

week3

November 20, 2021

1 3주차

1.1 Posterior Predictive Distribution

- After the data y have been observed, we can predict an unknown observable \tilde{y}
- The posterior predictive distribution of a future observation, \tilde{y} is:

$$\begin{aligned} p(\tilde{y}|y) &= \int p(\tilde{y}, \theta|y) d\theta \\ &= \int p(\tilde{y}|\theta, y) p(\theta|y) d\theta \\ &= \int p(\tilde{y}|\theta) p(\theta|y) d\theta \end{aligned}$$

- Assumed y and \tilde{y} are conditional independent given θ
- prior predictive distribution before y observed

1.1.1 Example 1 : Binomial model

$$y_i \stackrel{iid}{\sim} \text{Bern}(\theta)$$

$$Y \sim \text{Bin}(n, \theta), 0 \leq \theta \leq 1$$

$$\theta \sim \text{Unif}(0, 1)$$

$$\Rightarrow \theta|y \sim \text{Beta}(y+1, n-y+1)$$

* Posterior predictive distribution for $\tilde{y} = 1$

$$P(\tilde{y} = 1|y) = \frac{y+1}{n+2}$$

which is known as Laplace's law of succession.

$$p(\tilde{y} = 1|y) = \int_0^1 p(\tilde{y} = 1|\theta) p(\theta|y) d\theta = \int_0^1 \theta p(\theta|y) d\theta = E[\theta|y] = \frac{y+1}{n+2}$$

$$y = 0 \text{ (all failures)} \Rightarrow p(\tilde{y} = 1|y) = \frac{1}{n+2}$$

$$y = 1 \text{ (all successes)} \Rightarrow p(\tilde{y} = 1|y) = \frac{n+1}{n+2}$$

- prior predictive distribution

$$p(\tilde{y} = 1) = \int_0^1 p(\tilde{y} = 1|\theta)p(\theta)d\theta = \int_0^1 \theta d\theta = \frac{1}{2}$$

1.1.2 Example 2 : Poisson Model

- Data model: $y_i \stackrel{iid}{\sim} \text{Poisson}(\theta), i = 1, \dots, n$
- Prior distribution : $\theta \sim \text{Gamma}(\alpha, \beta)$
- Posterior distribution of θ given y

$$L(\theta) = \prod_{i=1}^n \frac{1}{y_i!} \theta^{y_i} e^{-\theta} = \left(\prod_{i=1}^n \frac{1}{y_i!} \theta^{\sum y_i} e^{-n\theta} \right)$$

- MLE for θ :

$$\begin{aligned} \log L(\theta) &= \log \left(\prod_{i=1}^n \frac{1}{y_i!} + \sum y_i \log \theta - n\theta \right) \\ \frac{\partial \log L}{\partial \theta} &= \frac{\sum y_i}{\theta} - n = 0 \\ \Rightarrow \hat{\theta}_{ML} &= \frac{1}{n} \sum_{i=1}^n y_i = \bar{y} \end{aligned}$$

- Posterior Distribution

$$\begin{aligned} p(\theta|y) &\propto p(y|\theta)p(\theta) \\ &= \left[\prod_{i=1}^n \frac{1}{y_i!} \theta^{y_i} e^{-\theta} \right] \frac{1}{\Gamma(\alpha)\beta^\alpha} \theta^{\alpha-1} e^{-\frac{\theta}{\beta}} \\ &\propto \theta^{\sum y_i + \alpha - 1} e^{-\theta \left(n + \frac{1}{\beta} \right)} \\ &= e^{\sum y_i + \alpha - 1} e^{-(n + \frac{1}{\beta})\theta} \\ &\int_0^\infty \theta^{\sum y_i + \alpha - 1} e^{-(n + \frac{1}{\beta})\theta} d\theta = 1 \\ \Rightarrow \theta|y &\sim \text{Gamma} \left(\sum y_i + \alpha, \left[n + \frac{1}{\beta} \right]^{-1} \right) \\ \Rightarrow p(\theta|y) &= \frac{(n + \frac{1}{\beta})^{\sum y_i + \alpha}}{\Gamma(\sum y_i + \alpha)} e^{\sum y_i + \alpha - 1} e^{-(n + \frac{1}{\beta})\theta} \\ E[\theta|y] &= \frac{\sum y_i + \alpha}{n + \frac{1}{\beta}} = \hat{\theta}_{Bayes} \\ &= \frac{n}{n + \frac{1}{\beta}} \left(\frac{\sum y_i}{n} \right) + \frac{\frac{1}{\beta}}{n + \frac{1}{\beta}} (\alpha\beta) \\ \Rightarrow &\text{Weighted average of sample mean and prior mean} \end{aligned}$$

- $n \uparrow \Rightarrow E[\theta|y] \rightarrow \hat{\theta}_{ML}$
- $n \downarrow \Rightarrow E[\theta|y] \rightarrow \alpha\beta$ (prior mean)

