10/1/20 Assume Well-separated "modes" in P(i) P(i), instead of model P(i) as sum of 2 truncated Gaussians: $\begin{cases} \frac{g_1}{\sigma_1 \sqrt{2\pi}} & -(i-\mu_1)^2/2\sigma_1^2 \\ \frac{g_2}{\sigma_1 \sqrt{2\pi}} & e \end{cases}$ $\begin{cases} \frac{g_2}{\sigma_2 \sqrt{2\pi}} & -(i-\mu_2)^2/2\sigma_2^2 \\ \frac{g_2}{\sigma_2 \sqrt{2\pi}} & e \end{cases}$ $\begin{cases} \frac{g_1}{\sigma_1 \sqrt{2\pi}} & \frac{g_2}{\sigma_2 \sqrt{2\pi}} \\ \frac{g_2}{\sigma_2 \sqrt{2\pi}} & e \end{cases}$ $\begin{cases} \frac{g_1}{\sigma_1 \sqrt{2\pi}} & \frac{g_2}{\sigma_2 \sqrt{2\pi}} \\ \frac{g_2}{\sigma_2 \sqrt{2\pi}} & \frac{g_2}{\sigma_2 \sqrt{2\pi}} \end{cases}$ $H(t) = -\sum_{i=0}^{t} P(i) \log \left[\frac{g_1}{\sigma_1 \sqrt{2\pi}} e^{-(i-\mu_1)^2/2\sigma_1^2} - \sum_{i=t+1}^{255} P(i) \log \left[\frac{g_2}{\sigma_2 \sqrt{2\pi}} e^{-(i-\mu_2)^2/2\sigma_2^2} \right] \right]$

STRAINDS \$178543773 STEERS OF 1885.52 STRAINDS - CARACTER STEERS ON GENERAL STEERS OF 1885.54 STRAINDS \$178543745 STEERS ON STEERS OF 1885.54

=
$$-\frac{t}{\sum_{i=0}^{2}} P(i) \left[\log q_1 - \log \sigma_1 - \log \sqrt{2\pi} - \frac{(i-\mu_1)^2}{2\sigma_1^2} \right]$$

 $-\frac{255}{\sum_{i=0}^{2}} P(i) \left[\log q_2 - \log \sigma_2 - \log \sqrt{2\pi} - \frac{(i-\mu_2)^2}{2\sigma_2^2} \right]$
 $i=t+1$

$$= \frac{1}{2} + \frac{1}{2} \log_{1} 2\pi - 9, \log_{1} - 9, \log_{2} - 9, \log_{2} 2$$

$$+ 9, \log_{1} + 9, \log_{2} = H(t)$$

Choose the t that minimizes H(t).

Steps:

- 1 Compute histogram, h(i)
- (2) Normalize: $P(i) = \frac{h(i)}{\#pixels}$
- 3 For each t, compute H(t).
- 4) Pick the t having the smallest H(t).

Recursive Calculation
$$\begin{array}{l}
g_{1}(0) = P(0), \quad g_{1}(t) = g_{1}(t-1) + P(t) \\
g_{2}(t) = 1 - g_{1}(t) \\
M_{1}(0) = 0 \quad & \text{if } P(i) \\
M_{1}(t) = \frac{1 - 0}{g_{1}(t)} & \text{if } P(i) \\
&= \frac{1 - 0}{g_{1}(t-1)} & \text{if } P(i) + P(t) \\
g_{1}(t) & \text{if } P(i) + P(t) \\
g_{1}(t) & \text{if } P(i) & \text{if } P(i) \\
M_{2}(t) = \frac{M - g_{1}(t)M_{1}(t-1) + P(i)}{g_{2}(t)} \\
M_{2}(0) = \frac{M}{g_{2}(0)}
\end{array}$$

$$M_{2}(0) = \frac{M}{g_{2}(0)}$$

$$\sigma_{1}^{2}(0) = 0$$

$$\sigma_{1}^{2}(t) = \frac{1}{g_{1}(t)} \left(g_{1}(t-1) + \left[\mu_{1}(t-1) - \mu_{1}(t)\right]^{2}\right)$$

$$+ P(t) \left[t - \mu_{1}(t)\right]^{2}$$

$$+ P(t) \left[t - \mu_{1}(t)\right]^{2}$$

$$\sigma_{2}^{2}(0) = \sum_{i=1}^{255} \left[i - \mu_{2}(0)\right]^{2} \frac{P(i)}{g_{2}(0)}$$

$$\sigma_{2}^{2}(t) = \frac{1}{g_{2}(t)} \left(g_{2}(t-1) - \mu_{2}(t)\right]^{2}$$

$$- P(t) \left[t - \mu_{2}(t)\right]^{2}$$

Note: • Use double-precision for M, g, o's

• Problem at ends of histogram:

log of = log 0 => bad

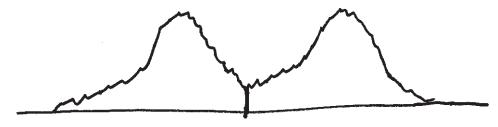
Similar prob. at rt end. with 52.

