I. Parametric Models: Prior Information II. From Models to Answers

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Machine Learning Jan 25, 2017



Recall: Your first consulting job

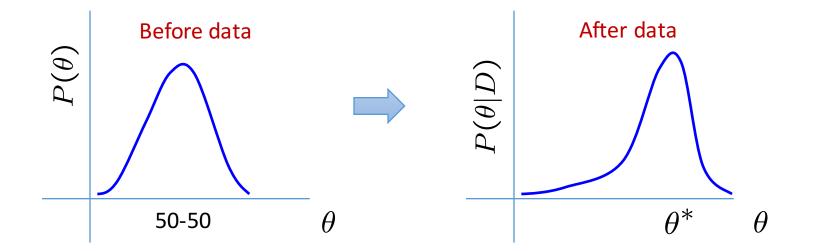
- A billionaire from the suburbs of Seattle asks you a question:
 - He says: I have a coin, if I flip it, what's the probability it will fall with the head up?
 - You say: Please flip it a few times:



- You say: The probability is: 3/5 because... frequency of heads in all flips
- He says: But can I put money on this estimate?
- You say: ummm.... Maybe not.
 - Not enough flips (less than sample complexity)

What about prior knowledge?

- Billionaire says: Wait, I know that the coin is "close" to 50-50.
 What can you do for me now?
- You say: I can learn it the Bayesian way...
- Rather than estimating a single θ , we obtain a distribution over possible values of θ



Bayesian Learning

Data

Use Bayes rule:

Parameters

$$P(\theta \mid \mathcal{D}) = \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

likelihood

prior

Bayes, Thomas (1763) An essay towards solving a problem in the doctrine of chances. *Philosophical Transactions of the Royal Society of*

London, 53:370-418

Bayesian Learning

Use Bayes rule:

$$P(\theta \mid \mathcal{D}) = \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

Or equivalently:

$$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)$$
 posterior likelihood prior



Bayes, Thomas (1763) An essay towards solving a problem in the doctrine of chances. *Philosophical Transactions of the Royal Society of London*, 53:370-418

AIDS test (Bayes rule)

Data

- Approximately 0.1% are infected
- Test detects all infections
- Test reports positive for 1% healthy people

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Probability of having AIDS if test is positive:

$$P(a = 1|t = 1) = \frac{P(t = 1|a = 1)P(a = 1)}{P(t = 1)}$$

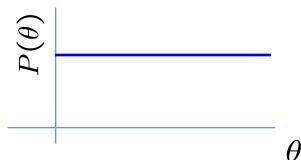
$$= \frac{P(t = 1|a = 1)P(a = 1)}{P(t = 1|a = 1)P(a = 1) + P(t = 1|a = 0)P(a = 0)}$$

$$= \frac{1 \cdot 0.001}{1 \cdot 0.001 + 0.01 \cdot 0.999} = 0.091$$

Only 9%!...

Prior distribution

- From where do we get the prior?
 - Represents expert knowledge (philosophical approach)
 - Simple posterior form (engineer's approach)
- Uninformative priors:
 - Uniform distribution



- Conjugate priors:
 - Closed-form representation of posterior
 - $P(\theta)$ and $P(\theta \mid D)$ have the same algebraic form as a function of \theta

Conjugate Prior

• $P(\theta)$ and $P(\theta \mid D)$ have the same form as a function of theta

Eg. 1 Coin flip problem

Likelihood given Bernoulli model:

$$P(\mathcal{D} \mid \theta) = \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

If prior is Beta distribution,

$$P(\theta) = \frac{\theta^{\beta_H - 1} (1 - \theta)^{\beta_T - 1}}{B(\beta_H, \beta_T)} \sim Beta(\beta_H, \beta_T)$$

