

The Front-Door Criterion in Theory and Practice

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Outline

MOTIVATION AND INTRODUCTION

FRONT-DOOR FOR ATT

FRONT-DOOR DIFFERENCE-IN-DIFFERENCES

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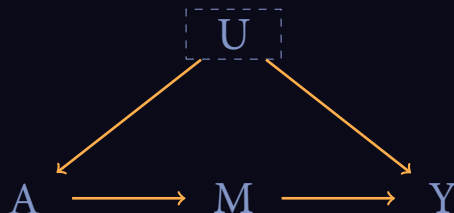
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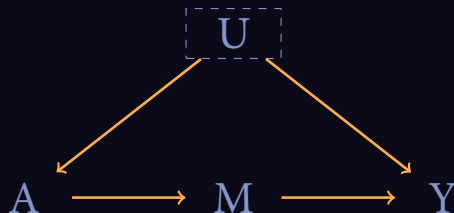
None of these techniques use information on the mechanisms by which the program is supposed to produce the effect (e.g. program attendance, learning, etc.).

The Front-door Approach



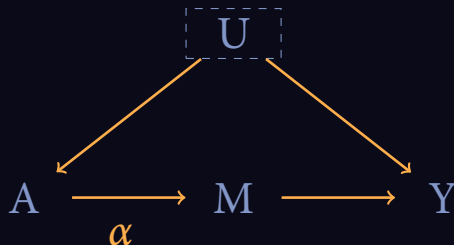
Pearl (1995) showed that post-treatment mechanism variables M can be used to non-parametrically identify causal effects when selection on observables doesn't hold.

Front-door Adjustment (easiest version)



With linear constant-effects models and singleton M , the front-door adjustment reduces to the following:

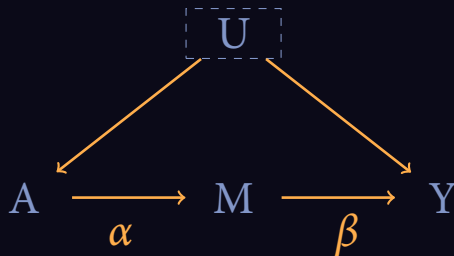
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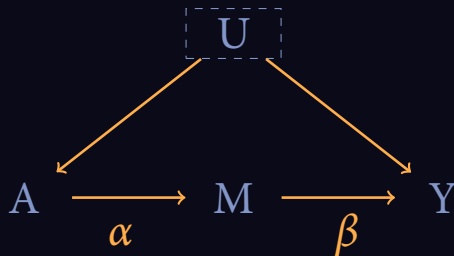
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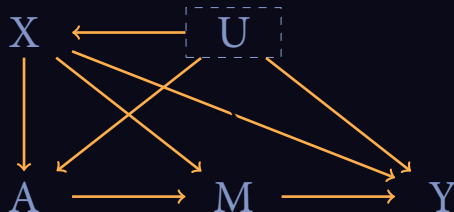
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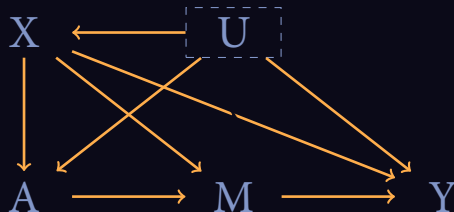
- ▶ Get coefficient on A from regression of M on A .
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- ▶ Multiply coefficients ($\alpha \cdot \beta$).

Front-door Adjustment (easiest, with covariates)



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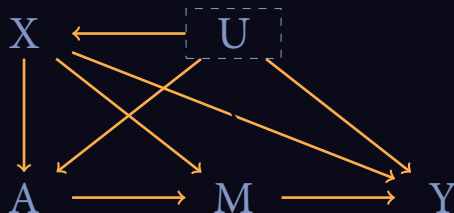
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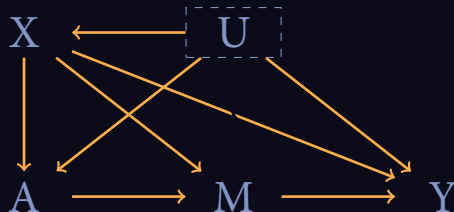
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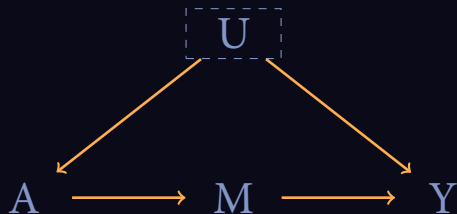
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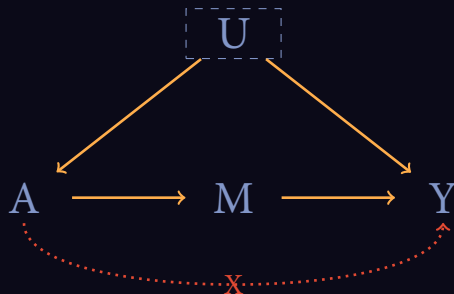
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- ▶ Multiply coefficients.

Front-door Criterion (Informal)

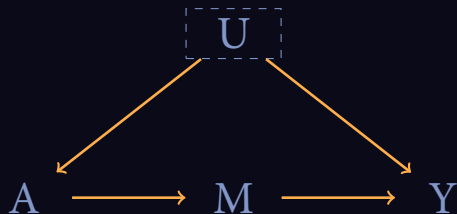


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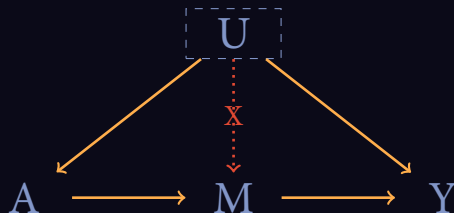


A affects Y only through M (e.g., M indicates program attendance).

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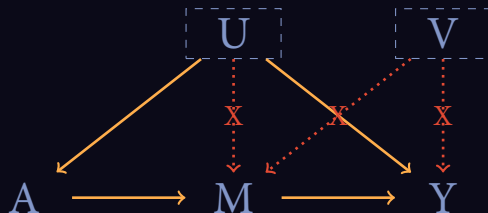


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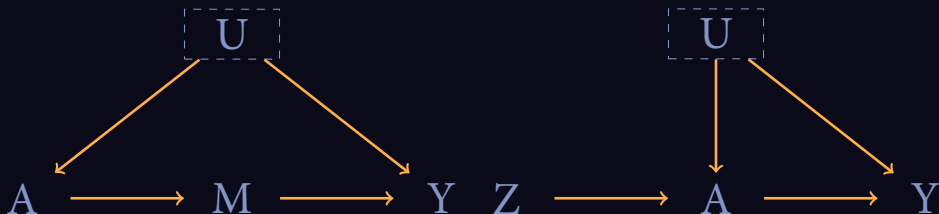
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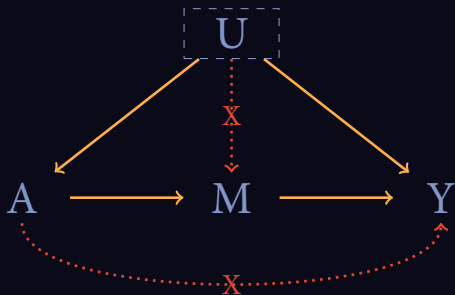
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Front-door vs Instrumental Variables

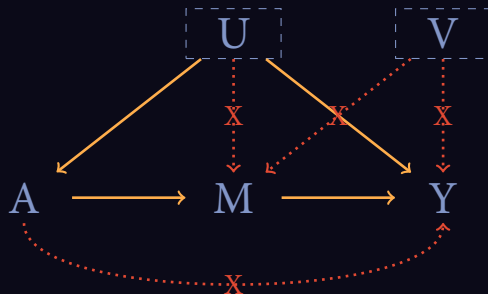


Why has front-door been so rarely used?

