# maoud: a Python package for Simulating Generalized Fading Channels

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Abstract—We present a well tested Python-based library for simulating and computing generalized fading channels, named maoud. We describe the applicability of maoud using examples in scenarios of communications channels impaired by generalized fading, namely: spectrum sensing, bit error rate computation, and fading estimation. For the latter, we develop an iterative algorithm using the Majorization-Minimization framework, which allows reliable estimation of the fading parameter. The development of maoud is open source and its code, along with the code for the examples presented in this paper, are available at http://github.com/mirca/maoud.

#### I. Introduction

The study of modern wireless communications systems heavily relies on fading channel simulation. By fading channel simulation, we refer to the generation of samples from a probability distribution that resembles the effects and impairments caused by real communications fading channels on the transmitted signal.

Although accurate and precise distributions for generalized fading have been estabilished in the literature, such as  $\alpha$ - $\mu$  [1],  $\kappa$ - $\mu$  [2] and  $\eta$ - $\mu$  [2], the generation of samples following these distributions is usually a time-consuming task. In [3], the authors built an efficient algorithm, based on the rejection method, for generation of samples from those distributions.

however, there are neither open nor closed source implementations available to the scientific community.

In this paper, we present an open source Python package, named maoud, for generation of samples following the  $\alpha$ - $\mu$ ,  $\kappa$ - $\mu$ , and  $\eta$ - $\mu$  distributions. The usefullness of maoud is illustrated through examples involving spectrum sensing, bit error rate (BER) computation, and fading estimation.

#### Notation

Scalars and random variables are denoted as italic, small-case letters e.g. x; sets and events are denoted as italic, capital letters e.g., A; vectors and random vectors are denoted as italic, boldface, small-case letters e.g. x. The n-th component of a vector x is denoted as  $x_n$ . A complex vector of length n is defined as  $x \in \mathbb{C}^{n \times 1}$ . All vectors are column vectors. Matrices are denoted as italic, boldface, capital letters as in X; the identity matrix of order n is denoted as  $I_n$ . We define a discrete-time circularly symmetric Gaussian process x as any collection of random variables x = x + y,

 $j \triangleq \sqrt{-1}$ , such that  $\boldsymbol{x}$  and  $\boldsymbol{y}$  are i.i.d. jointly Gaussian, with zero mean vector and covariance matrix given by  $\mathbb{E}\left(zz^{\dagger}\right)$ , in which  $z^{\dagger}$  means the conjugate transpose of z. The expectation value with respect to the probability distribution of a random variable x is denoted as  $\mathbb{E}_x$ . The probability of an event A is denoted as  $\mathbb{P}(A)$ . The indicator function is denoted as  $\mathbb{I}(\cdot)$ , it evaluates to one if its argument is true and zero otherwise. For any given two real functions f and g defined on the same domain D,  $f \cong g$  means that there exist a constant c such that f(x) = g(x) + c,  $\forall x \in D$ . The natural logarithm of a scalar x > 0 is denoted as  $\log x$ .

### II. REJECTION SAMPLING

#### III. EXAMPLES

A. Spectrum Sensing in Complex Generalized Fading Channels

The spectrum sensing problem consists in deciding whether or not a given channel frequency band is being occupied by a licensed (primary) user and, in case that such frequency band is available, how to opportuniscally allocate secondary users such that the interference on the primary user is negligible.

From a probabilistic point of view, the spectrum sensing problem may be framed as a decision theory problem, as follows

$$H_0: \ \boldsymbol{y} = \boldsymbol{w},\tag{1}$$

$$H_1: \ \boldsymbol{y} = h\boldsymbol{s} + \boldsymbol{w},\tag{2}$$

in which  $y \in \mathbb{C}^{n \times 1}$  is the received vector signal,  $w \in \mathbb{C}^{n \times 1}$  is complex Gaussian noise process with zero mean vector and covariance matrix given as  $\sigma^2 I_n$ , and h is the channel gain.

In [4], the authors have shown that the cumulative distribution function (cdf) of the energy statistic  $\tilde{y} \triangleq y^{\dagger}y$  conditioned on the knowledge of h, in case that s is an M-PSK signal such that every symbol has the same probability of occurrence,  $\mathbb{P}(s_n = s) = \frac{1}{M}$ , is given as

$$P(\tilde{y}|h, H_1) = 1 - Q_n \left( \sqrt{\frac{2n|h|^2 E_s}{\sigma^2}}, \sqrt{\frac{2\tilde{y}}{\sigma^2}} \right), \quad (3)$$

in which  $Q_n$  is the Marcum-Q function and  $E_s$  is the energy per symbol.

