

SOBOL' INDICES, SHAPLEY EFFECTS AND A (NEW) PATH TOWARDS HANDLING DEPENDENT INPUTS

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Sensitivity Analysis Discord Group

Online

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☞ **Before:** Ph.D. Candidate (2021-2024)

EDF R&D - Institut de Mathématiques de Toulouse

Nicolas Bousquet, Fabrice Gamboa, Bertrand Iooss, Jean-Michel Loubes

Interpretability methods for certifying machine learning models applied to critical systems

☞ **Now:** Postdoctoral Researcher (2024-2026+)

UQÀM - IID (ULaval)

Arthur Charpentier (UQÀM), Marie-Pier Côté (ULaval)

Interpretability, fairness and causal inference of black-box models

(Some) Topics of interest:

XAI • Uncertainty quantification • Sensitivity analysis • Industrial risks • Probabilistic modelling

• Statistics • Statistical Learning • Applied cooperative game theory • Functional analysis

Introduction

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How to go beyond mutually independent inputs

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Quick outline:

- ☞ From Hoeffding's decomposition to Sobol' indices
- ☞ Shapley effects and the "cooperative games" solution
- ☞ Generalized Hoeffding decomposition: disentangling interaction and dependence

Hoeffding's decomposition

- ☞ Let $D = \{1, \dots, d\}$ and let \mathcal{P}_D denote the **power-set** of D
- ☞ Let $X = (X_1, \dots, X_d)$ be a random vector of **mutually independent inputs**
- ☞ For every $A \in \mathcal{P}_D$, $A \neq \emptyset$, let X_A be a **subset of the inputs**
- ☞ Let G be a **given “black-box” model** such that $\mathbb{V}(G(X)) < \infty$

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Hoeffding (1948): We can **uniquely** write

$$G(X) = \sum_{A \in \mathcal{P}_D} G_A(X_A),$$

where G_\emptyset is a constant, and the **representants** are all **pairwise orthogonal**, i.e.,

$$\forall A, B \in \mathcal{P}_D, A \neq B, \quad \mathbb{E}[G_A(X_A)G_B(X_B)] = 0$$

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Moreover, we can characterize

$$G_A(X_A) = \sum_{B \in \mathcal{P}_A} (-1)^{|A|-|B|} \mathbb{E}[G(X) | X_B], \quad \forall A \in \mathcal{P}_D$$

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For example:

$$G(X) = X_1 + X_2 X_3, \quad X = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix} \right)$$

In this case, we have that

$$\begin{aligned} G_1(X_1) &= X_1 & G_2(X_2) &= 0, & G_3(X_3) &= 0, \\ G_{12}(X_{12}) &= 0, & G_{13}(X_{13}) &= 0, & G_{23}(X_{23}) &= X_2 X_3, \\ G_{123}(X_{123}) &= 0 \end{aligned}$$

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We can retrieve the full model by only having access to the **representants**

Sobol' indices

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An input is important \iff It **contributes** to the model's **uncertainty**

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Sobol (2001) proposed to use

$$S_A = \frac{\mathbb{V}(G_A(X_A))}{\mathbb{V}(G(X))} = \frac{1}{\mathbb{V}(G(X))} \times \sum_{B \in \mathcal{P}_A} (-1)^{|A|-|B|} \mathbb{V}(\mathbb{E}[G(X) | X_B])$$

based on the rationale of **Hoeffding's decomposition**

$$\mathbb{V}(G(X)) = \mathbb{V}\left(\sum_{A \in \mathcal{P}_D} G_A(X_A)\right) \stackrel{\perp}{=} \sum_{A \in \mathcal{P}_D} \mathbb{V}(G_A(X_A))$$

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The model's uncertainty is equal to the sum of its representant's uncertainty

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If we **keep the same formula** : $S_A = \frac{1}{V(G(X))} \times \sum_{B \in \mathcal{P}_A} (-1)^{|A|-|B|} V(\mathbb{E}[G(X) | X_B])$

$$G(X) = X_1 + X_2 X_3, \quad X = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 & \rho \\ 0 & 1 & 0 \\ \rho & 0 & 1 \end{pmatrix}\right)$$