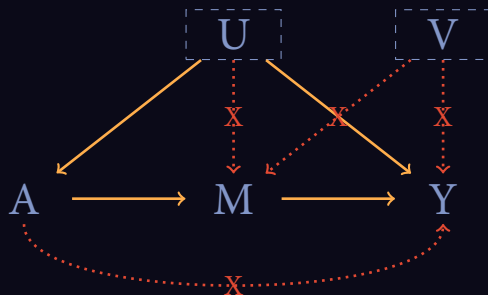
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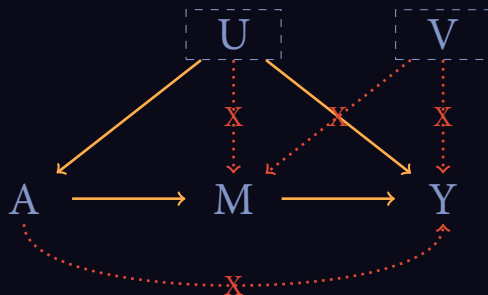


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Glynn and Kashin 2017 and 2018 show that front-door can be informative, even when conditions don't hold exactly.

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Front-door Estimator for ATT

Let x be a set of observed covariates, and u be a set of unobserved covariates, such that

$$\underbrace{E[Y(a_0)|a_1]}_{\text{Counterfactual}} = \sum_x \sum_u \underbrace{E[Y|a_0, x, u]}_{\text{Unobserved}} \cdot \underbrace{P(u|x, a_1)}_{\text{Unobserved}} \cdot P(x|a_1)$$

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Let m be a set of mediating variables. The front-door estimator for this counterfactual can be written as

$$\hat{E}_{fd}[Y(a_0)|a_1] = \sum_x \sum_m E[Y|a_1, x, m] \cdot P(m|x, a_0) \cdot P(x|a_1)$$

Identification

$$\sum_x \sum_u E[Y|a_0, x, u] \cdot P(u|x, a_1) \cdot P(x|a_1) = \sum_x \sum_m E[Y|a_1, x, m] \cdot P(m|x, a_0) \cdot P(x|a_1)$$

when the following assumptions hold

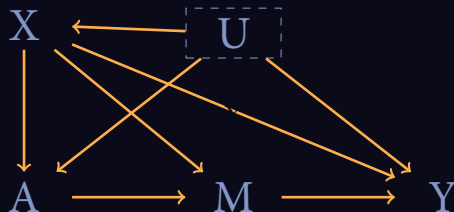
Assumption (1)

$$E[Y|a_0, x, u, m] = E[Y|a_1, x, u, m]$$

Assumption (2)

$$P(m|x, a_0) = P(m|x, a_0, u) \text{ and } P(u|x, a_1) = P(u|x, a_1, m)$$

Identification



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Develop intuition in a simple (canonical?) case

- ▶ M is binary variable that denotes receipt of treatment (e.g., m_1 denotes program attendance)
- ▶ Non-compliance for treated units

$$0 < P(m_0|a_1) < 1$$

- ▶ No non-compliance for control units

$$P(m_1|a_0) = 0$$

Standard and front-door estimators under one-sided noncompliance

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$$\widehat{ATT}_{\text{Standard}} = E[Y|a_1] - \sum_x \underbrace{E[Y|a_0, x]}_{\text{Controls}} \cdot P(x|a_1)$$

Standard and front-door estimators under one-sided noncompliance

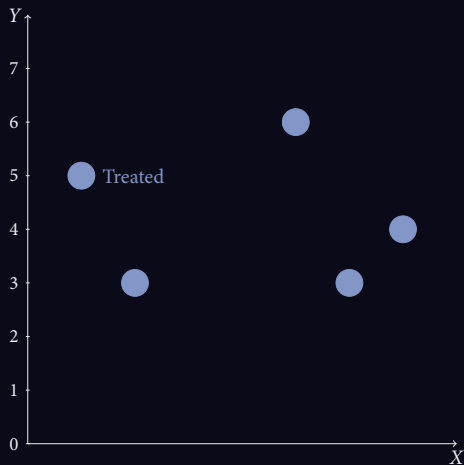
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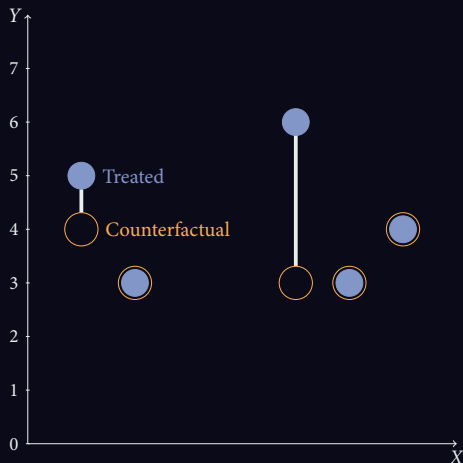
Glynn and Kashin (2018) shows that under these conditions the front-door estimator simplifies to the following:

$$\widehat{ATT}_{\text{Front-door}} = E[Y|a_1] - \sum_x \underbrace{E[Y|a_1, m_0, x]}_{\text{Treated non-compliers}} \cdot P(x|a_1)$$

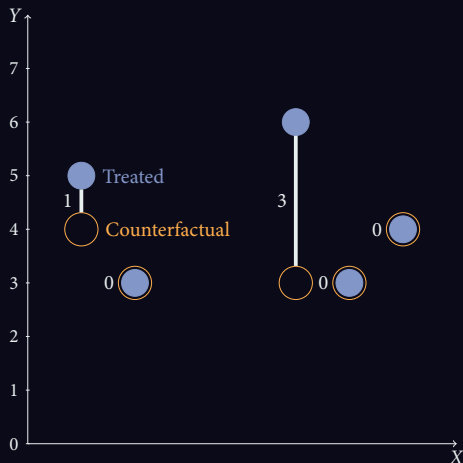
Visualizing the ATT



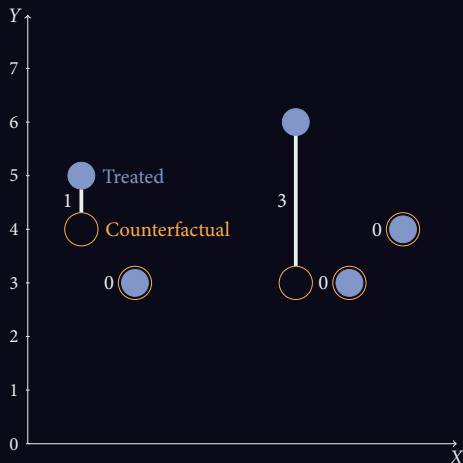
Each treated unit has an associated counterfactual



...and thus an individual causal effect

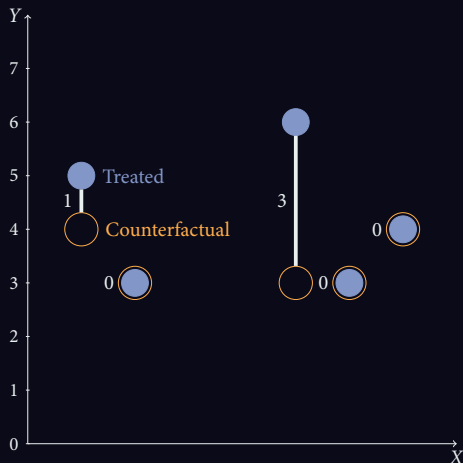


The ATT is an average of individual causal effects



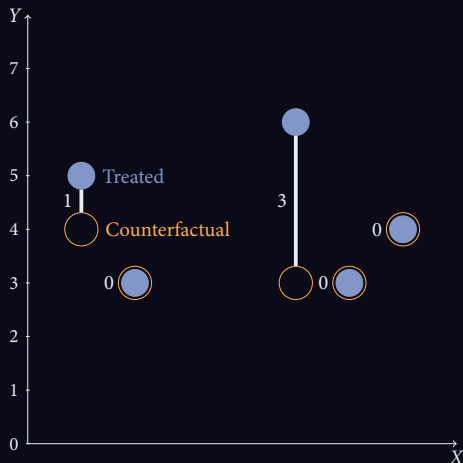
$$ATT = \frac{1 + 0 + 3 + 0 + 0}{5} = \frac{4}{5}$$

