

PS8

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Problem 1

*a)

To know if Pareto decay more quickly or not than exponential distribution, we have $\exp(x) = \lambda e^{-\lambda x}$ and $fPareto(x) = \frac{\beta \alpha^\beta}{x^{\beta+1}} I\{\alpha < x\}$. Therefore

$$\lim_{x \rightarrow \infty} \frac{\exp(x)}{fPareto(x)} = \lim_{x \rightarrow \infty} \frac{\lambda e^{-\lambda x}}{\frac{\beta \alpha^\beta}{x^{\beta+1}}}$$

As we can see numerator and denominator are 0 when x goes to infinity, so we can apply the L'Hospital's rule. We need to do is differentiate the numerator and differentiate the denominator and then take the limit. We can find the numerator() is a constant and the denominator goes to infinity,

$$\lim_{x \rightarrow \infty} \frac{\exp(x)}{fPareto(x)} = \lim_{x \rightarrow \infty} \frac{\lambda}{\lambda^{\beta+1} \beta \alpha^\beta e^{\lambda x}} = 0$$

Thus, the tail of the Pareto decay more slowly than that of an exponential distribution.

Problem1

b)

```
library(EnvStats)

##
## Attaching package: 'EnvStats'
## The following objects are masked from 'package:stats':
##
##   predict, predict.lm
## The following object is masked from 'package:base':
##
##   print.default

#Estimate E(X) and E(X^2) when m = 10000
ParetoSample <- rpareto(10000,location = 2, shape = 3)
weight <- dexp(ParetoSample-2)/dpareto(ParetoSample,location =2,shape =3)
X <- weight*ParetoSample
#Expectation of X, should close to 3
mean(X)

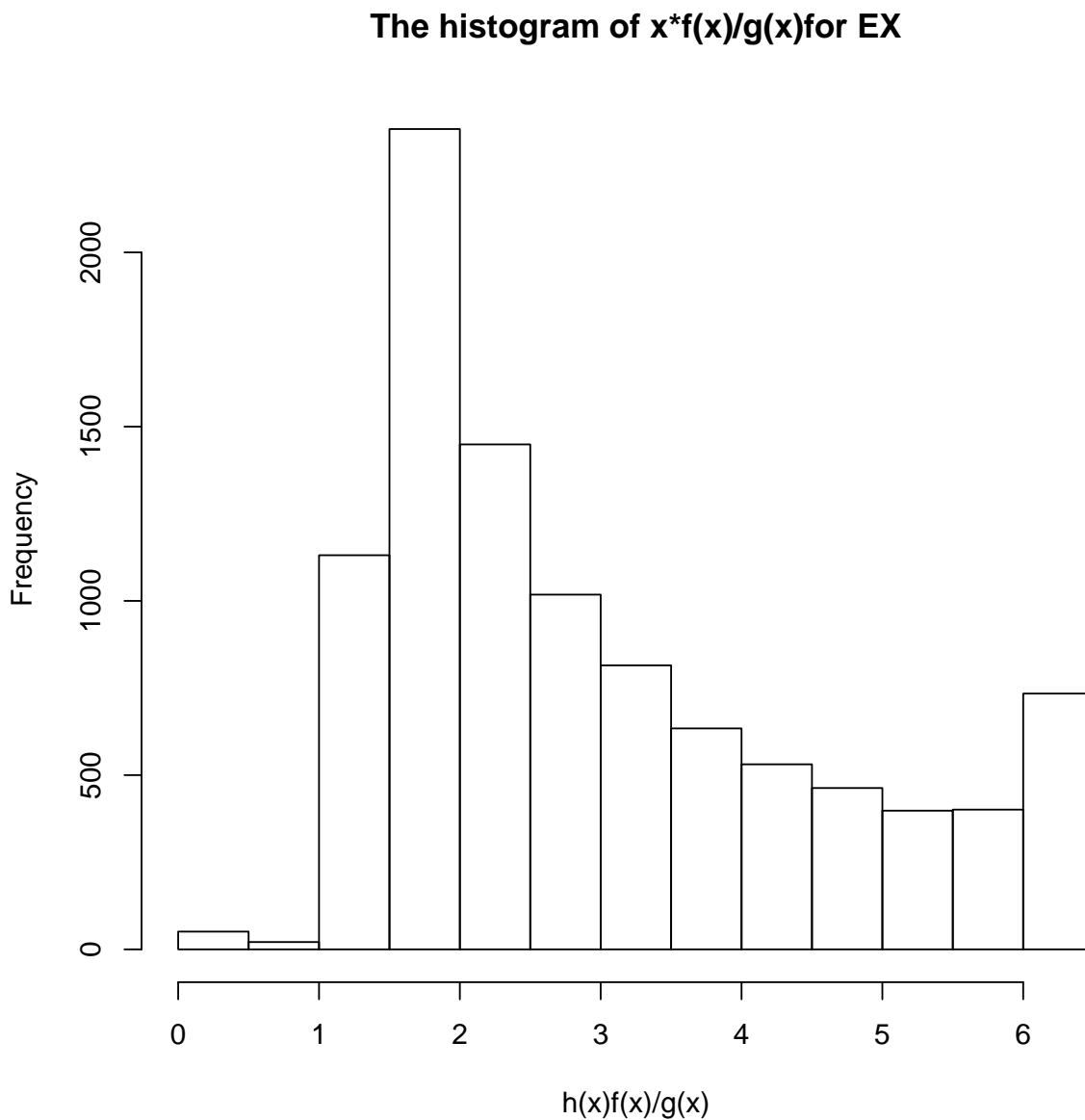
## [1] 3.022294

X_Square <- weight*ParetoSample^2
##Expectation of X_Square, should close to 10
mean(X_Square)

## [1] 10.08577
```

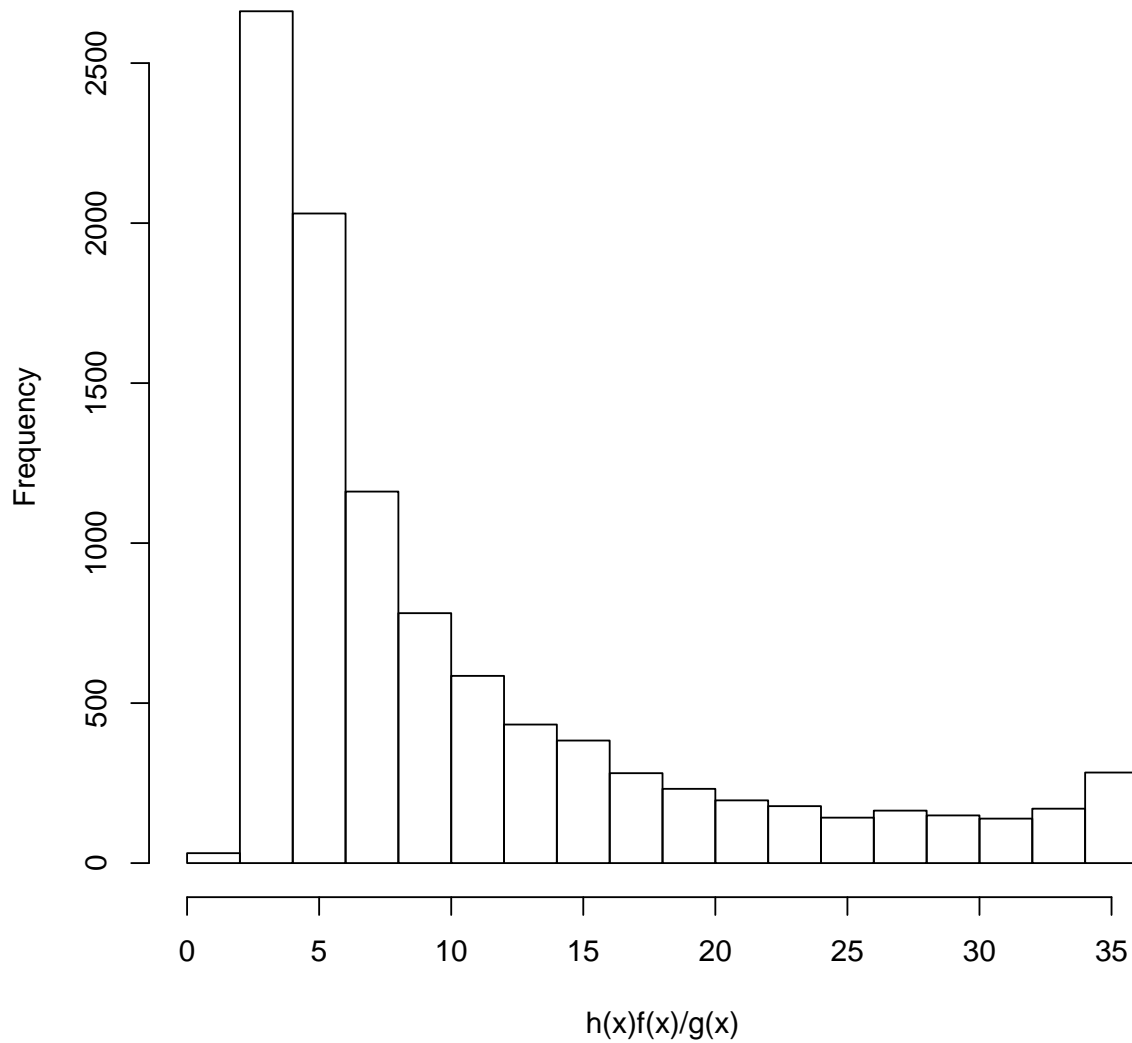
For exponential distribution with parameter =1 and shift by 2. $E(X) = 1 + 2 = 3$ and Variance remain the same. Therefore, the $E(X^2) = Var(X) + E(X)^2 = 1 + 9 = 10$

```
# Histogram of  $h(x)f(x)/g(x)$ 
hist(X, xlab = "h(x)f(x)/g(x)", main="The histogram of  $x*f(x)/g(x)$ for EX")
```