Mutually independent case ($\rho = 0$)

$$\begin{aligned} S_1 &= 0.5 & S_2 &= 0, & S_3 &= 0, \\ S_{12} &= 0, & S_{13} &= 0, & S_{23} &= 0.5, \\ S_{123} &= 0 \end{aligned}$$

Correlated case ($\rho \neq 0$)

$$\begin{aligned} S_1 &= 0.5 & S_2 &= 0, & S_3 &= \rho^2/2, \\ S_{12} &= \rho^2/2, & S_{13} &= -\rho^2/2, & S_{23} &= 0.5, \\ S_{123} &= -\rho^2/2 \end{aligned}$$

☞ They still **sum to 1** but they can be **negative**

But their overall interpretation as (pure) interaction effects does not hold anymore...

Sobol' indices computed on dependent inputs

There are (mainly) 2 problems:

1. **Negative effects** (dependence)
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But, there are solutions!

1. **Cooperative games' allocations** (e.g., Shapley values)
2. **Beyond Hoeffding's original decomposition**

Shapley effects

The Shapley effects (Owen 2014): **Egalitarian redistribution of the Sobol' indices.**

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The resulting indices can be written as

$$\forall i \in D, \quad \textcolor{orange}{Sh}_i = \sum_{A \in \mathcal{P}_D : i \in A} \frac{S_A}{|A|}$$

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But let's take a step back...

Cooperative game theory

The **Shapley effects** are derived from the **Shapley (1951) values**

Stems from **cooperative game theory**

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Let $D = \{1, \dots, d\}$ be a **set of players**, and \mathcal{P}_D the **set of coalitions**

Let $v : \mathcal{P}_D \rightarrow \mathbb{R}$ be a **chosen value function**

☞ (D, v) formally defines a **cooperative game**

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How to redistribute $v(D)$ to each of the d players?

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Hint: by using an **allocation**

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How can we define efficient and nonnegative allocations?

The Harsanyi set

The **Harsanyi (1963) dividends** of a cooperative game (D, v) are defined as:

$$D_v(A) = \sum_{B \in \mathcal{P}_A} (-1)^{|A|-|B|} v(B)$$

It is a mapping $D_v(A) : \mathcal{P}_D \rightarrow \mathbb{R}$

- ☞ They can be interpreted as the **added value** produced by each **coalition**
- ☞ They **always** sum-up to $v(D)$:

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The **Harsanyi set of allocations** (Vasil'ev and Laan 2001) are **aggregations of the Harsanyi dividends**:

$$\psi(i) = \sum_{A \in \mathcal{P}_D : i \in A} \lambda_i(A) \mathcal{D}_v(A), \quad \text{where} \quad \begin{cases} \forall i \in D, \forall A \in \mathcal{P}_D, \lambda_i(A) \geq 0, \\ \forall A \in \mathcal{P}_D, \sum_{i \in D} \lambda_i(A) = 1 \end{cases}$$

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- They are always efficient

- Nonnegative if v is monotonic

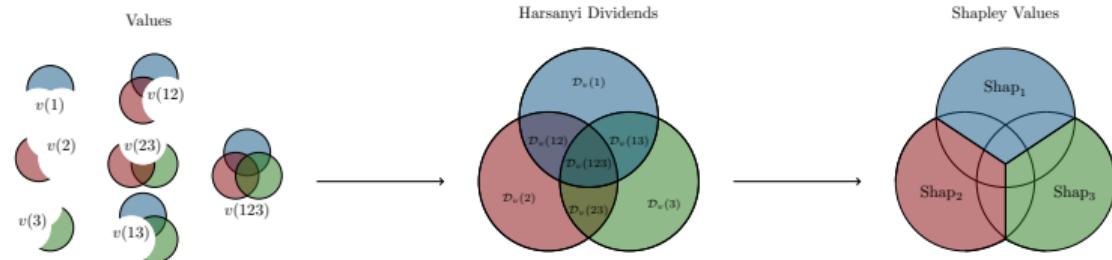
It is the case for $\mathbb{V}(G(X) | X_A)$

Egalitarian redistribution: the Shapley values

The **Shapley (1951) values** are the **egalitarian redistribution of the dividends**.

