

# Probability - Lecture notes - Unit 01

## Events and outcomes

### 01 Theory

#### 📖 Events and outcomes – informally

- An **event** is a *description* of something that can happen.
- An **outcome** is a *complete description* of something that can happen.

All outcomes are events. An event is usually a *partial* description. Outcomes are events given with a *complete* description.

Here ‘complete’ and ‘partial’ are within the context of the **probability model**.

- ⚠ It can be misleading to say that an ‘outcome’ is an ‘observation’.
  - ‘Observations’ occur in the *real world*, while ‘outcomes’ occur in the *model*.
  - To the extent the model is a good one, and the observation conveys *complete* information, we can say ‘outcome’ for the observation.

Notice:

- ⓘ Because outcomes are *complete*, no two distinct outcomes could *actually happen* in a run of the experiment being modeled.

When an event happens, the *fact* that it has happened constitutes **information**.

#### 📖 Events and outcomes – mathematically

- The **sample space** is the *set of possible outcomes*, so it is the set of the complete descriptions of everything that can happen.
- An **event** is a *subset* of the sample space, so it is a *collection of outcomes*.

- ⓘ For mathematicians: some “wild” subsets are not *valid* events. Problems with infinity and the continuum...

#### 🔖 Notation

- Write  $S$  for the set of possible outcomes,  $s \in S$  for a single outcome in  $S$ .
- Write  $A, B, C, \dots \subset S$  or  $A_1, A_2, A_3, \dots \subset S$  for some events, subsets of  $S$ .
- Write  $\mathcal{F}$  for the collection of all events. This is frequently a *huge* set!
- Write  $|A|$  for the **cardinality** or *size* of a set  $A$ , i.e. the *number of elements it contains*.

Using this notation, we can consider an *outcome itself as an event* by considering the “singleton” subset  $\{\omega\} \subset S$  which contains that outcome alone.

## 02 Illustration

### ≡ Example - Coin flipping

Flip a fair coin two times and record both results.

- *Outcomes*: sequences, like  $HH$  or  $TH$ .
- *Sample space*: all possible sequences, i.e. the set  $S = \{HH, HT, TH, TT\}$ .
- *Events*: for example:
  - $A = \{HH, HT\} =$  “first was heads”
  - $B = \{HT, TH\} =$  “exactly one heads”
  - $C = \{HT, TH, HH\} =$  “at least one heads”

With this setup, we may combine events in various ways to generate other events:

- *Complex events*: for example:
  - $A \cap B = \{HT\}$ , or in words:  
“first was heads” AND “exactly one heads” = “heads-then-tails”  
  
Notice that the last one is a *complete description*, namely the *outcome*  $HT$ .
  - $A \cup B = \{HH, HT, TH\}$ , or in words:  
“first was heads” OR “exactly one heads”  
= “starts with heads, else it’s tails-then-heads”

### ✂ Exercise - Coin flipping: counting subsets

Flip a fair coin five times and record the results.

How many elements are in the sample space? (How big is  $S$ ?)

How many events are there? (How big is  $\mathcal{F}$ ?)

[Solution](#)

## 03 Theory

### 📄 New events from old

Given two events  $A$  and  $B$ , we can form new events using set operations:


$$A \cup B \longleftrightarrow \text{“event } A \text{ OR event } B\text{”}$$

$$A \cap B \longleftrightarrow \text{“event } A \text{ AND event } B\text{”}$$

$$A^c \longleftrightarrow \text{not event } A$$

We also use these terms for events  $A$  and  $B$ :

- They are **mutually exclusive** when  $A \cap B = \emptyset$ , that is, they have *no elements in common*.
- They are **collectively exhaustive**  $A \cup B = S$ , that is, when they jointly *cover all possible outcomes*.

-  In probability texts, sometimes  $A \cap B$  is written “ $A \cdot B$ ” or even (frequently!) “ $AB$ ”.

### Rules for sets

#### Algebraic rules

- Associativity:  $(A \cup B) \cup C = A \cup (B \cup C)$ . Analogous to  $(A + B) + C = A + (B + C)$ .
- Distributivity:  $A \cap (B \cup C) = (A \cap B) \cup (A \cap C)$ . Analogous to  $A(B + C) = AB + AC$ .

#### De Morgan’s Laws

- $(A \cup B)^c = A^c \cap B^c$
- $(A \cap B)^c = A^c \cup B^c$

In other words: you can distribute “ $c$ ” but must simultaneously do a switch  $\cap \leftrightarrow \cup$ .

