

$$\vec{y} = \vec{A} \cdot \vec{x} + b = \vec{A} \cdot \vec{\hat{x}} = \begin{pmatrix} x_1 \\ x_2 \\ x_3 \end{pmatrix} \cdot \begin{bmatrix} a \\ b \end{bmatrix} \quad \begin{matrix} \hat{x} - \text{estimat} & x - \text{true} \\ \hat{x} - \text{linear estimator} \end{matrix}$$

$$E[\hat{\theta}] = (F^T F)^{-1} F^T E(y) = (F^T F)^{-1} F^T E(F\theta + \epsilon)$$

$$\textcircled{1} \text{ } y = F\theta + \epsilon \quad \text{lsq: } \min_{\theta} \epsilon^T \epsilon \quad \epsilon \sim N(0, I)$$

$$\text{sol.: } \hat{\theta} = (F^T F)^{-1} F^T y$$

assuming θ is deterministic, y is deterministic

$$E[\hat{\theta}] = E\left[\theta + (F^T F)^{-1} F^T \epsilon\right] = \theta + (F^T F)^{-1} F^T E[\epsilon] = \underline{\underline{\theta}}$$

$$\begin{aligned} \text{cov}(\hat{\theta}) &= E[(\hat{\theta} - \theta)(\hat{\theta} - \theta)^T] = E\left[\left(\theta + (F^T F)^{-1} F^T \epsilon - \theta\right)\left(\theta + (F^T F)^{-1} F^T \epsilon - \theta\right)^T\right] \\ &= E\left[(F^T F)^{-1} F^T \epsilon \epsilon^T F (F^T F)^{-1}\right] = (F^T F)^{-1} F^T E(\epsilon \epsilon^T) F (F^T F)^{-1} \end{aligned}$$

$E(\epsilon \epsilon^T) \rightarrow$ since $\epsilon \sim N(0, I)$, the off-diagonal elements will be 0, and the diagonal values will be 1.

$$\downarrow \\ E(\epsilon \epsilon^T) = I$$

$$\text{cov}(\hat{\theta}) = (F^T F)^{-1} F^T I F (F^T F)^{-1} = \underline{\underline{(F^T F)^{-1}}}$$

$$b) (F^T F) = \begin{bmatrix} 6 & 3 \\ 3 & 2 \end{bmatrix} \quad (F^T F)^{-1} = \begin{bmatrix} \frac{2}{3} & -1 \\ -1 & 2 \end{bmatrix}$$

$$(F^T F)^{-1} F^T = \begin{bmatrix} -\frac{1}{3} & \frac{1}{3} & \frac{2}{3} \\ 1 & 0 & -1 \end{bmatrix} \quad \hat{\Theta} = (F^T F)^{-1} F^T y = \begin{bmatrix} \frac{1}{3} \\ 1 \end{bmatrix}$$

$$y - F\hat{\Theta} = \begin{bmatrix} -\frac{1}{3} \\ \frac{1}{3} \\ -\frac{1}{3} \end{bmatrix} \quad \|y - F\hat{\Theta}\|_2^2 = \frac{1}{3}$$

$$c) \hat{\Theta} = [\hat{\Theta}_a \hat{\Theta}_b \hat{\Theta}_n]^T \quad \hat{\Theta}_a \in \mathbb{R} \quad \hat{\Theta}_b \in \mathbb{R} \quad \hat{\Theta}_n \in \mathbb{R}^{n-2}$$

split up $\hat{\Theta}$ just as F

$$\begin{bmatrix} 2 \\ 3 \end{bmatrix} = \begin{bmatrix} 1 & 1 & 4^T \\ 2 & 1 & 34^T \end{bmatrix} \cdot \hat{\Theta} \quad \hat{\Theta}_a + 2\hat{\Theta}_b = 3 \quad \hat{\Theta}_a = 3 - 2\hat{\Theta}_b$$

$$\epsilon = y - F\hat{\Theta} = y - 1_a \hat{\Theta}_a - 1_b \hat{\Theta}_b - M\hat{\Theta}_n = y - M\hat{\Theta}_n - 1_a(3 - 2\hat{\Theta}_b) -$$

