{i=1}^n (F(x_i) - F_0(x_i))^2 = \frac{1}{n} \sum_{i=1}^n F(x_i)^2 - \frac{1}{n} \left[ \sum_{i=1}^n F_0(x_i) \right]^2$$

$$F = \frac{1}{n} \int_1^n (F(x_i) - F_0(x))^2 dx = \frac{1}{n} \int_1^n F(x)^2 dx - \frac{1}{n} \left[ \int_1^n F_0(x) dx \right]^2$$

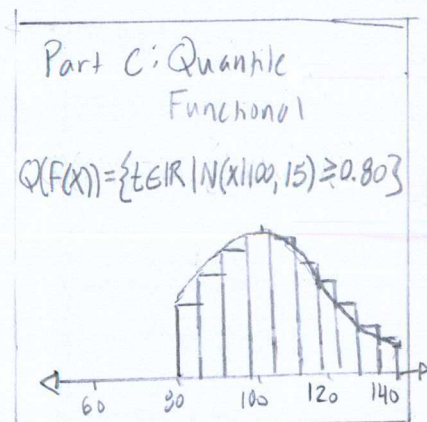
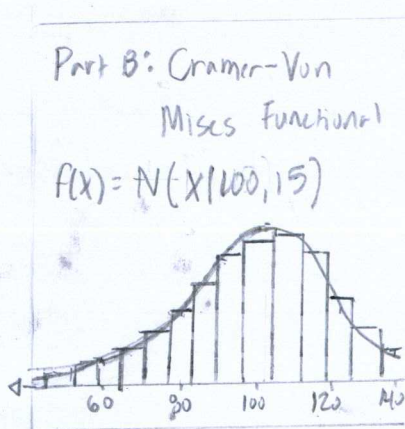
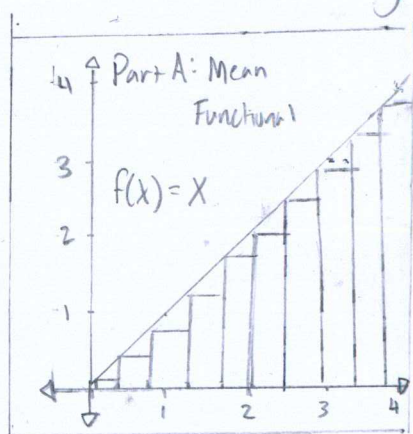
Cramer-von Mises asymptotically approaches zero as  $n$  rises.

Cramer (1928), (Plot B)

iii) The Quantile Functional:  $Q_n(F) = \inf \{ t \in \mathbb{R} \mid F(t) \geq \alpha \}$

$$F(t) = F(Q_n(F) + \epsilon) > \alpha \text{ and } F(t) = F(Q_n(F) - \epsilon) < \alpha$$

The quantile functional is continuous from the left and the right. (Plot c)



(Theorem 4.10: Rademacher Lower and Upper Bounds)

$$\|P_n - P\| \leq 2 \cdot R_n(F) + o(1) \text{ unless } R_n(F) = o(1), \text{ then } \|P_n - P\| \xrightarrow{a.s.} 0$$

$$\text{Where } F = \{1 \mid S \in S\}$$

$$4.2 \quad R_n(s) = \mathbb{E} \left[ \sup_{s \in S} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i [X_i] \right| \right] \geq 1/2 \text{ at a lower bound}$$

$$2R_n(s) + o(1) \geq 1 \text{ so the function does not asymptotically converge.}$$

$R_n(s)$  is one at 100% accumulative probability or 100% convergent probability in a distribution.

$$4.3a)i) \text{ Bernoulli: } p_\theta(x) = \frac{e^{\theta x}}{1 + e^{\theta x}} \text{ for } x \in \{0, 1\}$$

$$\begin{aligned} R(\theta, \hat{\theta}^*) &= \mathbb{E} \left[ \log \frac{p_{\hat{\theta}^*}(x)}{p_\theta(x)} \right] = \sum_0^1 \theta^* \cdot \log \frac{p_{\hat{\theta}^*}(x)}{p_\theta(x)} \\ &= \sum_0^1 \theta^* \log \left( \frac{e^{\theta^* x}}{1 + e^{\theta^* x}} \cdot \frac{1 + e^{\theta x}}{e^{\theta x}} \right) \\ &= \log \left( \frac{e^{x(1-\theta)} + e^x}{e^x + 1} \right) \end{aligned}$$

