

Lecture 2

Structural and Counterfactual Causal Models

An Introduction

Causal Inference Using Graphs

August 6, 2019

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

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Goals and Objectives for This Afternoon:

- Discuss relationship between nonparametric structural equation models (NPSEM) and DAGs
- Present back-door criterion for control variable selection
- Present SWIGs as an alternative to DAGs and BDC.
- Present disjunctive criterion for control variable selection

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Some Typical Advice on Control Variable Selection

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

In sum, if the variable we leave out is positively correlated with the variable we include and also has itself a positive effect on the dependent variable, then our misspecified model will produce an upward-biased estimator of the effect of X_1 on Y . [Ashenfelter, Levine, and Zimmerman, p. 189]

Some Typical Advice on Control Variable Selection

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

*The short answer to the questions posed above is that the OLS estimates based on the equation containing too many explanatory variables will indeed be unbiased estimates of the true parameters.
[Ashenfelter, Levine, and Zimmerman, p. 187]*

Some Typical Advice on Control Variable Selection

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

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SWIGs

Disjunctive Criterion

...we cannot be sure that treatment assignment is strongly ignorable given the observed covariates because there may remain unmeasured covariates that affect both outcomes and treatment assignment. [Rosenbaum and Rubin 1984, p. 522]

Some Typical Advice on Control Variable Selection

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

control for all pre- treatment covariates [Rubin 2009]

The Typical Advice can be Misleading

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

In the rest of this lecture, we will explore why this sort of typical advice can be misleading, and present some alternative criteria for covariate/control variable selection.

1 Nonparametric Structural Equation Models and DAGs

2 Back-door Criterion

3 SWIGs

4 Disjunctive Criterion

Definition (Deterministic Nonparametric Structural Equation Model (DNPSEM))

A DNPSEM is a triple

$$M = \langle \mathbf{U}, \mathbf{V}, \mathbf{g} \rangle$$

where:

- ① \mathbf{U} is a set of **exogenous** variables
- ② \mathbf{V} is a set of **endogenous** variables, that are **uniquely** determined by the variables in the model—that is, variables in $\mathbf{U} \cup \mathbf{V}$; and
- ③ \mathbf{g} is a set of functions (one for each endogenous variable) that maps $\mathbf{U} \cup \mathbf{V}$ onto \mathbf{V} .

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Note: we will often find it convenient to represent different variables in \mathbf{V} with the the subsets of variables \mathbf{X} , \mathbf{Y} , and \mathbf{Z} .

Example (Stroke Victims)

Suppose a group of stroke victims are admitted to a randomized placebo controlled trial where the active treatment is a blood thinning drug. After a period of time, patients are observed to be alive or dead.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Stroke Victims)

Suppose a group of stroke victims are admitted to a randomized placebo controlled trial where the active treatment is a blood thinning drug. After a period of time, patients are observed to be alive or dead.

For this example, a DNPSEM for each individual would include two exogenous variables: one for treatment randomization and one that specifies response to treatment.

$$U_1 = \begin{cases} \text{heads} \\ \text{tails} \end{cases}$$
$$U_2 = \begin{cases} \text{always recover} \\ \text{never recover} \\ \text{saved by drug} \\ \text{killed by drug} \end{cases}$$

In this case, the possible values of U_2 might correspond to whether a stroke victim suffers from leaking or clotting type strokes and the severity of the condition.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Stroke Victims)

... two endogenous variables: observed treatment (active/placebo) and observed outcome (alive/dead) and their respective functions.

$$X \leftarrow g_x(u_1)$$

$$Y \leftarrow g_y(x, u_2)$$

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

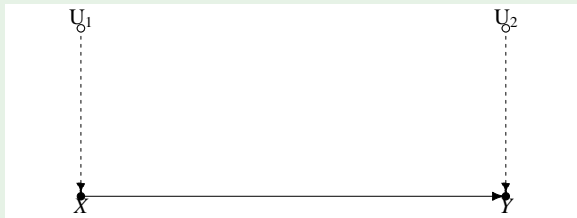
Example (Stroke Victims)

... two endogenous variables: observed treatment (active/placebo) and observed outcome (alive/dead) and their respective functions.

$$X \leftarrow g_x(u_1)$$

$$Y \leftarrow g_y(x, u_2)$$

We could express this model graphically as:



Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Stroke Victims)

Observed X and Y for the possible values of U_1 and U_2 .