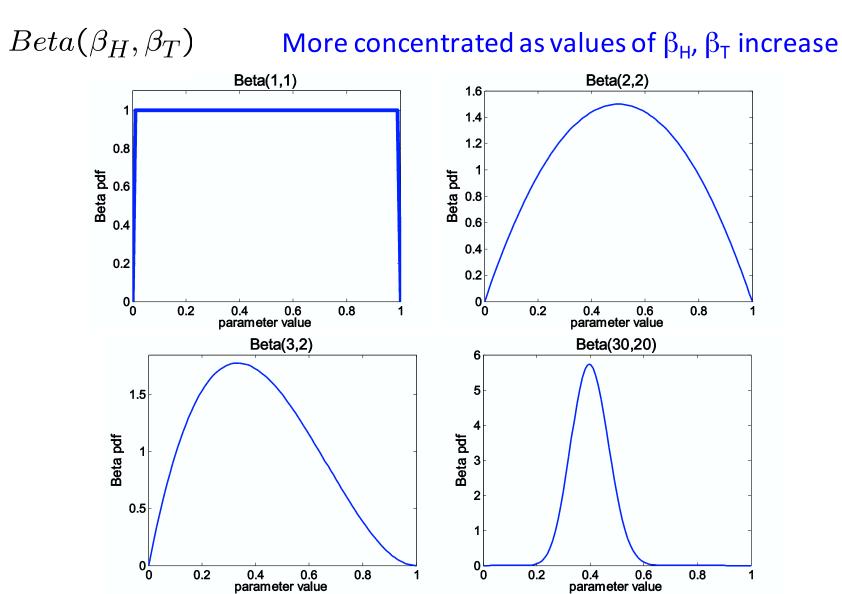
Then posterior is Beta distribution

$$P(\theta|D) \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

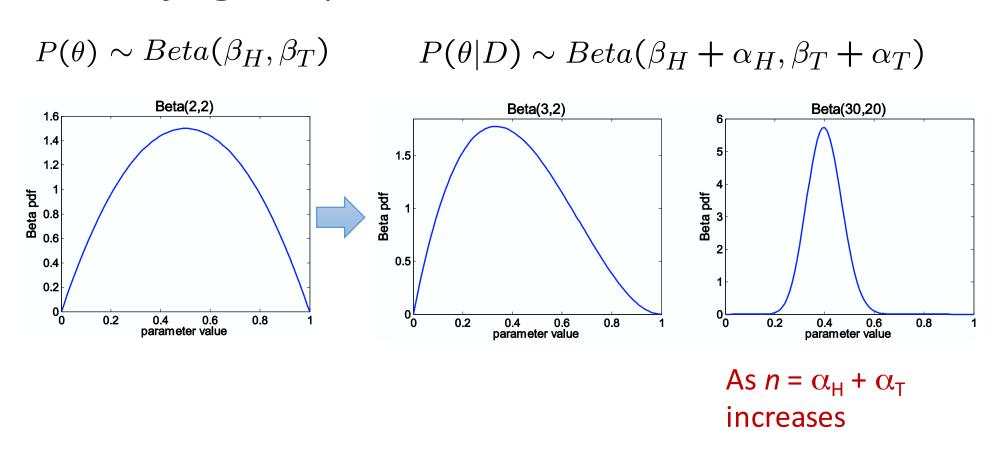
For Binomial, conjugate prior is Beta distribution.



Beta distribution



Beta conjugate prior



As we get more samples, effect of prior is "washed out"

Conjugate Prior

• $P(\theta)$ and $P(\theta \mid D)$ have the same form

Eg. 2 Dice roll problem (6 outcomes instead of 2)



Likelihood is ~ Multinomial($\theta = \{\theta_1, \theta_2, ..., \theta_k\}$)

$$P(\mathcal{D} \mid \theta) = \theta_1^{\alpha_1} \theta_2^{\alpha_2} \dots \theta_k^{\alpha_k}$$

If prior is Dirichlet distribution,

$$P(\theta) = \frac{\prod_{i=1}^{k} \theta_i^{\beta_i - 1}}{B(\beta_1, \dots, \beta_k)} \sim \text{Dirichlet}(\beta_1, \dots, \beta_k)$$

Then posterior is Dirichlet distribution

$$P(\theta|D) \sim \text{Dirichlet}(\beta_1 + \alpha_1, \dots, \beta_k + \alpha_k)$$

For Multinomial, conjugate prior is Dirichlet distribution.