$$- 1_b \hat{\Theta}_b = y - 31_a - (1_b - 21_a)\hat{\Theta}_b - M\hat{\Theta}_n$$

$$2 = 3 - 2\hat{\Theta}_b + \hat{\Theta}_b + 4^T \hat{\Theta}_n$$

$$\hat{\Theta}_b = 4^T \hat{\Theta}_n + 1$$

$$\epsilon = (y - 31_a) - (1_b - 21_a) \cdot (4^T \hat{\Theta}_n + 1) - M\hat{\Theta}_n =$$

$$= \underbrace{(y - 31_a - 1_b + 21_a)}_{y_{\text{aug}}} - \underbrace{(1_b 4^T - 21_a 4^T + M)}_{F_{\text{aug}}} \hat{\Theta}_n \rightarrow \text{final least-squares problem}$$

$$\hat{\Theta}_n = (F_{\text{aug}}^T F_{\text{aug}})^{-1} F_{\text{aug}}^T y_{\text{aug}}$$

$$\textcircled{2} \alpha, \quad p(\theta) = \frac{1}{\sqrt{(2\pi)^n |P_\theta|}} \cdot \exp\left(-\frac{1}{2}(\theta - \mu_\theta)^T P_\theta^{-1} (\theta - \mu_\theta)\right) \quad n = \dim(\theta)$$

$$p(y|\theta) = \frac{1}{\sqrt{(2\pi)^n |P_{y|\theta}|}} \cdot \exp\left(-\frac{1}{2}(y - \mu_{y|\theta})^T P_{y|\theta}^{-1} (y - \mu_{y|\theta})\right) \quad n = \dim(y)$$

$$p(\theta|y) = \frac{1}{\sqrt{(2\pi)^n |P_{\theta|y}|}} \exp\left(-\frac{1}{2}(\theta - \mu_{\theta|y})^T P_{\theta|y}^{-1} (\theta - \mu_{\theta|y})\right)$$

$$P_y^{-1} = (LL^T)^{-1} = W = W^T \quad P_\theta > 0 \rightarrow P_\theta = \begin{pmatrix} P_\theta^{-1} & F^T W F \\ 0 & 1 \end{pmatrix}$$

$$\mu_{y|\theta} = F\theta \quad \mu_{\theta|y} = \theta^T F^T$$

$$p(y|\theta) \cdot p(\theta) = \alpha \cdot \exp\left(-\frac{1}{2}(y^T - \theta^T F^T) W (y - F\theta) - \frac{1}{2}(\theta - \mu_\theta)^T P_\theta^{-1} (\theta - \mu_\theta)\right)$$

$$= \alpha \cdot \exp\left(-\frac{1}{2} y^T W y - \frac{1}{2} (\theta^T F^T W F \theta) + \frac{1}{2} \theta^T F^T W y + \frac{1}{2} y^T W F \theta - \frac{1}{2} \theta^T P_\theta^{-1} \theta - \frac{1}{2} \mu_\theta^T P_\theta^{-1} \mu_\theta + \frac{1}{2} \mu_\theta^T P_\theta^{-1} \theta + \frac{1}{2} \theta^T P_\theta^{-1} \mu_\theta\right)$$

$$= \alpha \cdot \exp\left(-\frac{1}{2} \theta^T (F^T W F + P_\theta^{-1}) \theta + \frac{1}{2} (\theta^T F^T W y + y^T W F \theta + \mu_\theta^T P_\theta^{-1} \theta + \theta^T P_\theta^{-1} \mu_\theta) - \frac{1}{2} y^T W y - \frac{1}{2} \mu_\theta^T P_\theta^{-1} \mu_\theta\right)$$

