### **STA 371G: Statistics and Modeling**

# Simple Linear Regression: Covariance and Correlation, Goodness of Fit

Mingyuan Zhou McCombs School of Business The University of Texas at Austin

http://mingyuanzhou.github.io/teaching

### Sample Mean and Sample Variance

Sample Mean: measure of centrality

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^{n} Y_i$$

Sample Variance: measure of spread

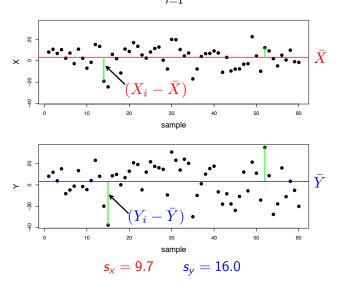
$$s_y^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

Sample Standard Deviation:

$$s_y = \sqrt{s_y^2}$$

### Example

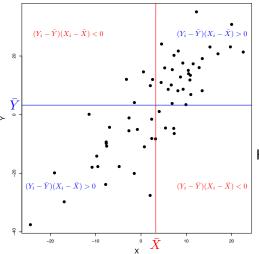
$$s_y^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$



#### Sample Covariance

Measure the direction and strength of the linear relationship between Y and X

$$Cov(Y, X) = \frac{\sum_{i=1}^{n} (Y_i - \bar{Y})(X_i - \bar{X})}{n-1}$$



- $s_v = 15.98, s_x = 9.7$
- $\sim Cov(X, Y) = 125.9$

How do we interpret that?

Correlation is the standardized covariance:

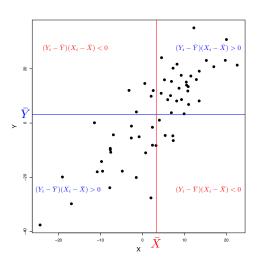
$$\operatorname{corr}(X,Y) = \frac{\operatorname{cov}(X,Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\operatorname{cov}(X,Y)}{s_x s_y}$$

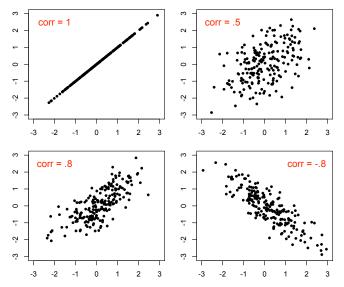
The correlation is scale invariant and the units of measurement don't matter: It is always true that  $-1 \le corr(X, Y) \le 1$ .

$$\operatorname{corr}(aX, bY) = \frac{\operatorname{cov}(aX, bY)}{s_{ax}s_{by}} = \frac{ab \operatorname{cov}(X, Y)}{|a||b| s_x s_y} = \operatorname{sign}(ab)\operatorname{corr}(X, Y)$$

This gives the direction (- or +) and strength  $(0 \rightarrow 1)$  of the linear relationship between X and Y.

$$corr(Y, X) = \frac{cov(X, Y)}{\sqrt{s_x^2 s_y^2}} = \frac{cov(X, Y)}{s_x s_y} = \frac{125.9}{15.98 \times 9.7} = 0.812$$

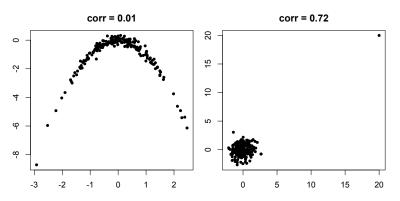




Only measures linear relationships:

corr(X, Y) = 0 does not mean the variables are not related!

Example: corr(X, Y)=0 if  $Y = X^2, X \sim \mathcal{N}(0, 1)$ .



Also be careful with influential observations. Excel Break: correl, stdev,...