Posterior Distribution

- The approach seen so far is what is known as a **Bayesian** approach
- Prior information encoded as a **distribution** over possible values of parameter
- Using the Bayes rule, you get an updated posterior distribution over parameters, which you provide with flourish to the Billionaire
- But the billionaire is not impressed
 - Distribution? I just asked for one number: is it 3/5, 1/2, what is it?
 - How do we go from a distribution over parameters, to a single estimate of the true parameters?

Maximum A Posteriori Estimation

Choose θ that maximizes a posterior probability

$$\widehat{\theta}_{MAP}$$
 = $\underset{\theta}{\operatorname{arg\,max}}$ $P(\theta \mid D)$
= $\underset{\theta}{\operatorname{arg\,max}}$ $P(D \mid \theta)P(\theta)$

MAP estimate of probability of head:

$$P(\theta|D) \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

$$\widehat{\theta}_{MAP} = rac{lpha_H + eta_H - 1}{lpha_H + eta_H + lpha_T + eta_T - 2}$$
 Mode of Beta distribution

MLE vs. MAP

Maximum Likelihood estimation (MLE)

Choose value that maximizes the probability of observed data

$$\widehat{\theta}_{MLE} = \arg \max_{\theta} P(D|\theta)$$

Maximum a posteriori (MAP) estimation
 Choose value that is most probable given observed data and prior belief

$$\widehat{\theta}_{MAP} = \arg \max_{\theta} P(\theta|D)$$

$$= \arg \max_{\theta} P(D|\theta)P(\theta)$$

When is MAP same as MLE?

MLE vs. MAP

$$\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$



What if we toss the coin too few times?

- You say: Probability next toss is a head = 0
- Billionaire says: You're fired! ...with prob 1 ©

$$\widehat{\theta}_{MAP} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2}$$

- Beta prior equivalent to extra coin flips
- As $n \rightarrow \text{infty}$, prior is "forgotten"
- But, for small sample size, prior is important!

MLE vs MAP

You are no good when sample is small



You give a different answer for different priors

MAP for Gaussian mean and variance

- Conjugate priors
 - Mean: Gaussian prior
 - Variance: Wishart Distribution

Prior for mean:

$$P(\mu \mid \eta, \lambda) = \frac{1}{\lambda \sqrt{2\pi}} e^{\frac{-(\mu - \eta)^2}{2\lambda^2}} = N(\eta, \lambda^2)$$

MAP for Gaussian Mean

$$\widehat{\mu}_{MLE} = \frac{1}{n} \sum_{i=1}^{n} x_i$$

$$\widehat{\mu}_{MAP} = \frac{\frac{1}{\sigma^2} \sum_{i=1}^n x_i + \frac{\eta}{\lambda^2}}{\frac{n}{\sigma^2} + \frac{1}{\lambda^2}}$$

MAP of Gaussian variance - Later

Prior Information

- In the Bayesian approach, the prior information is encoded through a prior distribution over the parameters
- Seems onerous: the distribution typically seems to be obtained from convenience (conjugate distribution)
- What other ways can we encode our prior knowledge about the parameters?
- A non-Bayesian approach is via constraints

Encoding prior information via constraints

MLE:

$$\max_{\theta} \log \mathbb{P}(D; \theta).$$

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Constrained MLE:

$$\max_{\theta} \log \mathbb{P}(D; \theta)$$

s.t.
$$\mathcal{R}(\theta) \ll C$$
.

Encoding prior information via constraints

MLE:

$$\max_{\theta} \log \mathbb{P}(D; \theta).$$

Constrained MLE:

$$\max_{\theta} \log \mathbb{P}(D; \theta)$$
s.t. $\mathcal{R}(\theta) \le C$.

