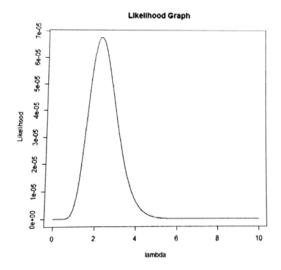
## Problem 5.4:

**1.** 
$$fy(y1 \dots yn|\lambda) = \prod_{i=0}^n \frac{e^{-\lambda}\lambda^{yi}}{y_{i!}} = \frac{e^{-n\lambda}\lambda^{n\bar{y}}}{\prod_{i=1}^n (y_i!)}$$

**2.** 
$$\bar{y} = \frac{1}{5}(2+5+1+0+3) = 2.2$$

$$L(\lambda) = \prod_{i=0}^{5} \frac{e^{-\lambda} \lambda^{yi}}{yi!} = \frac{e^{-5\lambda} \lambda^{5*2.2}}{2! \, 5! \, 1! \, 0! \, 3!} = \frac{e^{-5\lambda} \lambda^{11}}{1440}$$



3. 
$$\frac{\partial L(*)}{\partial x} = \frac{\partial}{\partial x} \left( \frac{e^{-5\lambda} \lambda^{11}}{1440} \right) = \frac{-5e^{-5\lambda} \lambda^{10} (-5\lambda + 11)}{1440} = 0$$

$$\frac{\boldsymbol{\partial}^2}{\boldsymbol{\partial} x^2} \left( \frac{e^{-5\lambda} \lambda^{11}}{1440} \right) < \mathbf{0}$$

Therefore,  $\hat{\lambda} = \frac{11}{5} = 2.2 = \bar{y}$  is the mle of  $\lambda$ .

```
t conf.level=0.95)

Exact Poisson test

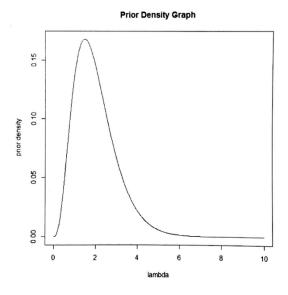
data: 11 time base: 5
number of events = 11, time base = 5, p-value = 0.02043
alternative hypothesis: true event rate is not equal to 1
95 percent confidence interval:
1.098232 3.936408
sample estimates:
event rate
2 2
```

> poisson.test(11,5,alternative = c("two.sided","less","greater"),

The R output has a same **mle of 2.2**. The 95% CI for  $\lambda$  is (1.0982, 3.9364)

## Problem 5.5:

- 1. The prior is a member of Gamma (4,2).
- **2.** The plot indicates that the president's prior belief about rate parameter is maximized around somewhere between 1.6 to 2.5. The majority of expected number of breakdowns are distributed around 0 to 5.



3. 
$$p(\lambda|y) = p(\lambda)L(\lambda) \propto \lambda^3 e^{-2\lambda} \left(\frac{e^{-5\lambda}\lambda^{11}}{1440}\right) = \frac{\lambda^{14}e^{-7\lambda}}{1440} = Gamma(15,7)$$

- **4.** Mean is 15/7 = 2.143
- **5.** Yes, it was. It's a Poisson distribution because they have the same likelihood as the one in problem 5.4.

## Problem 5.6:

| U). | $L(\lambda y) = \frac{e \cdot \lambda x^{y}}{y!} \qquad l(\lambda y) = -\lambda + y \log \lambda - \log(y!)$ $\frac{\partial l}{\partial \lambda} = -1 + \frac{y}{\lambda} \qquad ; \qquad \frac{\partial l}{\partial \lambda^{2}} = -\frac{y}{\lambda^{2}}$ |
|-----|--|
|     | $\Rightarrow I(\lambda y) = -E(\frac{\partial^2 I(\lambda y)}{\partial \lambda^2}) = E(\frac{y}{\lambda^2}) = \frac{1}{\lambda}$ $= I(\lambda y) = h \times I(\lambda y) = \frac{n}{\lambda}$  |
|     | PIN) & JILNIYS = / To of   |
| (>) | PINIY) = PIN) PYIN) ~ JT e-5Nx" = e-5Nx" = Bamma (11.5,5)  |
| (3) | $mean = \frac{11.5}{5} = 2.3$  |

The 95% credible set is (1.1689, 3.8076).