General expression for ATT



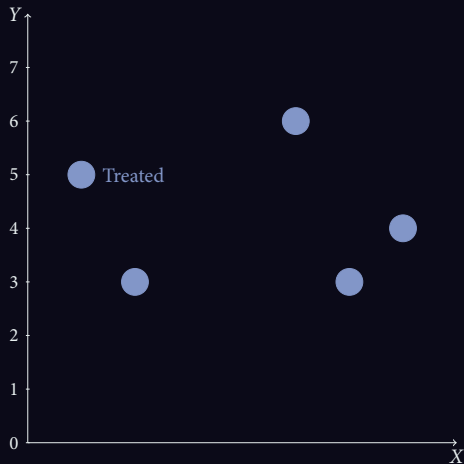
$$ATT = E[Y - Y(a_0)|a_1]$$

General expression for ATT

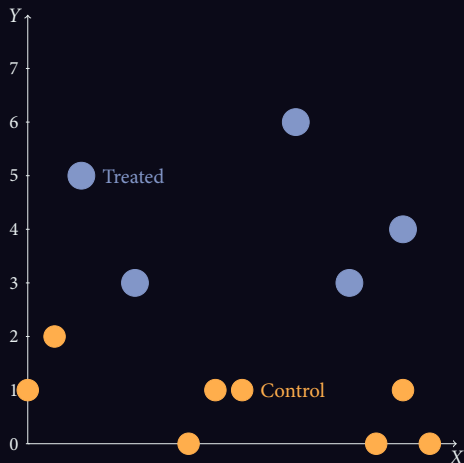


$$ATT = E[Y - Y(a_0)|a_1] = E[Y|a_1] - E[Y(a_0)|a_1]$$

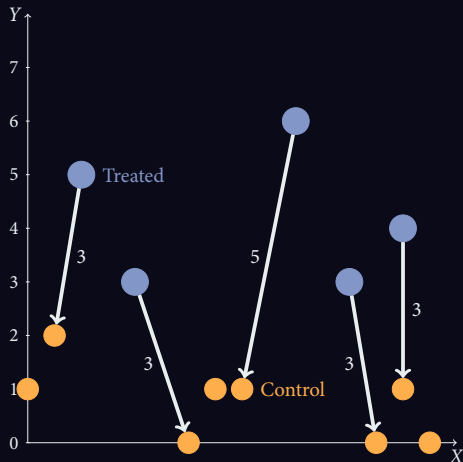
How does one usually estimate ATT?



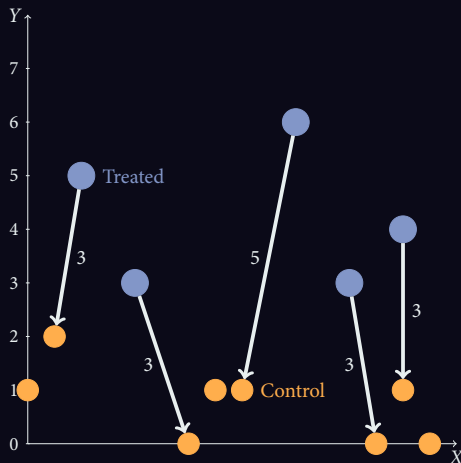
Find a group of control units



Perhaps match each treated unit to nearest control unit

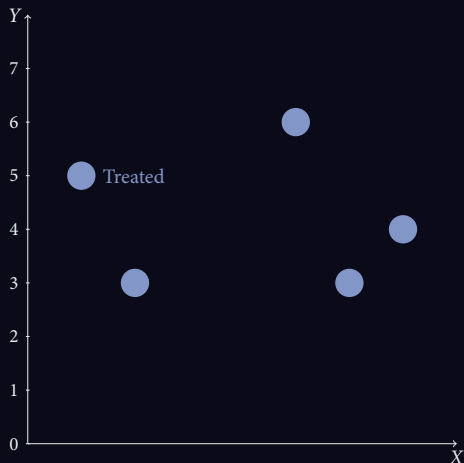


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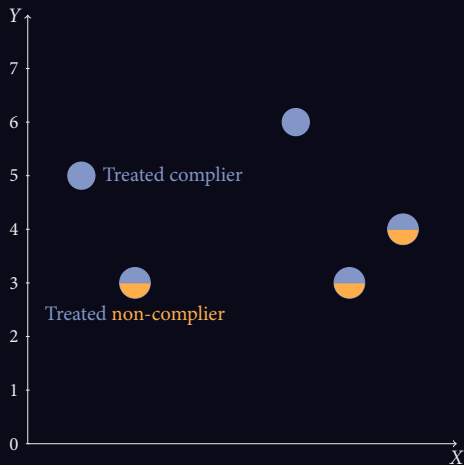


$$\widehat{ATT} = \frac{3 + 3 + 5 + 3 + 3}{5} = \frac{17}{5}$$

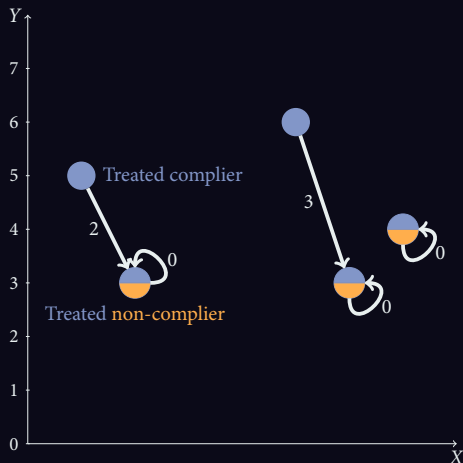
What if no (comparable) control group?



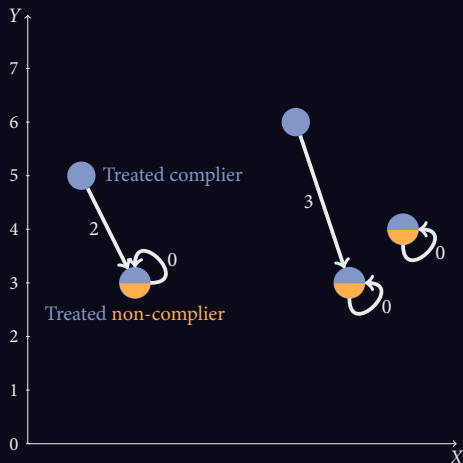
Observe compliance for each treated unit



Match each treated unit to nearest treated non-complier

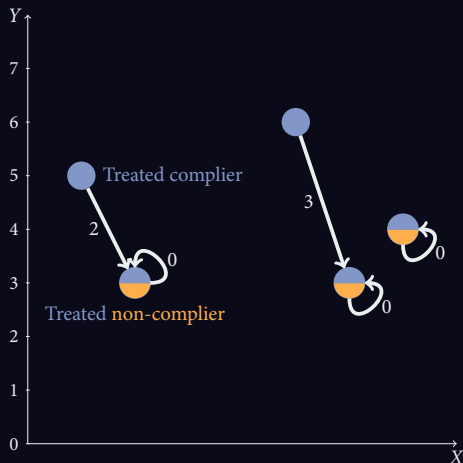


Front-door estimator (with one-sided noncompliance)



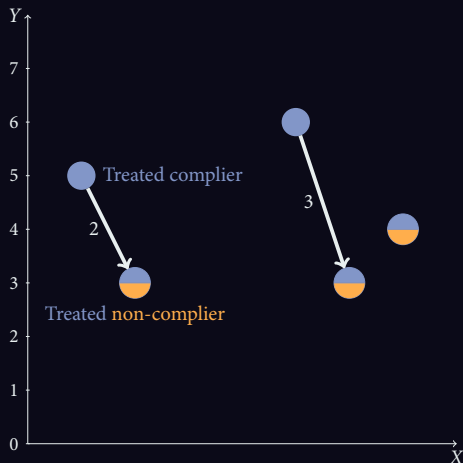
$$\widehat{ATT} = \frac{2 + 0 + 3 + 0 + 0}{5} = 1$$

