## STA1000F Summary

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## 1 Module 1: Probability

## 1.1 Work Unit 1: Introducing Probability

#### 1.1.1 Definitions

- 1. Random Experiment: A procedure whose outcome (result) in a particular performance (trial) cannot be predetermined.
- 2. Long Run Pattern (Average): the average result over a large number of trials
- 3. Random Variation: This implies that we never know the value of the next random event
- 4. Statistical Distributions: Distributions of data
- 5. **Fair Game:** No one wins or loses in the long run
- 6. **House Advantage:** The profit made by the house (casino)

#### 1.1.2 Formula

1.

$$Win\% = \frac{Total\ Payout\ for\ a\ winning\ number}{amount\ to\ be\ bet\ over\ all\ numbers} \times 100$$

2.

$$Win\% \times fair\ payout = payout$$

3.

$$fair\ payout - fp \times HA = payout$$

4.

$$fair\ payout = (probability\ of\ winning)^{-1} \times bet$$

5.

$$HA = 100 - Win\%$$

## 1.1.3 Examples

- 1. Die:  $S = \{1,2,3,4,5,6\} P(6) = \frac{1}{6} P(even) = \frac{3}{6} = \frac{1}{2}$
- 2. Odds: odds of  $\{6\} = 1.5$

# 1.2 Work Unit 2: Set Theorem, Probability Axioms and Theorems

#### 1.2.1 Definitions

- 1. Sets can be determined by a list of elements  $(A=\{e,f,g,1,2\})$  or a rule  $(B=\{x|1\leq x\leq 10,\ x\in \mathbb{Z}\})$
- 2. Order and repetition in sets is irrelevant  $\{1, 3, 4, 4, 2\} = \{1, 2, 3, 4\}$
- 3. Element member of set  $(e \in A)$  vs Element not member  $(e \notin B)$
- 4. Subsets: if  $G \subset H$  and  $G \supset H$  then G = H

- 5. Intersection:  $A \cap B = \{1, 2\}$
- 6. **Union:**  $A \cup B = \{e, f, g, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$
- 7. Complement:  $S = \{1, 2, 3, 4\}, C = \{1, 2\}, \bar{C} = \{3, 4\}$
- 8. Empty Set:  $\emptyset = \{\}$
- 9. Universal Set (Sample Sapce) S all possible outcomes of a random experiment
- 10. Mutually Exclusive (Disjoint): If  $L \cap M = \emptyset$
- 11. Pairwise Mutually Exclusive, Exhaustive Sets:  $A_1, A_2, \dots, A_n$  s.t.  $A_i \cap A_j = \emptyset$  if  $i \neq j$  and  $A_1 \cup A_2 \cup \dots \cup A_n = S$
- 12. **Event:** Subset of sample space (S = certain event,  $\emptyset$  = impossible event)
- 13. **Elementary Event:** Event with one member  $(A = \{3\})$  Always mutually exclusive,  $P(A) = \frac{n(A)}{n(S)}$ . NB not  $\emptyset$
- 14. A occurs if the outcome of the trial is a member of A
- 15. **Relative Frequency:**  $\sqrt[r]{n}$ , r= number of times A occurs, n = number of trials,  $0 \le r/n \le 1$ ,  $P(A) = \lim_{n \to \infty} r/n$

#### 1.2.2 Formula

1.

$$(A \cup B) \cap C = (A \cap C) \cup (B \cap C)$$

2.

$$A \cup B = A \cup (B \cap \bar{A})$$

3. Kolmogorov's Axioms of Probability  $S = \text{sample space}, \forall A \subset S, P(A) \in \mathbb{R} \text{ st}$ 

(a) 
$$0 \le P(A) \le 1$$

- (b) P(S) = 1
- (c) If  $A \cap B = \emptyset$  then  $P(A \cup B) = P(A) + P(B)$
- (d) Consequence:  $P(\emptyset) = 0$

4.

Let 
$$A \subset S$$
 then  $P(\bar{A}) = 1 - P(A)$ 

5.

If 
$$A \subset S$$
 and  $B \subset S$  then  $P(A) = P(A \cap B) + P(A \cap \overline{B})$ 

6.

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

7.