2nd part

$$p(\theta|y) = \beta \cdot \exp\left(-\frac{1}{2} \theta^T P_{\theta|y}^{-1} \theta + \frac{1}{2} (\theta^T P_{\theta|y}^{-1} \mu_{\theta|y} + \mu_{\theta|y}^T P_{\theta|y}^{-1} \theta) - \frac{1}{2} (\mu_{\theta|y}^T P_{\theta|y}^{-1} \mu_{\theta|y})\right)$$

$$\frac{1}{2} (\mu_{\theta|y}^T P_{\theta|y}^{-1} \mu_{\theta|y})$$

1st part

$\theta^T P_{\theta|y}^{-1} \mu_{\theta|y}$ may not be symmetric

$$b) \quad E[(\theta - \mu_0) \epsilon^T] = 0$$

$$y = F\theta + L\epsilon$$

$$\text{unbiased: } E[\mu_{0|y}] = \mu_0$$

$$\text{minimum-variance: } E[(\theta - \mu_{0|y})(\theta - \mu_{0|y})^T] \text{ is minimized}$$

$$\mu_{0|y} = [M \ N] \begin{bmatrix} y \\ \mu_0 \end{bmatrix}$$

μ_0 is deterministic

$$\rightarrow E[\mu_{0|y}] = E[M y] + N \mu_0 = E[M F \theta + M \epsilon] + N \mu_0 = M F E[\theta] + M E[\epsilon] + N \mu_0$$

$$+ N \mu_0 = M F \mu_0 + N \mu_0 = \mu_0 \Rightarrow M F + N = I \quad E[\epsilon] = 0$$

$$N = (I - M F)$$

$$\rightarrow (\theta - \mu_{0|y}) = \theta - M F \theta - M L \epsilon - N \mu_0 = (I - M F) \theta - (I - M F) \mu_0 - M L \epsilon =$$

$$= (I - M F)(\theta - \mu_0) - M L \epsilon$$

$$\rightarrow E[(\theta - \mu_{0|y})(\theta - \mu_{0|y})^T] = E[(I - M F)(\theta - \mu_0) - M L \epsilon][(I - M F)(\theta - \mu_0) - M L \epsilon]^T =$$

$$= E[(I - M F)(\theta - \mu_0)(\theta - \mu_0)^T(I - M F)^T - M L \epsilon(\theta - \mu_0)^T(I - M F) -$$

$$- (I - M F)(\theta - \mu_0) \epsilon^T L^T + M L \epsilon \epsilon^T L^T M] = (I - M F) E[(\theta - \mu_0)(\theta - \mu_0)^T] (I - M F)^T +$$

$$(I - M F)^T - M L E[\epsilon(\theta - \mu_0)^T] (I - M F) - (I - M F) E[(\theta - \mu_0) \epsilon^T] M + M L E[\epsilon \epsilon^T] M$$

$$(E[(\theta - \mu_0) \epsilon^T])^T = 0^T = 0 \quad (LL^T) = W^{-1}$$

$$\rightarrow E[(\theta - \mu_{0|y})(\theta - \mu_{0|y})^T] = (I - M F) P (I - M F)^T + M W^{-1} M^T =$$

$$= [I \ -M] \begin{bmatrix} P & P F^T \\ F P & F P F^T + W^{-1} \end{bmatrix} \begin{bmatrix} I \\ -M^T \end{bmatrix}$$

$$P > 0 \Rightarrow W^{-1} > 0 \quad P > 0 \Rightarrow F P F^T > 0 \Rightarrow F P F^T + W^{-1} > 0$$

since the bottom right entry of Q is positive definite the schur-decomposition can be applied.

