`yacoub`: a Python package for Simulating Generalized Fading Channels

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***Abstract*—We present a well tested Python-based library for simulating and computing generalized fading channels, named `yacoub`. We describe the applicability of `yacoub` using examples in recent communications systems challenges, namely: spectrum sensing, bit error rate computation, and parameter estimation in generalized fading channels. For the latter, a Majorization-Minimization algorithm is derived. The development of `yacoub` open source and its code is avaliable at `http://github.com/mirca/yacoub`.**

I. Introduction

A. Note on notation

Scalars and random variables are denoted as *italic* small-case letters *e.g.* x; vectors and random vectors are denoted as *italic*, boldface, small-case letters *e.g.* $\boldsymbol{x}$. The $n-$th component of a vector $\boldsymbol{x}$ is denoted as x_n. A complex vector of length n is defined as $\boldsymbol{x} \in \mathbb{C}^{n\times 1}$. All vectors are column vectors. Matrices are denoted as *italic*, boldface, capital letters as in $\boldsymbol{X}$; the identity matrix of order n is denoted as $\boldsymbol{I}_n$. We define a discrete-time circularly symmetric Gaussian process $\boldsymbol{z}$ as any (finite or infinite) collection of random varibles $\boldsymbol{z} = \boldsymbol{x} + j\boldsymbol{y}$, $j \triangleq \sqrt{-1}$, such that $\boldsymbol{x}$ and $\boldsymbol{y}$ are iid jointly Gaussian with zero mean vector and covariance matrix given by $\mathbb{E}\left[\boldsymbol{z}\boldsymbol{z}^{\dagger}\right]$, in which $\boldsymbol{z}^{\dagger}$ means the conjugate transpose of $\boldsymbol{z}$. The expectation value wrt a random variable x is denoted as $\mathbb{E}_x$. The probability of an event A is denoted as $\mathbb{P}(A)$. The indicator function is denoted as $\mathbb{I}(\cdot)$ and it evaluates to 1 if its argument is true and 0 otherwise.

II. The Acceptance-Rejection sampler in log-space

III. Examples

A. Spectrum Sensing in Complex Generalized Fading Channels

The spectrum sensing problem consists in deciding whether or not a given channel frequency band is being occupied by a licensed (primary) user and, in case that such frequency band is available, how to opportuniscally allocate secondary users such that the interference on the primary user is negligible.

From a probabilistic point of view, the spectrum sensing problem may be framed as a decision theory problem, as follows

$$H_0: \ \boldsymbol{y} = \boldsymbol{w}, \tag{1}$$

$$H_1: \ \boldsymbol{y} = h\boldsymbol{s} + \boldsymbol{w}, \tag{2}$$

in which $\boldsymbol{y} \in \mathbb{C}^{n\times 1}$ is the decoded received vector signal, $\boldsymbol{w} \in \mathbb{C}^{n\times 1}$ is complex Gaussian noise process with zero mean vector and covariance matrix given as $\sigma^2\boldsymbol{I}_n$, and h is the channel gain.

In [**?**], the authors have shown that the probability distribution of the energy statistic $\tilde{y} \triangleq \boldsymbol{y}^{\dagger}\boldsymbol{y}$ conditioned on the knowledge of h, in case that $\boldsymbol{s}$ is an M-PSK signal such that every symbol has the same probability of occurrence, $\mathbb{P}(s_n = s) = \frac{1}{M}$, is given as

$$p(\tilde{y}|h, H_1) = 1 - Q_n\left(\sqrt{\frac{2n|h|^2E_s}{\sigma^2}}, \sqrt{\frac{2\tilde{y}}{\sigma^2}}\right), \tag{3}$$

in which Q_n is the Marcum-Q function and E_s is the energy per symbol.

The pdf of $\tilde{y}$ can be written using the Law of Total Expectation

$$p(\tilde{y}|H_1) = \mathbb{E}_h\left[p(\tilde{y}|h, H_1)\right] = \int_{-\infty}^{+\infty} p(\tilde{y}|h, H_1)p(h)\,\mathrm{d}h. \tag{4}$$

Recall that the energy detection rule can be expressed as

$$d_{\delta}(\tilde{y}) = \mathbb{I}(\tilde{y} > \delta) \tag{5}$$

in which δ is a strictly positive real number known as energy threshold, and $d_{\delta}(\tilde{y}) = j, \ j \in \{0, 1\}$, means that the detector has decided in favor of the hypothesis H_j.

As a result, the probabilities of false alarm and miss detection can be written as

$$p_f \triangleq \mathbb{P}\left(d_{\delta}(\tilde{y}) = 1|H_0\right) = 1 - p(\delta|H_0), \tag{6}$$

$$p_d \triangleq \mathbb{P}\left(d_{\delta}(\tilde{y}) = 0|H_1\right) = \mathbb{E}_h\left[p(\delta, h|H_1)\right], \tag{7}$$

B. Parameter Estimation in α-μ fading

The α-μ fading probability density is given as

$$p(h) = \frac{\alpha\mu^{\mu}h^{\alpha\mu-1}}{\Gamma(\mu)}\exp\left(-\mu h^{\alpha}\right), \tag{8}$$

$h \in \mathbb{R}_+$, $\alpha \in \mathbb{R}_+$, and $\mu \in \mathbb{R}_+$.

Given a vector $\boldsymbol{h} \in \mathbb{R}^{n\times 1}$ of iid samples from the $\alpha - \mu$ distribution, we would like to estimate the value of α. The pdf of $\boldsymbol{h}$ is given as

$$p(\boldsymbol{h}) = \prod_{i=1}^{n} p(h_i) = \prod_{i=1}^{n} \frac{\alpha\mu^{\mu} h_i^{\alpha\mu-1}}{\Gamma(\mu)} \exp\left(-\mu h_i^{\alpha}\right) \tag{9}$$

$$= \left(\frac{\alpha\mu^{\mu}}{\Gamma(\mu)}\right)^{n} \exp\left(-\mu \sum_{i=1}^{n} h_i^{\alpha}\right) \prod_{i=1}^{n} h_i^{\alpha\mu-1} \tag{10}$$

and the log likelihood function (up to an additive constant that does not depend upon α) is given as

$$\log p(\boldsymbol{h}) \cong n \log \alpha - \mu \sum_{i=1}^{n} h_i^{\alpha} + \alpha\mu \sum_{i=1}^{n} \log h_i \tag{11}$$

Noting that

$$\sum_{i=1}^{n} h_i^{\alpha} \geq \sum_{i=1}^{n} h_i^{\alpha_t} + (\alpha - \alpha_t) \sum_{i=1}^{n} h_i^{\alpha_t} \log h_i \tag{12}$$

and therefore

$$\log p(\boldsymbol{h}) \leq n \log \alpha - \mu \left(\sum_{i=1}^{n} h_i^{\alpha_t} + (\alpha - \alpha_t) \sum_{i=1}^{n} h_i^{\alpha_t} \log h_i\right) + \alpha\mu \sum_{i=1}^{n} \log h_i \tag{13}$$

C. BER in Complex $\alpha - \mu$ Fading

IV. Conclusions

Acknowledgement

The authors would like to thank the Federal University of Campina Grande (UFCG) and the Institute for Advanced Studies in Communications (Iecom) for supporting this research.