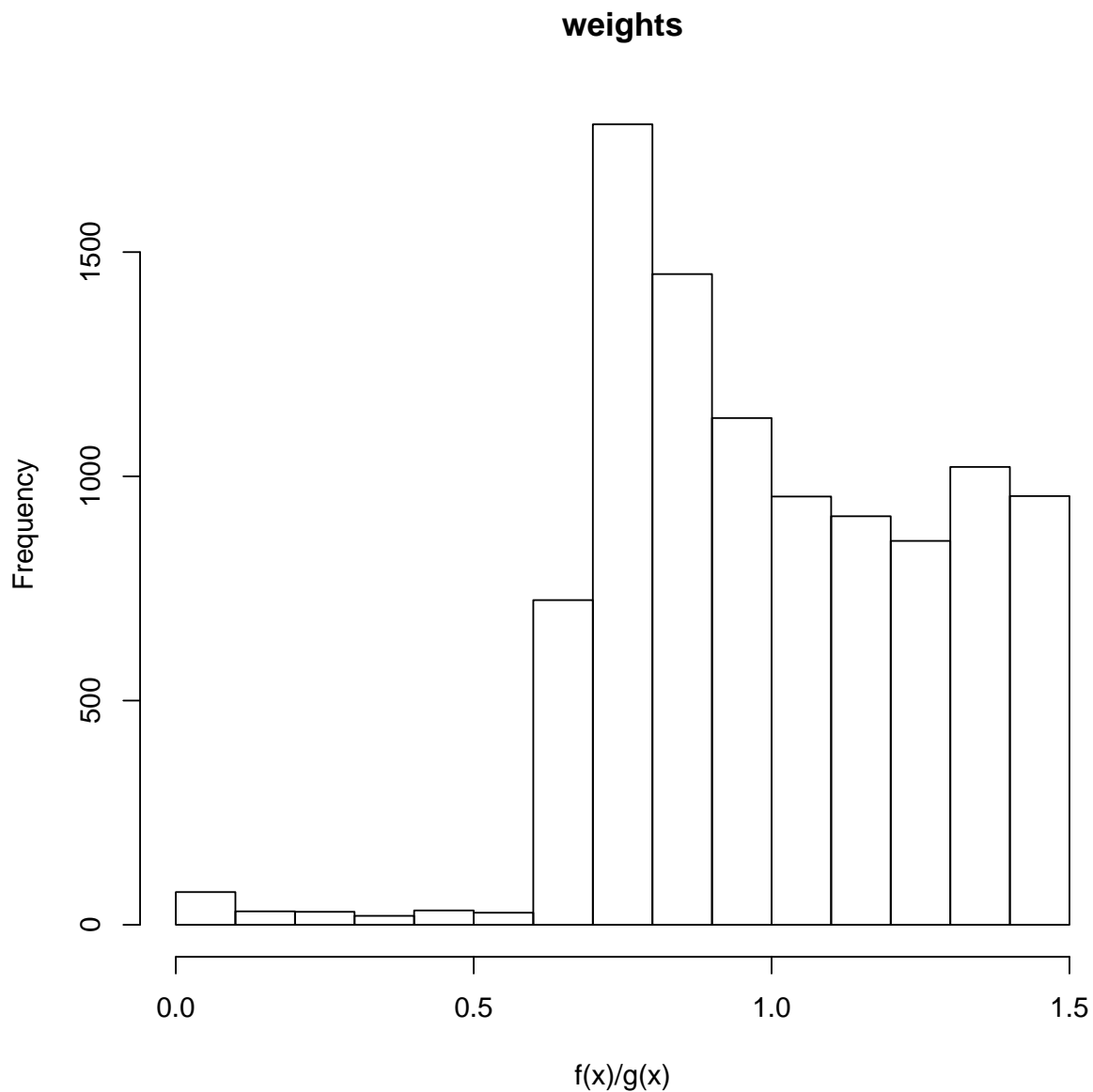


```
hist(X_Square, xlab = "h(x)f(x)/g(x)", main="The histogram of expression(  $h(x)*f(x)/g(x)$ for EX2")
```

The histogram of expression($h(x)*f(x)/g(x)$ for EX^2



```
#the weight  $f(x)/g(x)$   
hist(weight, xlab = "f(x)/g(x)", main= "weights")
```



As we can see from the histogram, the weights lie in the interval $(0.5, 1.5)$ and rest few lie in $(0, 0.5)$. Therefore, there is no strong influence on the estimated $h(X)$

