maoud: a Python package for Simulating Generalized Fading Channels

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Abstract—We present a well tested Python-based library for simulating and computing generalized fading channels, named maoud. We describe the applicability of maoud using examples in scenarios of communications channels impaired by generalized fading, namely: spectrum sensing, bit error rate computation, and fading estimation. For the latter, we develop an iterative algorithm using the Majorization-Minimization framework, which allows reliable estimation of the fading parameter. The development of maoud is open source and its code along with examples are avaliable at http://github.com/mirca/maoud.

I. INTRODUCTION

Notation

Scalars and random variables are denoted as italic, smallcase letters e.g. x; sets and events are denoted as italic, capital letters e.g., A; vectors and random vectors are denoted as *italic*, boldface, small-case letters e.g. x. The n-th component of a vector x is denoted as x_n . A complex vector of length n is defined as $x \in \mathbb{C}^{n \times 1}$. All vectors are column vectors. Matrices are denoted as italic, boldface, capital letters as in X; the identity matrix of order n is denoted as I_n . We define a discrete-time circularly symmetric Gaussian process z as any collection of random variables z = x + iy, $i \triangleq \sqrt{-1}$, such that x and y are i.i.d. jointly Gaussian, with zero mean vector and covariance matrix given by $\mathbb{E} |zz^{\dagger}|$, in which z^{\dagger} means the conjugate transpose of z. The expectation value w.r.t. to the probability distribution of a random variable x is denoted as \mathbb{E}_x . The probability of an event A is denoted as $\mathbb{P}(A)$. The indicator function is denoted as $\mathbb{I}(\cdot)$, it evaluates to one if its argument is true and zero otherwise. For any given two real functions f and g defined on the same domain D, $f \cong g$ means that there exist a constant c such that f(x) = g(x) + c,

II. THE ACCEPTANCE-REJECTION SAMPLER IN LOG-SPACE

III. MAJORIZATION-MINIMIZATION ALGORITHMS

IV. EXAMPLES

A. Spectrum Sensing in Complex Generalized Fading Channels

The spectrum sensing problem consists in deciding whether or not a given channel frequency band is being occupied by a licensed (primary) user and, in case that such frequency band is available, how to opportuniscally allocate secondary users such that the interference on the primary user is negligible.

From a probabilistic point of view, the spectrum sensing problem may be framed as a decision theory problem, as follows

$$H_0: \ \boldsymbol{y} = \boldsymbol{w},\tag{1}$$

$$H_1: \ \mathbf{y} = h\mathbf{s} + \mathbf{w},\tag{2}$$

in which $y \in \mathbb{C}^{n \times 1}$ is the decoded received vector signal, $w \in \mathbb{C}^{n \times 1}$ is complex Gaussian noise process with zero mean vector and covariance matrix given as $\sigma^2 I_n$, and h is the channel gain.

In [1], the authors have shown that the probability distribution of the energy statistic $\tilde{y} \triangleq y^{\dagger}y$ conditioned on the knowledge of h, in case that s is an M-PSK signal such that every symbol has the same probability of occurrence, $\mathbb{P}(s_n = s) = \frac{1}{M}$, is given as

$$p(\tilde{y}|h, H_1) = 1 - Q_n \left(\sqrt{\frac{2n|h|^2 E_s}{\sigma^2}}, \sqrt{\frac{2\tilde{y}}{\sigma^2}} \right), \quad (3)$$

in which Q_n is the Marcum-Q function and E_s is the energy per symbol.

The pdf of \tilde{y} can be written using the Law of Total Expectation

$$p(\tilde{y}|H_1) = \mathbb{E}_h \left[p(\tilde{y}|h, H_1) \right] = \int_{-\infty}^{+\infty} p(\tilde{y}|h, H_1) p(h) \, dh.$$
 (4)

Recall that the energy detection rule can be expressed as

$$d_{\delta}(\tilde{y}) = \mathbb{I}(\tilde{y} > \delta) \tag{5}$$

in which δ is a strictly positive real number known as energy threshold, and $d_{\delta}(\tilde{y}) = j, \ j \in \{0,1\}$, means that the detector has decided in favor of the hypothesis H_{j} .