1.2 Posterior Predictive Distribution of Poisson Model

- Posterior predictive distribution, $p(\tilde{y}|y)$:

$$\begin{aligned}
 p(\tilde{y}|y_1, \dots, y_n) &= \int_0^\infty p(\tilde{y}|\theta)p(\theta|y_1, \dots, y_n)d\theta \\
 &= \int_0^\infty \frac{1}{\tilde{y}!} e^{-\theta} \theta^{\tilde{y}} \frac{(n + \frac{1}{\beta})^{\sum y_i + \alpha}}{\Gamma(\sum y_i + \alpha)} \theta^{\sum y_i + \alpha - 1} e^{-(n + \frac{1}{\beta})\theta} d\theta \\
 &= \left[\frac{1}{\tilde{y}!} \frac{(n + \frac{1}{\beta})^{\sum y_i + \alpha}}{\Gamma(\sum y_i + \alpha)} \right] \int_0^\infty \theta^{\tilde{y} + \sum y_i + \alpha - 1} e^{-(n + \frac{1}{\beta} + 1)\theta} d\theta \\
 &= \left[\frac{1}{\tilde{y}!} \frac{(n + \frac{1}{\beta})^{\sum y_i + \alpha}}{\Gamma(\sum y_i + \alpha)} \right] \left[\frac{\Gamma(\tilde{y} + \sum y_i + \alpha)}{(n + \frac{1}{\beta} + 1)^{\tilde{y} + \sum y_i + \alpha}} \right] \\
 &= \frac{\Gamma(\tilde{y} + \sum y_i + \alpha)}{\Gamma(\sum y_i + \alpha) \tilde{y}!} \left(\frac{n + \frac{1}{\beta}}{n + \frac{1}{\beta} + 1} \right)^{\sum y_i + \alpha} \left(\frac{1}{n + \frac{1}{\beta} + 1} \right)^{\tilde{y}} \\
 &\Rightarrow \theta|y \sim \text{NegBin}(\sum y_i + \alpha, \frac{n + \frac{1}{\beta}}{n + \frac{1}{\beta} + 1})
 \end{aligned}$$

\tilde{y} = the number failures until r th successes

1.3 Normal Model with a Single Observation

- Normal model with unknown mean θ and known variance σ^2

$$y \sim N(\theta, \sigma^2)$$

- Prior distribution : $\theta \sim N(\mu, \tau^2)$
- Posterior distribution of θ given y

$$\begin{aligned}
 p(\theta|y) &\propto p(y|\theta)p(\theta) \\
 &= \left[\frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(y-\theta)^2} \right] \left[\frac{1}{\sqrt{2\pi}\tau} e^{-\frac{1}{2\tau^2}(\theta-\mu)^2} \right] \\
 &\propto \exp\left[-\frac{1}{2\sigma^2}(y^2 - 2y\theta + \theta^2) - \frac{1}{2\tau^2}(\theta^2 - 2\theta\mu + \mu^2)\right] \\
 &\propto \exp\left[-\frac{1}{2\sigma^2}(\theta^2 - 2y\theta) - \frac{1}{2\tau^2}(\theta^2 - 2\theta\mu)\right] \\
 &= \exp\left[-\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right)\theta^2 + 2\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{\mu}{\tau^2}\right)\theta\right] \\
 &= \exp\left[-\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right)\left(\theta^2 - 2\frac{\frac{y}{\sigma^2} + \frac{\mu}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}\theta\right)\right] \\
 &\propto \exp\left[-\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right)\left(\theta - \frac{\frac{y}{\sigma^2} + \frac{\mu}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}\right)^2\right] \\
 \theta|y &\sim N\left(\frac{\frac{y}{\sigma^2} + \frac{\mu}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}, \left[\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right]^{-1}\right)
 \end{aligned}$$

$$E[\theta|y] = \frac{\frac{1}{\sigma^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} y + \frac{\frac{1}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} \mu$$

- τ^2 means prior variance
- If $\tau^2 \uparrow \Rightarrow$ Little information $\Rightarrow E[\theta|y] \rightarrow y$ (sample mean)
- If $\tau^2 \downarrow \Rightarrow$ Much information $\Rightarrow E[\theta|y] \rightarrow \mu$
- posterior variance = $[\frac{1}{\sigma^2} + \frac{1}{\tau^2}]^{-1}$
- precision = $\frac{1}{\text{variance}}$ (reciprocal of variance)
- $\frac{1}{\sigma^2}$: precision of data model
- $\frac{1}{\tau^2}$: precision of prior
- posterior precision = prior precision + data precision

1.4 Normal Model with Multiple Observations

- Normal model with unknown mean θ and known variance σ^2

$$y_i \stackrel{iid}{\sim} N(\theta, \sigma^2), i = 1, \dots, n$$

- Prior distribution : $\theta \sim N(\mu, \tau^2)$
- Posterior distribution of θ given y_1, \dots, y_n

$$\begin{aligned} p(\theta|y) &\propto p(y|\theta)p(\theta) \\ &\propto \left[\prod_{i=1}^n e^{-\frac{1}{2\sigma^2}(y_i - \theta)^2} \right] e^{-\frac{1}{2\tau^2}(\theta - \mu)^2} \\ &= e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \theta)^2} e^{-\frac{1}{2\tau^2}(\theta - \mu)^2} \\ &\propto \exp\left[-\frac{1}{2}\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)\left(\theta - \frac{\frac{\sum y_i}{\sigma^2} + \frac{\mu}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}\right)^2\right] \\ \Rightarrow \theta|y &\sim N\left(\frac{\frac{\sum y_i}{\sigma^2} + \frac{\mu}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}, \left[\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right]^{-1}\right) \end{aligned}$$

- Posterior mean

$$E[\theta|y] = \frac{\frac{n}{\sigma^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}} \left(\frac{\sum y_i}{n}\right) + \frac{\frac{1}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}} (\mu)$$

$$n \uparrow \Rightarrow E[\theta|y] \rightarrow \hat{\theta}_{ML} = \bar{y}$$

$$\tau^2 \uparrow (\text{no information}) \Rightarrow E[\theta|y] \rightarrow \bar{y}$$

$$\tau^2 \downarrow (\text{much information}) \Rightarrow E[\theta|y] \rightarrow \mu$$

- Posterior precision = sample precision + prior precision