Front-door estimator (with one-sided noncompliance)

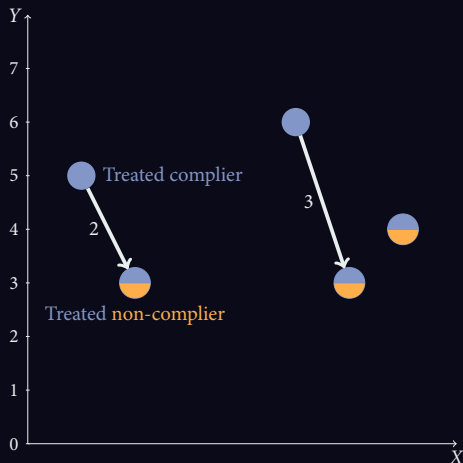


$$\widehat{ATT} = E[Y|a_1] - \sum_x E[Y|a_1, m_0, x] \cdot P(x|a_1)$$

“Effect of attendance” conceptualization

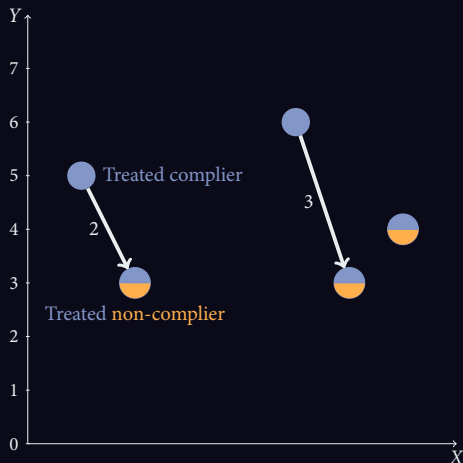


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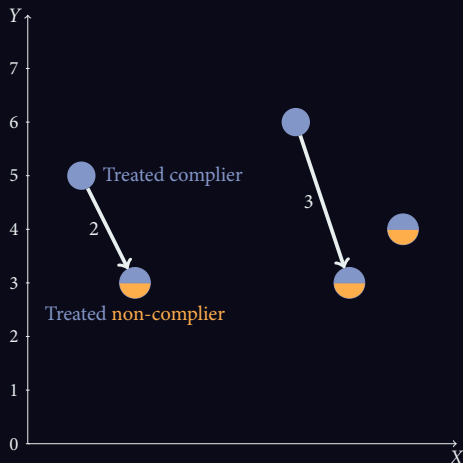
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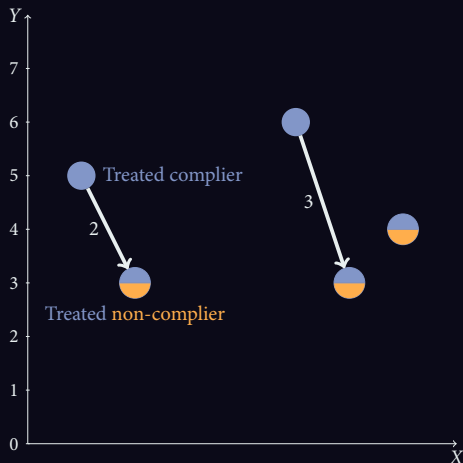
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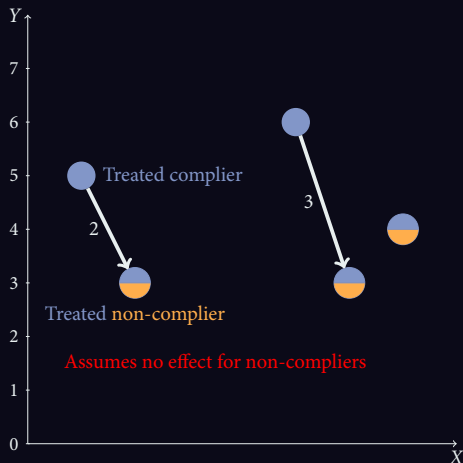
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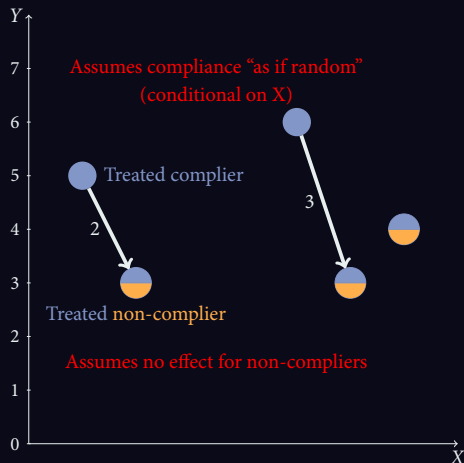
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EXAMPLES

SEE IF WE CAN RECOVER EXPERIMENTAL BENCHMARK.

The National JTPA Study

Job training evaluation program with experimental treatments and controls, compliance information, covariates, and an observational control group:

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Job training evaluation program with experimental treatments and controls, compliance information, covariates, and an observational control group:

- ▶ **Treatment:** allowed to receive JTPA services or not

Sign up

The National JTPA Study

Job training evaluation program with experimental treatments and controls, compliance information, covariates, and an observational control group:

- ▶ **Treatment:** allowed to receive JTPA services or not
- ▶ **Outcome:** 18-month post-program earnings

Sign up  Earnings

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Job training evaluation program with experimental treatments and controls, compliance information, covariates, and an observational control group:

- ▶ **Treatment:** allowed to receive JTPA services or not
- ▶ **Outcome:** 18-month post-program earnings
- ▶ **Compliance:** program participation (one-sided noncompliance)

Sign up —————> Show up —————> Earnings

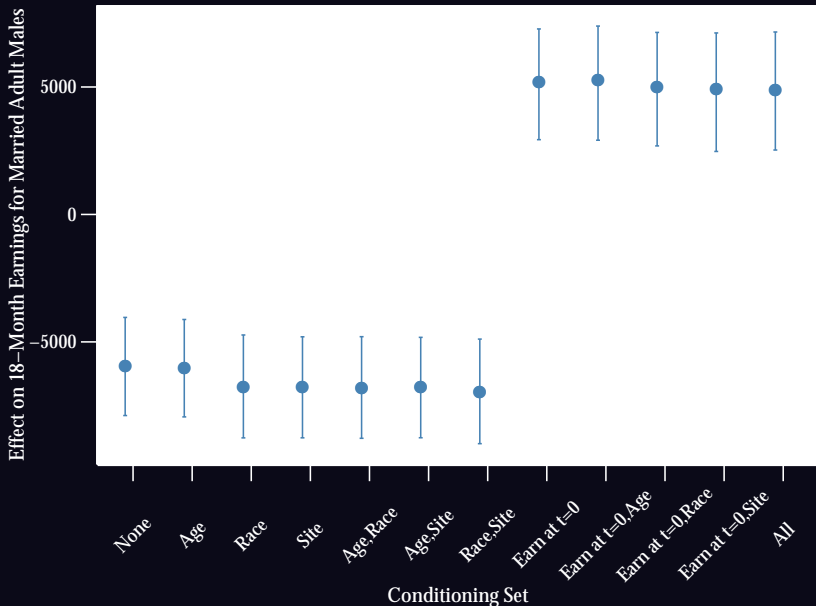
The National JTPA Study

Job training evaluation program with experimental treatments and controls, compliance information, covariates, and an observational control group:

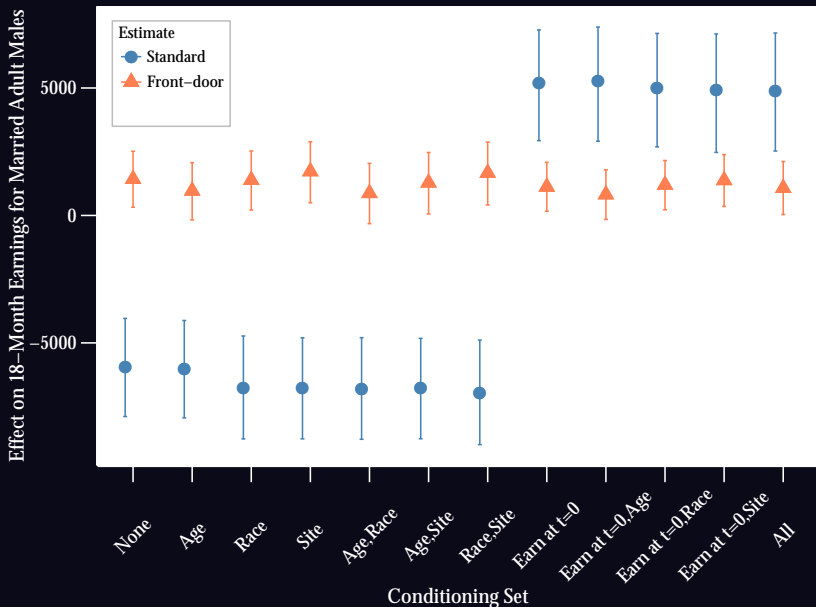
- ▶ **Treatment:** allowed to receive JTPA services or not
- ▶ **Outcome:** 18-month post-program earnings
- ▶ **Compliance:** program participation (one-sided noncompliance)
- ▶ **Group of Interest:** married adult males

Sign up —————> Show up —————> Earnings

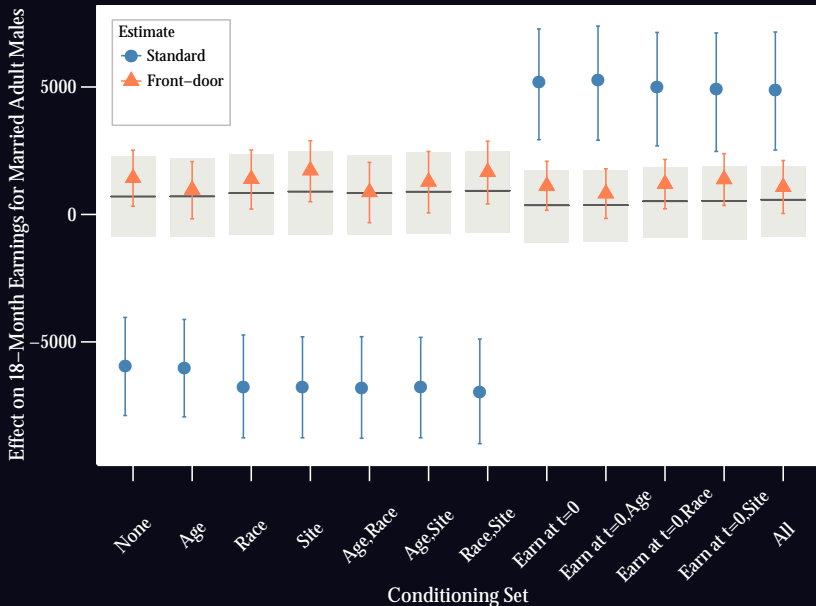
Are the standard estimates believable?



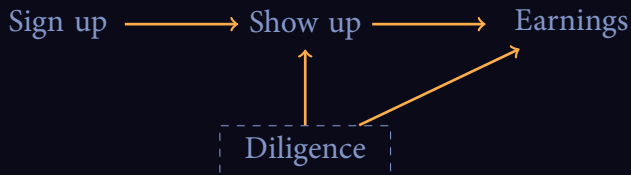
Which would you choose?



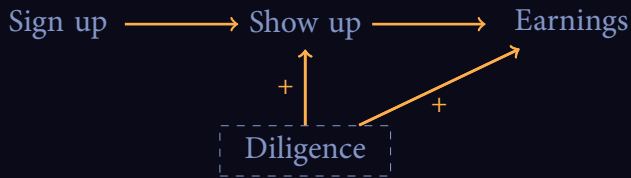
Front-door outperforms standard approach



Not surprising that front-door exhibits small positive bias



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Get-Out-the-Vote (GOTV) Phone Studies

Phone GOTV evaluation studies with experimental treatments and controls and compliance information:

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- **Treatment:** phone call encouraging turnout

Call

Get-Out-the-Vote (GOTV) Phone Studies

Phone GOTV evaluation studies with experimental treatments and controls and compliance information:

- ▶ **Treatment:** phone call encouraging turnout
- ▶ **Outcome:** vote vs not vote



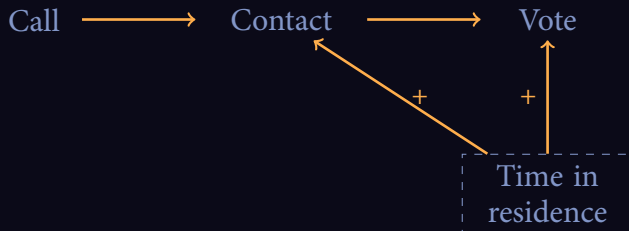
Get-Out-the-Vote (GOTV) Phone Studies

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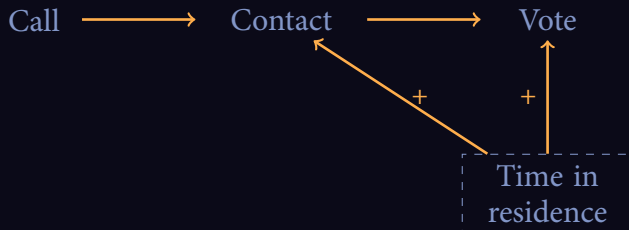
- ▶ **Treatment:** phone call encouraging turnout
- ▶ **Outcome:** vote vs not vote
- ▶ **Compliance:** contact (one-sided noncompliance)



Front-door for GOTV studies

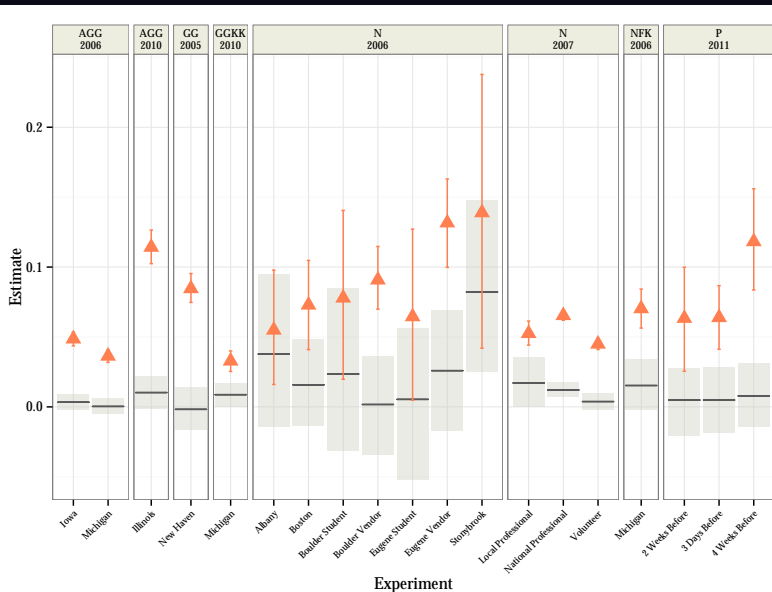


Front-door for GOTV studies



Front-door likely to exhibit positive bias.

