

Continuous Random Variables (I)

Let X be a continuous random variable.

i.e. the cdf F_X is a continuous function.

Let $f_X(x) \rightarrow$ probability density function (pdf) of X

$$\text{satisfies: } F_X(x) = \int_{-\infty}^x f_X(t) dt.$$

Recall:

- ① $f_X(x)$ must satisfy:
 - (i) $f_X(x) \geq 0$ for all $x \in (-\infty, \infty)$
 - (ii) $\int_{-\infty}^{\infty} f_X(x) dx = 1$.

② The expected value of X is

$$E(X) = \mu_X = \int_{-\infty}^{\infty} x f_X(x) dx$$

③ The variance of X is

$$V(X) = \sigma_X^2 = \int_{-\infty}^{\infty} (x - \mu_X)^2 f_X(x) dx.$$

④ The standard deviation of X is

$$\sigma_X = \sqrt{\sigma_X^2}$$

Important Identities

$$\text{① } E(h(X)) = \int_{-\infty}^{\infty} h(x) f_X(x) dx$$

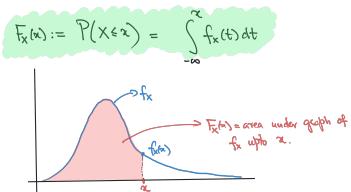
$$E(aX+b) = aE(X)+b, \quad V(aX+b) = a^2 V(X)$$

②

$$V(X) = E(X^2) - (E(X))^2$$

Calculating Probabilities

The cdf is defined as



Note:

$$\begin{aligned} \text{① } P(a \leq X \leq b) &= \int_a^b f_X(x) dx \quad (a < b) \\ &= \int_{-\infty}^b f_X(x) dx - \int_{-\infty}^a f_X(x) dx \\ &= F_X(b) - F_X(a) \end{aligned}$$

$$\text{② } P(X=c) = P(c \leq X \leq c) = F_X(c) - F_X(c) = 0$$

Probability that a continuous rr take a particular value is zero !!!

$$\text{③ } P(X > c) = 1 - P(X \leq c) = 1 - F_X(c)$$

$$\begin{aligned} P(X \leq c) &= P(X < c) + P(X=c) \\ &= P(X < c) \end{aligned}$$

We can use the cdf F_X to calculate $f_X(x)$

If the derivative $F'_X(x)$ exist at x

$$F'_X(x) = f_X(x)$$

Let $p \in (0,1) \rightarrow$ the " $(100p)\text{th}$ " Percentile for the distribution of X

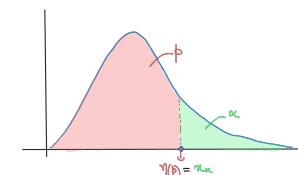
$\eta(p)$ and satisfies

$$p = \int_{-\infty}^{\eta(p)} f_X(x) dx$$

Let $\alpha \in (0,1) \rightarrow$ the α^{th} critical value for the distribution of X is

α and satisfies

$$\alpha = P(X > \alpha) = 1 - F_X(\alpha)$$



Note: $(100p)\text{th}$ percentile \rightarrow some α \rightarrow $(1-p)^{th}$ critical value.

Example: Uniform Continuous Distribution

Let $[A,B] \subseteq \mathbb{R}$ be a closed interval. ($A < B$)

$$\text{Define: } f(x) = \begin{cases} \frac{1}{B-A} & x \in [A,B] \\ 0 & \text{otherwise} \end{cases} \quad \text{pdf of the uniform cont distribution}$$

then

$$F(x) = \begin{cases} 0 & \text{if } x < A \\ \frac{x-A}{B-A} & \text{if } x \in [A,B] \\ 1 & \text{if } x > B. \end{cases} \quad \text{cdf of the uniform cont distribution.}$$

Let X have the uniform distribution on $[A,B]$

$$\begin{aligned} \text{then } E(X) &= \int_{-\infty}^{\infty} x f(x) dx = \int_A^B x \frac{1}{B-A} dx \\ &= \frac{(B+A)}{2} \\ V(X) &= \int_{-\infty}^{\infty} (x - \mu_X)^2 f(x) dx = \int_A^B (x - \frac{B+A}{2})^2 \frac{1}{B-A} dx \\ &= \frac{(B-A)^2}{12} \end{aligned}$$

If $a, b \in [A,B]$ and $a < b$

$$P(a \leq X \leq b) = \int_a^b \frac{1}{B-A} dx = \frac{(b-a)}{(B-A)}.$$

Suppose X is a r.v.

$F_X \rightarrow$ cdf of X

$$F_X(x) = P(X \leq x)$$

If $F_X(x)$ is a continuous function

↓

X is a continuous r.v.