### Back to Least Squares

#### 1. Intercept:

$$b_0 = \bar{y} - b_1 \bar{x} \Rightarrow \bar{y} = b_0 + b_1 \bar{x}$$

- ▶ The point  $(\bar{x}, \bar{y})$  is on the regression line!
- Least squares finds the point of means and rotate the line through that point until getting the "right" slope

#### 2. Slope:

$$b_1 = corr(X, Y) \times \frac{s_y}{s_x} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$
$$= \frac{Cov(X, Y)}{var(X)}$$

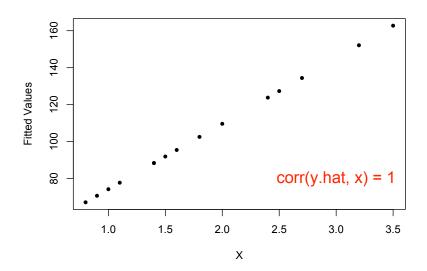
► So, the right slope is the *correlation coefficient* times a *scaling* factor that ensures the proper units for b<sub>1</sub> 9

#### More on Least Squares

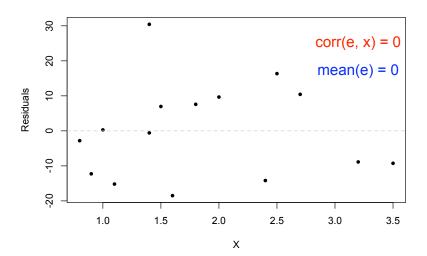
From now on, terms "fitted values"  $(\hat{y}_i)$  and "residuals"  $(e_i)$  refer to those obtained from the least squares line.

The fitted values and residuals have some special properties. Lets look at the housing data analysis to figure out what these properties are...

#### The Fitted Values and X

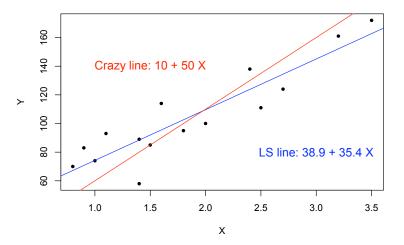


#### The Residuals and X



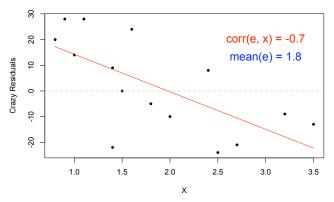
### Why?

What is the intuition for the relationship between  $\hat{Y}$  and e and X? Lets consider some "crazy" alternative line:



#### Fitted Values and Residuals

This is a bad fit! We are underestimating the value of small houses and overestimating the value of big houses.



Clearly, we have left some predictive ability on the table!

#### Fitted Values and Residuals

As long as the correlation between e and X is non-zero, we could always adjust our prediction rule to do better.

We need to exploit all of the predictive power in the X values and put this into  $\hat{Y}$ , leaving no "Xness" in the residuals.

### In Summary: $Y = \hat{Y} + e$ where:

- $\hat{Y}$  is "made from X";  $\operatorname{corr}(X, \hat{Y}) = 1$ .
- e is unrelated to X;  $\operatorname{corr}(X, e) = 0$ .
- $(\bar{X}, \bar{Y})$  is on the regression line.
- $\bar{e} = \frac{\sum_{i=1}^{n} e_i}{n} = 0.$
- $ightharpoonup corr(\hat{Y}, e) = ?$

#### Decomposing the Variance

How well does the least squares line explain variation in Y?

Remember that  $Y=\hat{Y}+e$ ;  $\hat{Y}$  and e are uncorrelated, i.e.  $\mathrm{corr}(\hat{Y},e)=0$ ; and  $\bar{e}=0$ .

$$\sum_{i=1}^{n} (Y_i - \bar{Y})^2 = \sum_{i=1}^{n} [(Y_i - \hat{Y}_i) + (\hat{Y}_i - \bar{Y})]^2$$

$$= \sum_{i=1}^{n} [e_i + (\hat{Y}_i - \bar{Y})]^2$$

$$= \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^{n} e_i^2 + 2 \sum_{i=1}^{n} e_i (\hat{Y}_i - \bar{Y})$$

$$= \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^{n} e_i^2$$

#### Decomposing the Variance

$$\sum_{\substack{i=1\\ \text{Total Sum of }\\ \text{Squares }\\ \text{SST}}}^{n} (Y_i - \overline{Y})^2 = \sum_{\substack{i=1\\ \text{Regression SS}\\ \text{SSR}}}^{n} (\hat{Y}_i - \overline{Y})^2 + \sum_{\substack{i=1\\ \text{Error SS}\\ \text{SSE}}}^{n} e_i^2$$

SSR: Variation in Y explained by the regression line.