> qgamma(c(0.025,0.975),11.5,rate=5)
[1] 1.168855 3.807563

- **4.** The Bayesian point estimate is **larger** than the classical estimate (2.3 > 2.2). In addition, the interval estimate is **wider** than the classical interval estimate.
- **5.** The employee's estimate is larger than that of the president since the employee has larger  $\lambda$ .
- **6.** Employee:

```
> pgamma(2,11.5,rate=5,lower.tail=FALSE)
[1] 0.6419118
```

**7.** President:

```
> pgamma(2,15,rate=7,lower.tail=FALSE)
[1] 0.5704367
```

(problems in Section 6 are on next page)

Problem 6.1.

$$\frac{\sum_{i=1}^{n} (y_{i} - u_{i})^{2}}{2ut^{2}} = \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2}}{2ut^{2}} = \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2}}{2ut^{2}} = \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2}}{2ut^{2}} + \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2}}{2ut^{2}} + \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2}}{2ut^{2}} + \frac{\sum_{i=1}^{n} (y_{i} - y_{i})^{2}}{2ut^{2}} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2} = \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2} = \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2} = \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2} = \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y - u_{i})^{2} + y_{i}y_{i} = 0$$

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$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

$$= \sum_{i=1}^{n} (y_{i} - y_{i})^{2} + n(y_{i} - y_{i})^{2} + y_{i}y_{i} = 0$$

13) 
$$L(y|\sigma^2) = \frac{1}{(\sigma^2)^{\frac{n}{2}}} \cdot e^{-\frac{\sum (y_i - M)^2}{2\sigma^2}}$$

$$\Rightarrow P(\sigma^2|y) \propto Prior \cdot L = \frac{1}{(\sigma^2)^{\alpha + \frac{n}{2} + 1}} \cdot e^{-\frac{\sum (y_i - M)^2}{2\sigma^2}}$$

$$\sim \frac{16}{(\alpha + \frac{n}{2})} \cdot \frac{\beta + \sum (y_i - M)^2/2}{\beta + \sum (y_i - M)^2/2}$$

Z(41-11)= 106-51)+158-51)+140-51)+ -- +(50-51)= 368 => P(03/4) ~ 2G (38+10, 444+368/2) = 7G (48,628)

(4) Upost = 
$$\frac{\beta^*}{\alpha^*-1} = \frac{628}{48-1} = \frac{13.362}{13.362}$$

Varpost =  $\frac{(\beta^*)^2}{(\alpha^*-1)^2(\alpha^*-2)} = \frac{628^2}{47^246} = \frac{3.8812}{47^246}$ 

151. We assume knowing the true pop, mean of candy-making process. It's unleasonable because the properties of machines are very different.

Therefore, it's not applicable to this case if we want it to be accurate.

(Extra Credit) Problem 6.3:

$$f_{r}(y) = f_{x}(x) \cdot \left| \frac{dx}{dy} \right| = f_{(x)}(\frac{1}{y}) \left| \frac{d}{dy}(\frac{1}{y}) \right|$$

$$= \frac{\beta^{\alpha}}{P(\alpha)} \cdot \left(\frac{1}{y}\right)^{\alpha-1} - \beta(\frac{1}{y}) \cdot \left| \frac{1}{y^{\alpha}} \right|$$

$$= \frac{\beta^{\alpha}}{P(\alpha)} \cdot \left(\frac{1}{y}\right)^{\alpha-1+2} \cdot e^{-\frac{\beta}{y}}$$

$$= \frac{\beta^{\alpha}}{P(\alpha)} \cdot \frac{1}{y^{\alpha+1}} \cdot e^{-\frac{\beta}{y}} \sim \frac{1}{1} \cdot (\alpha, \beta) \square$$