Note: A possible typo exists for  $a(i)$  in the denominator as by Bernoulli

$$ii) \text{ Poisson: } p_\theta(x) = \frac{e^{\theta x} e^{-\exp(\theta)}}{x!} \text{ for } x \in \{0, 1, 2, \dots\}$$



$$\begin{aligned}
 R(\theta, \theta^*) &= \mathbb{E}_{\theta^*} \left[ \log \frac{P_{\theta^*}(x)}{P_{\theta}(x)} \right] \\
 &= \sum_{\theta=0}^{\infty} \theta^* \log \left( \frac{e^{\theta^* x} e^{-\exp(\theta^*)}}{x!} \cdot \frac{x!}{e^{\theta x} e^{-\exp(\theta)}} \right) \\
 &= \sum_{\theta=0}^{\infty} \theta^* [x(\theta^* - \theta) - e^{\theta^*} + e^{\theta}] \\
 &= \text{Divergent risk!!!}
 \end{aligned}$$

ii) Multi-variate Gaussian:  $P_{\theta}(x) = N(\theta, \Sigma)$

$$\begin{aligned}
 R(\theta, \theta^*) &= \mathbb{E}_{\theta^*} \left[ \log \frac{P_{\theta^*}(x)}{P_{\theta}(x)} \right] \\
 &= \sum_{\theta=-\infty}^{\infty} \theta^* \left[ \log \frac{N(\theta^*, \Sigma)}{N(\theta, \Sigma)} \right] \\
 &= \sum_{\theta=-\infty}^{\infty} \theta^* \left[ -\frac{(x - \theta^*)^2}{2\Sigma^{-1}\Sigma} + \frac{(x - \theta)^2}{2\Sigma^{-1}\Sigma} \right] \\
 &= \text{Divergent risk!!!}
 \end{aligned}$$

The expectation wants an integral or summation about  $\theta^*$  and not  $x$ .

b) i)  $\mathbb{E}(\hat{\theta}, \theta^*) = R(\hat{\theta}, \theta^*) - \inf_{\theta \in \Omega} R(\theta, \theta^*)$

Bernoulli:  $R(\arg\min_{\theta} \{R(\theta, \theta^*)\}, \theta^*) - \inf_{\theta \in \Omega} R(\theta, \theta^*)$

Where  $R(\theta, \theta^*) = \log \left( \frac{e^{x(1-\theta)} + e^x}{e^x + 1} \right)$

Poisson:  $R(\arg\min_{\theta} \{R(\theta, \theta^*)\}, \theta^*) - \inf_{\theta \in \Omega} R(\theta, \theta^*)$

Where  $R(\theta, \theta^*) = \sum_{\theta=0}^{\infty} \theta^* [x(\theta^* - \theta) - e^{\theta^*} + e^{\theta}]$

Multivariate:  $R(\arg\min_{\theta} \{R(\theta, \theta^*)\}, \theta^*) - \inf_{\theta \in \Omega} R(\theta, \theta^*)$

Gaussian

Where  $R(\theta, \theta^*) = \sum_{\theta=-\infty}^{\infty} \theta^* \left[ -\frac{(x - \theta^*)^2}{2\Sigma^{-1}\Sigma} + \frac{(x - \theta)^2}{2\Sigma^{-1}\Sigma} \right]$

The solutions describe cases when  $\theta^* \in \Omega_0$  and not a subset of  $\Omega$ .

ii) The upper limit on excess risk is not a minimal likelihood estimate. A Rademacher complexity contains no minimizing, and in risk decomposition, not the minimal (infimum), nor expectation, but likely a supremum.

4.4

a) Equation 4.17:  $\mathbb{E}[\|P_n - P\|] \leq \mathbb{E}_{X,Y} \left[ \sup \left| \frac{1}{n} \sum_{i=1}^n f(X_i) - f(Y_i) \right| \right]$   
 $\sup \mathbb{E}[g(X)] \leq \mathbb{E}[\sup |g(X)|]$ , when the supremum and expectation do not commute. A real world case depends on acceptable error or experimental error, such that  $\sup \mathbb{E}[g(X)] - \mathbb{E}[\sup |g(X)|] \leq \text{Error}$ .

Jensen's proof about inequalities also justifies the statement when  $g(X)$ , a convex function.

b) If  $\phi$  is convex and non-decreasing, then Jensen's inequality,  $\phi(\mathbb{E}[X]) \leq \mathbb{E}[\phi(X)]$  relates.

$\sup_{g \in G} \phi(\mathbb{E}[|g(X)|]) \leq \mathbb{E}[\phi(\sup_{g \in G} |g(X)|)]$  because the inequality generalizes a secant above: a concave function as reasonable argument for "less than." Proposition 4.11 applies Jensen's theory by the assumption, "Expected law of large numbers value is a minimum with upper and lower bounds about the local convexity."



