

Project nr 2

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1 Task

Fix the parameters: $r = 0.05$, $\sigma^2 = 0.125$ (thus $\mu^* = r - \sigma^2/2 = -0.0125$), $S(0) = 100$, and $K = 100$.

Estimate the $I_{n,C}$ given in (3.3) using

- a) Crude Monte Carlo estimator.
- b) Stratified estimator.
- c) For $n = 1$: Antithetic estimator. You may take (Z_{2i-1}, Z_{2i}) with $Z_{2i} = -Z_{2i-1}$, where Z_{2i-1} , $i = 1, \dots, R/2$, are i.i.d. standard normal $\mathcal{N}(0, 1)$.
- d) For $n = 1$: Control variate estimator. As a control variate, you may take $X = B(1)$.

Compare the results. For the case $n = 1$, compare estimations with the exact value using the Black-Scholes formula (3.4). For stratified estimators, consider proportional and optimal allocation schemes. Provide a report in a .pdf file and the working implementation you used. Test your results for at least two different values of C . In the report grading, the following will be taken into account:

1. Code (0-4 pts) - 1 pt for results reproducibility, 3 pts for code quality, i.e. code structure, appropriate comments, lack of redundancy. Code generator by AI tools needs to be clearly flagged.
2. Report (0-6 pts) - structure of the report, readability and visualizations
3. Methodology (0-10 pts) - completion of all parts of the task, discussion on parameter choices and comments on the comparison of European call options and discretized binary up-and-out options.

2 Brownian motion and geometric Brownian motion

DISCLAIMER: THIS IS BY NO MEANS A FULL INTRODUCTION TO BROWNIAN MOTION. IT IS A *MINIMALIST* INTRODUCTION FOR THE PURPOSES OF THIS PROJECT.

2.1 Brownian motion

Roughly speaking, a stochastic process $\mathbf{B} = (B(t))_{t \leq T}$ is a **Brownian motion** if $B(t_0) = 0$ at $t_0 = 0$, and for any $0 \leq t_1 < \dots < t_n \leq T$, the vector $(B(t_1), \dots, B(t_n))$ is a zero-mean multivariate normal random variable $\mathcal{N}(\mathbf{0}, \Sigma)$ with covariance matrix

$$\Sigma(i, j) = \text{Cov}(B(t_i), B(t_j)) = \min(t_i, t_j), \quad i, j = 1, \dots, n.$$

In this project, we consider $T = 1$ and equally spaced time points $(t_1, t_2, \dots, t_n) = \left(\frac{1}{n}, \frac{2}{n}, \dots, 1\right)$.

2.2 Stratified sampling of a multivariate normal $\mathcal{N}(\mathbf{0}, \Sigma)$ random variable

Suppose we want to sample a random variable $\mathbf{B} = (B_1, \dots, B_n)^T \sim \mathcal{N}(\mathbf{0}, \Sigma)$ using m strata. Let $\mathbf{Z} = (Z_1, \dots, Z_n)^T$ be a multivariate standard normal random variable. The strata will be defined by ascending rings A^1, \dots, A^m , which are determined by balls A'_i centered at $(0, \dots, 0)$ with suitable radii such that $\mathbb{P}(\mathbf{Z} \in A^i) = 1/m$. Thus, let

- A'_1 be an n -dimensional ball such that $\mathbb{P}(\mathbf{Z} \in A'_1) = 1/m$;
- A'_2 be a ball such that $\mathbb{P}(\mathbf{Z} \in A'_2 \setminus A'_1) = 1/m$;
- etc.

Set $A^1 = A'_1$, $A^2 = A'_2 \setminus A'_1$, \dots , $A^m = A'_m \setminus A'_{m-1}$.

Let \mathbf{A} be such that $\Sigma = \mathbf{A}\mathbf{A}^T$ (Cholesky decomposition).