Front-door estimates exhibit positive bias



Outline

MOTIVATION AND INTRODUCTION

FRONT-DOOR FOR ATT

FRONT-DOOR DIFFERENCE-IN-DIFFERENCES

General form of front-door difference-in-differences

Suppose we have a “differencing group” of observations for which we assume there is no treatment effect (e.g., pre-treatment outcomes for compliers and non-compliers)

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Within levels of X :

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2. “effect of attendance” for the differencing group
3. Proportion of compliers in group of interest

$$\text{FD-DID} = (3) \cdot [(1) - (2)]$$

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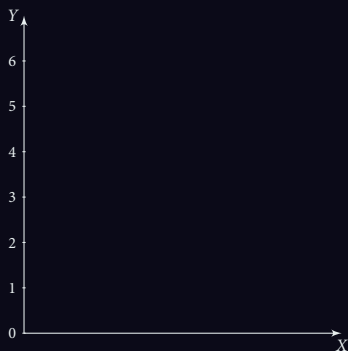
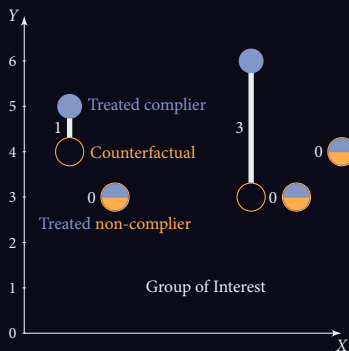
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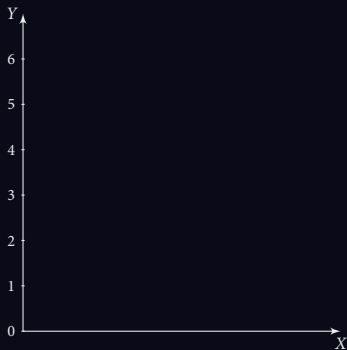
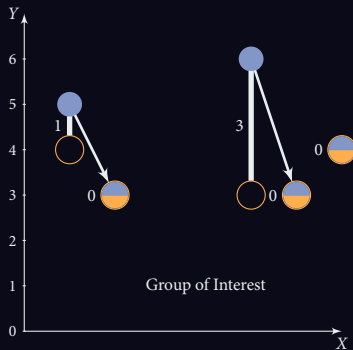
$$\text{FD-DID} = (3) \cdot [(1) - (2)]$$

Assumes bias due to “not as if random” the same for group of interest and differencing group.

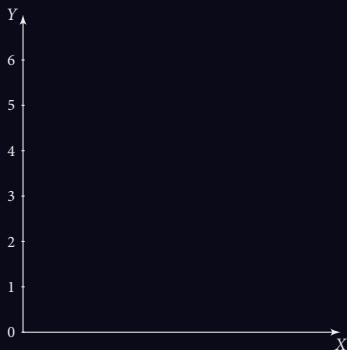
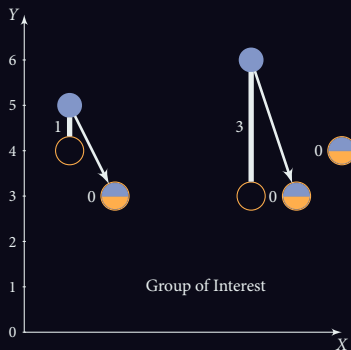
Revisiting the stylized example



Positive bias in “effect of attendance”

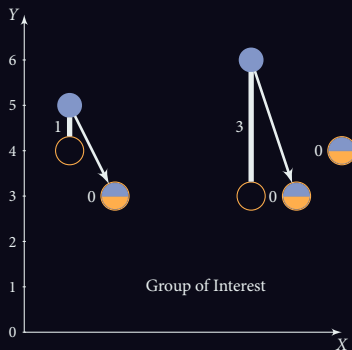


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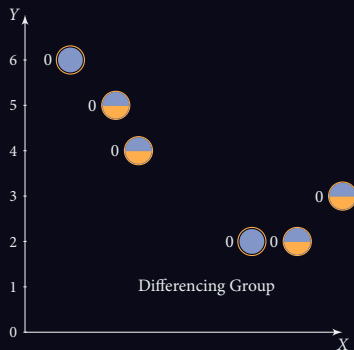


$$\frac{2+3}{2}$$

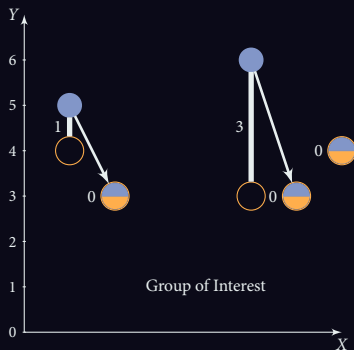
Find differencing group where believe effect is zero



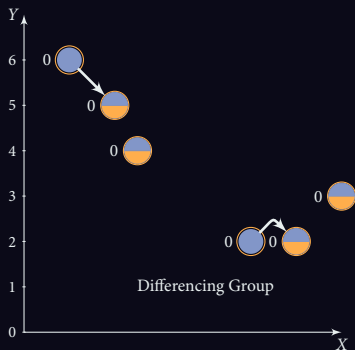
$$\frac{2+3}{2}$$



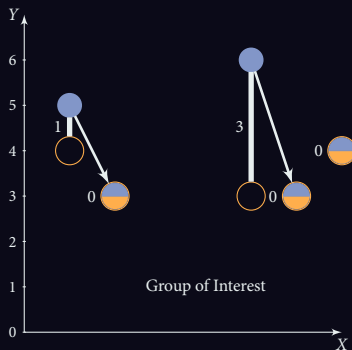
Non-zero “effect of attendance” estimate is evidence of bias



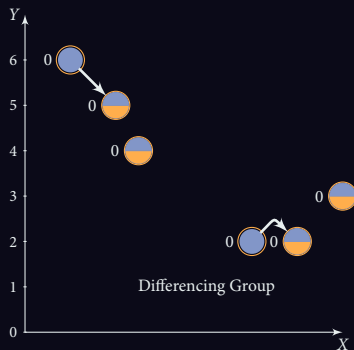
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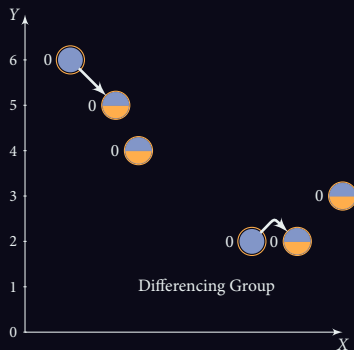
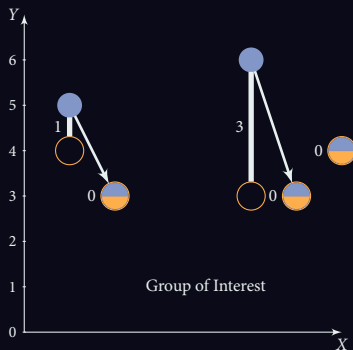


$$\frac{2+3}{2}$$



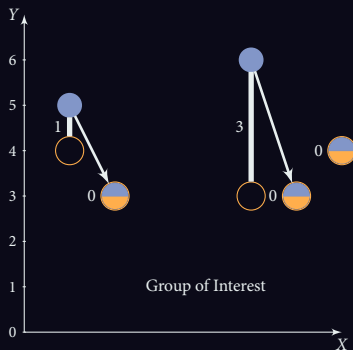
$$\frac{1+0}{2}$$

Remove bias from group of interest

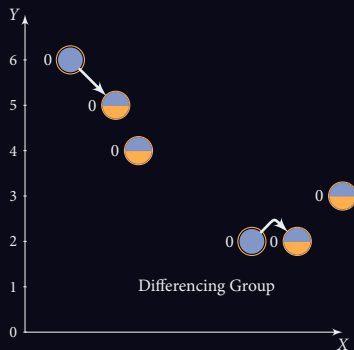


$$\frac{2}{5} \cdot \left[\frac{2+3}{2} - \frac{1+0}{2} \right]$$

Front-door diff-in-diff enables point identification



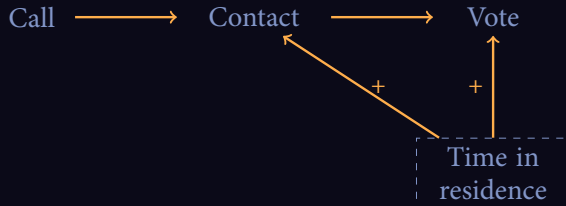
$$\frac{2}{5} \cdot \left[\frac{2+3}{2} \right]$$



$$- \left[\frac{1+0}{2} \right] = \frac{4}{5}$$

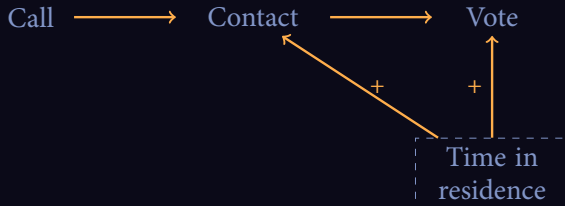
Want to remove bias due to unmeasured confounder