## Probability models

### 04 Theory

#### Axioms of probability


A **probability measure** is a function  $P : \mathcal{F} \rightarrow \mathbb{R}$  satisfying:

#### Kolmogorov Axioms:

- **Axiom 1:**  $P[A] \geq 0$  for every event  $A$   
(probabilities are not negative!)
- **Axiom 2:**  $P[S] = 1$   
(probability of “anything” happening is 1)
- **Axiom 3:** additivity for any *countable collection* of *mutually exclusive* events:

$$P[A_1 \cup A_2 \cup A_3 \cup \dots] = P[A_1] + P[A_2] + P[A_3] + \dots$$

when:  $A_i \cap A_j = \emptyset$  for all  $i \neq j$

-  Notation: we write  $P[A]$  instead of  $P(A)$ , even though  $P$  is a function, to emphasize the fact that  $A$  is a set.

## Probability model

A **probability model** or **probability space** consists of a triple  $(S, \mathcal{F}, P)$ :

- $S$  the sample space
- $\mathcal{F}$  the set of valid events, where every  $A \in \mathcal{F}$  satisfies  $A \subset S$
- $P : \mathcal{F} \rightarrow \mathbb{R}$  a probability measure satisfying the Kolmogorov Axioms

## Finitely many exclusive events

It is a consequence of the Kolmogorov Axioms that additivity also works for finite collections of events:

$$P[A \cup B] = P[A] + P[B]$$

$$P[A_1 \cup \dots \cup A_n] = P[A_1] + \dots + P[A_n]$$

## Inferences from Kolmogorov

A probability measure satisfies these rules.

They can be deduced from the Kolmogorov Axioms.

- **Negation:** Can you find  $P[A^c]$  but not  $P[A]$ ? Use negation:

$$P[A] = 1 - P[A^c]$$

- **Monotonicity:** Probabilities grow when outcomes are added:

$$A \subset B \implies P[A] \leq P[B]$$

- **Inclusion-Exclusion:** A trick for resolving unions:

$$P[A \cup B] = P[A] + P[B] - P[A \cap B]$$

(even when  $A$  and  $B$  are *not exclusive!*)

## Inclusion-Exclusion

The principle of inclusion-exclusion generalizes to three events:

$$P[A \cup B \cup C] =$$

$$P[A] + P[B] + P[C] - P[A \cap B] - P[A \cap C] - P[B \cap C] + P[A \cap B \cap C]$$

The same pattern works for any number of events!

The pattern goes: “include singles” then “exclude doubles” then “include triples” then ...

Include, exclude, include, exclude, include, ...

## 05 Illustration

### ≡ Example - Lucia is Host or Player

The professor chooses three students at random for a game in a class of 40, one to be Host, one to be Player, one to be Judge. What is the probability that Lucia is either Host or Player?

#### ≡ Solution ∨

##### 1. ≡ Set up the probability model.

- Label the students 1 to 40. Write  $L$  for Lucia's number.
- **Outcomes:** assignments such as  $(H, P, J) = (2, 5, 8)$   
These are ordered triples with *distinct* entries in 1, 2, ..., 40.
- **Sample space:**  $S$  is the collection of all such distinct triples
- **Events:** any subset of  $S$
- **Probability measure:** assume all outcomes are equally likely, so  $P[(i, j, k)] = P[(r, l, p)]$  for all  $i, j, k, r, l, p$
- In total there are  $40 \cdot 39 \cdot 38$  triples of distinct numbers.
- Therefore  $P[(i, j, k)] = \frac{1}{40 \cdot 39 \cdot 38}$  for any *specific* outcome  $(i, j, k)$ .
- Therefore  $P[A] = \frac{|A|}{40 \cdot 39 \cdot 38}$  for any event  $A$ . (Recall  $|A|$  is the *number* of outcomes in  $A$ .)

##### 2. ⇔ Define the desired event.

- Want to find  $P[\text{“Lucia is Host or Player”}]$
- Define  $A = \text{“Lucia is Host”}$  and  $B = \text{“Lucia is Player”}$ . Thus:

$$A = \{(L, j, k) \mid \text{any } j, k\}, \quad B = \{(i, L, k) \mid \text{any } i, k\}$$

- So we seek  $P[A \cup B]$ .

