## UNIVERSITÀ DEGLI STUDI DI PADOVA

Dipartimento di Fisica e Astronomia "Galileo Galilei"

Applicazione della teoria del trasporto ottimo alla modellizzazione ecologica e per la inferenza della rete di interazione tra specie.

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Abstract:

# Indice

1	Mathematical Modeling Basics	1
	1.1 A general class of models	1
	1.2 Types of ecological interactions, with some real world examples	1
	1.3 Ecological interactions can be context- dependent	
	1.3.1 Some real world examples of context - dependency	
	1.3.2 Community matrix in the gLV model has fixed sign pattern for feasible equilibria	
<b>2</b>	Analysis of the MICEGUT dataset	5
	2.1 Dataset Description	5
	2.2 Summary Statistics	6
	2.2.1 Rank- Abundance distributions	6
	2.2.2 Sample composition across subjects	7
	2.3 Species abundance summary data TO BE COMPLETED	9
	2.4 Gallery of timeseries plots	13
3	Inference through forward linear regression	15
	3.1 Results with simulated data	16
	3.2 Results on mice data	18
	3.3 Performance evaluation with simulated data	18
	3.4 Results on mice data	19
4	Bayesian inference via the Ising Model	21

## Mathematical Modeling Basics

#### 1.1 A general class of models

The class of models we will be concerned about takes this form;

$$\dot{x}_i(t) = x_i(t) \cdot f_i[x(t)] \tag{1.1}$$

The reason behind the choice to factor out  $x_i$  is to highlight that in this class of models invasion of the ecosystem is not allowed: if a species is not present at time t = 0, it cannot enter the ecosystem at later times.

#### 1.2 Types of ecological interactions, with some real world examples

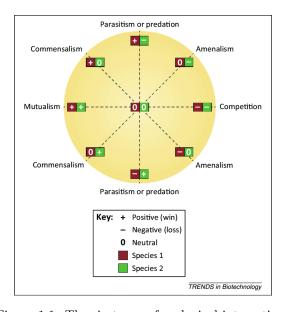


Figura 1.1: The six types of ecological interactions

• Syntrophy: is a specific type of mutualistic (++) relationship where two or more species cooperate to degrade a substrate that neither species can break down alone. The metabolic products of one organism are essential for the other organism to thrive, and vice versa. Syntrophy often occurs when the degradation of a compound requires a sequential or coupled process that depends on the interaction of different species. The classic example is methanogenic syntrophy in anaerobic environments: certain bacteria break down complex organic matter into simpler compounds like hydrogen  $(H_2)$  and acetate. However, high levels of  $H_2$  can inhibit their activity. Methanogens (archaea) consume  $H_2$ , converting it into methane  $(CH_4)$ , which lowers the  $H_2$  concentration, allowing the bacteria to continue their breakdown process. Without the methanogens, the bacteria couldn't degrade the organic matter efficiently. This makes the interaction beneficial for both sides (++).

• Cross-Feeding: is a broader term that refers to any process where one organism consumes the metabolic by-products of another organism, regardless of whether this interaction is obligatory for either species. It can be mutualistic (+ +), but it can also be commensal (+ 0), where one species benefits without affecting the other. A well-known example is lactate cross-feeding in the human gut microbiome: Bifidobacteria ferment dietary fibers into lactate. Certain butyrate-producing bacteria, such as Eubacterium hallii, consume lactate and convert it into butyrate, a short-chain fatty acid beneficial for gut health. This relationship benefits the butyrate producers but is not strictly required for the bifidobacteria.

#### 1.3 Ecological interactions can be context-dependent

There is experimental evidence that the type (the sign) and the strength of ecological interactions between species can change, depending on biotic and abiotic <sup>1</sup> factors in the environment. This fact is commonly referred to as *context-dependency*. That is, there are situations where the type of interaction (mutualism, commensalism, competition etc..) cannot be assumed as fixed, and species can behave in different ways with respect to each other if conditions change in the environment, for example switching from midly competing to strongly competing, or even from competing to mutualistic and so on.

In mathematical terms, this means that the sign pattern of the Jacobian matrix at equilibrium (the community matrix) may change across the different steady (or equilibrium) states of the system. The community matrix details the effect of increasing the density of one species on any other species around the equilibrium point.

$$\dot{x}_i(t) = x_i(t) \cdot f_i[x(t)] \tag{1.2}$$

$$\overline{x}$$
 steady state  $J(\overline{x})_{i,j} = \frac{\partial f_i}{\partial_j}$  community matrix at equilibrium  $\overline{x}$  (1.3)

This fact adds complexity to the task of building models for species evolution.