$g_x(u_1)$	
U_1	X
heads	active
tails	placebo

$g_y(g_x(u_1), u_2)$	Y	
U_2	$U_1 = \text{heads}$	$U_1 = \text{tails}$
always recover		
never recover		
saved by drug		
killed by drug		

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

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$g_x(u_1)$	
U_1	X
heads	active
tails	placebo

$g_y(g_x(u_1), u_2)$	Y	
U_2	$U_1 = \text{heads}$	$U_1 = \text{tails}$
always recover	alive	alive
never recover	dead	dead
saved by drug	alive	dead
killed by drug	dead	alive

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Definition (Submodel)

Let M be a DNPSEM, and let $\mathbf{X} \subset \mathbf{V}$ be a set of *endogenous* intervention variables. A **submodel** $M_{\mathbf{x}}$ of M is the DNPSEM

$$M_{\mathbf{x}} = \langle \mathbf{U}, \mathbf{V}, \mathbf{g}_{-\mathbf{x}} \rangle$$

where

$$\mathbf{g}_{-\mathbf{x}} = \{g_w : w \notin \mathbf{x}\}$$

Effectively, a submodel makes the previously endogenous intervention variables exogenous.

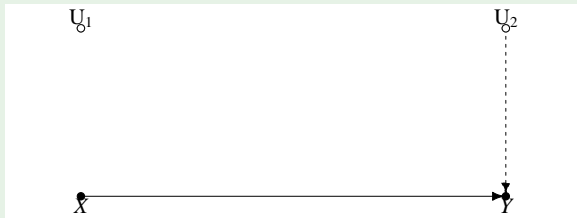
Example (Stroke Victims)

In our stroke victim example, we could create a submodel by breaking the random treatment assignment mechanism and forcing the subject to take one of the two values.

$$X \leftarrow x$$

$$Y \leftarrow g_y(x, u_2)$$

We could express this submodel graphically as:



Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Definition (Potential Outcome)

Let M be a DNPSEM, and let $\mathbf{X} \subset \mathbf{V}$ be a set of *endogenous* intervention variables and let $\mathbf{Y} \subset \mathbf{V}$ be a disjoint set of outcome variables. A **potential outcome** $\mathbf{Y}(\mathbf{x}, \mathbf{u})$ is the unique output of the functions

$$\mathbf{g}_{\mathbf{y}} = \{g_w : w \in \mathbf{y}\}$$

in the submodel $M_{\mathbf{x}}$.

This quantity can be seen as analogous to the potential outcomes that are taken to be primitive in the Neyman-Rubin causal model.

Example (Stroke Victims)

In the stroke victim example, each individual will have two potential outcomes corresponding to the two levels of treatment.

Potential Outcomes U_2	$Y(x, u_2)$	
	$X = 1$ (active)	$X = 0$ (placebo)
always recover	1 (alive)	1 (alive)
never recover	0 (dead)	0 (dead)
saved by drug		
killed by drug		

Individual (Total) Causal Effects

Definition (Individual Total Causal Effects (ICEs))

Let M be a DNPSEM, and index individual units by $i = 1, \dots, n$. Let $Y_i(X_i = x, \mathbf{u})$ be the potential outcome for unit i when X_i is set to the value x . We define the individual (total) causal effect for unit i to be:

$$Y_i(X_i = x, \mathbf{u}) - Y_i(X_i = x', \mathbf{u})$$

Notice the following about this definition:

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

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Notice the following about this definition:

- 1 \mathbf{u} is not indexed by i , hence it represents the background variables for all units. This leaves open the possibility for a type of dependence across units.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

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Notice the following about this definition:

- 1 \mathbf{u} is not indexed by i , hence it represents the background variables for all units. This leaves open the possibility for a type of dependence across units.
- 2 The intervention variable X_i is indexed by i , but the functional definition of the potential outcomes does not preclude the possibility that intervening on unit i will affect Y_i though the other units.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Stroke Victims)

In the stroke victim example, each individual will have two potential outcomes corresponding to the two levels of treatment (assume all variables are indexed by i).