For a player $i \in D$,

$$\text{Shap}_i = \sum_{A \in \mathcal{P}_D : i \in A} \frac{\mathcal{D}_v(A)}{|A|}.$$

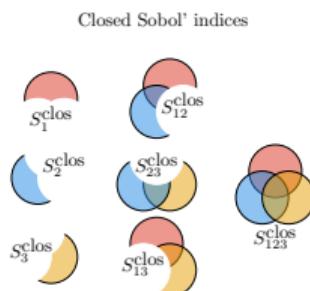
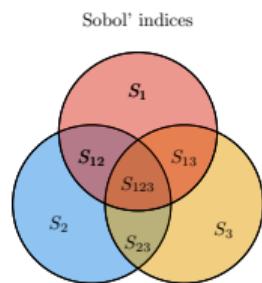
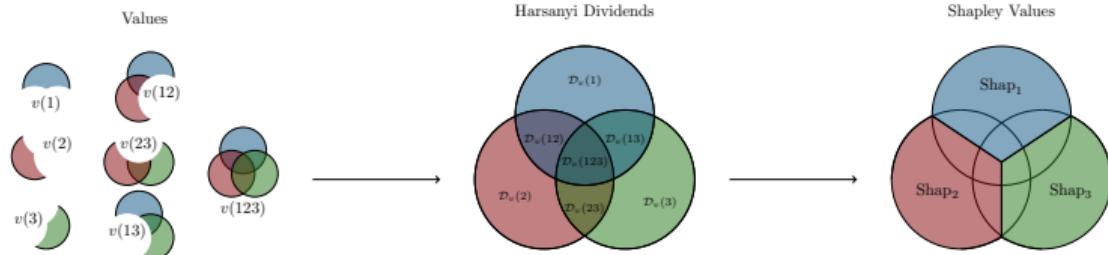


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Shapley effects (Owen 2014):

Shapley values with $v(A) = \mathbb{V}(\mathbb{E}[G(X) | X_A]) = S_A^{\text{clos}}$

$$\text{Sh}_i = \sum_{A \in \mathcal{P}_D : i \in A} \frac{S_A}{|A|}$$

The **Harsanyi dividends** become S_A : the **Sobol' indices**

Exogeneity detection and Shapley's joke

However, the **Shapley effects** have a **practical drawback**

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An **exogenous input** can have a
non-zero share of importance.

We call it **Shapley's joke**

(looss and Prieur 2019)

$$G(X) = X_1 + X_2, \quad X = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 & \rho \\ 0 & 1 & 0 \\ \rho & 0 & 1 \end{pmatrix} \right)$$

$$\text{Sh}_1 = 0.5 - \rho^2/4, \quad \text{Sh}_2 = 0.5, \quad \underline{\text{Sh}_3 = \rho^2/4}$$

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To solve this issue, we proposed to use **a proportional redistribution of the dividends**
(Ortmann 2000)

It led to the definition of the **proportional marginal effects (PME)** (Herin et al. 2023)

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$$\text{Sh}_1 = 0.5 - \rho^2/4, \quad \text{Sh}_2 = 0.5, \quad \underline{\text{Sh}_3 = \rho^2/4}$$

To solve this issue, we proposed to use **a proportional redistribution of the dividends**
(Ortmann 2000)

It led to the definition of the **proportional marginal effects (PME)** (Herin et al. 2023)

Proposition (*Exogeneity detection*). Under mild assumptions on the probabilistic structure of X ,

$$\text{PME}_i = 0 \iff X_i \text{ is exogenous.}$$

In practice, they tend to “discriminate more” than the **Shapley values** when the inputs are **highly correlated**

Estimation

Estimating the PME/Shapley effects \iff Estimating $v(A)$ for every $A \in \mathcal{P}_D$.

Two settings:

- You can sample your model at will (Monte Carlo): Requires a number proportional to $d!(d - 1)$ model evaluations (Song, Nelson, and Staum 2016).
The estimation cost can be substantially lowered by giving-up precision.
- You only have access to an i.i.d. sample (Given-data): The nearest-neighbor procedure requires 2^d estimates (Broto, Bachoc, and Depecker 2020a).

These methods are time-consuming and scale exponentially with the number of inputs, but the estimates can be recycled to compute both indices at once.