**Key Idea** Species populations across different environmental conditions or community memberships generate distinct interspecies interaction networks, the comparison of which may provide an idea of how interactions are modulated by the impact of abiotic and biotic factors (es: how will the climate change affect the ecosystems?)

The generalized Lokta Volterra model does not account for context-dependency of interactions. Instead, one class of models that can account for context-dependency is the **Consumer - Resource models (CRM)** class. A paradigmatic CRM is the Mac Arthur's consumer-resource model (MCRM).

A general consumer-resource model consists of M resources whose abundances are  $R_1, \ldots, R_M$  and S consumer species whose populations are  $N_1, \ldots, N_S$ . A general consumer-resource model is described by the system of coupled ordinary differential equations,

$$\frac{dN_i}{dt} = N_i g_i(R_1, \dots, R_M), \quad i = 1, \dots, S,$$

$$\frac{dR_{\alpha}}{dt} = f_{\alpha}(R_1, \dots, R_M, N_1, \dots, N_S), \quad \alpha = 1, \dots, M,$$

where  $g_i$ , depending only on resource abundances, is the per-capita growth rate of species i, and  $f_{\alpha}$  is the growth rate of resource  $\alpha$ . An essential feature of CRMs is that species growth rates and populations are mediated through resources and there are no explicit species-species interactions. Through resource

<sup>&</sup>lt;sup>1</sup>Biotic means living and abiotic means non - living. For example, considering the microbiota ecosystem, some abiotic factors include the host's diet, the host's immune system response, the body temperature. Biotic factors, instead, include for example the metabolic products of the species and the species population densities.

interactions, there are emergent inter-species interactions.(source: Wikipedia Eng Consumer-Resource Models).

#### 1.3.1 Some real world examples of context - dependency

Some of the most common factors that may shift the interaction type and strength are resource availability and population density.

- Plant-fungi relationships (mutualism to commensalism to parasitism):
   Many plants form mycorrhizal associations with fungi, where fungi enhance nutrient uptake for the plant (a mutualistic relationship). However, in nutrient-rich soils, allocating resources to the fungi can become disadvantageous for the plant. Thus, the fungi effectively becomes a parasite. [Jonhson 2008].
- Quorum sensing virulence expression:
   Bacteria often communicate using quorum sensing, a mechanism where they release signaling molecules into their environment to detect population density. Many pathogenic bacteria are harmless when their population densities are low. At low densities, they might exist in a commensal relationship with the host or other bacteria. However, when their density surpasses a certain threshold (as detected via quorum sensing), they might collectively express virulence factors, turning from harmless commensals into aggressive pathogens.
- Glucose alters the symbiotic relationships between the host and the gut microbiota. For instance, glucose metabolism by Escherichia coli increases the production of organic acids, which lowers intestinal pH resulting in reduced survival of Vibrio cholerae in the host. This represents a mutualistic interaction between host glucose and a commensal (E. coli) to combat a potential pathogen (V. cholerae). Also, high glucose availability in the jejunum, cecum, and colon can result in increased glucose metabolism by certain pathogenic bacteria. This increased bacterial glucose flux can reduce virulence and tip the balance from parasitism to commensalism. [Anhê, Barra e Schertzer 2020].

# 1.3.2 Community matrix in the gLV model has fixed sign pattern for feasible equilibria.

The classic generalized lokta volterra is the simplest possible model of species interactions. It assumes fixed sign and strength of interactions. As a consequence of this assumption, the sign pattern of the community matrix is fixed across all feasible equilibria [Allesina 2020].