$$\rightarrow Q = \begin{bmatrix} A & B \\ C & D \end{bmatrix} = \begin{bmatrix} I & BD^T \\ 0 & I \end{bmatrix} \begin{bmatrix} A - BD^T C & 0 \\ 0 & D \end{bmatrix} \begin{bmatrix} I & 0 \\ D^T C & I \end{bmatrix}$$

$$\rightarrow Q = \begin{bmatrix} I & PF(FPF^T + W^{-1})^{-1} \\ 0 & I \end{bmatrix} \begin{bmatrix} P - PF(FPF^T + W^{-1})^{-1}FP & 0 \\ 0 & FPF^T + W^{-1} \end{bmatrix} \begin{bmatrix} I & 0 \\ (FPF^T + W^{-1})^{-1}FP & I \end{bmatrix}$$

$$\begin{bmatrix} I & 0 \\ (FPF^T + W^{-1})^{-1}FP & I \end{bmatrix}$$

\rightarrow substituting it back into $E(\Theta - \Theta_0)(\Theta - \Theta_0)^T$

$$\rightarrow E[(\Theta - \Theta_0)(\Theta - \Theta_0)^T] = \begin{bmatrix} I & -M \end{bmatrix} \begin{bmatrix} I & PF \\ 0 & FPF^T + W^{-1} \end{bmatrix} =$$

$$= \begin{bmatrix} I & PF(FPF^T + W^{-1})^{-1} - M \end{bmatrix} \begin{bmatrix} P - PF(FPF^T + W^{-1})^{-1}FP & 0 \\ 0 & (FPF^T + W^{-1}) \end{bmatrix} \begin{bmatrix} I \\ (FPF^T + W^{-1})^{-1}FP - M^T \end{bmatrix} =$$

$$= \begin{bmatrix} P - PF(FPF^T + W^{-1})^{-1}FP & PF^T - M(FPF^T + W^{-1}) \end{bmatrix}$$

$$\begin{bmatrix} I \\ (FPF^T + W^{-1})^{-1}FP - M^T \end{bmatrix} = \begin{bmatrix} P - PF(FPF^T + W^{-1})^{-1}FP \\ PF^T - M(FPF^T + W^{-1}) \end{bmatrix} +$$

$$+ \begin{bmatrix} PF^T - M(FPF^T + W^{-1}) \\ (FPF^T + W^{-1})^{-1}FP - M^T \end{bmatrix} \cdot (FPF^T + W^{-1}) \begin{bmatrix} I \\ (FPF^T + W^{-1})^{-1}FP - M^T \end{bmatrix}$$

since the goal is to minimize the above expression one can set either $(PF^T(FPF^T + W^{-1})^{-1} - M)$ or its transpose to 0

$$S_0 \Rightarrow M = P F^T (F P F^T + W)^{-1}$$

$$\begin{aligned} \rightarrow \mu_{0/y} &= P F^T (F P F^T + W)^{-1} y + \mu_0 - P F^T (F P F^T + W)^{-1} F \mu_0 = \\ &= \underline{P F^T (F P F^T + W)^{-1} (y - F \mu_0) + \mu_0} \end{aligned}$$

If $(P F^T (F P F^T + W)^{-1} - M)$ is not lo 0 the remaining term in the expression for $E[(\hat{\theta} - \mu_{0/y})(\hat{\theta} - \mu_{0/y})^T] =$

$$= (P - P F^T (F P F^T + W)^{-1} F P) = P_{0/y}$$

c, since $P > 0$, and the Matrix inversion lemma \downarrow

$$(A + D C D^T)^{-1} = A^{-1} - A^{-1} B (C^T + D A^{-1} B) D A^{-1}$$

$$\rightarrow (F P F^T + W)^{-1} = W^{-1} - W^{-1} F (P^{-1} + F^T W F)^{-1} F^T W, \mu_{0/y} \text{ becomes}$$

$$\begin{aligned} \rightarrow P F^T (W^{-1} - W^{-1} F (P^{-1} + F^T W F)^{-1} F^T W) &= P (F^T W^{-1} - F^T W F (P^{-1} + F^T W F)^{-1} F^T W) = \\ &= P (I - F^T W F (P^{-1} + F^T W F)^{-1}) F^T W = P ((P^{-1} + F^T W F) \cdot (P^{-1} + F^T W F)^{-1} - F^T W F (P^{-1} F W F)) \\ &F^T W = P ((P^{-1} + F^T W F - F^T W F) (P^{-1} + F^T W F)) F^T W = \underline{(P^{-1} + F^T W F) F^T W} \end{aligned}$$