Recall that the energy detection rule can be expressed as

$$d_{\delta}(\tilde{y}) = \mathbb{I}(\tilde{y} > \delta) \tag{4}$$

in which  $\delta$  is a strictly positive real number known as energy threshold, and  $d_{\delta}(\tilde{y}) = j, \ j \in \{0,1\}$ , means that the detector has decided in favor of the hypothesis  $H_{j}$ .

As a result, the probabilities of false alarm and miss detection can be written as

$$\mathbb{P}\left(d_{\delta}(\tilde{y}) = 1 | H_{0}\right) = 1 - P(\delta | H_{0}) = 1 - \gamma\left(n, \frac{\delta}{\sigma^{2}}\right), \quad (5)$$

$$\mathbb{P}\left(d_{\delta}(\tilde{y}) = 0 | H_{1}\right) = \mathbb{E}_{h}\left(P(\delta | h, H_{1})\right)$$

$$= \int_{-\infty}^{+\infty} P(\delta | h, H_{1}) p(h) \, \mathrm{d}h, \quad (6)$$

in which  $\gamma$  is the regularized lower incomplete Gamma function, p(h) is the pdf of the fading, and  $\delta$  is the energy detection threshold.

The performance of detection schemes can be measured by computing the Receiver Operating Characteristic (ROC), which consists in varying  $\delta$  and computing the pairs of probability of false alarm and miss detection, as illustrated in Fig. 1.

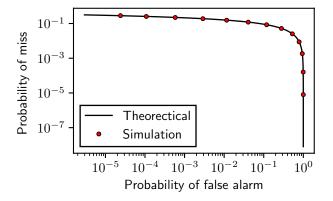


Fig. 1. Receiver Operating Characteristic for the energy detector in  $\alpha\text{-}\mu$  fading channel with  $\alpha=2$  and  $\mu=1$ , i.e., Nakagami-m, m=1. The solid curve represents the theorectical probabilities as stated on (5) and (6) for different values of  $\delta$ , whereas bullets represent Monte Carlo simulations acquired with  $10^6$  realizations. The input signal s consists in a vector of length n=25 in which each entry represents a symbol of a 64-PSK constellation. Symbols are assumed to be equiprobable. The signal-to-noise ratio is set to 5 dB.

### B. BER in $\alpha$ - $\mu$ Fading

Consider the system

$$y = hs + w \tag{7}$$

in which  $s \in \{0, a\}$ ,  $a \in \mathbb{R}_+$ , is a transmitted signal, h is an  $\alpha$ - $\mu$  random variable and  $\boldsymbol{w}$  is a Gaussian random variable with zero mean vector and variance  $\sigma^2$ , and y is the received signal.

Assume that the binary symbols are equiprobable, then the probability of one bit error is given as

$$p_e = \frac{1}{2} \left( \mathbb{P} \left( \hat{y} = 0 | s = a \right) + \mathbb{P} \left( \hat{y} = 1 | s = 0 \right) \right).$$
 (8)

Further, assume that the decoded bit  $\hat{y}$  is estimated using the minimum distance decoding rule, i.e.,

$$\hat{y} = \mathbb{I}(|y - a|^2 < |y|^2) \tag{9}$$

therefore

$$\mathbb{P}\left(\hat{y}=1|s=0\right) = \mathbb{P}\left(|w-a|^2 - |w|^2 < 0\right)$$
$$= \mathbb{P}\left(w > \frac{a}{2}\right) = 1 - \Phi\left(\frac{a}{2\sigma}\right) \tag{10}$$

and likewise

$$\mathbb{P}(\hat{y} = 0|s = a) = \mathbb{P}\left(ha + w < \frac{a}{2}\right)$$
$$= \mathbb{E}_h\left(\Phi\left(\frac{a(1-2h)}{2\sigma}\right)\right). \tag{11}$$

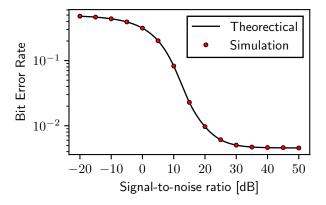


Fig. 2. BER

## IV. CONCLUSIONS REFERENCES

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