Generalized deterministic Lokta Volterra

$$\dot{x}_i(t) = x_i(t) \cdot \left[ r_i + \sum_j a_{i,j} x_j \right]$$
(1.4)

where the  $r_i$  are the intrinsic growth rates,  $a_{i,j}$  are the interaction terms. Usually, growth rates are taken as negative to account for a limited carrying capacity of the environment (logistic growth). In compact matrix form:

$$\dot{x} = \operatorname{diag}(x) \cdot [r + A \cdot x] = f(x) \tag{1.5}$$

Then lets compute the entries of the community matrix at the feasible equilibrium  $\bar{x}$ :

$$\begin{cases} \frac{\partial f_i}{\partial x_j} = a_{i,j} \,\overline{x}_i & i \neq j\\ \frac{\partial f_i}{\partial x_j} = [r + A\overline{x}]_i + a_{i,i} \overline{x}_i & i = j \end{cases}$$
(1.6)

: A feasible equilibrium, by definition, is an equilibrium in the positive open subset of  $\mathbb{R}^n$ ,  $\mathbb{R}^n_+ = \{x | x_i > 0 \,\forall i\}$ , that is an equilibrium where all the species coexist with positive abundances. Since in a feasible equilibrium diag $(\overline{x})$  has all positive entries,  $\overline{x}$  must satisfy

$$\begin{cases} r + A\overline{x} = 0\\ \overline{x}_i > 0 \quad \forall i \end{cases} \tag{1.7}$$

Solutions to the first equation depends on rk[A|r]. Let us suppose solutions exist (rk[A|r] = rk[A]). The community matrix at the feasible equilibrium satisfies

$$J_{i,j} = a_{i,j} \,\overline{x}_i \quad \forall x_{i,j} \tag{1.8}$$

In particular, in case multiple feasible equilibria do exist (rk[A] < n and rk[A|r] = rk[A]), the sign patter of their community matrix is always the same and is equal to the sign pattern of the interaction matrix A.

# Analysis of the MICEGUT dataset

### 2.1 Dataset Description

#### to be completed

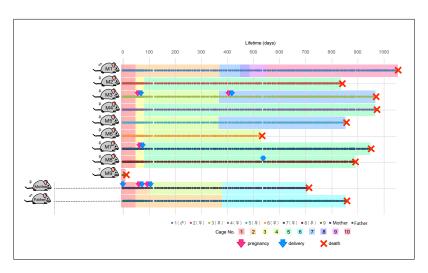


Figura 2.1

## 2.2 Summary Statistics

#### 2.2.1 Rank- Abundance distributions

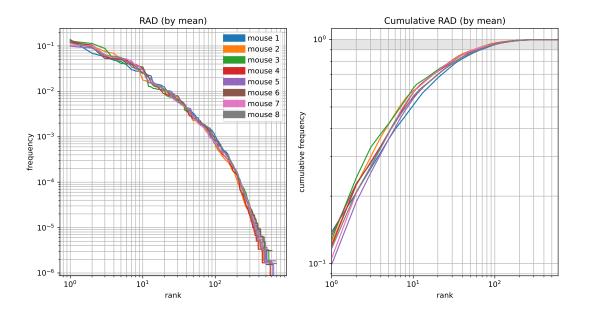


Figura 2.2: Caption

#### 2.2.2 Sample composition across subjects

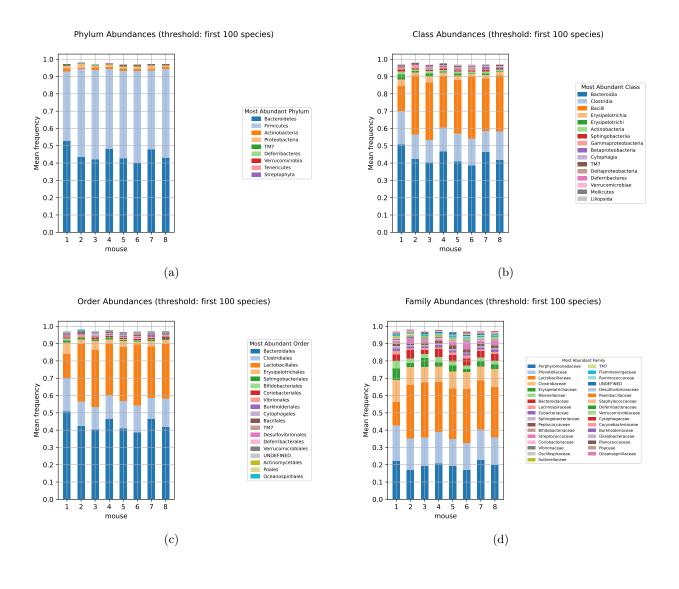
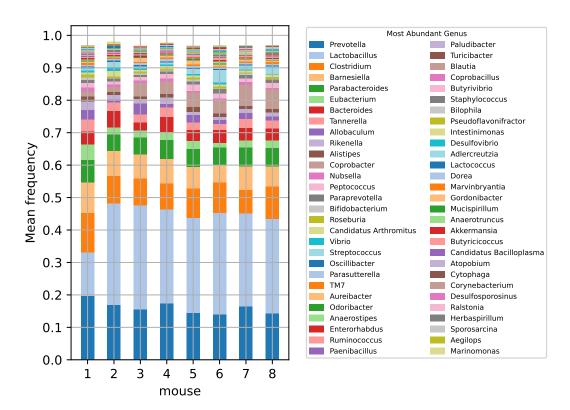