c)

```
#we exchange the f and g
#Estimate E(X) and E(X^2) when m = 10000
ExponetialSample <- rexp(10000) + 2
weight_new <- dpareto(ExponetialSample, location = 2, shape = 3) / dexp(ExponetialSample - 2)
X_new <- weight_new * ParetoSample
#Expectation of X, should close to 3
mean(X_new)

## [1] 2.981714
```

```

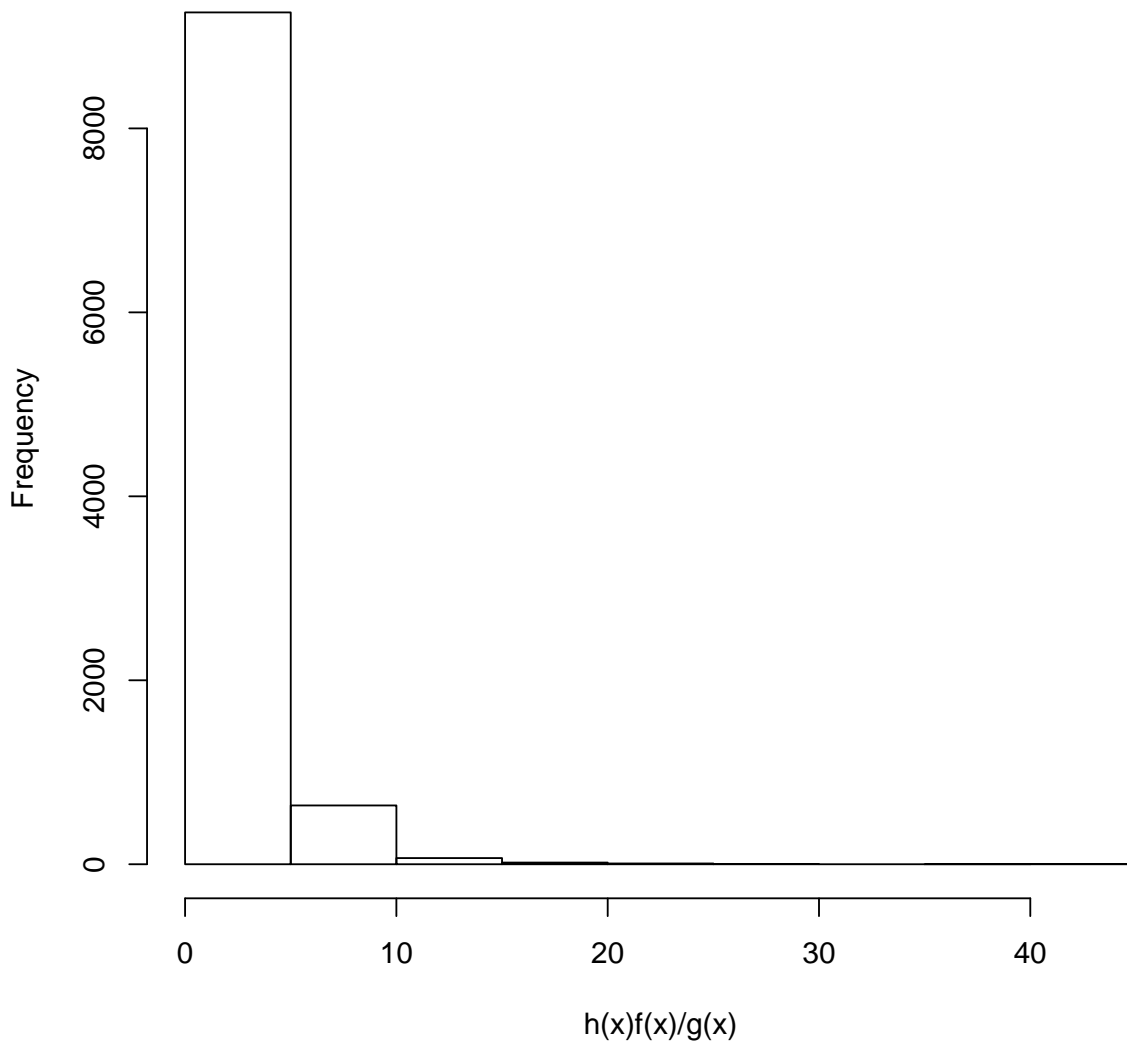
X_Square_new <- weight_new*ParetoSample^2
##Expectation of X_Square, should close to 10
mean(X_Square_new)

## [1] 11.67727

# Histogram of  $h(x)f(x)/g(x)$ 
hist(X_new, xlab = "h(x)f(x)/g(x)", main="The histogram of x*f(x)/g(x)")

```

The histogram of $x*f(x)/g(x)$

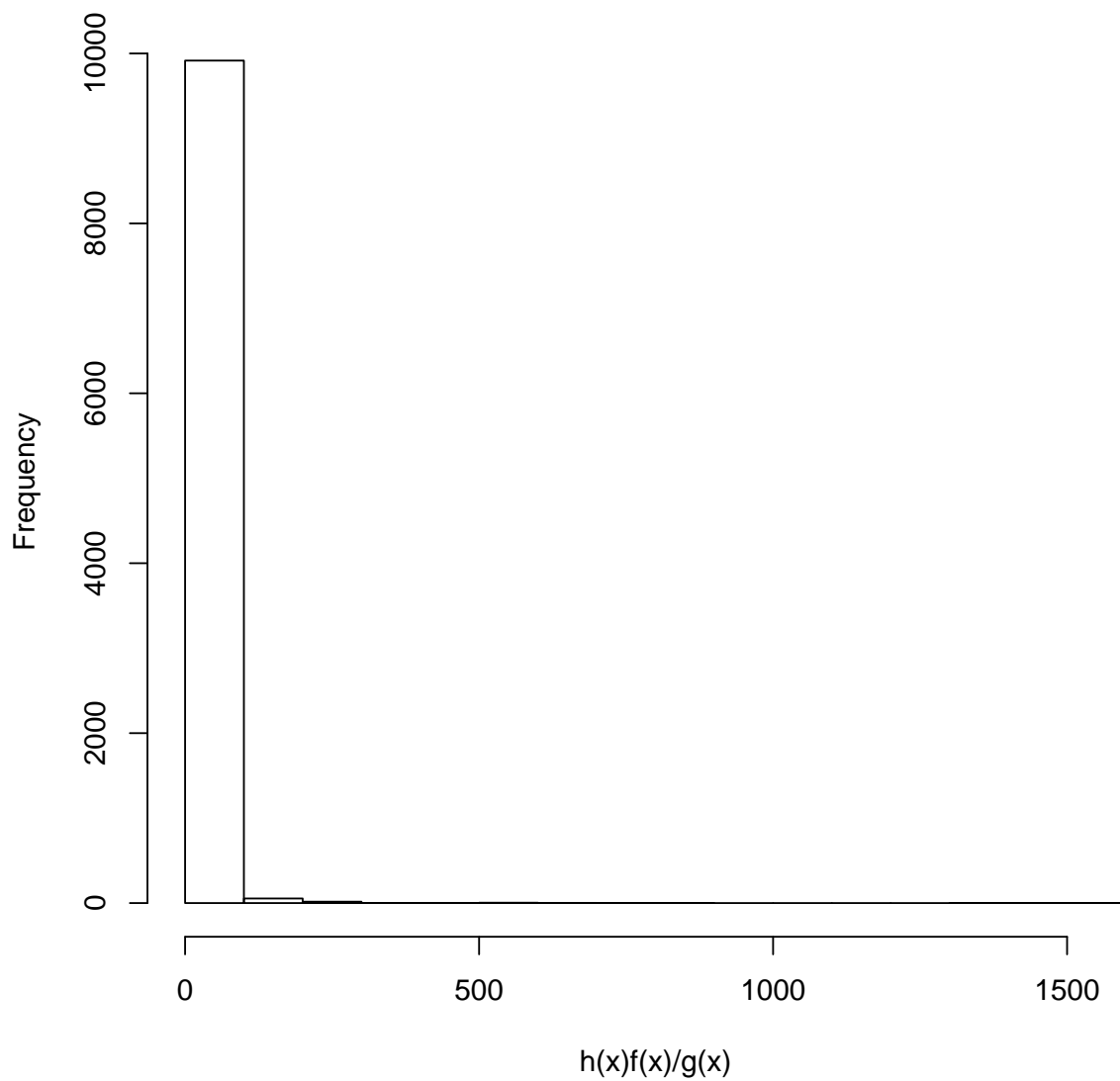


```

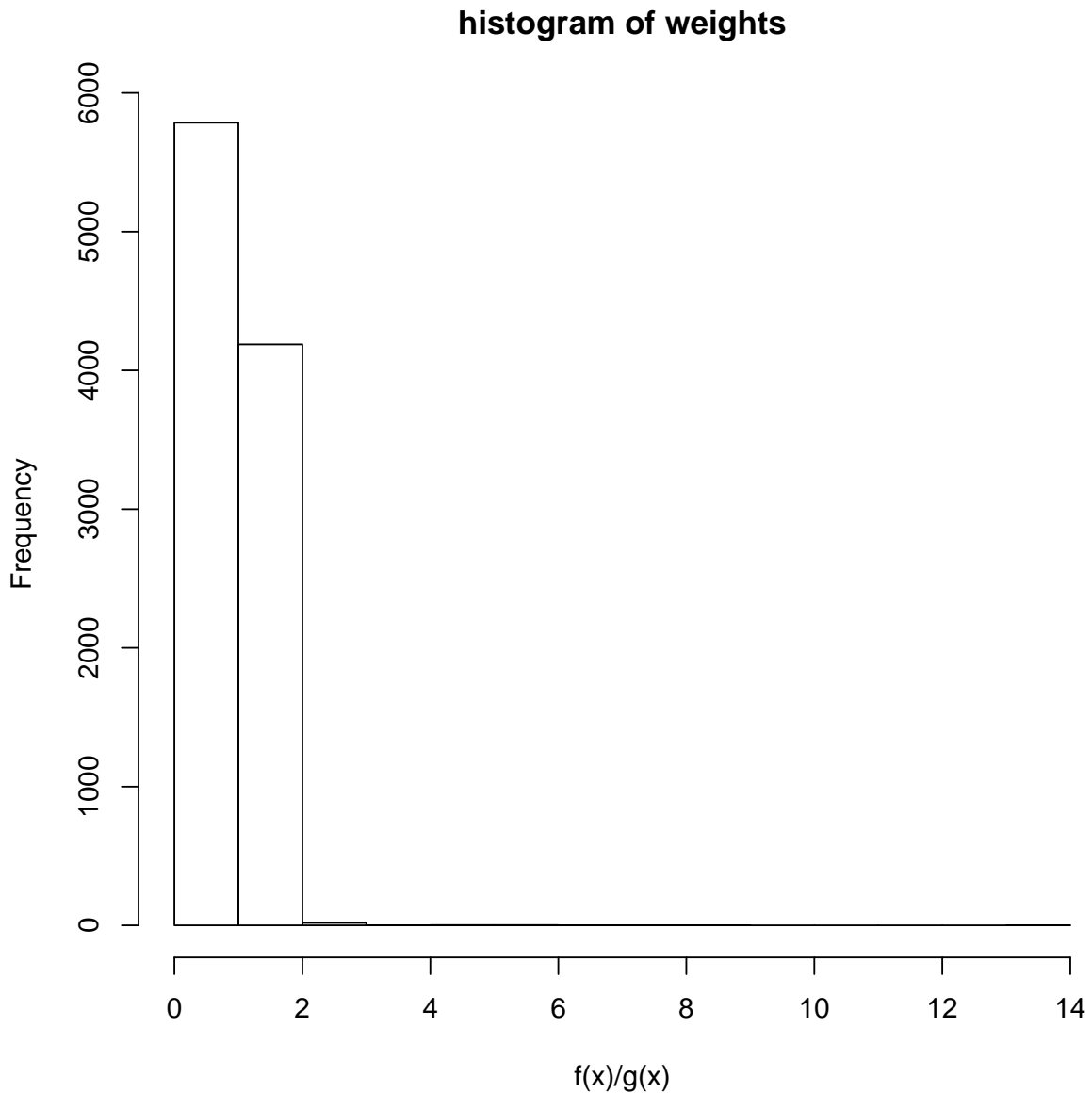
hist(X_Square_new, xlab = "h(x)f(x)/g(x)", main="The histogram of expression(x^2*f(x)/g(x)")