Define the i -th stratum by $S^i = \{\mathbf{Az} : \mathbf{z} \in A^i\}$.

Assume that $\mathbf{Z}^i \stackrel{D}{=} (\mathbf{Z} | \mathbf{Z} \in A^i)$. Then $\mathbf{B}^i = \mathbf{Az}^i$ is from stratum S^i .

It remains to show how to sample $\mathbf{Z}^i \stackrel{D}{=} (\mathbf{Z} | \mathbf{Z} \in A^i)$. For $n = 2$ and $m = 1$, the method was presented in the lecture (which *de facto* is the Box-Muller method). For general $n \geq 2$, let ξ_1, \dots, ξ_n be i.i.d. standard normal $\mathcal{N}(0, 1)$ random variables. Denote $\boldsymbol{\xi} = (\xi_1, \dots, \xi_n)^T$. Let $D > 0$. Then the vector

$$\left(D \frac{\xi_1}{\|\boldsymbol{\xi}\|}, \dots, D \frac{\xi_n}{\|\boldsymbol{\xi}\|}\right)^T$$

has a uniform distribution on a sphere with radius D . We have the following proposition:

Proposition 1 *Let $\mathbf{Z} = (Z_1, \dots, Z_n)$ be a standard multivariate normal random variable. Then the square of the length of \mathbf{Z} is $D^2 = Z_1^2 + \dots + Z_n^2$ and has a χ_n^2 distribution (χ^2 with n degrees of freedom).*

Recall that the density and c.d.f. of χ_n^2 are as follows:

$$f_{\chi_n^2}(r) = \frac{1}{2^{n/2}\Gamma(n/2)} r^{n/2-1} e^{-r/2}, \quad F_{\chi_n^2}(r) = \frac{1}{\Gamma(n/2)} \gamma_{n/2}(r/2),$$

where Γ is the gamma function, and γ is the *incomplete gamma function*.¹ For $n = 2$, the random variable D has the Rayleigh distribution. Admittedly, there is no explicit formula for the inverse function of $F_{\chi_n^2}(r)$ for general n , but numerically this inverse is available in several libraries.²

Summing up, sampling $\mathbf{B}^i \stackrel{D}{=} (\mathbf{B} \mid \mathbf{B} \in A^i)$ is as follows:

1. Perform Cholesky decomposition: $\Sigma = \mathbf{A}\mathbf{A}^T$.
2. Sample $\boldsymbol{\xi} = (\xi_1, \dots, \xi_n)^T$, where $\xi_i \sim \mathcal{N}(0, 1)$ i.i.d. Set

$$\mathbf{X} = (X_1, \dots, X_n)^T = \left(\frac{\xi_1}{\|\boldsymbol{\xi}\|}, \dots, \frac{\xi_n}{\|\boldsymbol{\xi}\|} \right)^T.$$

3. Sample $U \sim \mathcal{U}(0, 1)$. Set

$$D^2 = F_{\chi_n^2}^{-1} \left(\frac{i-1}{m} + \frac{1}{m} U \right).$$

4. Set $\mathbf{Z} = (Z_1, \dots, Z_n) = (DX_1, \dots, DX_n)$.
5. Return $\mathbf{B}^i = \mathbf{A}\mathbf{Z}$.

2.2.1 Stratified sampling of a Brownian motion

We can simply use the procedure described in Section 2.2. Recall that $\mathbf{B} = (B(1/n), B(2/n), \dots, B(1))$ is a multivariate normal random variable $\mathcal{N}(\mathbf{0}, \Sigma)$ with the covariance matrix

$$\Sigma(i, j) = \frac{1}{n} \min(i, j).$$

We can perform the Cholesky decomposition $\Sigma = \mathbf{A}\mathbf{A}^T$, where

$$\mathbf{A}(i, j) = \begin{cases} \frac{1}{\sqrt{n}} & \text{if } j \leq i \\ 0 & \text{otherwise.} \end{cases}$$

In Figure 1, 5000 points within 4 strata were simulated using the above method.