Figura 2.3: Caption

#### Genus Abundances (threshold: first 100 species)



#### Species Abundances (threshold: first 20 species)

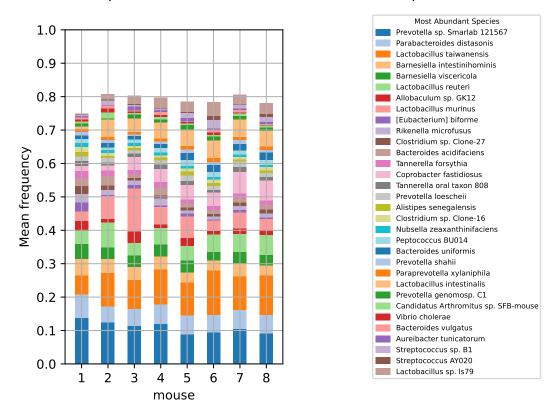
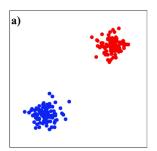


Figura 2.3: Taxonomic ranks from bottom to top are: Species  $\subset$  Genus  $\subset$  Family  $\subset$  Order  $\subset$  Class  $\subset$  Phylum.

#### 2.3 Species abundance summary data TO BE COMPLETED.

MISSING: Estimation of inter-class and intra-class variance using Linear Mixed Models



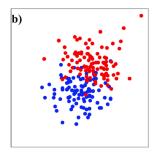


Figura 2.4: As a reminder. Inter-class and Intra-class variances concept. a) Low intra-class variance and high inter-class variance: compact well separated clusters. b) High intra-class variance and low inter-class variance: wide clusters without a clear frontier.

Tabella 2.1: The table shows the median counts and the mean counts for 136 species, taken as the union of the 100 most abundant species by median counts in all mice. The mean counts entry in the table refers to the mean of the mean counts in each mouse.

Rank	Species	Mean counts	Median counts
1	Prevotella sp. Smarlab 121567	328.61	320.23
2	Lactobacillus taiwanensis	300.42	263.12
3	Lactobacillus murinus	170.44	73.31
4	Parabacteroides distasonis	167.60	171.12
5	Lactobacillus reuteri	153.84	130.00
6	Lactobacillus intestinalis	131.43	104.62
7	Coprobacter fastidiosus	119.23	93.81
8	Barnesiella intestinihominis	111.12	111.52
9	Barnesiella viscericola	104.04	103.69
10	Lactobacillus sp. ls79	77.73	61.69
11	Allobaculum sp. GK12	56.04	15.19
12	Bacteroides acidifaciens	51.08	32.25
13	Bacteroides uniformis	47.93	33.31
14	Rikenella microfusus	44.60	37.44
15	Tannerella forsythia	44.11	42.38
16	Peptococcus BU014	39.93	27.56
17	Prevotella loescheii	39.85	32.38
18	Tannerella oral taxon 808	34.71	30.81
19	Clostridium sp. Clone-27	34.23	13.56
20	Streptococcus sp. B1	33.42	0.00
21	Paraprevotella xylaniphila	29.45	19.69
22	Prevotella genomosp. C1	28.07	26.27
23	Clostridium sp. Clone-16	27.81	14.81
24	[Eubacterium] biforme	26.30	0.38
25	Prevotella shahii	25.90	21.19
26	Alistipes senegalensis	25.76	15.12
27	Nubsella zeaxanthinifaciens	24.62	23.25
28	Clostridium sp. Clone-9	23.00	11.88
29	Candidatus Arthromitus sp. SFB-mouse	22.84	9.00
30	Prevotella sp. oral taxon 317	21.48	20.29
31	Aureibacter tunicatorum	19.81	16.12
32	Clostridium hathewayi	18.46	10.56
Continued on next pag			nued on next page