```

The histogram of expression $x^2 f(x)/g(x)$



```
#the weight  $f(x)/g(x)$   
hist(weight_new, xlab = "f(x)/g(x)", main= "histogram of weights")
```



From the histogram we can find the very strong variance of the weights and it influences the estimated $h(X)$ indirectly. Because the exponential distribution has fast decaying tail (lighter tail), the variance becomes large when we sampling from exponential distribution.

Problem 2

```
theta <- function(x1,x2) atan2(x2, x1)/(2*pi)

f <- function(x) {
  f1 <- 10*(x[3] - 10*theta(x[1],x[2]))
  f2 <- 10*(sqrt(x[1]^2 + x[2]^2) - 1)
  f3 <- x[3]
  return(f1^2 + f2^2 + f3^2)
}
```

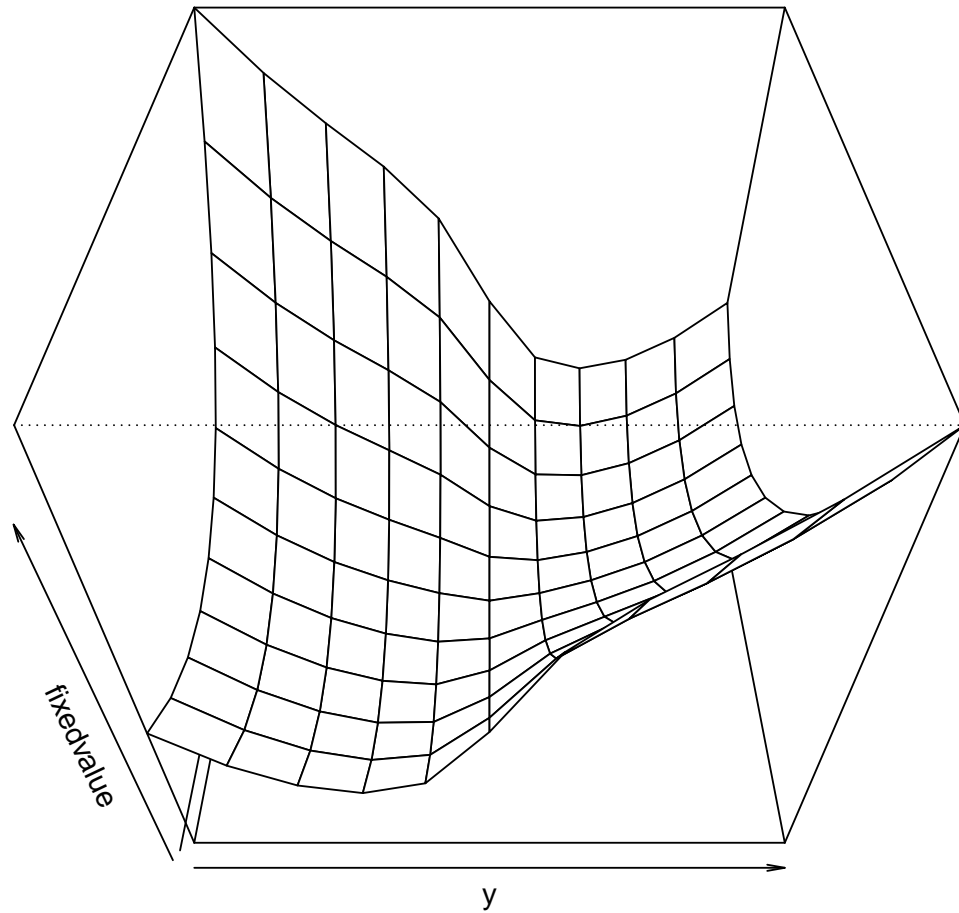
```

}
# fix one variable, the other two range from -5 to 5
# fixed the first variable as constant
Fixed_1 <- function(y,z){
  return(f(c(1,y,z)))
}
#to make the Fixed_1 able to calculate, we need to use vectorize
vectorized <- Vectorize(Fixed_1 ,vectorize.args = c("y","z"))
matrix_1<- outer(-5:5, -5:5, vectorized)

#dataframe_1 <- data.frame(expand.grid(x=-5:5, y=-5:5),value = c(t(matrix_1)))

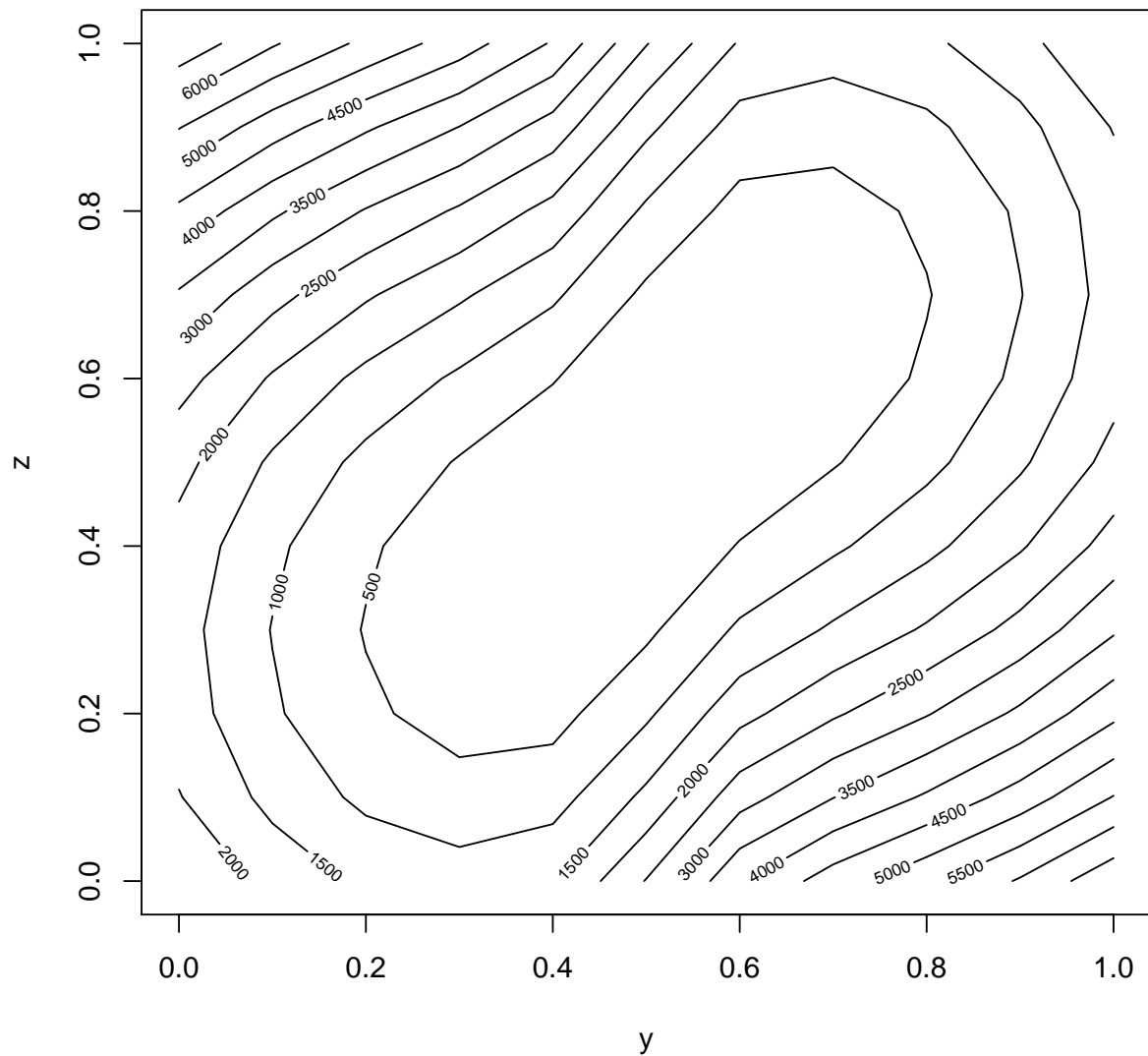
#plots of a surface over the y-z plane
persp(matrix_1, phi = 45, zlab = "fixedvalue", xlab = "y", ylab = "z")