If 
$$B \subset A$$
 then  $P(B) \leq P(A)$ 

8. If  $A_1, \ldots, A_n$  are pairwise mutually exclusive then:  $P(\bigcup_{i=1}^n A_i) = \sum_{i=1}^n P(A_i)$ 

## 1.3 Work Unit 3: Permutations and Combinations

## 1.3.1 Counting Rules

1. Permutation: order matters, repetition not allowed

n!

2. Permutation: order matters, repetition not allowed

$$(n)_r = \frac{n!}{(n-r)!}$$

3. Permutations: order matters, repetition allowed

 $n^r$ 

4. Combinations: Order doesn't matter, repetition not allowed

$$\binom{n}{r} = \frac{n!}{r!(n-r)!}$$

# 1.4 Work Unit 4: Conditional Probability and Independent Events

1. Conditional Probability:

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

2. Complement of Conditional Probability:

$$P(P|D) = 1 - P(\bar{P}|D)$$

3. Baye's Theorem:

$$P(D|P) = \frac{P(P|D)P(D)}{P(P|D)P(D) + P(P|\bar{D})P(\bar{D})}$$

4. Baye's Theorem for mutually exclusive exhaustive events:

$$P(A_i|B) = \frac{P(B|A_i)P(A_i)}{\sum_i P(B|A_i)P(A_i)}$$

5. Independent Events: (never mutually exclusive) A and B independent

$$P(A|B) = P(A)$$

$$P(A \cap B) = P(A)P(B)$$

$$P(A_1 \cap ... \cap A_n) = P(A_1) \times ... \times P(A_n)$$

## 2 Module 2: Exploring Data

## 2.1 Work Unit 1: Graphical Summaries

- Pie Charts: Categories add up to 100%; Units comparable; Slice reflect proportion; order biggest to smallest
- Bar Charts: order biggest to smallest
- Histograms: chose a number between ½L and 2L; all class intervals must be the same, Bimodal has 2 peaks

#### 2.1.1 Definitions

- Qualitative Data (Catagorical or Nominal Data): e.g. Nationality/hair colour (use Pie Chart or Bar Chart)
- 2. Quantitative Data (Fully Numeric): Count/measure captured on scale (no fixed 0) or a ratio (explicit 0) e.g. Weight/height/distance (Use histogram, scatter plot, box plot)
- 3. Ordinal Data: Ranked or ordered data, steps don't have to be the same size, e.g. level of satisfaction/education

#### 2.1.2 Formula

1. Interval Width

$$L = \frac{X_{max} - X_{min}}{\sqrt{n}}$$

or

$$L = \frac{X_{max} - X_{min}}{1 + loq_2 n} = \frac{X_{max} - X_{min}}{1 + 1.44loq_e n}$$

2. Trendline (linear regression)

$$y = a + bx$$

a = intercept, b = slope

3. Explainatory Value  $\mathbb{R}^2$  The amount of variation in Y that can be explained by X

# 2.2 Work Unit 2: Summary Measures of Location and Spread

#### 2.2.1 Definitions

- 1. **Statistic:** quantity calculated from the data values of a sample (subset of population)
- 2. Parameter: statistic calculated on the population

- 3. **5 Number Summary:** min; lower quartile; median; upper quartile; max
- 4. **Fences:** largest and smallest observations that aren't strays; whiskers in box-and-whisker plot go up to these when you also show outliers and strays

#### 2.2.2 Measures of Location

- 1. Median: Robust; not affected by big outliers/strays
- 2. **Mean:** sensitive to outlying values; useful for symmetric distributions

## 2.2.3 Measures of Spread

- 1. Range: Unreliable/sensitve
- 2. IQR: Robust
- 3. **Standard Deviation:**  $(\bar{x} s, \bar{x} + s)$  contains  $\frac{2}{3}$  of your observations

#### 2.2.4 Formula

- 1. Standard Deviation =  $\sqrt{Variance}$
- 2. Variance

$$s^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (x_{i} - \bar{x})^{2} = \frac{1}{n-1} \left[ \sum_{i=1}^{n} x_{i}^{2} - \frac{1}{n} \left( \sum_{i=1}^{n} x_{i} \right)^{2} \right]$$

3. Update Variance:

$$s* = \sqrt{\frac{1}{n}[(n-1)s^2 + n(\bar{x} - \bar{x}*)^2 + (x_{n+1} - \bar{x}*)^2]}$$