9

Tabella 2.1 – continued from previous page

Rank	Tabella 2.1 – continued from Species	Mean counts	Median counts
33	Candidatus Prevotella conceptionensis	17.26	16.06
34	Vibrio cholerae	16.72	0.00
35	Bacteroides vulgatus	16.58	9.31
36	Clostridium sp. ID4	15.93	6.06
37	Parasutterella excrementihominis	14.24	11.88
38	TM7 oral taxon 351	13.97	12.25
39	Streptococcus AY020	13.06	0.62
40	Eubacterium sp. F1	12.94	7.81
41	Clostridium sp. ASF502	12.62	5.25
$\begin{vmatrix} 11 \\ 42 \end{vmatrix}$	Bifidobacterium pseudolongum	12.18	0.00
43	Clostridium disporicum	11.41	2.62
44	Eubacterium coprostanoligenes	10.21	6.25
45	Enterorhabdus caecimuris	9.82	8.12
46	[Eubacterium] cylindroides	9.21	3.75
47	Eubacterium ventriosum	8.86	4.69
48	Staphylococcus aureus	8.75	0.00
49	Clostridium indolis	8.42	3.50
50	Odoribacter splanchnicus	8.23	5.00
51	Prevotella oulorum	8.11	7.85
$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	Clostridium fusiformis	8.06	5.00
53	Clostridium sp. 6-44	8.02	5.25
54	Pseudoflavonifractor capillosus	7.82	5.94
55	Prevotella oris	7.77	7.49
56		7.60	3.62
1	Clostridium sp. Culture-27	7.45	
57	Roseburia sp. 1120		0.25
58	Oscillibacter valericigenes	7.27	5.06
59	Clostridium sp. CRIB	6.88	2.12
60	Prevotella IK062	6.49	0.38
61	Clostridium aminophilum	6.45	2.69
62	Lactococcus garvieae	6.09	2.25
63	Desulfovibrio desulfuricans	5.88	2.31
64	Anaerostipes caccae	5.84	0.62
65	Ruminococcus lactaris	5.78	2.88
66	Paludibacter propionicigenes	5.62	5.25
67	Oscillibacter sp. G2	5.60	3.38
68	Clostridium sp. Clone-46	5.46	1.62
69	Adlercreutzia equolifaciens	5.38	4.25
70	Clostridium sp. Culture-46	5.31	2.38
71	Clostridium scindens	5.24	2.25
72	Clostridium sp. Culture-57	5.18	0.81
73	Dorea longicatena	5.18	2.56
74	Alistipes putredinis	5.12	4.25
75	Clostridium phytofermentans	4.96	0.88
76	Turicibacter sp. LA62	4.89	0.62
77	Clostridium sp. Clone-26	4.67	1.75
78	Clostridium sp. 619	4.51	3.25
79	Clostridium sp. SY8519	4.38	1.50
80	Eubacterium plexicaudatum	4.33	1.00
81	Roseburia hominis	4.28	2.25
82	Clostridium sp. Culture-23	4.26	2.06
		Conti	nued on next page

Tabella 2.1 – continued from previous page

Rank	Species	Mean counts	Median counts
83	Clostridium sp. Clone-44	4.24	1.75
84	Intestinimonas butyriciproducens	4.14	2.31
85	Coprobacillus sp. $8_2_54BFAA$	4.13	1.50
86	Clostridium sp. 826	4.05	1.62
87	Lactobacillus johnsonii	3.91	2.00
88	Ruminococcus flavefaciens	3.66	1.31
89	Clostridium sp. cTPY-12	3.60	1.12
90	Clostridium sp. ASF356	3.56	1.50
91	Parabacteroides merdae	3.29	2.50
92	[Clostridium] cocleatum	3.19	0.25
93	Gordonibacter pamelaeae	3.15	2.56
94	Akkermansia muciniphila	3.15	0.62
95	Roseburia sp. 499	2.92	0.88
96	Clostridium sp. YIT 12070	2.70	0.19
97	Clostridium clostridioforme	2.66	1.12
98	Clostridium saccharolyticum	2.47	1.19
99	Clostridium sp. Clone-47	2.28	0.88
100	Atopobium parvulum	2.26	0.75