But there are many strategies to drive the computational burden down at the expense of precision

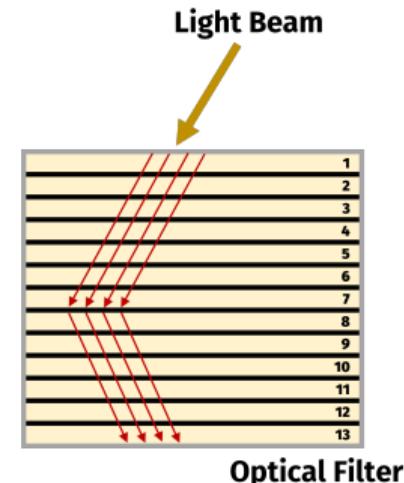
Optical filter transmittance - Feature selection

Transmittance performance of an **optical filter** composed of 13 consecutive layers (Vasseur et al. 2010).

The inputs I_1, \dots, I_{13} represent the **refractive index error** of each layer.

These errors are (highly) correlated due to the manufacturing process.

The numerical model computes the **transmittance error w.r.t. the “perfect filter”** over several wavelengths.



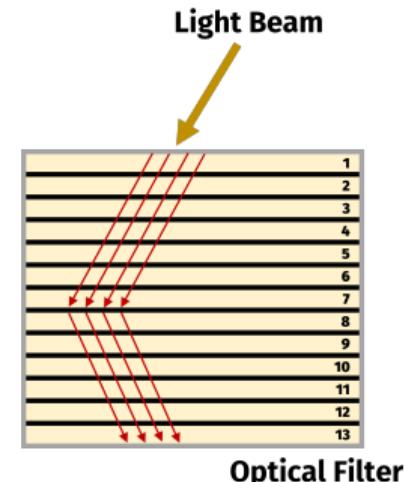
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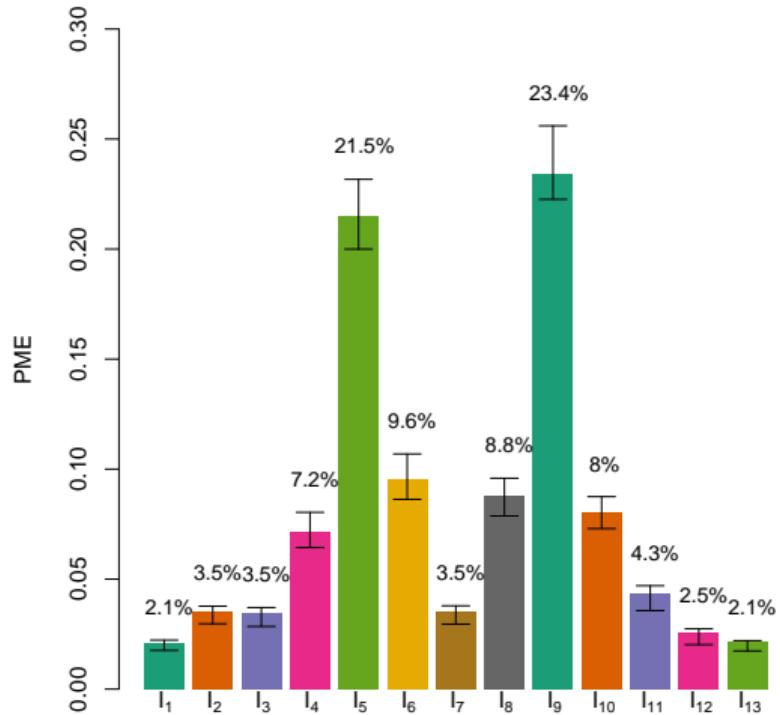
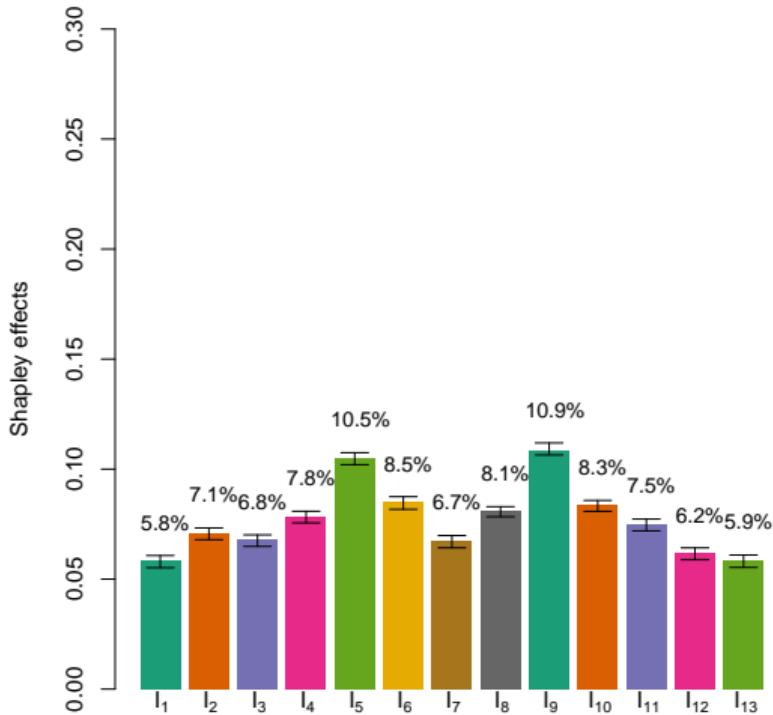
☞ We only have access to an i.i.d. input-output sample ($n = 1000$).