Group of interest:

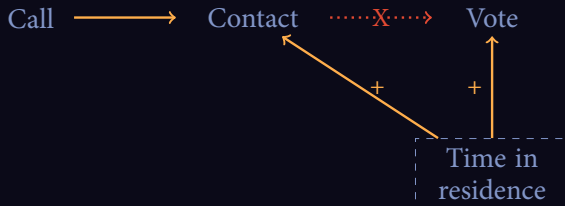


Find differencing group where believe effect is zero

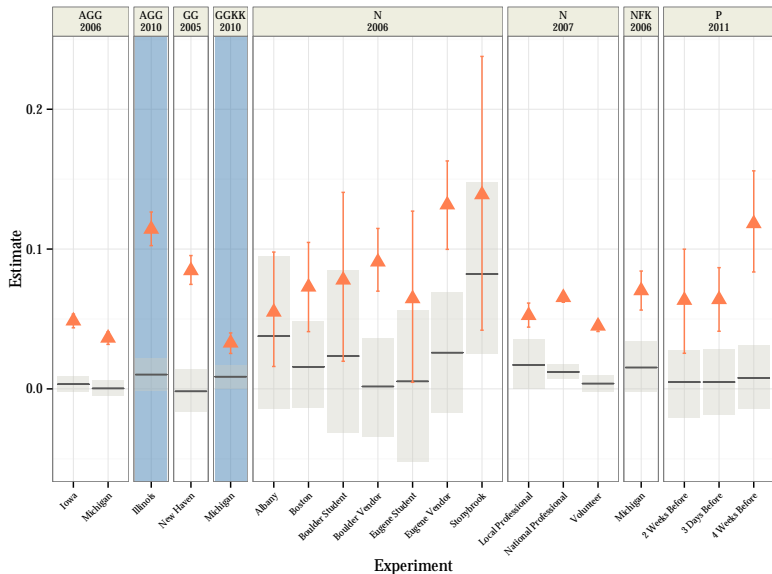
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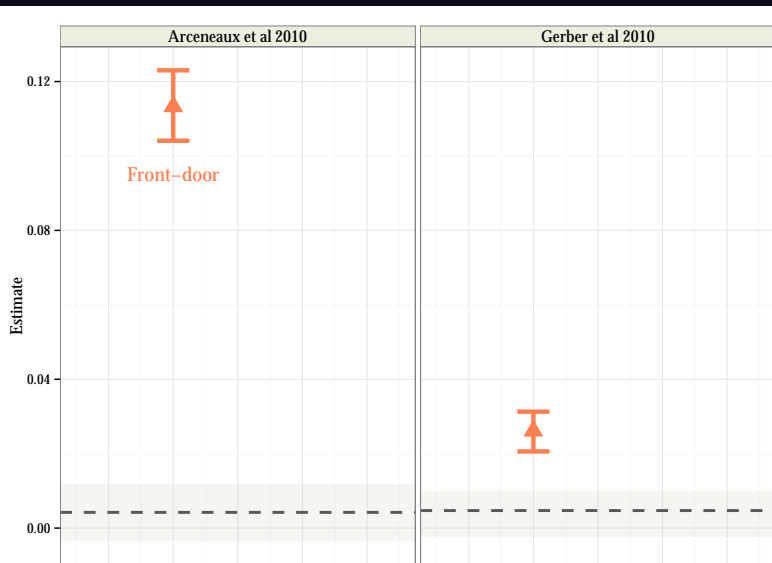
Differencing group:



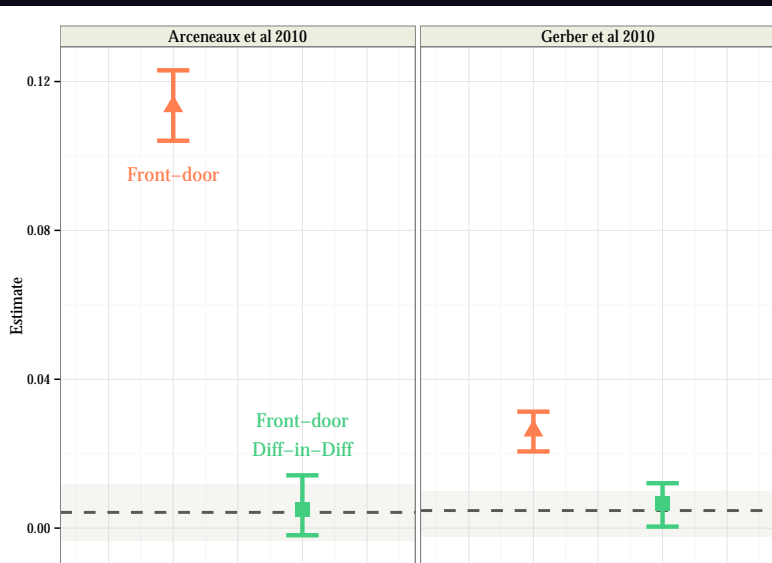
Placebo as prototypical differencing group



Prototypical FDDID as proof of concept



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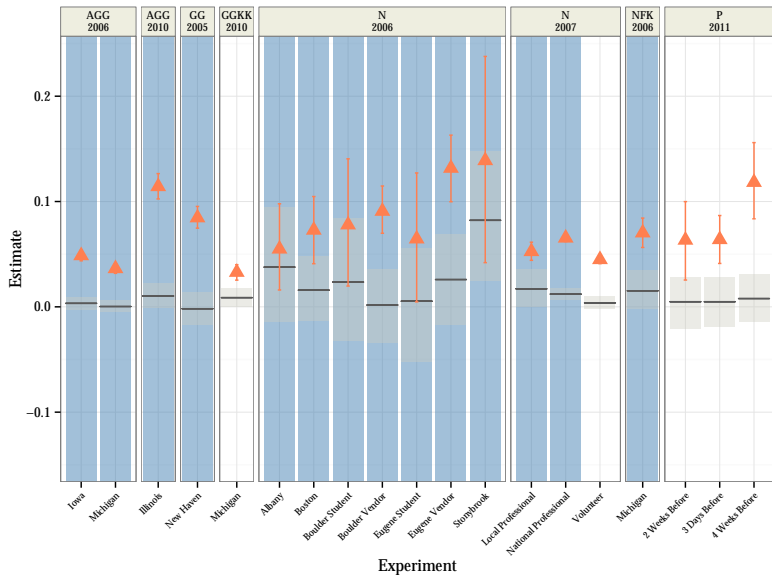
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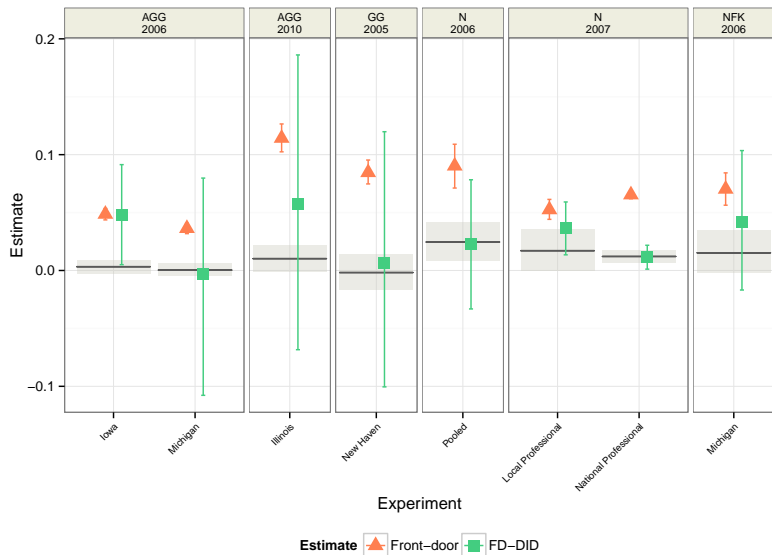
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Assumption of no effect can be relaxed for bounding analysis.