4. Mean

$$\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$$

5. Update Mean:

$$\bar{x}* = \frac{n\bar{x} + x_{n+1}}{n+1}$$

6. Strays

$$< Median - 3 \times (Median - LQ)$$

$$> Median + 3 \times (UQ - Median)$$

7. Outliers

$$< Median - 6 \times (Median - LQ)$$

$$> Median + 6 \times (UQ - Median)$$

- 8. Median =  $X_{(m)}$  where m = n + 1/2
- 9. LQ =  $X_{(l)}$  where l = [m] + 1/2
- 10. UQ =  $X_{(u)}$  where u = n l + 1
- 11. Range =  $X_{max} X_{min} = X_n X_1$
- 12.  $IQR = x_{(u)} x_{(l)}$
- 13. Half Rank:  $X_{(r+1/2)} = x_{(r)} + x_{(r+1)}/2$

## 2.3 Module 3: Random Variables

# 2.4 Work Unit 1: Probability Mass and Density Functions

- P(X = x) = P(x), where x is specific value of random variable X
- you can assign numbers to qualitative events

#### 2.4.1 Definitions

- 1. **Discrete Random Variable:** set of possible values is finite or countably infinite
  - (a) Probability Mass Function: p(x)
  - (b) Defined for all values of x, but non-zero at finite (or countably infinite) subset of these values
  - (c)  $0 \le p(x) \le 1 \ \forall x$
  - (d)  $\sum p(x) = 1$
  - (e)  $P(a \le x < b) = \sum_{x=a}^{b-1} p(x)$
- 2. Continuous Random Variable:  $\{x | a < x < b\}$ 
  - (a) Probability Density Function: f(x)
  - (b) Defined for all values of x
  - (c)  $0 \le f(x) \le \infty \ \forall x$
  - (d)  $\int_{-\infty}^{\infty} p(x) = 1$  check on non zero interval only
  - (e)  $P(a < X \le b) = P(a \le X < b) = P(a \le X \le b) = P(a < X < b) = \int_a^b p(x)$
  - (f) P(X = a) = 0
  - (g) can be measured to any degree of accuracy

## 2.5 Work Unit 2: Working with Random Variables

#### 2.5.1 Definitions

- Long Run Average: Expected value of random variable X (E(X)), weighted sum of possible values of X, also called mean, can have theoretical value (value not in S e.g. 3.5 is E(X) for dice)
- 2.  $\mu = E(X)$  and  $\sigma^2 = Var(X)$

#### 3. Discrete Random Variable

(a) 
$$E(X) = \sum xp(x)$$

(b) 
$$Var(X) = \sum (x - E(X))^2 p(x) = (\sum x^2 p(x)) - E(X)^2$$

(c) 
$$E(X^r) = \sum x^r p(x)$$

### 4. Continuous Random Variable

(a) 
$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

(b) 
$$Var(X) = E(X^2) - E(X)^2 = \int_a^b (x - E(X))^2 f(x) dx = (\int_a^b x^2 f(x) dx) - E(X)^2$$

(c) 
$$E(X^r) = \int_a^b x^r f(x) dx$$

## 5. Expected Value:

(a) 
$$E(A+B) = E(A) + E(B)$$

(b) 
$$E(A - B) = E(A) - E(B)$$

(c) 
$$E(cA) = cE(A)$$
 where c is constant

(d) 
$$Y = aX + b$$
 then  $E(Y) = aE(X) + b$ 

#### 6. Variance:

(a) 
$$Var(A+B) = Var(A) + Var(B)$$

(b) 
$$Var(A - B) = Var(A) + Var(B)$$

(c) 
$$Var(cA) = c^2 Var(A)$$
 where c is constant

(d) 
$$Y = aX + b$$
 then  $Var(Y) = a^2 Var(X)$ 

## 7. Coefficient of Variation (CV) =

$$\frac{\sqrt{Var(X)}}{E(X)}$$

only if lower limit of X is 0

- 8. in graphs of pdf, pmf peaked = small var, flat = large var
- 9. -vely skewed peak on right, symmetric peak in middle, +vely skewed peak on left
- 10. **Heavy-tailed Distributions:** Probability of observations far from the mean is relatively large
- 11. **Light-tailed distributions:** observations far from the mean are unlikely