The indices are computed using a **nearest-neighbors approach** (Broto, Bachoc, and Depecker 2020b).

Parallelized implementation using the R package `sensitivity` (~ 4 min runtime, 8 cores).

Arbitrarily chosen number of neighbors: 6.

Optical filter transmittance - Feature selection



Optical filter transmittance - Feature selection

Scenario: We want to build a surrogate model (Gaussian process*) of this numerical model.

Using the whole dataset: $Q^2 = 99.48\%$.

Feature selection:

- First threshold: 2.5% importance.
 - **Shapley effects:** No features removed.
 - **PME:** I_1 and I_3 are removed, $Q^2 = 99.14\%$.
- Second threshold: 5% importance.
 - **Shapley effects:** No features removed.
 - **PME:** 7 inputs are removed, $Q^2 = 98.79\%$.

* 5/2 Matérn covariance kernel, constant trend.

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But what about the second issue?

Generalizing Hoeffding's decomposition to dependent inputs

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But we believe we found an interesting approach to tackle it!

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Under these two assumptions **we can actually decompose the space of outputs**

Consequences

Under these two assumptions, any real-valued **random output** $G(X)$ with finite second moment can be **uniquely** decomposed:

$$G(X) = \sum_{A \in \mathcal{P}_D} G_A(X_A),$$

where $G_A(X_A)$ are **hierarchically orthogonal**

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They can be characterized using **oblique projections** $\mathbb{M}_A[G(X)]$ which **change w.r.t. the dependence structure**:

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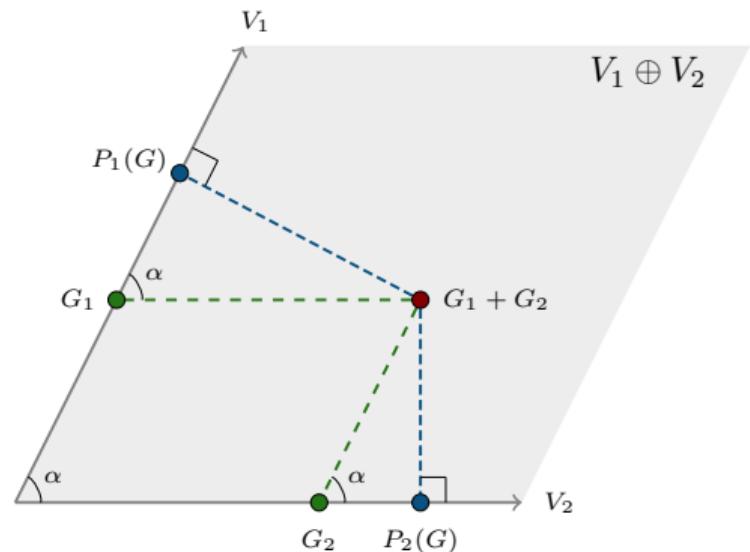
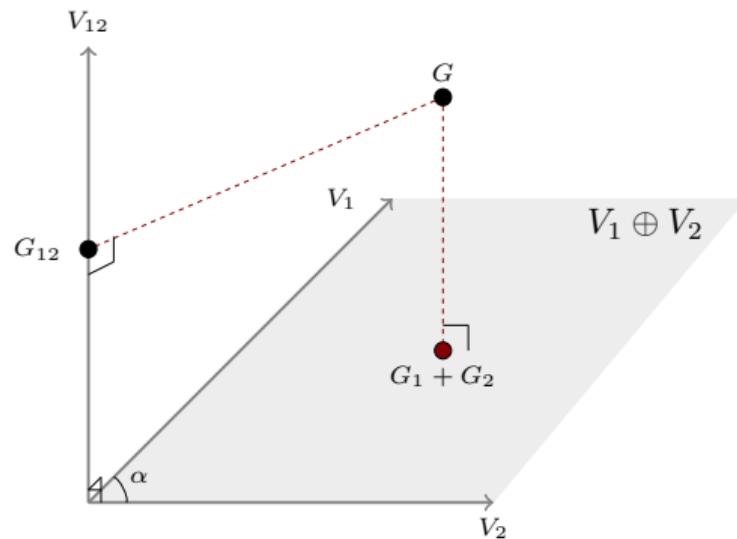
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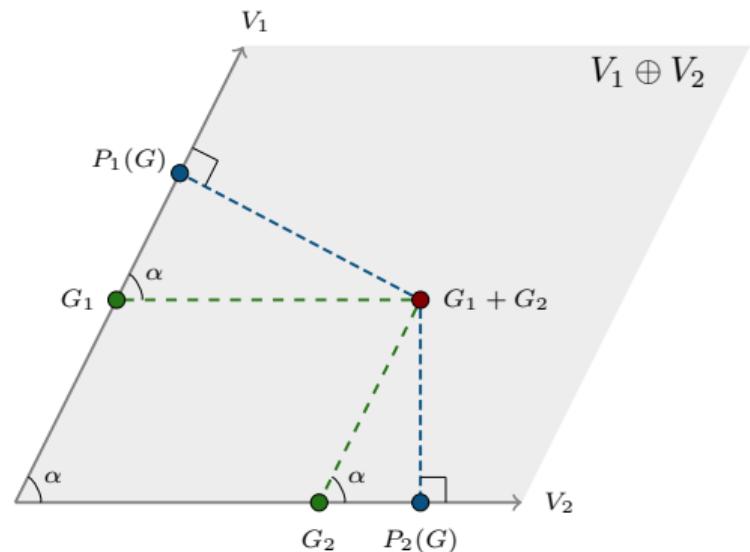
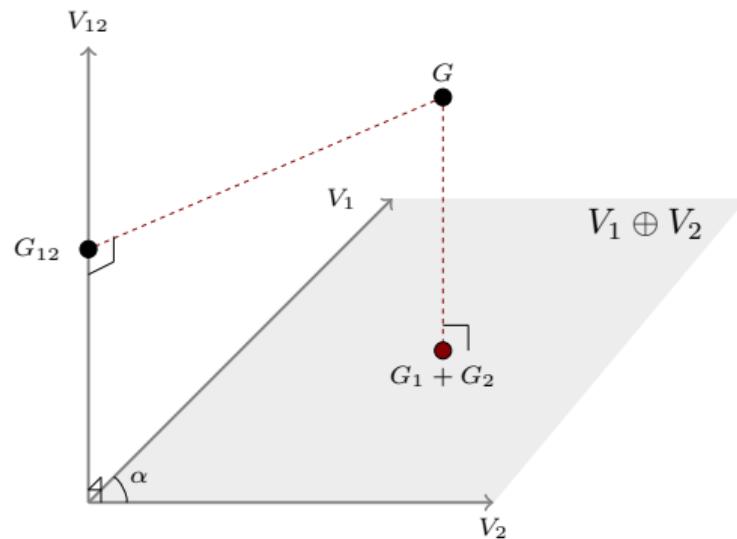
☞ We recover Hoeffding's classical decomposition

Illustration*: bivariate function



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The dependence structure of X is **unwanted**, and one wishes to study its effects.