```

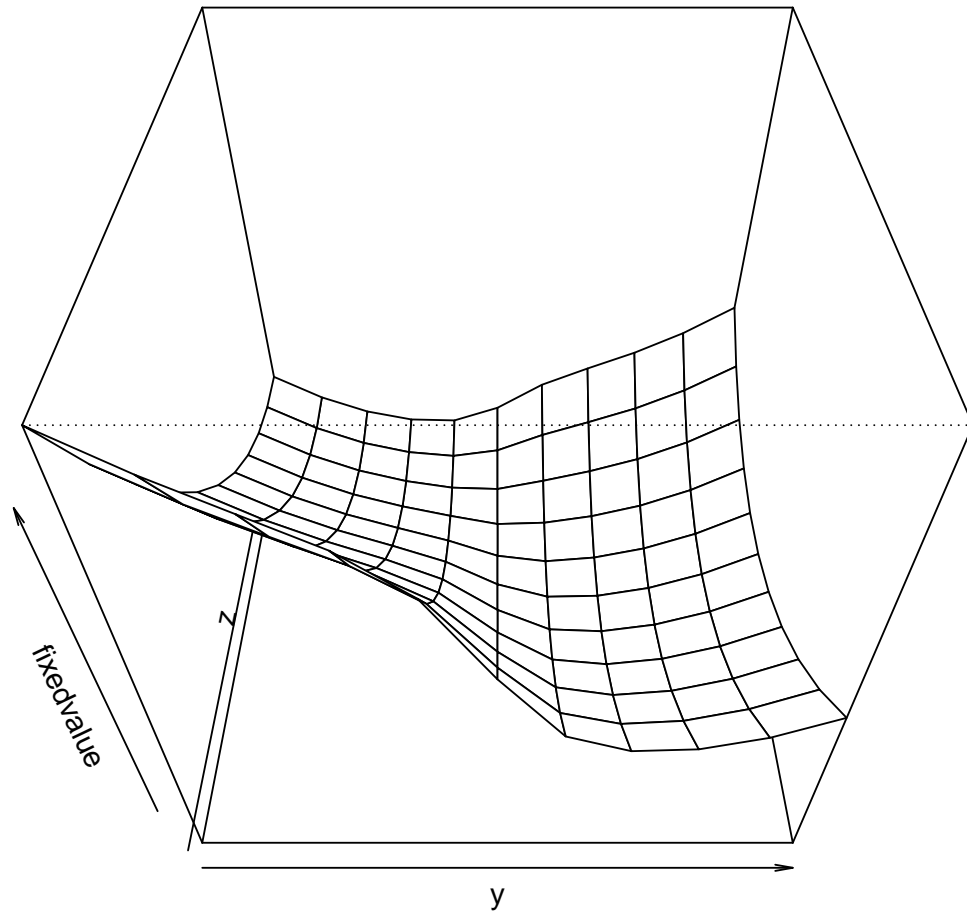
```
# making a contour plot  
contour(matrix_1, xlab = "y", ylab = "z", main = "contour plot of fix first variable")
```