When $\mathcal{R}(\theta)$ is convex, constrained MLE is equivalent to regularized MLE:

$$\max_{\theta} \left\{ \log \mathbb{P}(D; \theta) + \lambda \mathcal{R}(\theta) \right\}.$$

Regularized MLE

Regularized MLE:

$$\max_{\theta} \left\{ \log \mathbb{P}(D; \theta) + \lambda \mathcal{R}(\theta) \right\}.$$

Trades off maximizing the log-likelihood (i.e. fit to data), against the "prior" constraints encoded by regularization (which do not involve the data at all).

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The MAP estimator can be seen to be a special case by simply setting

$$\lambda \mathcal{R}(\theta) = \log P(\theta).$$

Here, the tradeoff between likelihood and prior is naturally captured by setting the regularization function equal to the log of the prior distribution.

Popular Regularization functions

• ℓ_2 regularization:

$$\mathcal{R}(\theta) = \|\theta\|_2^2 = \sum_{j=1}^p \theta_j^2.$$

This regularization encodes the prior information that the parameter values are not too large (where how large is determined by the regularization tradeoff parameter λ).

This regularization is thus a "general purpose" regularization function (who wants their parameters to be very large?)

Popular Regularization functions

• ℓ_1 regularization:

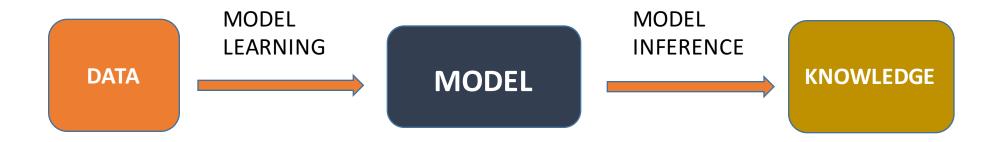
$$\mathcal{R}(\theta) = \|\theta\|_1 = \sum_{j=1}^p |\theta_j|.$$

This regularization encodes the prior information that the parameter values be **sparse**: i.e. with many zero values.

This is a very important prior constraint in big data settings: with very large number of parameters, we expect the true model to depend on only a few non-zero parameters.

Widely used in high-dimensional model learning: called LASSO when used with linear regression models.

Recall: Model-based ML



- Learning: From data to model
 - A model thus is a summary of the data
 - But also a story of how the data was generated
 - Could thus be used to describe how future data can be generated
 - E.g. given (symptoms, diseases) data, a model explains how symptoms and diseases are related
- Inference: From model to knowledge
 - Given the model, how can we answer questions relevant to us
 - E.g. given (symptom, disease) model, given some symptoms, what is the disease?

Model to Knowledge

- We now know how to learn a model from data, with guarantees
- How do we go from model to knowledge?

- i.e. How do we get the answers we seek from the model?
- E.g. the Billionaire might be really after answers to questions such as:
 - Which side is more likely in the next flip?
 - If a bookie gives 3 to 5 odds on tails, should he take the bet?

Model to Knowledge: Plugin Estimates

- In most cases, the knowledge we seek is a fixed function f(P) of the distribution of the data
 - E.g. is the coin fair? Does the coin have better odds than 3/5, etc.
- Once we learn a model, we have an estimate of the distribution of the data: $P_{\widehat{\theta}}$
- So we can simply "plugin" the model for the distribution to get our answers: $f(P_{\widehat{\theta}})$
- Is the coin fair: $\mathbb{I}(\theta == 1/2)$
 - Plugin Estimate: $\mathbb{I}(\widehat{\theta} == 1/2)$

Specification of Knowledge

- In the previous, the specification of what knowledge we were seeking was through an explicit function of the distribution
 - E.g. is the coin fair? Does the coin have better odds than 3/5, etc.
- But such an explicit specification is not always possible
- An important construct in machine learning is a language for an implicit specification of task/what knowledge we seek
 - Through "performance measures"
 - Whenever you encounter a task, you should automatically think about the appropriate performance measure

Supervised Learning Prediction Task

Task:

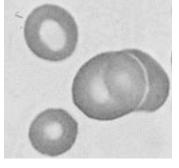
Given $X \in \mathcal{X}$, predict $Y \in \mathcal{Y}$.