4.5. (Proposition 4.12)

$$\|P_n - P\| \geq \frac{1}{2} R_n(F) - \frac{\sup |E[F]|}{2\sqrt{n}} - \delta \quad \text{with } IP \text{ at least } 1 - e^{-\frac{n\delta^2}{2b^2}}$$

a) A recentered function class:  $\bar{F} = \{F - E[F] \mid F \in \mathcal{F}\}$

$$\|S_n\| = \sup \left| \frac{1}{n} \sum \epsilon_i F(x) - \frac{1}{n} \sum \epsilon_i E[F(x)] \right| \geq \sup \left| \frac{1}{n} \sum \epsilon_i F(x) \right| - \left| \sum \epsilon_i \right| \frac{\sup |E[F(x)]|}{n}$$

$$E[\|S_n\|] \geq E[\|S_n\|_F] - E\left[\left|\sum \epsilon_i\right|\right] \frac{\sup |E[F(x)]|}{n}$$

With Chebyshev's inequality,

$$E[\|S_n\|] \geq E[\|S_n\|_F] - \sqrt{n} \frac{\sup |E[F(x)]|}{\sqrt{n}}$$

$$\geq E[\|S_n\|_F] - \frac{\sup |E[F(x)]|}{\sqrt{n}}$$

b) Equation 4.21:  $\frac{1}{2} E[\|S_n\|] \leq E[\|P_n - P\|] \leq 2 E[\|S_n\|]_F$

$$\frac{1}{2} \left( R(F) - \frac{\sup |E[F(x)]|}{\sqrt{n}} \right) \leq \frac{1}{2} E[\|S_n\|] \leq E[\|P_n - P\|]$$

$$\frac{1}{2} \left( R(F) - \frac{\sup |E[F(x)]|}{2\sqrt{n}} \right) \leq \|P_n - P\|$$

4.6,  $\mathcal{F} := [x \mapsto \text{sign}(\langle \theta, x \rangle) \mid \theta \in \mathbb{R}^n, \|\theta\|_2 = 1]$

$$R(F(x_i^n) | n) = E\left[\sup \left| \frac{1}{n} \sum \epsilon_i \text{sign}(\epsilon_i) \right| \right] = 1 \quad \text{because } n$$

Rademacher complex above  $1/2$  is fit for empirical risk minimization, an overfit at 1.0

(Converse Inequality)

$$a < b \iff b > a$$

$$\text{and } a \leq b \iff b \geq a \quad \text{for } a, b \in \mathbb{R}$$

(Triangle Inequality)

$$\|x + y\| = \|x\| + \|y\| \quad \text{for } a, b \in \mathbb{R}$$

4.7a) If  $(F \text{ and } \bar{F}) \in \text{conv}(F)$ , then

$$\left| \frac{\sum \epsilon_i \bar{F}(x_i)}{n} \right| \leq \left| \frac{\sum \epsilon_i F(x_i)}{n} \right| \leq \sup \left| \frac{\sum \epsilon_i F(x_i)}{n} \right|$$

$$R_n \left[ \frac{\sum \epsilon_i \bar{F}(x_i)}{n} \right] \leq R_n \left[ \frac{\sum \epsilon_i F(x_i)}{n} \right] \leq \sup R_n \left[ \frac{\sum \epsilon_i f(x_i)}{n} \right]$$

$$R_n[\bar{F}] \leq R_n[F] \leq R_n[\text{conv.}(F)]$$

$$b) \|F+G\| \leq \|F\| + \|G\|$$

$$R[\|F+G\|] \leq R[\|F\|] + R[\|G\|]$$

$$c) \left\| \frac{1}{n} \sum \epsilon_i F(x) + \frac{1}{n} \sum \epsilon_i g(x) \right\| \leq \left\| \frac{1}{n} \sum \epsilon_i F(x) \right\| + \left\| \frac{1}{n} \sum \epsilon_i g(x) \right\|$$

$$R[\|F+g\|] \leq R[\|F\|] + \left\| \sum \epsilon_i \right\| \frac{g(x)}{n}$$

$$\leq R[\|F\|] + \frac{g(x)}{\sqrt{n}} \quad \text{by Cauchy-Schwartz inequality}$$

(Vapnik-Chervonenski Dimension)

"VC": a measure about the maximum points on algorithm handles without "shatter."