contour plot of fix first variable



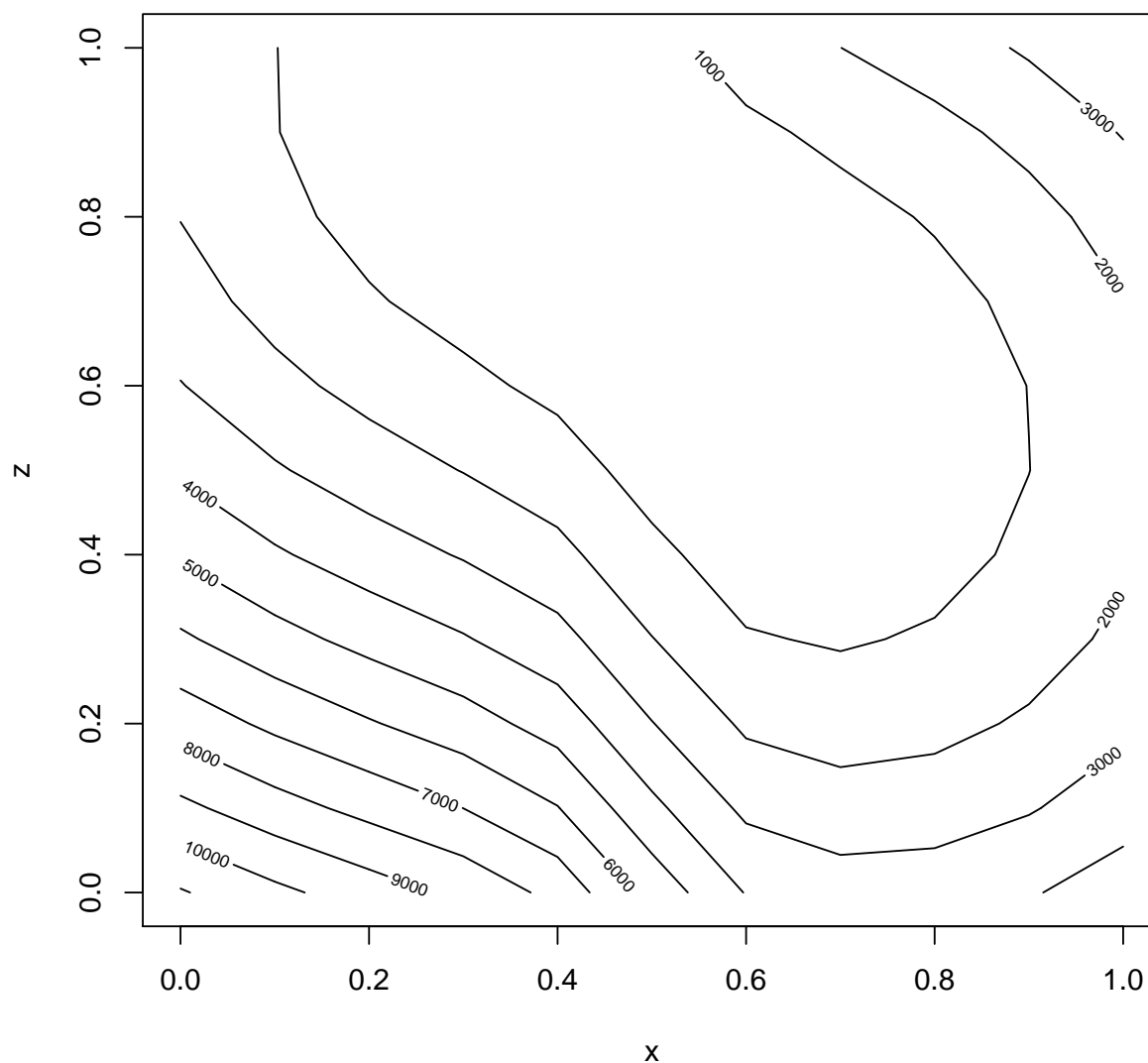
```
# fixed the second variable as constant
Fixed_2 <- function(x,z){
  return(f(c(x,1,z)))
}
#to make the Fixed_1 able to calculate, we need to use vectorize
vectorized <- Vectorize(Fixed_2 ,vectorize.args = c("x","z"))
matrix_2<- outer(-5:5, -5:5, vectorized)

#plots of a surface over the y-z plane
persp(matrix_2, phi = 45, zlab = "fixedvalue", xlab = "y", ylab = "z")
```



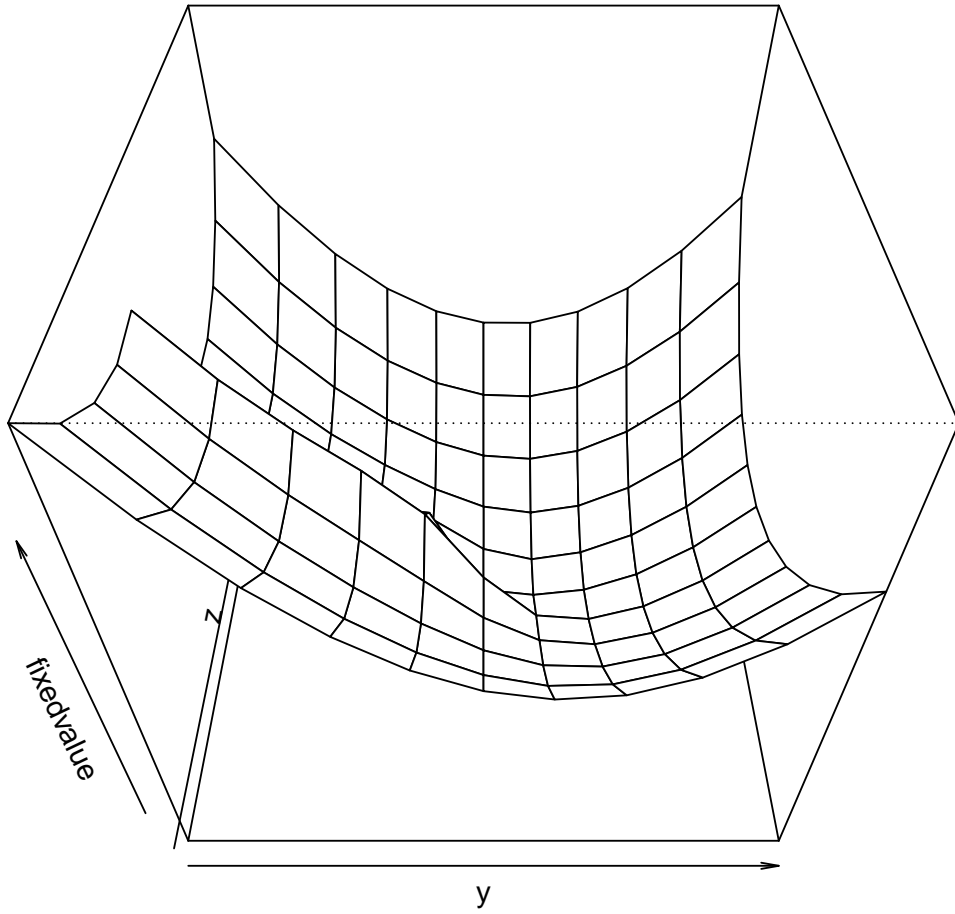
```
# making a contour plot  
contour(matrix_2, xlab = "x", ylab = "z", main = "contour plot of fix second variable")
```