As one can see by looking at the table 2.1, there are some species which are high in mean abundance while low in median abundance: these are species which are present during the first days of mice's lives and then drop rapidly to zero counts, meaning they go extinct or very rare. Some examples of such species are Streptococcus AY020, Vibrio cholerae, Eubacterium biforme, Candidatus Arthromitus sp. SFB-mouse, Streptococcus sp. B1 [Figure: 2.5]. In particular, some strains of Vibrio cholerae are pathogenic and responsible for cholera disease. The data shows in that this bacterium is initially present in neonatal mice and then gets suppressed. "Candidatus Arthromitus" sp. strain SFB-mouse-NL (SFB, segmented filamentous bacteria) is a commensal bacterium necessary for inducing the postnatal maturation of homeostatic innate and adaptive immune responses in the mouse gut [Bolotin A 2014].

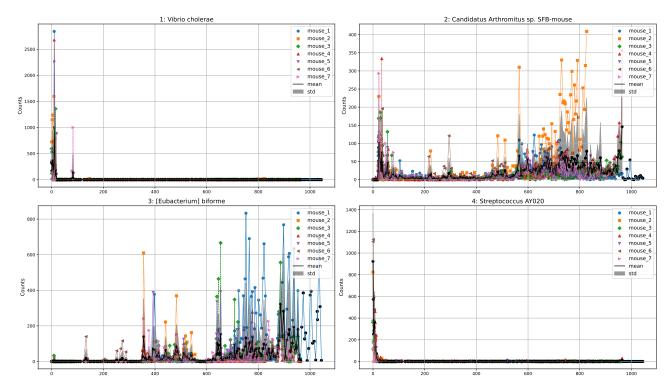


Figura 2.5: Caption

2.4 Gallery of timeseries plots.

## Inference through forward linear regression

Let us suppose that the observed data are produced by an autonomous system of equations of the form, with a multiplicative stochastic noisec term:

$$\dot{x}_i = F(x) + x_i \cdot \sigma \cdot \eta_i$$

where the model f(x) is unknown and  $\eta_i$  is a normally distributed random variable with zero mean and unit variance. Let us also make the (strong) assumption that the observed data evolves closely to a steady state  $x^*, f(x^*) = 0$ . Under the latter assumption, we can linearize the system around the equilibrium:

$$\dot{x}_i(t) = J_F(x^*) \cdot (x - x^*) + o(||x - x^*||) + +x_i^* \cdot \sigma \cdot \eta_i$$

the jacobian  $J_F(x^*)$  is the so-called community matrix at the equilibrium  $x^*$ .

This method essentially consists in inferring the i-th row of the community matrix through a least squares regression of the derivative  $\dot{x_i}$  (covariate variable) versus the abundances  $\{(x_j - x_j^*)\}_{j=1}^N$ . The compositional constraint on the abundances  $(\sum_i x_i(t) \equiv 1 \quad \forall t)$ , poses a technical difficulty, as the design matrix is singular and hence, the usual least squares regression has no unique solution. To bypass this problem, a forward step regression can be performed instead, which also has the benefit of promoting sparsity. Forward stepwise regression consists in adding one regressor at each step, up to a maximum of (N-1) regressors. By doing this, the columns of the design matrix are always linearly independent. The dataset is first divided into two disjoint subsets: a training dataset and a test dataset. Then, one step of the algorithm works like this:

- 1. For each possible new regressor, the least squares fit is performed on the training dataset
- 2. the performance (MSE) of the fits are evaluated on the test dataset
- 3. The new regressor is chosen as the one for which the fit gives the best performance on the test dataset

This basic step is repeated until either the fit improvement falls below a pre-specified threshold, or the maximum number of regressors is reached.

Computational cost can be a concern for this method: in the worst case scenario, the inference of the i-th row requires repeating the least sugares regression (N-1) times. So the overall cost scales as  $\sim N^2$  cost of one single least squares.