The probability density function, $f_X(x)$ satisfies

$$F_X(x) = \int_{-\infty}^x f_X(t) dt$$

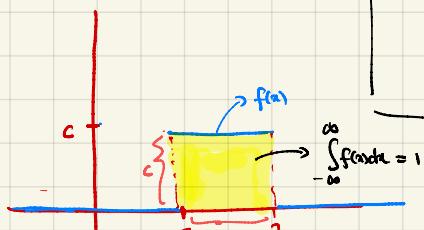
$f_X(x)$ satisfies (i) $f_X(x) \geq 0 \forall x \in (-\infty, \infty)$

$$(ii) \quad \int_{-\infty}^{\infty} f_X(x) dx = 1$$

Example: ① Uniform continuous distribution

Suppose $[3, 7] \subseteq \mathbb{R}$

$$f(x) = \begin{cases} c & x \in [3, 7] \\ 0 & \text{otherwise} \end{cases}$$



If f has to be a pdf:

(i) $f(x) \geq 0$ for all $x \in (-\infty, \infty)$

$$(ii) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\text{Area} = 4C = \int_{-\infty}^{\infty} f(x) dx$$

$$\therefore \int_{-\infty}^{\infty} f(x) dx = 1 \Rightarrow 4C = 1$$

$$\Rightarrow C = \frac{1}{4}$$

$$f(x) = \begin{cases} \frac{1}{4} & x \in [3, 7] \\ 0 & \text{otherwise.} \end{cases}$$

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^3 0 dx + \int_3^7 c dx + \int_7^{\infty} 0 dx \\ &= \int_3^7 c dx = c(7-3) = \underline{\underline{4c}} \end{aligned}$$

(i) from 4.1 Devore

$$f(x) = \begin{cases} \frac{x}{\theta^2} \cdot e^{-x^2/2\theta^2} & x > 0 \\ 0 & \text{otherwise. } (x \leq 0) \end{cases}$$

Use R to plot this curve

raylength ← function ($x, theta$) {

return(π/θ .) }

(i) Show that $f(x)$ is a pdf.

(a) $f(x) \geq 0 \quad \forall x \in (-\infty, \infty)$

✓ $x \sim \text{exp}(1, 100, \text{length.out}=1000)$
 $y = \text{raylength}(x, \theta=40)$

$$(b) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^0 0 dx + \int_0^{\infty} \frac{x}{\theta^2} \cdot e^{-x^2/2\theta^2} dx$$

improper integral of type II.

$$= \lim_{c \rightarrow \infty} \int_0^c \left(\frac{x}{\theta^2} e^{-x^2/2\theta^2} \right) dx$$

$$\int_0^c \frac{x}{\theta^2} e^{-x^2/2\theta^2} dx$$

$$u = \frac{x^2}{2\theta^2} \quad \begin{array}{l} u=0 \\ u=c \end{array}$$

$$du = \frac{2x}{2\theta^2} dx \quad \begin{array}{l} x=0 \\ x=c \end{array}$$

$$\int_0^{c/2\theta^2} e^{-u} du = \frac{e^{-u}}{-1} \Big|_0^{c^2/2\theta^2} = \frac{e^{-c^2/2\theta^2}}{2\theta^2}$$

$$= \underline{\underline{(1 - e^{-c^2/2\theta^2})}}$$

$$\int_{-\infty}^{\infty} f(x) dx = \lim_{c \rightarrow \infty} \left(1 - e^{-\frac{c^2}{2\sigma^2}} \right)$$

= 1

Example

$$E(X) = \int_{-\infty}^{+\infty} x \cdot \frac{1}{4} dx$$

$$= \frac{1}{4} \cdot \left[\frac{x^2}{2} \right]_{-3}^{+3} = \frac{1}{4} \cdot \frac{(7^2 - 3^2)}{2}$$

$$= \frac{(7-3)(7+3)}{4 \cdot 2} = \frac{4 \cdot (7+3)}{4 \cdot 2} = \frac{7+3}{2} = \underline{\underline{5}}$$

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx \longrightarrow \text{expected value of } X$$

Useful formula

$$V(X) = E(X^2) - (E(X))^2$$

$$\text{Var}(X) = \int_{-\infty}^{\infty} (x - \mu_X)^2 \cdot f(x) dx \longrightarrow \text{Variance of } X.$$

$f(x) \rightsquigarrow \text{pdf of } X.$

Example:

$$f(x) = \begin{cases} \frac{1}{4} & x \in [3, 7] \\ 0 & \text{otherwise.} \end{cases}$$

$$F_X(x) = \int_{-\infty}^x f(t) dt = \int_{-\infty}^3 0 dt + \int_3^x \frac{1}{4} dt$$

$$= \frac{x-3}{4} = \underline{\underline{F_X(x)}}$$