¹https://en.wikipedia.org/wiki/Incomplete_gamma_function

²E.g., `scipy.stats.chi2.ppf` in Python or `chi2inv` in Matlab

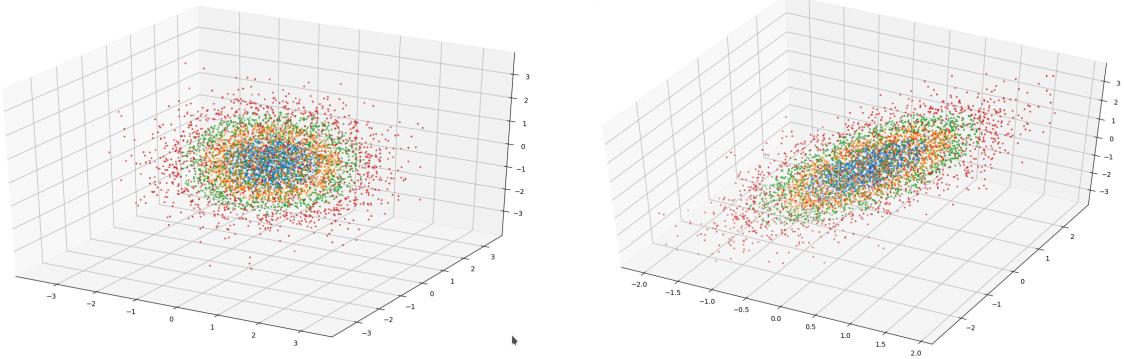


Figure 1: 5000 points from a 3-dimensional standard normal distribution obtained using stratified (4 strata) sampling (left). Points from a 3-dimensional normal distribution with covariance matrix $\Sigma(i, j) = \min(i, j)/3$ (right).

2.3 Geometric Brownian motion

The evolution of stocks (assets) is often modeled as geometric Brownian motion, denoted $\text{GBM}(\mu, \sigma)$, which is defined by

$$S(t) = S(0) \exp \left(\left(r - \frac{\sigma^2}{2} \right) t + \sigma B(t) \right), \quad 0 \leq t \leq T, \quad (2.1)$$

where $B(t)$ ($0 \leq t \leq T$) is Brownian motion. In computing option prices, often the interest rate r and volatility σ are known; we then make computations for $\text{GBM}(r, \sigma)$. Denote $\mu^* = r - \sigma^2/2$. Then we have

$$S(t) = S(0) \exp (\mu^* t + \sigma B(t)), \quad 0 \leq t \leq T. \quad (2.2)$$

3 European and Barrier call options

We are interested in estimating the following (called an *option*, with discounted payoff at time 1) with price given by the formula

$$A_{n,C} = \begin{cases} e^{-r}(S(1) - K)_+, & \text{when } \forall i \in 1, 2, \dots, n, S(i/n) < C, \\ 0, & \text{otherwise} \end{cases}, \quad (3.3)$$

where $S(t)$ is given in (2.2) and

$$I_{n,C} = E[A_{n,C}]$$

In the case $C = \infty$, this is called a **European call option**; otherwise, it is called a **discrete barrier up-and-out call option**.

3.1 Black-Scholes formula

In the case $C = \infty$ (i.e., a European call option), the exact value of $E(A_{1,\infty} - K)_+ = E(S(1) - K)_+$ is provided by the Black-Scholes formula (where Φ is the c.d.f. of $\mathcal{N}(0, 1)$):

$$E(S(1) - K)_+ = S(0)\Phi(d_1) - Ke^{-r}\Phi(d_2), \quad (3.4)$$

where

$$d_1 = \frac{1}{\sigma} \left[\log \left(\frac{S(0)}{K} \right) + r + \frac{\sigma^2}{2} \right],$$

and

$$d_2 = d_1 - \sigma.$$