Also, one should account for the error-in-variables problem and the collinearity of the covariate and the regressors. The covariates  $x_i(t)$  can be estimated through finite differences:

$$\dot{x}_i = \frac{x_i(t + \Delta t) - x_i(t)}{\Delta t}$$

or

$$\dot{x}_i = \frac{x_i(t + \Delta t) - x_i(t - \Delta t)}{2 \, \Delta t}$$

or higher order formulas. Clearly, for these numerical estimations to be precise, the time interval $\Delta t$ between consecutive measures needs to be sufficiently small. Otherwise, the data should be interpolated, but this introduces a bias.
3.1 Results with simulated data
The data
Una prova su dati simulati che effettivamente evolvono attorno all'equilibri, quindi soltanto per controllare se l'algoritmo funziona come deve. La regressione con le derivate inferisce la matrice jacobiana all'equilibrio (community matrix), che per il LV è $J_{i,j} = a_{i,j} \cdot \overline{x}_i$ , metren la regressione con il modello proposto da fisher inferisce direttamente la matrice delle interazioni, moltiplicata per l'intervallo tra due misure consecutive: $a_{i,j} \cdot \Delta_t$ .

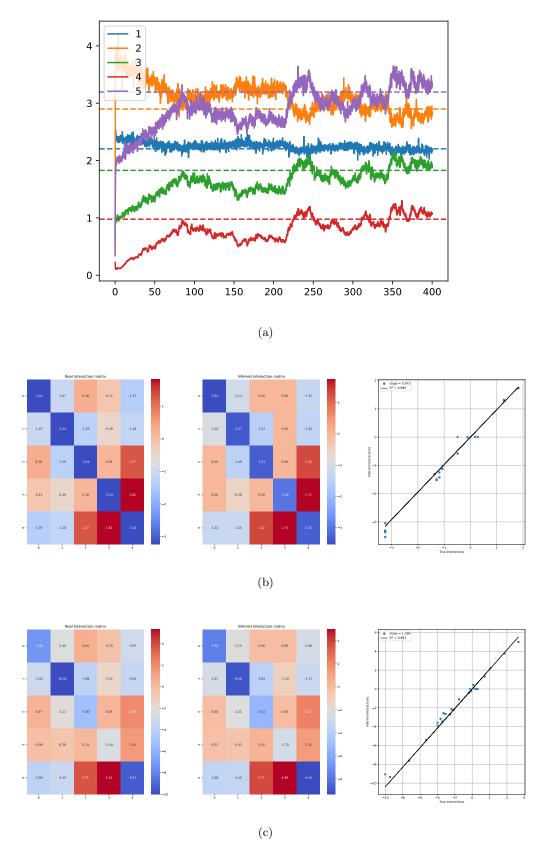


Figura 3.1: Fig (a): Qui ci sono 5 specie, giorni 400, sampling rate 100 misure al giorno, sigma = 5. Soglia impostata a 0.1% e numbero di bagging 100. (Fig. b) metodo di fisher, matrice delle interazioni  $a_{i,j}$  reale (sx) e inferita (dx), (Fig. c) classica regressione sulle derivate, matrice di comunità  $(J_{i,j} = a_{i,j} \cdot \overline{x}_i)$  reale e inferita. Performance buone e grossomodo paragonabili. Ci mancherebbe! Se non funzionasse nemmeno su questi dati sarebbe completamente inutile.

#### 3.2 Results on mice data

The method proposed in [C. Fisher 2014] assumes that:

- 1. the observed data is well described by a generalized Lokta-Volterra model, with discrete timesteps and a stochastic noise term,
- 2. in particular, the observed data is a trajectory in the neighbourhood of a **stable** (locally? globally? not clear from the paper, but not crucial) and **feasible** equilibrium  $\overline{x}$  The equilibrium values  $\overline{x_i}$  can be directly inferred from the data by taking the median (or mean) of the abundaces.

The discrete stochastic Lokta Volterra equation is the following:

$$x_i(t + \Delta t^{update}) = \eta_i(t) \cdot x_i(t) \cdot \exp\left[\Delta t^{update} \sum_{j=1}^{N} C_{ij} \left[ x_j(t) - \bar{x}_j \right] \right]$$
(3.1)

 $\eta_i(t)$  is a lognormally distributed variable, representing a stochastic noise. The reason why it is assumed lognormal is that when taking the logaritm we get a normally distributed noise  $\eta_i(t) \sim \mathcal{N}(0, \sigma)$ .

$$\ln\left[x_i(t + \Delta t^{update})\right] - \ln\left[x_i(t)\right] = \zeta_i(t) + \Delta t^{update} \sum_{j=1}^{N} C_{ij} \left[x_j(t) - \bar{x}_j\right]$$
(3.2)