Let "first" = index of first-occupied hist. bin. "last" = Index of last-occupied hist bin. Assume hist [first+1] = 0 hist [lest-] 70 In the iteration, only try cardidate threshold values in first+1 = t = lest-2 If we begin with t = first + 1, then $\mu_i(first+1) = \sum_{i=1}^{n} i P(i) / g_i(first+1)$

where 9. (first+1) = P(first) + P(first+1)

Can determine similar init. For other parms.

Kittles's model: sum of truncated Gaussians:



M, in Kittler's alg is computed using a distribution whose right tail was truncated. So, M, is biased too small. Similarly, Mz is biased too large.

Cho's correction terms:

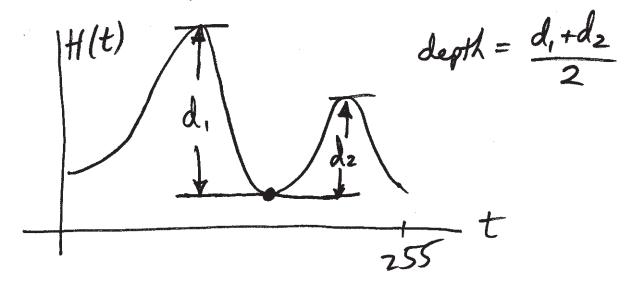
$$\hat{\mu}_{i} = \mu_{i} + \frac{\sigma_{i}}{\phi\left(\frac{t-\mu_{i}}{\sigma_{i}}\right)} \cdot \frac{1}{\sqrt{2\pi}} e^{-\left(t-\mu_{i}\right)^{2}/2\sigma_{i}^{2}}$$
where $\phi(x) = 1 - \text{erfc}(x)$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{u^{2}}{2}} du$$

 $\hat{\mu}_{2} = \mu_{2} - \frac{\sigma_{2}}{1 - \phi \left(\frac{t - \mu_{2}}{\sigma_{2}}\right)} \frac{1}{\sqrt{2\pi}} e^{-\left(t - \mu_{2}\right)^{2}/2\sigma_{2}^{2}}$

Cho: Minimum Finding

For more robustness, they found the deepest valley instead of the global min. of H(t).



Multi-Level Thresholding See Kittler.

Other Histogram-Based Segm. Methods

Li, Gong, & Chen

Pattern Recog, 1997

They found that a hist. of local ave.

gray level gave better thresholding

performance.

Chow & Kaneko

1972 - See Haralick & Shapiro,

Computer & Robot Vision, v. 1.

Local histogramming (adaptive thresh.)

- · divide imp into disjoint blocks (e.s., 33x33 or 65x65)
- · compute local hist & block
- a determine a thresh. For each block
- . Associate the thresholds with the block centers.
- . Interpolate the thresholds. to get thresholds for other pixels.

We52ka J.S, RN. Nagel, & A. Rosenfeld

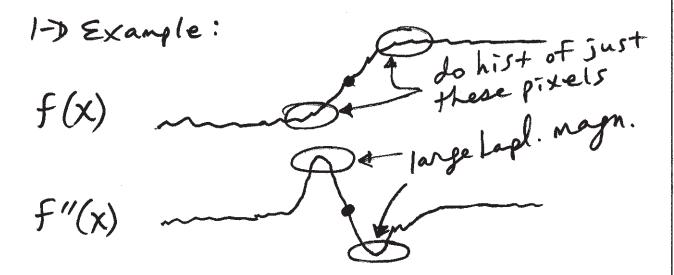
"A Threshold Selection Techn."

IEEE T-Computers V. C-23, 1974, pp 1322-6.

Laplacian-based method.

To List, thresholding for only the pixels

Do hist. thresholding for only the pixels with large Laplacian magnitude.



- · This includes the fixels adj. to the edges, but not the edges themselves.
- . The resulting hist should have a sharper valley.

orig hist

hist of the pixels selected pixels

of fy & bg pixels, making the modes about the same size.

Watanabe, S. CGIP, v.3, 1974, pp 350-8.

Gradient-based scheme.

Choose the threshold that maximizes
the sum of gradients at all pixels
whose gray level equal the thresh.
i.e., grad. magn.

Weska & Rosenfeld

T-SMC, v. 8, 1978, P622-9.

Busyness-based scheme.

· H candidate thresh, compute busyness = % of pixels that are border pixels (either fg or by border pixels). Choose the

f(x) ~~~~

· Choose the thresh that minimized busyness.

Kohler's Method

"A Segm. Sys. Based on thresholding;" CGIP, v. 15, 1981, pp 319-38.

Contrast-based scheme.

- · Let T = thresh.
- · Define edges to be

$$E(T) = \{[(i,j),(k,l)]:$$

pixels (i, j) & (k, l) are neighbors,

and min $\{I(i,j), I(k,l)\} \leq T \leq$

 $\max \{I(i,j), I(k,l)\}$

· Défine total contrast to be

$$C(T) = \underbrace{S'}_{min} \sum_{j=1}^{min} I(i,j)-T_{j},$$

$$\underbrace{[(i,j),(K,l)]}_{E(T)} |I(K,l)-T|$$

- · Ave contrast of edges detected by threshold T is $\frac{C(T)}{\#E(T)}$
- · Choose the thresh that meximizes are contrast.