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Organic variance decomposition: separate **pure interaction effects** to **dependence effects**.
The dependence structure of X is **unwanted**, and one wishes to study its effects.

Orthocanonical variance decomposition: the dependence structure of X is **inherent in the uncertainty modeling** of the studied phenomenon. It amounts to quantify **structural** and **correlative** effects.

Organic variance decomposition: Pure interaction

The notion of pure interaction is intrinsically linked with the notion of mutual independence.

Let $\tilde{X} = (\tilde{X}_1, \dots, \tilde{X}_d)^\top$ be the random vector such that

$$\tilde{X}_i \stackrel{d}{=} X_i, \quad \text{and } \tilde{X} \text{ is mutually independent.}$$

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Definition (Pure interaction). For every $A \in \mathcal{P}_D$, define the **pure interaction of X_A on $G(X)$** as

$$S_A = \frac{\mathbb{V}(P_A(G(\tilde{X})))}{\mathbb{V}(G(\tilde{X}))} \times \mathbb{V}(G(X)).$$

These indices are the **Sobol' indices** computed on the mutually independent version of X .

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This approach **strongly resembles the “independent Sobol’ indices”** proposed by Mara, Tarantola, and Annoni (2015).

(see, also, Lebrun and Dutfoy (2009a, 2009b))

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To define **dependence effects**, we measure **the distance between the orthogonal and oblique projections**

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$$S_A^D = \mathbb{E} \left[(Q_A(G(X)) - P_A(G(X)))^2 \right].$$

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Probably the *overall effect of the dependence on the model's uncertainty!*

Canonical variance decomposition

The structural effects represent the variance of each of the $G_A(X_A)$. It amounts to perform a **covariance decomposition** (Hart and Gremaud 2018; Da Veiga et al. 2021).

Definition (*Structural effects*).

For every $A \in \mathcal{P}_D$, define the **structural effects of X_A on $G(X)$** as

$$S_A^U = \mathbb{V}(G_A(X_A)).$$

The **correlative effects** represent the part of variance that is due to the correlation between the representants $G_A(X_A)$

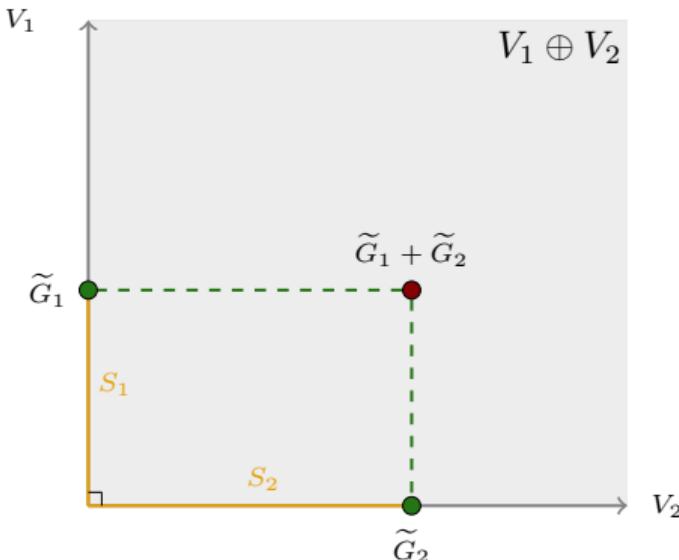
Definition (*Correlative effects*).