##### 3. ≡ Compute the desired probability.

- Importantly,  $A \cap B = \emptyset$  (mutually exclusive).  
There are no outcomes in  $S$  in which Lucia is *both* Host and Player.
- By *additivity*, we infer  $P[A \cup B] = P[A] + P[B]$ .
- Now compute  $P[A]$ .
  - There are  $39 \cdot 38$  ways to choose  $j$  and  $k$  from the students besides Lucia.
  - Therefore  $|A| = 39 \cdot 38$ .
  - Therefore:

$$P[A] \gg \gg \frac{|A|}{40 \cdot 39 \cdot 38} \gg \gg \frac{39 \cdot 38}{40 \cdot 39 \cdot 38} \gg \gg \frac{1}{40}$$

- Now compute  $P[B]$ . It is similar:  $P[B] = \frac{1}{40}$ .
- Finally compute that  $P[A] + P[B] = \frac{1}{20}$ , so the answer is:

$$P[A \cup B] \gg \gg P[A] + P[B] \gg \gg \frac{1}{20}$$

### ≡ Example - iPhones and iPads

At Mr. Jefferson's University, 25% of students have an iPhone, 30% have an iPad, and 60% have neither.

What is the probability that a randomly chosen student has *some* iProduct? (Q1)

What about *both*? (Q2)

#### ≡ Solution ∨

##### 1. ≡ Set up the probability model.

- A student is chosen at random: an *outcome* is the chosen student.
- *Sample space*  $S$  is the set of all students.
- Write  $O$  = "has iPhone" and  $A$  = "has iPad" concerning the chosen student.
- All students are equally likely to be chosen: therefore  $P[E] = \frac{|E|}{|S|}$  for any event  $E$ .
- Therefore  $P[O] = 0.25$  and  $P[A] = 0.30$ .
- Furthermore,  $P[O^c A^c] = 0.60$ . This means 60% have "not iPhone AND not iPad".

##### 2. ≡ Define the desired event.

- Q1: desired event =  $O \cup A$
- Q2: desired event =  $OA$

##### 3. ≡ Compute the probabilities.

- We do not believe  $O$  and  $A$  are exclusive.
- Try: apply inclusion-exclusion:

$$P[O \cup A] = P[O] + P[A] - P[OA]$$

- We know  $P[O] = 0.25$  and  $P[A] = 0.30$ . So this formula, with given data, RELATES Q1 and Q2.
- Notice the complements in  $O^c A^c$  and try *Negation*.
- *Negation*:

$$P[(OA)^c] = 1 - P[OA]$$

DOESN'T HELP.

- Try again: *Negation*:

$$P[(O^c A^c)^c] = 1 - P[O^c A^c]$$

- And De Morgan (or a Venn diagram!):

$$(O^c A^c)^c \gg \gg O \cup A$$

- Therefore:

$$P[O \cup A] \gg \gg P[(O^c A^c)^c]$$

$$\gg \gg 1 - P[O^c A^c] \gg \gg 1 - 0.6 = 0.4$$

- We have found Q1:  $P[O \cup A] = 0.40$ .
- Applying the RELATION from inclusion-exclusion, we get Q2:

$$P[O \cup A] = P[O] + P[A] - P[OA]$$

$$\gg \gg 0.40 = 0.25 + 0.30 - P[OA]$$

$$\gg \gg P[OA] = 0.15$$

## Conditional probability

### 06 Theory

#### Conditional probability

The **conditional probability** of “ $B$  given  $A$ ” is defined by:

$$P[B \mid A] = \frac{P[B \cap A]}{P[A]}$$

This conditional probability  $P[B \mid A]$  represents the probability of event  $B$  taking place *given the assumption* that  $A$  took place. (All within the given probability model.)

By letting the actuality of event  $A$  be taken as a fixed hypothesis, we can define a *conditional probability measure* by plugging events into the slot of  $B$ :

$$P[- \mid A] = \frac{P[- \cap A]}{P[A]}$$

It is possible to verify each of the Kolmogorov axioms for this function, and therefore  $P[- \mid A]$  itself defines a bona fide *probability measure*.

#### Conditioning

What does it really mean?

Conceptually,  $P[B \mid A]$  corresponds to *creating a new experiment* in which we run the old experiment and record data *only those times that A happened*. Or, it corresponds to finding ourselves with *knowledge* or *data* that  $A$  happened, and we seek our best estimates of the likelihoods of other events, based on our existing model and the actuality of  $A$ .

Mathematically,  $P[B \mid A]$  corresponds to *restricting* the probability function to outcomes in  $A$ , and *renormalizing* the values (dividing by  $p[A]$ ) so that the total probability of all the outcomes (in  $A$ ) is now 1.

The definition of conditional probability can also be turned around and reinterpreted:

### 🔲 Multiplication rule

$$P[AB] = P[A] \cdot P[B \mid A]$$

“The probability of  $A$  AND  $B$  equals the probability of  $A$  *times* the probability of  $B$  -given- $A$ .”