Potential Outcomes	$Y(x, u_2)$		
U_2	$X = 1$	$X = 0$	ICE
always recover	1 (alive)	1 (alive)	
never recover	0 (dead)	0 (dead)	
saved by drug	1 (alive)	0 (dead)	1
killed by drug	0 (dead)	1 (alive)	

Individual Joint Causal Effects

Definition (Individual Partial Causal Effects (IJCE))

Let M be a DNPSEM and index individual units by $i = 1, \dots, n$. Let $Y_i(X_i = x, Z_i = z, \mathbf{u})$ be the potential outcome for unit i when X_i is set to the value x and Z_i is set to the value z and let $Y_i(X_i = x', Z_i = z', \mathbf{u})$ be the potential outcome for unit i when X_i is set to the value x' and Z_i is set to the value z' . We define the individual joint causal effect for unit i :

$$Y_i(X_i = x, Z_i = z, \mathbf{u}) - Y_i(X_i = x', Z_i = z', \mathbf{u})$$

Notice the following about this definition:

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

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$$Y_i(X_i = x, Z_i = z, \mathbf{u}) - Y_i(X_i = x', Z_i = z', \mathbf{u})$$

Notice the following about this definition:

- 1 A controlled direct causal effect could be defined by fixing the second intervention variable at a particular value ($z = z'$) while changing the first intervention variable.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Individual Joint Causal Effects

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$$Y_i(X_i = x, Z_i = z, \mathbf{u}) - Y_i(X_i = x', Z_i = z', \mathbf{u})$$

Notice the following about this definition:

- 1 A controlled direct causal effect could be defined by fixing the second intervention variable at a particular value ($z = z'$) while changing the first intervention variable.
- 2 We could expand this definition to include many intervention variables and to allow changes over multiple intervention variables.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

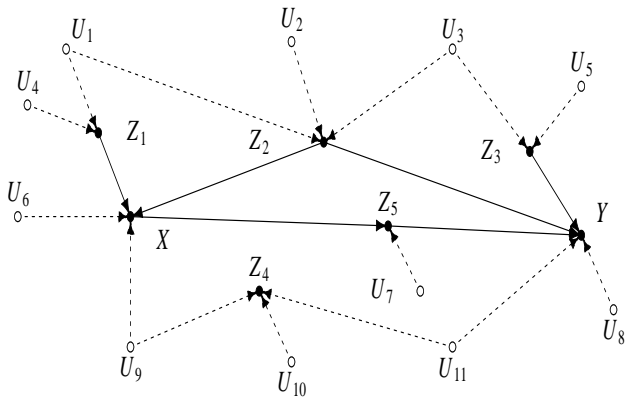
1 Nonparametric Structural Equation Models and DAGs

2 Back-door Criterion

3 SWIGs

4 Disjunctive Criterion

Example (Complicated Model)



Goals and Objectives

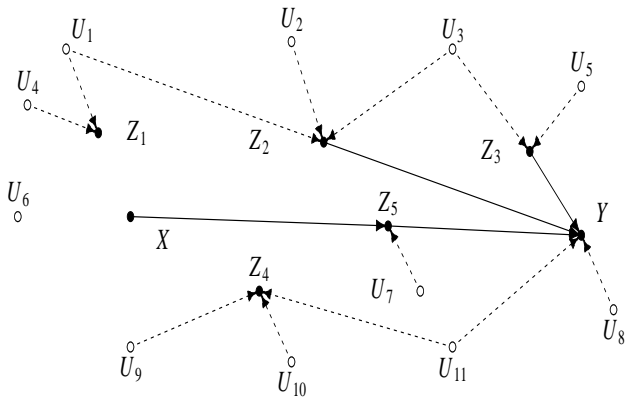
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Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Complicated Submodel)



Goals and Objectives

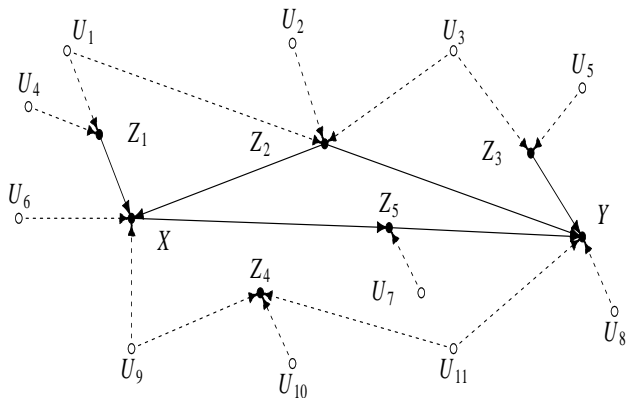
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Directed Acyclic Graphs



Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

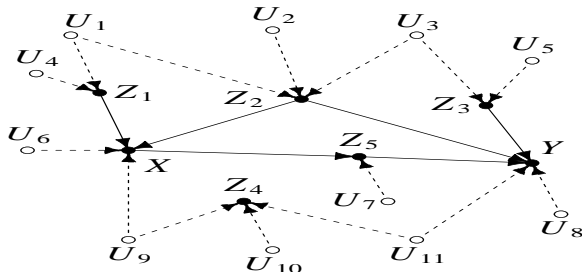
SWIGs

Disjunctive Criterion

Choosing Conditioning Variables

Definition (Back-Door Criterion (Pearl (2000, p. 79)))

- 1 none of the conditioning variables is post- X ; and
- 2 X is “ d -separated” from Y by the conditioning variables in the graph formed by deleting all arrows out of X .



Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

BD1: Don't condition on post-treatment variables

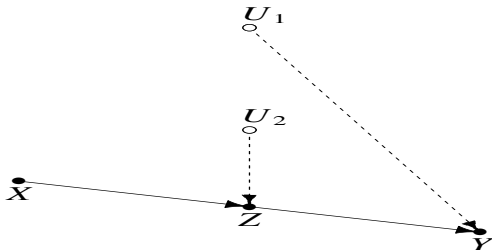
Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

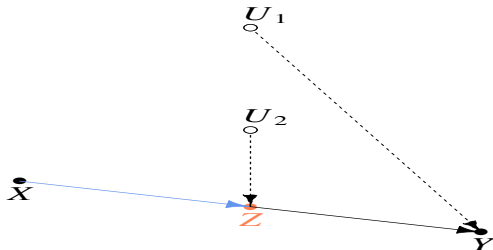
Back-door Criterion

SWIGs

Disjunctive Criterion



BD1: Don't condition on descendants of X



Goals and Objectives

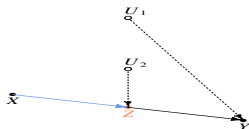
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

BD1: Don't condition on descendants of X



In sum, if the variable we leave out is positively correlated with the variable we include and also has itself a positive effect on the dependent variable, then our misspecified model will produce an upward-biased estimator of the effect of X_1 on Y . [Ashenfelter, Levine, and Zimmerman, p. 2003]

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

BD2: X is d -separated from Y by the conditioning variables in the graph formed by deleting all arrows out of X .

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

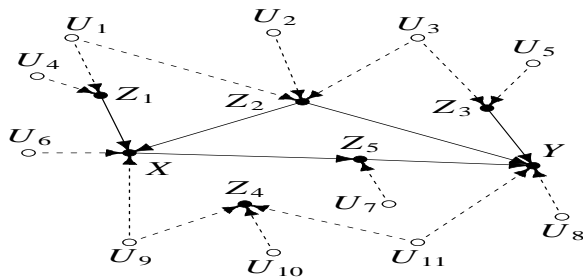
Disjunctive Criterion

Definition (d -Separation (Pearl (2000, p. 16)))

- 1 X is d -Separated from Y if all paths are blocked
- 2 paths are blocked by “unconditioned colliders” or conditioned non-colliders

d -Separation allows conditional independence relations to be read from a DAG consistent with a probabilistic model.

Example: Sufficient Conditioning Sets



Remove arrows out of X .

Goals and Objectives

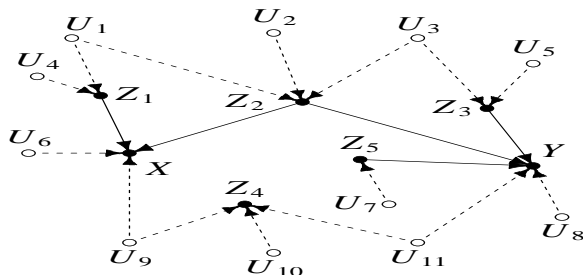
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example: Sufficient Conditioning Sets



Recall that paths are blocked by “unconditioned colliders” or conditioned non-colliders