FD-DID USING 18-19 yo differencing group



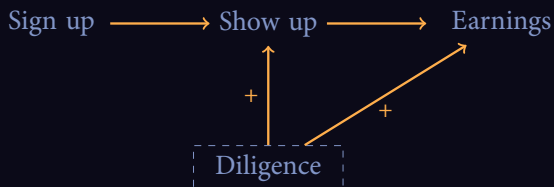
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Revisiting the National JTPA Study

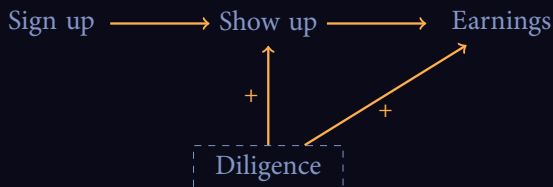
Revisiting the National JTPA Study

Married adult males (group of interest):

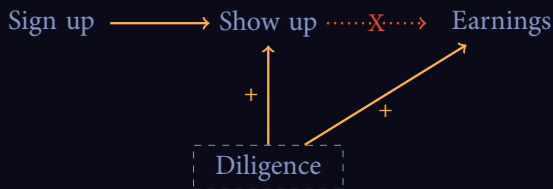


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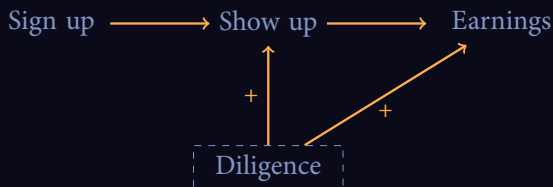


Single adult males (differencing group):

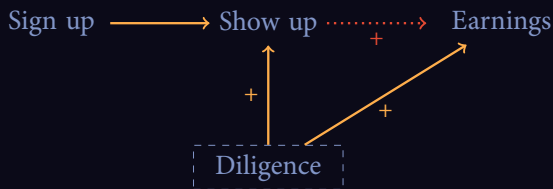


Lower bound if small positive effect for differencing group

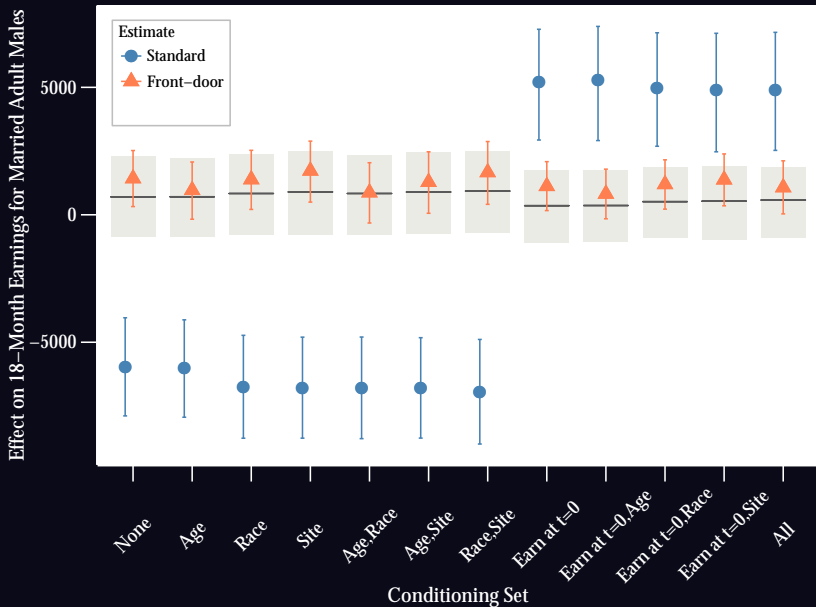
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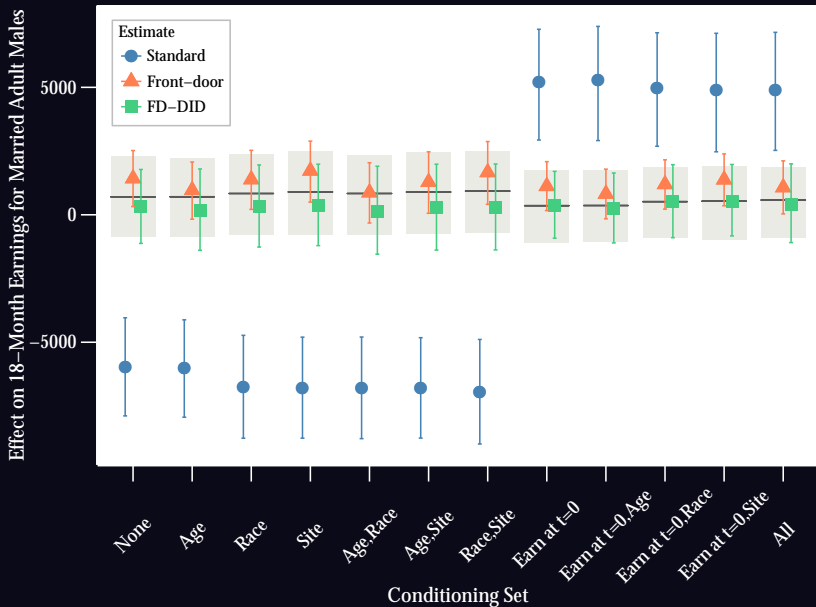
Single adult males (differencing group):



Front-door estimates exhibit slight positive bias



Front-door diff-in-diff provides lower bound



Summary

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- ▶ Can be used in concert with front-door difference-in-differences to bracket.
- ▶ In the canonical example (binary M with one-sided non-compliance), only treated units are used.

Thank You

Thank You

References:

Glynn, A.N. and Kashin, K. (2017) “Front-door Difference-in-Differences Estimators.” *American Journal of Political Science*. 61 (4): 989–1002.

Glynn, A.N. and Kashin, K. (2018) “Front-door Versus Back-door Adjustment with Unmeasured Confounding: Bias Formulas for Front-door and Hybrid Adjustments with Application to a Job Training Program,” *Journal of the American Statistical Association*. 113 (523): 1040–1049.

Proof

$$\begin{aligned}
B_{a_1}^{fd} &= \mu_{0|a_1}^{fd} - \mu_{0|a_1} \\
&= \sum_x \sum_m P(m|a_0, x) \cdot E[Y|a_1, m, x] \cdot P(x|a_1) \\
&\quad - \sum_x \sum_u E[Y|a_0, x, u] \cdot P(u|x, a_1) \cdot P(x|a_1) \\
&= \sum_x \sum_m P(m|a_0, x) \sum_u E[Y|a_1, m, x, u] \cdot P(u|a_1, m, x) \cdot P(x|a_1) \\
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&= \sum_x P(x|a_1) \sum_m \sum_u P(m|a_0, x) \cdot E[Y|a_1, m, x, u] \cdot P(u|a_1, m, x) \\
&\quad - \sum_x P(x|a_1) \sum_m \sum_u P(m|a_0, x, u) \cdot E[Y|a_0, m, x, u] \cdot P(u|a_1, x)
\end{aligned}$$

Program impact likely greater for married males

- ▶ Cross-sectional evidence that marriage increases male wages (Blackburn and Korenman, 1994; Korenman and Neumark, 1991; Krashinsky, 2004; Shoeni, 1995)
 - ▶ Studies generally estimate 10-40% marriage premium
- ▶ Some evidence that positive impact of marriage of productivity (Korenman and Neumark, 1991; Hellerstein et al., 1999; Mehay and Bowman, 2005)