This principle generalizes to any events in sequence:

### ☰ Generalized multiplication rule

$$P[A_1 A_2 A_3] = P[A_1] \cdot P[A_2 \mid A_1] \cdot P[A_3 \mid A_1 A_2]$$

$$P[A_1 \cdots A_n] = P[A_1] \cdot P[A_2 \mid A_1] \cdot P[A_3 \mid A_1 A_2] \cdots P[A_n \mid A_1 \cdots A_{n-1}]$$

The generalized rule can be verified like this. First substitute  $A_2$  for  $B$  and  $A_1$  for  $A$  in the original rule. Now repeat, substituting  $A_3$  for  $B$  and  $A_1 A_2$  for  $A$  in the original rule, and combine with the first one, and you find the rule for triples. Repeat again with  $A_4$  and  $A_1 A_2 A_3$ , combine with the triples, and you get quadruples.

## 07 Illustration

### ✍ Exercise - Simplifying conditionals

Let  $A \subset B$ . Simplify the following values:

$$P[A \mid B], \quad P[A \mid B^c], \quad P[B \mid A], \quad P[B \mid A^c]$$

[Solution](#)

### ☰ Coin flipping: at least 2 heads

Flip a fair coin 4 times and record the outcomes as sequences, like  $HHTH$ .



Let  $A_{\geq 2}$  be the event that there are at least two heads, and  $A_{\geq 1}$  the event that there is at least one heads.

First let's calculate  $P[A_{\geq 2}]$ .

Define  $A_2$ , the event that there were exactly 2 heads, and  $A_3$ , the event of exactly 3, and  $A_4$  the event of exactly 4. These events are exclusive, so:

$$P[A_{\geq 2}] = P[A_2 \cup A_3 \cup A_4] \gg \gg P[A_2] + P[A_3] + P[A_4]$$

Each term on the right can be calculated by counting:

$$P[A_2] = \frac{|A_2|}{2^4} \gg \gg \frac{\binom{4}{2}}{16} \gg \gg \frac{6}{16}$$

$$P[A_3] = \frac{|A_3|}{2^4} \gg \gg \frac{\binom{4}{1}}{16} \gg \gg \frac{4}{16}$$

$$P[A_4] = \frac{|A_4|}{2^4} \gg \gg \frac{\binom{4}{0}}{16} \gg \gg \frac{1}{16}$$

Therefore,  $P[A_{\geq 2}] = \frac{11}{16}$ .

Now suppose we find out that “at least one heads definitely came up”. (Meaning that we know  $A_{\geq 1}$ .) For example, our friend is running the experiment and tells us this fact about the outcome.

Now what is our estimate of likelihood of  $A_{\geq 2}$ ?

The formula for conditioning gives:

$$P[A_{\geq 2} \mid A_{\geq 1}] = \frac{P[A_{\geq 2} \cap A_{\geq 1}]}{P[A_{\geq 1}]}$$

Now  $A_{\geq 2} \cap A_{\geq 1} = A_{\geq 2}$ . (Any outcome with at least two heads automatically has at least one heads.) We already found that  $P[A_{\geq 2}] = \frac{11}{16}$ . To compute  $P[A_{\geq 1}]$  we simply **add** the probability  $P[A_1]$ , which is  $\frac{4}{16}$ , to get  $P[A_{\geq 1}] = \frac{15}{16}$ .

Therefore:

$$P[A_{\geq 2} \mid A_{\geq 1}] = \frac{11/16}{15/16} \gg \gg \frac{11}{15}$$

### ≡ Example - Flip a coin, then roll dice

Flip a coin. If the outcome is heads, roll two dice and add the numbers. If the outcome is tails, roll a single die and take that number. What is the probability of getting a tails AND a number at least 3?

#### ≡ Solution

This “two-stage” experiment lends itself to a solution using conditional probability.

1.  $\equiv$  Label the events of interest.

- Let  $H$  and  $T$  be the events that the coin showed heads and tails, respectively.
- Let  $A_1, \dots, A_{12}$  be the events that the final number is  $1, \dots, 12$ , respectively.
- The value we seek is  $P[TA_{\geq 3}]$ .

2.  $\equiv$  Observe known (conditional) probabilities.

- We know that  $P[H] = 0.5$  and  $P[T] = 0.5$ .
- We know that  $P[A_5 | T] = \frac{1}{6}$ , for example, or that  $P[A_1 | H] = \frac{1}{12}$ .

3.  $\Rightarrow$  Apply “multiplication” rule.

- This rule gives:

$$P[TA_{\geq 3}] = P[T] \cdot P[A_{\geq 3} | T]$$

- We know  $P[T] = 0.5$  and can see by counting that  $P[A_{\geq 3} | T] = 0.5$ .
- Therefore  $P[TA_{\geq 3}] = 0.25$ .