 \equiv Construct **prediction rule** $f: \mathcal{X} \rightarrow \mathcal{Y}$





"Anemic cell (0)"





"Healthy cell (1)"

Performance Measure:

loss(Y, f(X)) - Measure of closeness between true label Y and prediction f(X)

 $loss(Y, f(X)) = 1_{\{f(X) \neq Y\}}$

0/1 loss

Performance:

Measure:

loss(Y, f(X)) - Measure of closeness between true label Y and prediction f(X)

X	Share price, Y	f(X)	loss(Y, f(X))
Past performance, trade volume etc. as of Sept 8, 2010	"\$24.50"	"\$24.50"	0
		"\$26.00"	1?
		"\$26.10"	2?

 $loss(Y, f(X)) = (f(X) - Y)^2$ square loss

Performance:

Measure:

loss(Y, f(X)) - Measure of closeness between true label Y and prediction f(X)

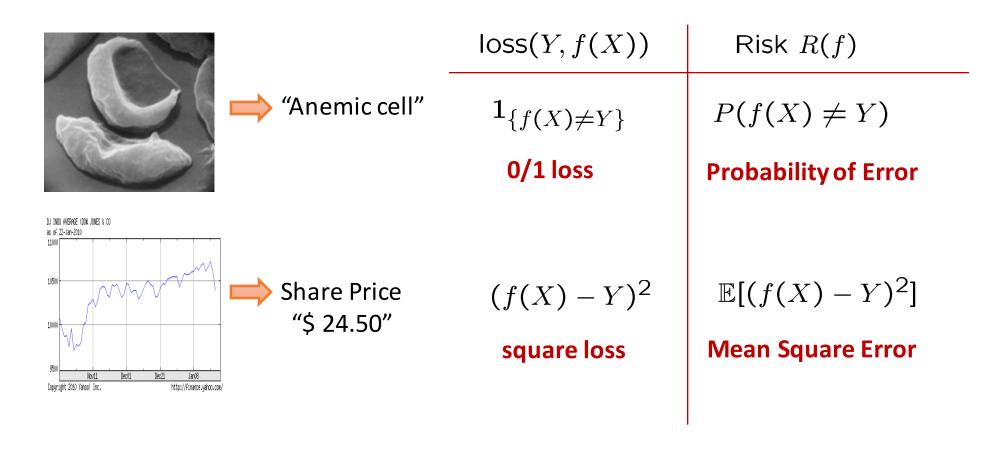
Don't just want label of one test data (cell image), but any cell image $X \in \mathcal{X}$

$$(X,Y) \sim P_{XY}$$

Given a cell image drawn randomly from the collection of all cell images, how well does the predictor perform on average?

Risk
$$R(f) \equiv \mathbb{E}_{XY} [loss(Y, f(X))]$$

Performance: Risk $R(f) \equiv \mathbb{E}_{XY} \left[loss(Y, f(X)) \right]$ Measure:



Bayes Optimal Rule

Knowledge
That we seek:

Construct **prediction rule** $f^*: \mathcal{X} \to \mathcal{Y}$

$$f^*(P) = \arg\min_f \mathbb{E}_{(X,Y)\sim P}[\log(Y, f(X))]$$

Bayes optimal rule

Best possible performance:

Bayes Risk
$$R(f^*) \leq R(f)$$
 for all f

Bayes Optimal Rule

Knowledge That we seek: Construct **prediction rule** $f^*: \mathcal{X} \to \mathcal{Y}$

 $f^*(P) = \arg\min_f \mathbb{E}_{(X,Y)\sim P}[\log(Y,f(X))]$ Bayes optimal rule

loss(Y, f(X))