contour plot of fix second variable



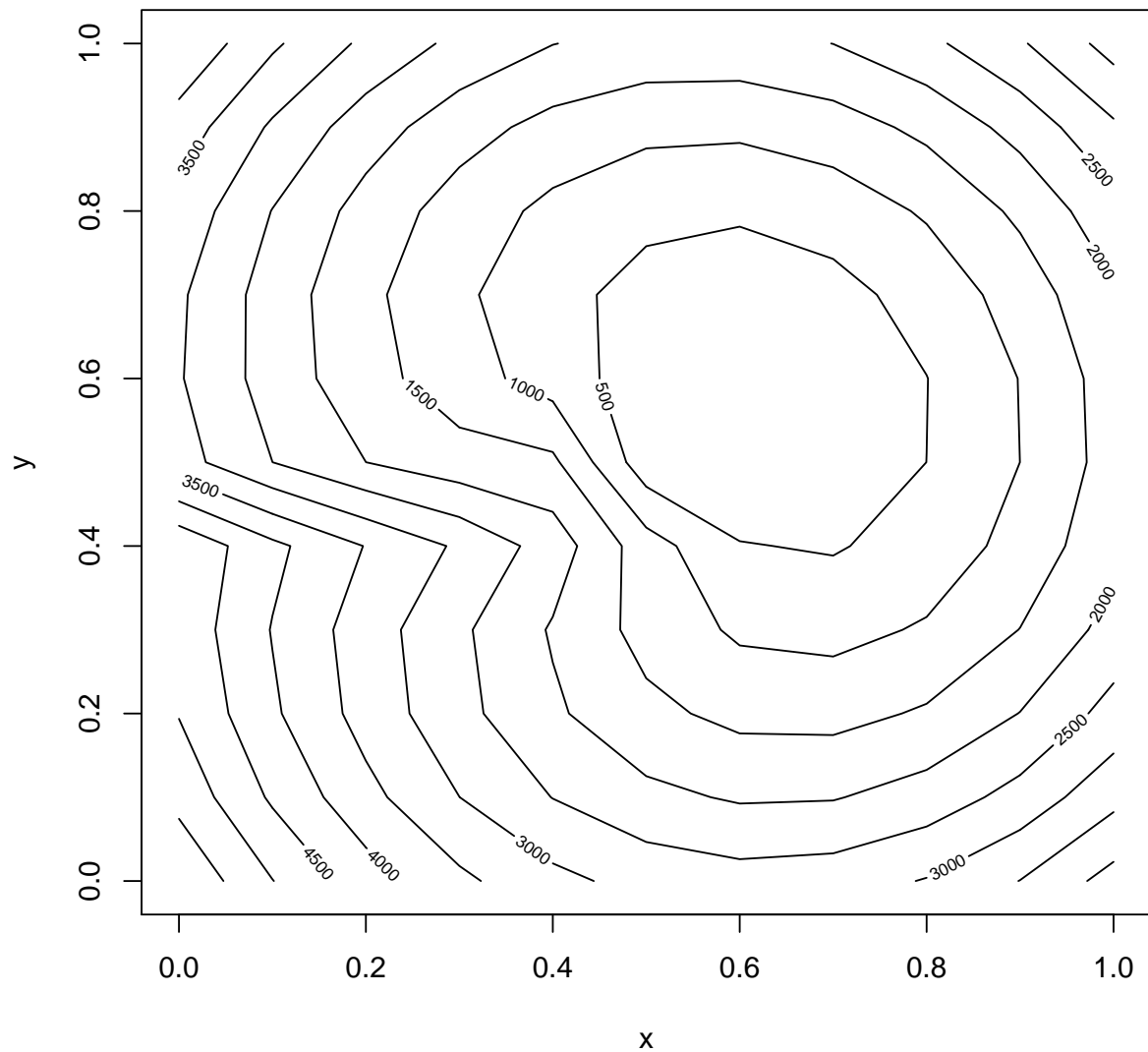
```
# fixed the third variable as constant
Fixed_3 <- function(x,y){
  return(f(c(x,y,1)))
}
#to make the Fixed_1 able to calculate, we need to use vectorize
vectorized <- Vectorize(Fixed_3 ,vectorize.args = c("x","y"))
matrix_3<- outer(-5:5, -5:5, vectorized)

#plots of a surface over the y-z plane
persp(matrix_3, phi = 45, zlab = "fixedvalue", xlab = "y", ylab = "z")
```



```
# making a contour plot  
contour(matrix_3, xlab = "x", ylab = "y", main = "contour plot of fix second variable")
```

contour plot of fix second variable



From above plot we can find that if we fix one variable to 1, the function tend to increases as the rest two variables' absolute values increase. However, near the center of the two unfixed variable, the function are usually to be small.

```
#minimum by optim function.
optim(c(1,1,1), f)$par

## [1] 0.9999779414 -0.0001349269 -0.0001927127

optim(c(0,1,0), f)$par

## [1] 0.9998882225 -0.0002040018 -0.0001412321

optim(c(0,0,1), f)$par

## [1] 0.999711245 0.001178555 0.001921611
```

```

optim(c(0,0,0), f)$par
## [1] 0.999978292 0.002730698 0.004284640

optim(c(-1,-1,-1), f)$par
## [1] 0.9997731731 0.0005162083 0.0009457322

optim(c(-5,-5,-5), f)$par
## [1] 1.0000614299 -0.0004356825 -0.0009109466

optim(c(0.6,0.6,0.6), f)$par
## [1] 9.999539e-01 -3.039159e-05 -2.508388e-05

```

```

## problem 3

set.seed(1)
n <- 100
beta0 <- 1
beta1 <- 2
sigma2 <- 6

x <- runif(n)
yComplete <- rnorm(n, beta0 + beta1*x, sqrt(sigma2))

## parameters chose such that signal in data is moderately strong
## estimate divided by std error is ~ 3
mod <- lm(yComplete ~ x)
summary(mod)$coef

##              Estimate Std. Error t value    Pr(>|t|)
## (Intercept) 0.5607442  0.5041346  1.112290 0.268734381
## x           2.7650812  0.8657927  3.193699 0.001889262

set.seed(0)
initial <- runif(3,-50, 50)
optim(par = initial, fn = f, control = list(trace= TRUE))

## Nelder-Mead direct search function minimizer
## function value for initial parameters = 217653.451008
## Scaled convergence tolerance is 0.00324329
## Stepsize computed as 3.966972
## BUILD           4 250164.956010 201295.454430
## EXTENSION        6 217653.451008 137712.813236
## LO-REDUCTION     8 209669.914815 137712.813236
## EXTENSION       10 201295.454430 86792.293623
## EXTENSION       12 148081.612739 22489.608231
## LO-REDUCTION    14 137712.813236 22489.608231
## REFLECTION      16 86792.293623 4471.532463
## REFLECTION      18 24789.301328 4188.613708
## REFLECTION      20 22489.608231 3725.104639
## HI-REDUCTION    22 4471.532463 2728.136602
## HI-REDUCTION    24 4188.613708 1351.920889

```

## LO-REDUCTION	26	3725.104639	358.565738
## HI-REDUCTION	28	2728.136602	358.565738
## HI-REDUCTION	30	1351.920889	358.565738
## LO-REDUCTION	32	1058.050465	285.049719
## HI-REDUCTION	34	470.535467	46.128439
## REFLECTION	36	358.565738	40.213717
## HI-REDUCTION	38	285.049719	40.213717
## HI-REDUCTION	40	92.636019	40.213717
## HI-REDUCTION	42	66.190678	24.409538
## HI-REDUCTION	44	46.128439	7.939230
## HI-REDUCTION	46	40.213717	7.939230
## HI-REDUCTION	48	24.409538	7.731708
## HI-REDUCTION	50	12.117106	6.777890
## HI-REDUCTION	52	7.939230	2.855432
## HI-REDUCTION	54	7.731708	2.048930
## HI-REDUCTION	56	6.777890	1.519195
## LO-REDUCTION	58	2.855432	1.519195
## REFLECTION	60	2.048930	1.318377
## HI-REDUCTION	62	1.834436	0.666956
## HI-REDUCTION	64	1.519195	0.666956
## HI-REDUCTION	66	1.318377	0.666956
## HI-REDUCTION	68	0.850947	0.666956
## HI-REDUCTION	70	0.736255	0.602566
## HI-REDUCTION	72	0.684598	0.579415
## HI-REDUCTION	74	0.666956	0.571048
## HI-REDUCTION	76	0.602566	0.545298
## HI-REDUCTION	78	0.579415	0.545298
## LO-REDUCTION	80	0.571048	0.543037
## REFLECTION	82	0.553831	0.527255
## HI-REDUCTION	84	0.545298	0.527255
## REFLECTION	86	0.543037	0.524988
## HI-REDUCTION	88	0.530530	0.524988
## EXTENSION	90	0.527255	0.509661
## LO-REDUCTION	92	0.526737	0.509661
## REFLECTION	94	0.524988	0.506376
## REFLECTION	96	0.516290	0.499993
## LO-REDUCTION	98	0.509661	0.499993
## EXTENSION	100	0.506376	0.489413
## LO-REDUCTION	102	0.503390	0.489413
## EXTENSION	104	0.499993	0.472072
## LO-REDUCTION	106	0.493135	0.472072
## LO-REDUCTION	108	0.489413	0.472072
## EXTENSION	110	0.477277	0.445537
## LO-REDUCTION	112	0.474771	0.445537
## LO-REDUCTION	114	0.472072	0.445537
## EXTENSION	116	0.450120	0.403716
## LO-REDUCTION	118	0.445690	0.403716
## LO-REDUCTION	120	0.445537	0.403716
## REFLECTION	122	0.419721	0.395192
## EXTENSION	124	0.408460	0.365805
## LO-REDUCTION	126	0.403716	0.365805
## LO-REDUCTION	128	0.395192	0.365805
## EXTENSION	130	0.372157	0.316378


```

## LO-REDUCTION      132 0.368095 0.316378
## LO-REDUCTION      134 0.365805 0.316378
## EXTENSION         136 0.327440 0.233956
## LO-REDUCTION      138 0.317164 0.233956
## LO-REDUCTION      140 0.316378 0.233956
## EXTENSION         142 0.262181 0.164547
## EXTENSION         144 0.237841 0.144745
## EXTENSION         146 0.233956 0.101887
## HI-REDUCTION      148 0.164547 0.101887
## EXTENSION         150 0.149986 0.077142
## HI-REDUCTION      152 0.144745 0.077142
## EXTENSION         154 0.105432 0.029256
## HI-REDUCTION      156 0.101887 0.029256
## LO-REDUCTION      158 0.077142 0.029256
## EXTENSION         160 0.067772 0.001110
## LO-REDUCTION      162 0.039595 0.001110
## LO-REDUCTION      164 0.029256 0.001110
## LO-REDUCTION      166 0.016071 0.001110
## LO-REDUCTION      168 0.010525 0.001110
## Exiting from Nelder Mead minimizer
##      170 function evaluations used
## $par
## [1] 0.998334136 -0.004430611 -0.008296619
##
## $value
## [1] 0.0004951965
##
## $counts
## function gradient
##      170      NA
##
## $convergence
## [1] 0
##
## $message
## NULL

optim(par = initial, fn = f, control = list(trace= TRUE), method = 'BFGS' )

## initial value 217653.451008
## iter 10 value 816.973113
## iter 20 value 6.060495
## iter 30 value 0.248793
## iter 40 value 0.000000
## final value 0.000000
## converged
## $par
## [1] 1.000000e+00 2.849319e-13 4.334081e-13
##
## $value
## [1] 3.504477e-25
##
## $counts
## function gradient