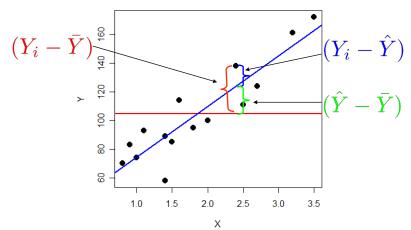
SSE: Variation in Y that is left unexplained.

$$SSR = SST \Rightarrow perfect fit.$$

Be careful of similar acronyms; e.g. SSR for "residual" SS.

#### Decomposing the Variance

$$(Y_i - \overline{Y}) = \hat{Y}_i + e_i - \overline{Y}$$
  
=  $(\hat{Y}_i - \overline{Y}) + e_i$ 



#### A Goodness of Fit Measure: $R^2$

The coefficient of determination, denoted by  $R^2$ , measures goodness of fit:

$$R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$$

- ▶  $0 < R^2 < 1$ .
- ▶ The closer R<sup>2</sup> is to 1, the better the fit.

## A Goodness of Fit Measure: $R^2$ (Optional)

An interesting fact:  $R^2 = r_{xy}^2$  (i.e.,  $R^2$  is squared correlation).

$$R^{2} = \frac{\sum_{i=1}^{n} (\hat{Y}_{i} - \bar{Y})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}}$$

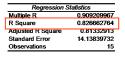
$$= \frac{\sum_{i=1}^{n} (b_{0} + b_{1}X_{i} - b_{0} - b_{1}\bar{X})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}}$$

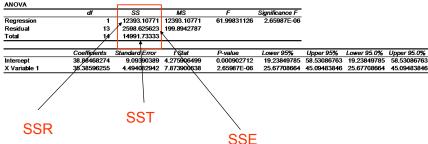
$$= \frac{b_{1}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}} = \frac{b_{1}^{2} s_{x}^{2}}{s_{y}^{2}} = r_{xy}^{2}$$

No surprise: the higher the sample correlation between X and Y, the better you are doing in your regression.

#### Back to the House Data

#### SUMMARY OUTPUT





$$R^2 = \frac{SSR}{SST} = 0.83 = \frac{12393}{12393 + 2599}$$

58 53086763

```
> HouseFit=lm(formula=Price~Size)
> summary(HouseFit)
Call:
lm(formula = Price ~ Size)
Residuals:
           10 Median 30 Max
   Min
-30.425 -8.618 0.575 10.766 18.498
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 38.885 9.094 4.276 0.000903 ***
           35.386 4.494 7.874 2.66e-06 ***
Size
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 14.14 on 13 degrees of freedom
Multiple R-squared: 0.8267. Adjusted R-squared: 0.8133
F-statistic: 62 on 1 and 13 DF, p-value: 2.66e-06
> anova(HouseFit)
Analysis of Variance Table
Response: Price
         Df Sum Sa Mean Sa F value Pr(>F)
         1 12393.1 12393.1 61.998 2.66e-06 ***
Residuals 13 2598.6 199.9
```

$$R^2 = \frac{SSR}{SST} = 0.83 = \frac{12393}{12393 + 2599}$$

```
> HouseFit=lm(formula=Price~Size)
> summary(HouseFit)
Call:
lm(formula = Price ~ Size)
Residuals:
           10 Median 30 Max
   Min
-30.425 -8.618 0.575 10.766 18.498
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 38.885 9.094 4.276 0.000903 ***
           35.386 4.494 7.874 2.66e-06 ***
Size
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 14.14 on 13 degrees of freedom
Multiple R-squared: 0.8267. Adjusted R-squared: 0.8133
F-statistic: 62 on 1 and 13 DF, p-value: 2.66e-06
> anova(HouseFit)
Analysis of Variance Table
Response: Price
         Df Sum Sa Mean Sa F value Pr(>F)
         1 12393.1 12393.1 61.998 2.66e-06 ***
Residuals 13 2598.6 199.9
```

$$R^2 = \frac{SSR}{SST} = 0.83 = \frac{12393}{12393 + 2599}$$

### Interpretation of $R^2$

- $ightharpoonup R^2$  measures the proportion of variation in Y explained by X.
- R<sup>2</sup> measures the reduction of SSE from
   Running regression without predictors X to
   Running regression with predictors X

#### Back to Baseball

Three very similar, related ways to look at a simple linear regression... with only one X variable, life is easy!

	$R^2$	corr	SSE
OBP	0.88	0.94	0.79
SLG	0.76	0.87	1.64
AVG	0.63	0.79	2.49