As a result, the probabilities of false alarm and miss detection can be written as

$$p_f \triangleq \mathbb{P}\left(d_\delta(\tilde{y}) = 1|H_0\right) = 1 - p(\delta|H_0),\tag{6}$$

$$p_d \triangleq \mathbb{P}\left(d_\delta(\tilde{y}) = 0 | H_1\right) = \mathbb{E}_h\left[p(\delta, h | H_1)\right],\tag{7}$$

B. Parameter Estimation in Nakagami-m fading

The Nakagami-m density is given as

$$p(\mathbf{h}) = \prod_{i=1}^{n} \frac{2m^m}{\Gamma(m)\Omega^m} h_i^{2m-1} \exp\left(-\frac{mh_i^2}{\Omega}\right)$$
$$= \left(\frac{2m^m}{\Gamma(m)\Omega^m}\right)^n \exp\left(-\frac{m\sum_{i=1}^{n} h_i^2}{\Omega}\right) \prod_{i=1}^{n} h_i^{2m-1} \tag{8}$$

And the log likelihood function (up to an additive constant) is given as

$$\log p(\boldsymbol{h}) \cong n \left(m \left(\log m - \log \Omega \right) - \log \Gamma(m) \right)$$

$$- m \left(\frac{\sum_{i=1}^{n} h_i^2}{\Omega} - 2 \sum_{i=1}^{n} \log h_i \right)$$
 (9)

A direct maximum likelihood estimator (MLE) for (9) has been investigated to be infeasible [2].

Therefore, we use a Majorization-Minimization algorithm to find smooth and easy to optimize lower bounds for $\log p(\boldsymbol{h})$. More precisely, we need to find lower bounds for $m\log m$ and $-\log\Gamma(m)$ for $m\geq\frac{1}{2}$. The former function is convex for $m\geq 0$, hence it can be lower bounded by its first order Taylor series as

$$m\log m \ge m(1+\log m_t) - m_t,\tag{10}$$

The function $-\log \Gamma(m)$ is concave, therefore, it can be lower bounded by its second order Taylor series expansion as

$$-\log \Gamma(m) \ge -\log \Gamma(m_t) - \psi(m_t)(m - m_t) - \frac{\psi'(\frac{1}{2})}{2}(m - m_t)^2, \ m \ge \frac{1}{2}, \ m_t \ge \frac{1}{2},$$
(11)

in which $\psi(x)=\frac{\Gamma'(x)}{\Gamma(x)}$ is known as the digamma function. For both inequalities presented above, equality is achieved at $m=m_t$.

Substituting (10) and (11) into (9), a lower bound for $\log p(h)$ is obtained in (17). Due to the simple form of $g(m|m_t)$, its maximizer can be found in closed form, and an updating rule for the MLE can be written as

$$m_{t+1} = 1 + \frac{1}{\psi'(\frac{1}{2})} \left(1 + \log \frac{m_t}{\Omega} - \psi(m_t) + \frac{1}{n} \left(2 \sum_{i=1}^n \log h_i - \frac{\sum_{i=1}^n h_i^2}{\Omega} \right) \right)$$
(12)

Additionally, note that m_{t+1} is an estimator of the true parameter value m. Its expected value, conditioned on the knowledge of m_t , is given as

$$\mathbb{E}(m_{t+1}|m_t) = 1 + \frac{1}{\psi'(\frac{1}{2})} \left(\log \frac{m_t}{\Omega} - \psi(m_t) + \frac{2}{n} \sum_{i=1}^n \mathbb{E}(\log h_i) \right)$$
(13)

Its variance, also conditioned on the knowledge of m_t , is given as

$$var(m_{t+1}|m_t) = \left(\frac{1}{n\psi'(\frac{1}{2})}\right)^2 var\left(2\sum_{i=1}^n \log h_i - \frac{\sum_{i=1}^n h_i^2}{\Omega}\right).$$
(14)

The Cramér-Rao Lower Bound for any unbiased estimator of m, say \hat{m} , is given as [2]

$$var(\hat{m}) = \frac{1}{n\left(\psi'(m) - \frac{1}{m}\right)}$$
 (15)

C. BER in Complex $\alpha - \mu$ Fading

Consider the system

$$y = hs + w \tag{17}$$

in which $s \in \mathbb{C}^{n \times 1}$ is a complex On-Off Keying (OOK) signal, h is a complex $\alpha - \mu$ random variable and w is a complex Gaussian process with zero mean vector and covariance matrix equals $\sigma^2 I_n$, and y is the received complex vector signal.

Assume that the OOK symbols are equiprobable and that there exist no interference between the in-phase and quadrature components, then the probability of one bit error is given as

$$p_e = \frac{1}{2} \left(\mathbb{P} \left(\hat{y}_i = 0 | s_i = 1 \right) + \mathbb{P} \left(\hat{y}_i = 1 | s_i = 0 \right) \right).$$
 (18)

Assume that the decoded vector \hat{y} is estimated using the minimum distance decoding rule, i.e.,

V. CONCLUSIONS

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$$g(m|m_t) = n \left(-m \log \Omega + m(1 + \log m_t) - m_t - \log \Gamma(m_t) - \psi(m_t)(m - m_t) - \frac{\psi'\left(\frac{1}{2}\right)}{2}(m - m_t)^2 \right) - m \left(\frac{\sum_{i=1}^n h_i^2}{\Omega} - 2\sum_{i=1}^n \log h_i \right)$$
(16)