```

```
##      113      42
##
## $convergence
## [1] 0
##
## $message
## NULL

nlm(f, p= initial)

## $minimum
## [1] 1.700918e-08
##
## $estimate
## [1] 0.9999994970 -0.0000822318 -0.0001300803
##
## $gradient
## [1] 4.612353e-08 3.302012e-08 -2.128825e-08
##
## $code
## [1] 1
##
## $iterations
## [1] 28
```

Problem 3

b)

In order to find a reasonable starting value, we can ignore the censored data and run the regression of the uncensored values. We used the estimated parameter we got from the regression as our starting values.

c)

```
set.seed(1)
n <- 100
beta0 <- 1
beta1 <- 2
sigma2 <- 6

x <- runif(n)
yComplete <- rnorm(n, beta0 + beta1*x, sqrt(sigma2))

## parameters chose such that signal in data is moderately strong
## estimate divided by std error is ~ 3
mod <- lm(yComplete ~ x)
summary(mod)$coef

##              Estimate Std. Error  t value    Pr(>|t|)
## (Intercept) 0.5607442   0.5041346  1.112290 0.268734381
## x           2.7650812   0.8657927  3.193699 0.001889262
```

```

summary(mod)$sigma

## [1] 2.305117

ycomplete1 <- yComplete

myfunction <- function(x, y, p, criteria){
  #censore the data based on the percentile we want to use
  ycomplete1[yComplete > quantile(yComplete, p)] <- unname(quantile(yComplete, p))
  #calculate the initial parameter value
  initial <- lm(ycomplete1~x)
  beta1 <- unname(initial$coefficients[2])
  beta0 <- unname(initial$coefficients[1])
  sigma <- summary(initial)$sigma
  n <- length(yComplete)

  # Threshold
  tau <- unname(quantile(yComplete, p))
  difference = 10
  i = 0
  while( difference > criteria){
    #print(c(beta1,beta0,sigma, i))
    # Calculate expected value and variance of truncated normal distribution
    mu <- beta0 + beta1*x
    tau2 <- (tau - mu) / sigma
    rho <- dnorm(tau2) / (1 - pnorm(tau2))
    Expectation <- mu + sigma*rho
    Var <- sigma^2 *(1 + tau2*rho - (rho)^2)

    #step of maximization
    ycomplete1[yComplete > tau]<- Expectation[yComplete > tau]
    #Do the regression again after the modification to calculate the coefficient beta
    Update <- lm(ycomplete1 ~ x)
    beta1_new <- unname(Update$coefficients[2])
    beta0_new <- unname(Update$coefficients[1])
    sigma_new <- sqrt(sum(Var[yComplete > tau])/n + sum((ycomplete1 - beta0_new - beta1_new * x)^2)/n)

    #calculate the difference of the parameter to check if the parameter converge
    diff_beta1 <- abs(beta1 - beta1_new)
    diff_beta0 <- abs(beta0 - beta0_new)
    diff_sigma <- abs(sigma - sigma_new)

    difference <- (diff_beta1 + diff_beta0 + diff_sigma)

    #assign the new parameter to the parameter will be used in next iteration
    beta1 <- beta1_new
    beta0 <- beta0_new
    sigma <- sigma_new

    i = i + 1
  }
}

```

```

}
return(c(beta0_new, beta1_new, sigma_new, i-1))
}

result1 <- myfunction(x,y,0.2,0.00000001)
cat("Beta0: ",result1[1],'Beta1: ', result1[2], 'sigma: ', result1[3], 'iteration: ', result1[4])

## Beta0:  0.3126394 Beta1:  2.87922 sigma:  1.960093 iteration:  208

result2 <- myfunction(x,y,0.8,0.00000001)
cat("Beta0: ",result2[1],'Beta1: ', result2[2], 'sigma: ', result2[3], 'iteration: ', result2[4])

## Beta0:  0.4566128 Beta1:  2.824108 sigma:  2.149157 iteration:  16

```

The complete data's beta0, beta1, and sigma are 0.5607422, 2.7650812, and 2.305117 respectively. when a modest 20 percent proportion of exceedances expected, beta0, beta1, and sigma are 0.3126394, 2.87922, and 1.960093 respectively. I did 208 times iteration. when a high 80 percent proportion of exceedances expected, beta0, beta1, and sigma are 0.4566128, 2.824108, 2.149157 respectively. I only did 16 times iteration. I choose the difference criteria as 0.0000001, because I think it small enough.

d)

```

myfunction_logL <- function(x, y, p=0.8){
  ycomplete1[yComplete > quantile(yComplete, p)] <- unname(quantile(yComplete, p))
  #calculate the initial parameter value
  initial <- lm(ycomplete1~x)
  beta1 <- unname(initial$coefficients[2])
  beta0 <- unname(initial$coefficients[1])
  sigma <- summary(initial)$sigma
  n <- length(x)
  tau <- unname(quantile(yComplete, p))
  mu <- rep(beta0, n) + beta1*x
  y <- (yComplete[1:(n*(1-p))] - mu[1:(n*(1-p))]) / sqrt(sigma)
  return(n*(1-p)*log(sqrt(sigma)) + sum(y^2)/2 - sum(log(1 - pnorm((tau - mu[(n*(1-p))+1):n]) / sqrt(s
}
  ycomplete1[yComplete > quantile(yComplete, 0.8)] <- unname(quantile(yComplete, 0.8))
  #calculate the initial parameter value
  initial <- lm(ycomplete1~x)
  par <- c(unname(initial$coefficients[2]), unname(initial$coefficients[1]),summary(initial)$sigma)
  optim(par, fn = myfunction_logL, method = "BFGS", x = x,y =ycomplete1)

## Error in quantile.default(yComplete, p): 'probs' outside [0,1]

```