Goals and Objectives

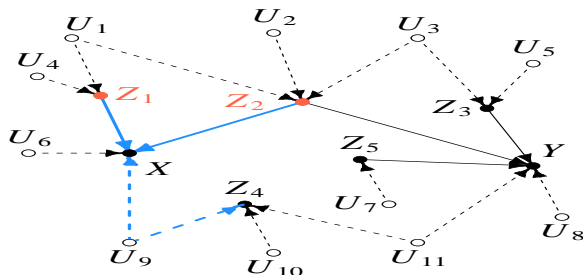
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example: Sufficient Conditioning Sets



Recall that paths are blocked by “unconditioned colliders” or conditioned non-colliders

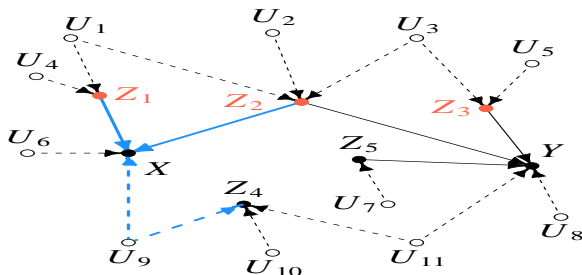
Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

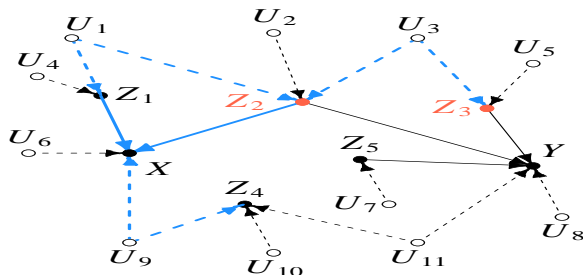
SWIGs

Disjunctive Criterion



Recall that paths are blocked by “unconditioned colliders” or conditioned non-colliders

Example: Sufficient Conditioning Sets



Recall that paths are blocked by “unconditioned colliders” or conditioned non-colliders

Goals and Objectives

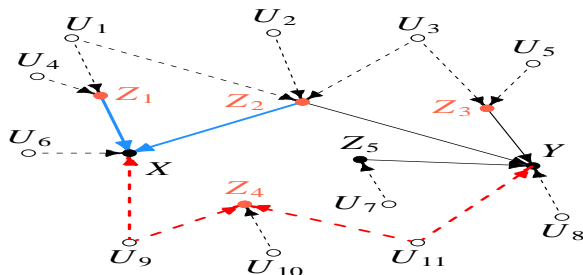
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example: Non-sufficient Conditioning Sets



Recall that paths are blocked by “unconditioned colliders” or conditioned non-colliders

Goals and Objectives

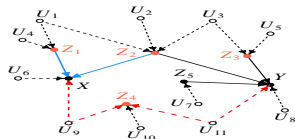
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example: Non-sufficient Conditioning Sets



The short answer to the questions posed above is that the OLS estimates based on the equation containing too many explanatory variables will indeed be unbiased estimates of the true parameters.

[Ashenfelter, Levine, and Zimmerman, p. 187]

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Section Conclusions

- Knowledge of the temporal order of observed variables is not sufficient to correctly specify **Z**
 - Adjusting for “pre-treatment in time” variables can induce bias
 - Adjusting for “post-treatment in time” variables can eliminate bias
- There can be multiple *non-nested* sufficient conditioning sets
- Obviously, some models will have no sufficient observable conditioning set

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

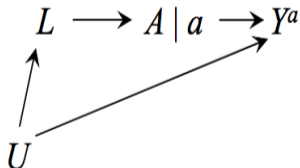
SWIGs

Disjunctive Criterion

Definition (Single World Intervention Graph)

See Richardson and Robins (2013)

Example



SWIG Example 2

Goals and Objectives

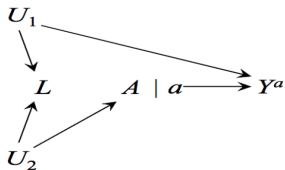
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example



An Easier Criterion

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

With a single graphical structure, the BDC and SWIGs allow us to find all possible sufficient sets, or if a set of plausible graphs is proposed, the BDC and SWIGs allow us to see if there exists a set that is sufficient for all graphs.

What if we don't want to propose graphs?