0/1 loss

 $1_{\{f(X)\neq Y\}}$

Risk R(f)

$$P(f(X) \neq Y)$$

Probability of Error

Bayes Optimal Rule $f^*(P)$

$$f^*(P) = \mathbb{I}(P(Y=1|X) > 1/2)$$

$$(f(X)-Y)^2$$

square loss

 $\mathbb{E}[(f(X) - Y)^2]$

Mean Square Error

$$f^*(P) = \mathbb{E}(Y|X)$$

Bayes Optimal Rule

Knowledge
That we seek:

Construct **prediction rule** $f^*: \mathcal{X} \to \mathcal{Y}$

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Bayes optimal rule

Best possible performance:

Bayes Risk
$$R(f^*) \leq R(f)$$
 for all f

BUT... Optimal rule is not computable

- depends on unknown distribution P over (X,Y)!

Use a model for P_{XY} !

Model-free Methods

<u>Knowledge</u>
<u>That we seek:</u>

Construct **prediction rule** $f^*: \mathcal{X} \to \mathcal{Y}$

$$f^*(P) = \arg\min_f \mathbb{E}_{(X,Y)\sim P}[\log(Y, f(X))]$$

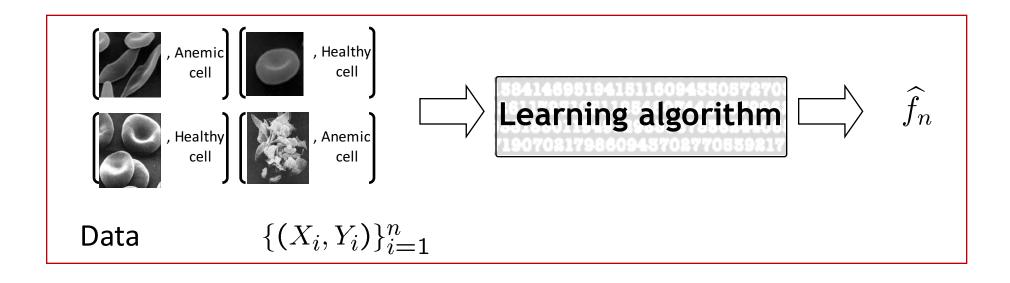
Bayes optimal rule

We could use a model for P_{XY} !

But can we estimate the knowledge through some learning algorithm that does not go through a model?

A model-free approach for ML

Model-free Methods



$$\widehat{f}_n$$
 is a mapping from $\mathcal{X} o \mathcal{Y}$ \widehat{f}_n $=$ "Anemic cell" Test data X

Popular Approach for model-free ML: **Empirical Risk Minimization**

Knowledge

Construct **prediction rule** $f^*: \mathcal{X} \to \mathcal{Y}$

That we seek:

$$f^*(P) = \arg\min_f \mathbb{E}_{(X,Y)\sim P}[\log(Y,f(X))]$$
 Bayes optimal rule

Given $\{X_i, Y_i\}_{i=1}^n$, **learn** prediction rule $\widehat{f}_n:\mathcal{X} o\mathcal{Y}$

Empirical Risk

Minimizer:
$$\widehat{f}_n = \arg\min_{f} \frac{1}{n} \sum_{i=1}^{n} [loss(Y_i, f(X_i))]$$

$$\frac{1}{n} \sum_{i=1}^{n} \left[\mathsf{loss}(Y_i, f(X_i)) \right] \xrightarrow{\mathsf{Law} \ \mathsf{of} \ \mathsf{Large}} \mathbb{E}_{XY} \left[\mathsf{loss}(Y, f(X)) \right]$$

Consistency and Rate of Convergence

 How does the performance of the algorithm compare with ideal performance?

Excess Risk
$$\mathbb{E}_{D_n}\left[R(\widehat{f}_n)\right] - R(f^*)$$

- Consistent algorithm if Excess Risk $\rightarrow 0$ as n $\rightarrow \infty$
- Rate of Convergence