A sets sizes growth by  $2^n$  "shatters" an algorithm.

$$4.8a) S^c = \{S^c | S^c \in S\}$$

$$TUS^c = \{T: S^c | T: S^c \in T \text{ or } T: S^c \in S\}$$

$$TUS^c = U \quad \text{and} \quad VC(TUS) = VC(U)$$

$$b) SNT = \{SNT | S \in S, T \in T\}$$

$$SNT = U \quad \text{and} \quad VC(SNT) = VC(U)$$

$$\text{card}(SNT(x_i^n)) \leq (n+1)^{VC(SNT)}$$

$$c) SUT = \{SUT | S \in S, T \in T\}$$

$$SUT = U \quad \text{and} \quad VC(SUT) = VC(U)$$

$$\text{card}(SUT(x_i^n)) \leq (n+1)^{VC(SUT)}$$

Notes: A set growth by  $2^n$  is positively above a set growth at  $n, n^2, n^3$ , or  $n^4$



$$4.9. R(F(x_i^n)|n) = \mathbb{E}[\sup |\frac{1}{n} \sum \epsilon_i f(x_i)|] \leq 4 \mathbb{E}[\sup |\frac{1}{n} \sum \epsilon_i f(x_i)|] \log(n+1)^V$$

by the Cauchy-Schwarz

$$R(F(x_i^n)|n) \leq 4 \sup \sqrt{\frac{\sum f^2(x)}{n}} \sqrt{\frac{V \log(n+1)}{n}}$$

where  $V$  is VC's number.

$$4.10. \binom{n-1}{k} + \binom{n-1}{k-1} = \frac{(n-1)!}{k!(n-1-k)!} + \frac{(n-1)!}{(k-1)!(n-1-k+1)!}$$

$$= \frac{(n-1)!}{k!(n-1-k)!} \frac{(n-1-k+1)}{(n-1-k+1)} + \frac{k}{k} \frac{(n-1)!}{(k-1)!(n-1-k+1)!}$$

$$= \frac{(n-1)!(n-k) + k(n-1)!}{k(n-k)!}$$

$$= \frac{n!}{k(n-k)!}$$

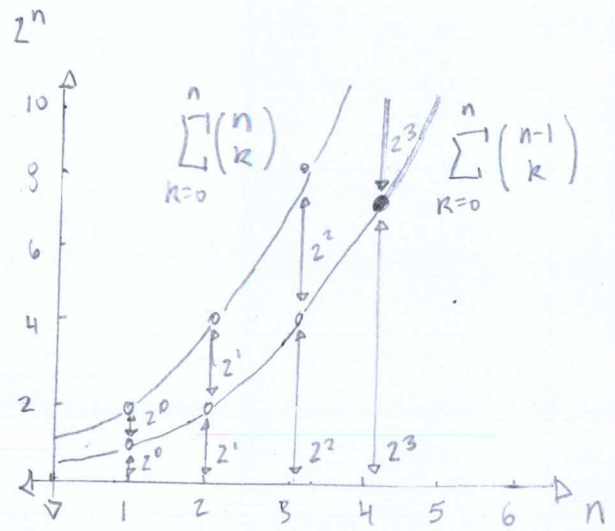
$$= \binom{n}{k}$$

$$\text{Card}(S(p)) \leq \sum_{k=0}^n \binom{n-1}{k} + \sum_{k=0}^n \binom{n-1}{k-1}$$

$$\leq \sum_{k=0}^n \binom{n-1}{k} + \sum_{k=0}^n \binom{n-1}{k-1}$$

$$\leq 2^{n-1} + 2^{n-1}$$

$$\leq 2^n = 2 \binom{n}{k}$$



(Vapnik-Chervonenkis, Sauer and Shelah)

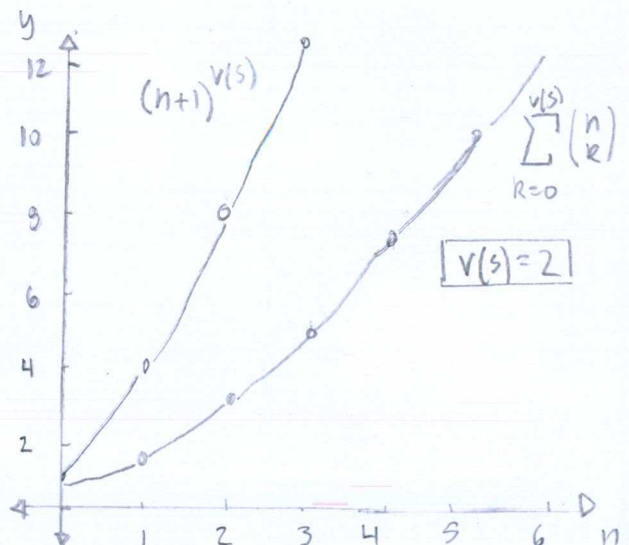
$$\text{Card}(S(p)) \leq \sum_{k=0}^{v(s)} \binom{n}{k} \leq (n+1)^{v(s)}, \quad n \geq v(s)$$