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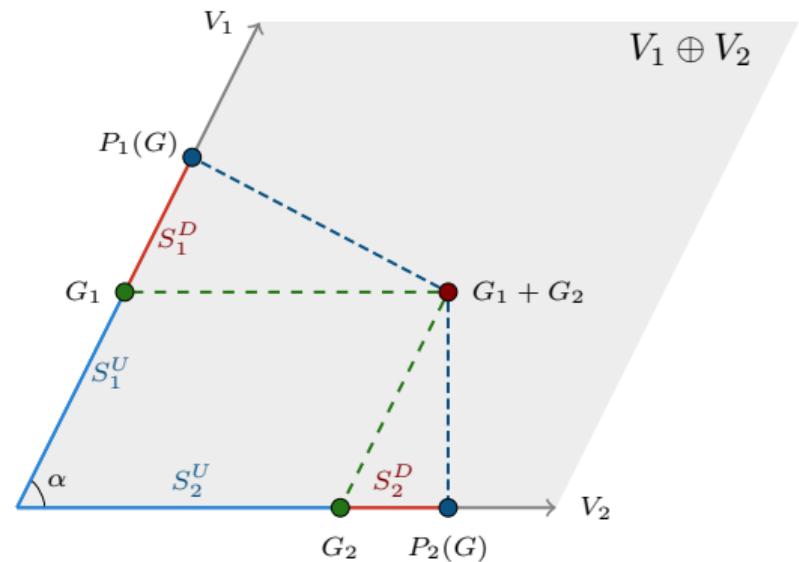
$$S_A^C = \text{Cov} \left(G_A(X_A), \sum_{B \in \mathcal{P}_D : B \neq A} G_B(X_B) \right).$$

Variance decomposition: Intuition

Pure interaction effects

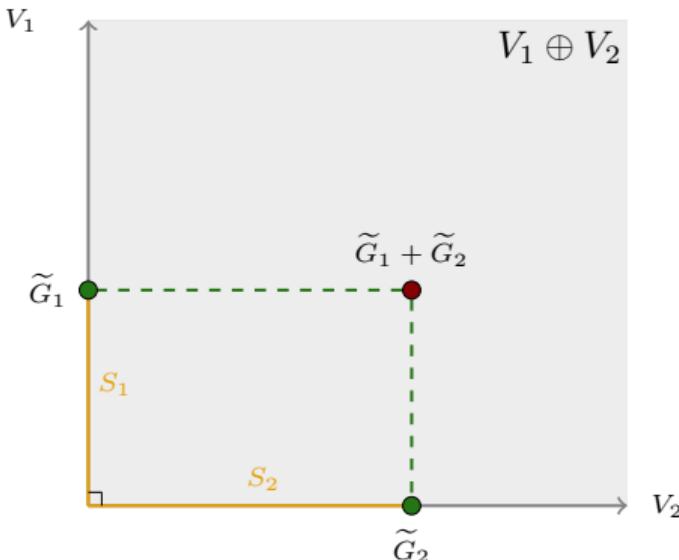


Structural and dependence effects

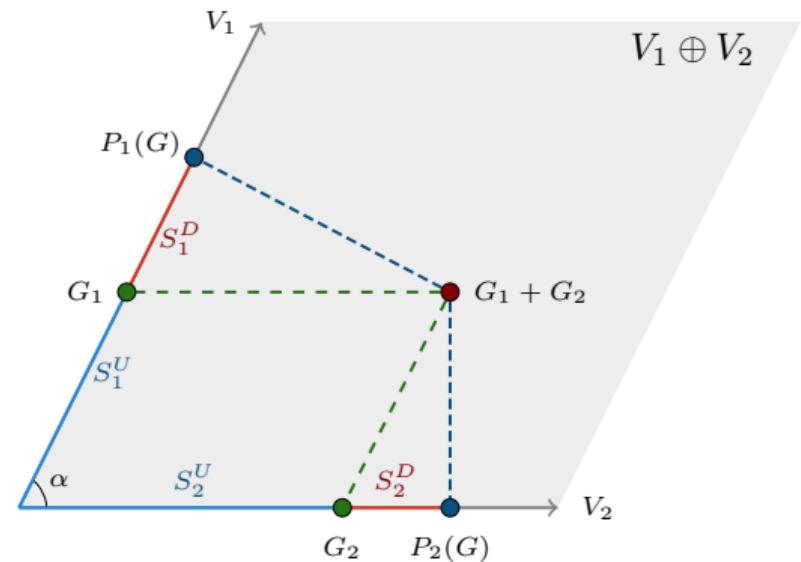


Variance decomposition: Intuition

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Structural and dependence effects



☞ Maybe a good candidate to solve our second issue!

Some perspectives

Main challenge: Estimating the oblique projections from observations

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☞ **Beyond hierarchical orthogonality**

e.g., Köhler, Rügamer, and Schmid (2024) with “stacked orthogonality” conditions

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THANK YOU FOR YOUR ATTENTION!

ANY QUESTIONS?

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