Definition (Disjunctive Criterion)

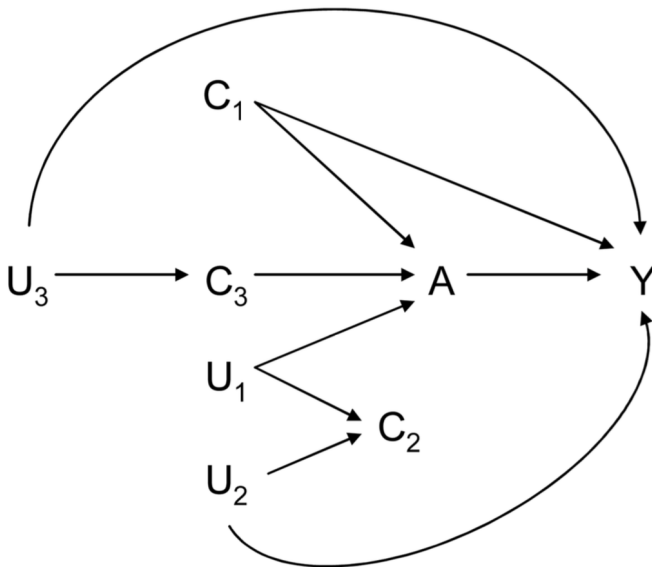
Let C be a set of measured pretreatment covariates. Let $S \subseteq C$ be the subset of C whose elements are either causes of A or of Y or of both. If there is any set $W \subseteq C$ such that W blocks all backdoor paths from A to Y then S does also and thus $Y_a \perp\!\!\!\perp A|S$.

Definition (Disjunctive Criterion)

Let C be a set of measured pretreatment covariates. Let $S \subseteq C$ be the subset of C whose elements are either causes of A or of Y or of both. If there is any set $W \subseteq C$ such that W blocks all backdoor paths from A to Y then S does also and thus $Y_a \perp\!\!\!\perp A|S$.

Note: while the disjunctive criterion does not require the specification of a graph or graphs, it does require assumptions about the true DAG.

Example (Figure 1 VS)



Note: ??

Goals and Objectives

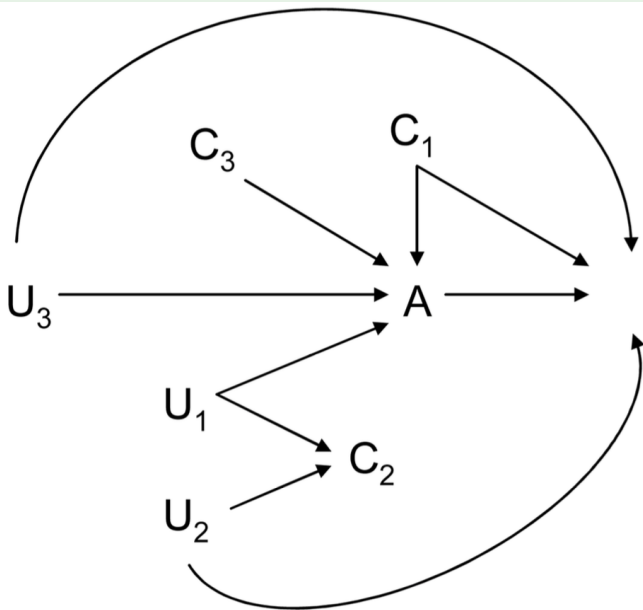
Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Example (Figure 2 VS)



Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Propositions 1 and 2

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

Definition (Propositions 1 and 2)

Suppose that for some set S , $Y_a \perp\!\!\!\perp A|S$ and that under some ordering of the elements of S , (S_1, \dots, S_n) , and for some k , $Y \perp\!\!\!\perp S_{i+1}|(A, S_1, \dots, S_i)$ for $i = k, \dots, n - 1$, then $Y_a \perp\!\!\!\perp A|(S_1, \dots, S_k)$.

Suppose that for some set S , $Y_a \perp\!\!\!\perp A|S$, that the distribution of (Y, A, S) is faithful to the underlying causal diagram G , and that under some ordering of the elements of S , (S_1, \dots, S_n) , and for some k , $Y \perp\!\!\!\perp S_i|(A, S_1, \dots, S_k)$ for $i = k + 1, \dots, n$ then $Y_a \perp\!\!\!\perp A|(S_1, \dots, S_k)$.

Goals and Objectives

Nonparametric
Structural Equation
Models and DAGs

Back-door Criterion

SWIGs

Disjunctive Criterion

End of slides for Day 1. Examples on board for remainder of day.