4.11

a) A plot visible toward

the right shows

$$\sum_{k=0}^{v(s)} \binom{n}{k} \leq (n+1)^{v(s)} \quad \text{where } n \geq v(s)$$



b) When  $n \geq v$ , prove:  $\text{card}(S(x_1^n)) \leq \left(\frac{en}{v}\right)^v$

$$e^v = \left(1 + \frac{x}{n}\right)^n \geq \sum_{k=0}^v \binom{n}{k} \left(\frac{v}{n}\right)^k$$

$$\left(\frac{n}{v}\right)^v e^v \geq \sum_{k=0}^v \binom{n}{k} = \text{card}(S(x_1^n))$$

4.12.  $S_{\text{left}}^d := \{(-\infty, t_1] \times (-\infty, t_2] \times \dots \times (-\infty, t_d) \mid (t_1, \dots, t_d) \in \mathbb{R}^d\}$

Each subset is countable, such as

$$S_{\text{left}}^1 := \{(-\infty, t_1] \mid t_1 \in \mathbb{R}\}, \text{ so } S_{\text{left}}^d \text{ and}$$

$$\text{VC}(S_{\text{left}}^d) = d.$$

4.13.

a)  $S_{\text{sphere}}^2 := \{S_{a,b} \mid (a,b) \in \mathbb{R}^2 \times \mathbb{R}_+\}$

$$\text{where } S_{a,b} := \{x \in \mathbb{R}^2 \mid \|x-a\|_2 \leq b\}$$

$$\text{and } b \geq 0 \text{ at } a = (a_1, a_2)$$

$$S^2 = S \cdot S = \|x-a\| \|x-a\| = \|x-a\|^2 \approx (n+1)^d$$

Although a real sphere with three dimensions shatters growth in the

$$\text{VC-measurement. } (x-a)^d = (n+1)^d$$

$$(x-a)^{d+1} \geq (n+1)^d$$

b) (D. Fitzpatrick, A. Iosevich, B. McDonald, E. Wyman, 2021)

Ultimately, a unique bound in shatter limit analysis.

The set size  $\{(n+1)^d\}$  in an algorithm remains

with solutions and unbreakable limits when  $d \leq 2$ .

A set of points growing at  $d=3$ , "shatters" the

algorithms... but the authors propose  $d \leq 2$  and  $d \geq 4$

as collapsible problem sets.

$$e^x = \lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n$$

Proof: Taylor Expansion

$$(1+x)^n = 1 + nx + \frac{n(n-1)}{2!}x^2 + \dots$$

$$\left(1 + \frac{x}{n}\right)^n = 1 + n\left(\frac{x}{n}\right) + \frac{n(n-1)}{2!}\left(\frac{x}{n}\right)^2 + \dots$$

$$e^x = 1 + nx + \frac{n(n-1)}{2!}x^2 + \dots$$

$$e^x = \left(1 + \frac{x}{n}\right)^n$$



4.14  $d \geq 2: h_s: \{0,1\}^d \rightarrow \{0,1\}$  of the form  $h_s(x_1, \dots, x_d) = \begin{cases} 1 & \text{if } x_j = 1 \text{ for all } j \\ 0 & \text{otherwise} \end{cases}$

The problem has a lower and upper bound.

A lower bound is a set size of 0, because  $\{0,1\}^d \rightarrow \{0,1\}$ ,  
While an upper at  $VC(h_s(x_1, \dots, x_d)) = d$  where each  
element a one.

(Convex Set) of all close

If a set joins subsets by a single intersection or line segment.

(Closed set)

A set whose complement is an open set; a set containing all  
limit points.

4.15:  $C_{cc}^d$  of all closed and convex sets in  $\mathbb{R}^n$ .

The set cardinality,  $\text{card}(C_{cc}^d)$  is not possible with  
the set of closed and convex sets. A convex set  
joined by intersection leaves out points when  
also closed set. The p...

4.16 The VC-dimension of a set of all polygons in  $\mathbb{R}^2$   
is  $2n+1$ , for a quadrilateral on a circle,  $2(4)+1=9$ .

4.17:  $f_t(x) = \text{sign}(\sin(tx))$ ;  $t \in \mathbb{R}$ ;  $f_t: [-1,1] \rightarrow \mathbb{R}$  [ $t \in \mathbb{R}$ ]

If  $x = 2^{-n}$ , then a binary expansion about  
 $\sin(tx)$  is infinite through "Cantor's diagonal  
argument," so  $VC\text{-measure}(f_t(x)) = \infty$ .