Equations 3.1 simplify to:

$$\dot{x}_i(t) = x_i(t) \cdot \left[ \sum_j c_{i,j} (x_j - \overline{x}_j) \right]$$

in the limit case  $\Delta t^{(sampling)} \to 0$  and  $\eta_i(t) \equiv 0$ . One can recast these equations into the usual Lokta Volterra equations  $\dot{x}_i(t) = x_i(t) \cdot \left[ r_i + \sum_j a_{i,j} x_j \right]$ , by identifying

$$a_{i,j} \equiv c_{i,j}$$
 and  $r_i \equiv -\sum_j c_{i,j} \, \overline{x}_j$ 

 $[Ax + r]\overline{x} = 0$  is indeed a necessary condition that must be satisfied by a feasible equilibrium.

#### 3.3 Performance evaluation with simulated data

Parameters of the simulation are: sampling matrix C, sampling rate m, noise  $\sigma$ . In the code, the entries of the interaction matrix are normalized to unity. In all the following simulations, the entries of the interaction matrix,  $c_{i,j}$  are normalized to unity. Re-scaling the interactions is perfectly equivalent to rescaling the update interval, keeping the other parameters fixed. A possibility must be considered that the update time scale of the evolution process is different from the time scale of the sampling, most likely smaller. In fact, the experimental measures might be taken at consecutive steps  $\Delta t^{(sampling)} = m \cdot \Delta t^{update}$ , where m > 0 is an integer multiplying  $\Delta t$ . For example, it might that the data is well described by a gdLV assuming a timescale is in the order of magnitude of hours, while the measurements are only made each few days. In this case, the accuracy of the fit might well be low. A partial solution, in order to still apply the method when  $\Delta t^{(sampling)} > \Delta t^{(update)}$  is to interpolate the missing experimental values in order to obtain a timeseries where m = 1. Of course, the more the two timescales differ, the least this approximation will work. Since we do not know the process timescale in advance when approaching data, we shall try by guess and then compare the results. Question for professors: what is the expected time scale of bacteria reproduction?

One issue arising with relative abundances data is that a value of zero counts is ambiguous: one cannot known if the species is truly absent, or just so low in abundance that it does not apppear in the sample.

The fact that the covariate in the regression is a logarithm  $y_i = log(x_it + 1) - log(x_it)$  induces the need to discard all the timepoints where there are zero counts. This makes the inference of interactions impractical for species like those plotted in 2.5. For this method is hence reasonable to select the most abundant species with respect to the median counts instead of the mean counts.

Thing to check in simulations are:

- 1. performance at variying threshold on fit improvement,
- 2. performance at varying species number N,
- 3. performance at varying stochasticity level  $\sigma$

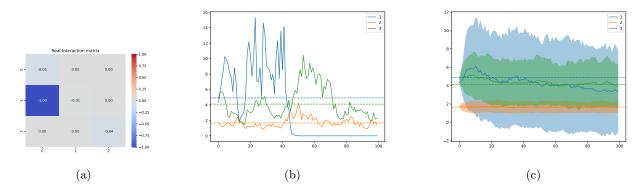


Figura 3.2: Parameters are  $\sigma = 0.2$ . The figure (c) is the mean and std deviation of process repeated 200 times with fixed initiali conditions. Figure (b) is a typical realization of the process.

#### 3.4 Results on mice data

4

## Bayesian inference via the Ising Model

For rare species, i.e. those with abundances below the sampling threshold, we cannot hope to use a model like the Lokta Volterra, simply because we do not have the necessary data. However, from how little we know, we cannot exclude that a species, however low in abundance, can excert a big influence on other species in the ecosystem and prove, indeed, of key global importance. This is why investigation of the interactions among the pletora of rare species in the ecosystem is equally important.

The supplementary issue arising in this case is that the quality of our data is much worse. In fact, unless improvements in the experimental accuracy, the data we can hope for is only binary: presence/absence.

An inferential approach of the kind showed in ?? cannot be applied. Here, instead, we try to employ a framework based upon Bayesian Inference that was developed recently Peixoto 2024. The key assumption of this scheme is that the time series of binary data is well described by assuming an Ising model. However, this is only the simplest option. Indeed, the framework developed by the author is much more general and could be extended in future work to other cases as well.

Will the interactions obtained with this kind of data be coherent with those obtained through the LV? Most probably not, but for now we just make preliminary tentatives.

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