## 2.6 Module 4: Probability Distributions

## 2.7 Work Unit 1: Uniform Distribution

$$X \sim U(a,b)$$

PDF:

$$f(x) = \left\{ \begin{array}{ll} \frac{1}{b-a} & a \leq x \leq b \\ 0 & elsewhere \end{array} \right.$$

Integrate this to get the probability

Use this when you have a continuous random variable that is equally likely to lie between a and b and impossible to lie outside this interval.

$$E(X) = \frac{1}{2}(b+a)$$

$$Var(X) = \frac{(b-a)^2}{12}$$

Distribution Function:

$$F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \le x \le b \\ 1 & x > b \end{cases}$$

### 2.8 Work Unit 2: Binomial Distribution

$$X \sim B(n, p)$$

PMF:

$$p(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & x = 0, 1, 2...n \\ 0 & elsewhere \end{cases}$$

If we are observing the number of successes in n (fixed number) independent trials of an experiment in which the outcome of each trial can only be success or failure with constant probability

$$E(X) = np$$
$$Var(X) = np(1-p)$$

## 2.9 Work Unit 3: Poisson and Exponential Distributions

#### 2.9.1 Poisson Distribution

$$X \sim P(\lambda = average)$$

PMF:

$$p(x) = \begin{cases} \frac{e^{-\lambda}\lambda^x}{x!} & x = 0, 1, 2... \\ 0 & elsewhere \end{cases}$$

Models number of events, occurring randomly with an average rate of occurence per unit of time/space/distance

$$E(X) = Var(X) = \lambda$$

## 2.9.2 Exponential Distribution

$$X \sim E(\lambda = average/unit)$$

PDF:

$$f(x) = \left\{ \begin{array}{ll} \lambda e^{-\lambda x} & x > 0 \\ 0 & elsewhere \end{array} \right.$$

Models space/distance between events, occurring randomly, with an average rate of occurrence

$$E(X) = \frac{1}{\lambda}$$
$$Var(X) = \frac{1}{\lambda^2}$$

### 2.10 Work Unit 4: Normal Distribution

Pattern of averages: If the random variable X is the sum of a large number of random increments then X has a normal distribution (Central Limit Theorem)

EG: Height of trees, amount of stuff in a jar Continuous so PDF

$$X \sim N(\mu, \sigma^2)$$
 
$$p(x < \mu) = 0.5 \quad p(x > \mu) = 0.5$$

PDF:

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-1}{2}(\frac{x-\mu}{\sigma})^2}$$
$$E(X) = \mu$$
$$Var(X) = \sigma^2$$

Can't integrate analytically so use the tables in introstat. first convert to z:

$$z = \frac{x - \mu}{\sigma}$$

Tables give P(0 < Z < z) so convert accordingly knowing that P(Z < 0) = 0.5 and the distribution is symmetric about 0.

Sum:

$$X_i \sim N(\mu_i, \sigma_i^2)$$
  $Y = \sum X_i$  then  $Y \sim N(\sum \mu_i, \sum \sigma_i^2)$ 

Difference:

$$X_1 \sim N(\mu_1, \sigma_1^2)$$
  $X_2 \sim N(\mu_2, \sigma_2^2)$   $Y = X_1 - X_2$  
$$Y \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$$

Multiplying by Constant:

$$X \sim N(\mu, \sigma^2)$$
  $Y = aX + b$   $Y \sim N(a\mu + b, a^2\sigma^2)$ 

If you want to find  $P(X>z^{(p)})=p$ , search for p or closest value in table.  $z^{(0.1)}=$  upper 10%, lower  $10\%=-z^{(0.1)}=z^{(0.9)}=$  upper 90%

## 2.11 Module 5: Hypothesis Testing

## 2.12 Work Unit 1: Sampling Distribution

Population

- Parameter greek
- mean =  $\mu$
- variance =  $\sigma^2$

vs Sample

- Statistic roman
- mean =  $\bar{x}$
- variance =  $s^2$
- statistic easier to measure
- can draw inference about population

#### Steps

- 1. Draw Random Sample
- 2. Measure sample statistic
- 3. use this as an estimate of the true unknown population parameter

#### Sample must be:

- 1. **Representative:** similar in structure to the population it is drawn from
- 2. **Random:** Every member of the population has an equal chance of being chosen as part of the sample

Statistics will vary due to random sample so a statistic is a random variable so  $\bar{x}$  is a random variable with a probability distribution called a sampling distribution.

 $\sum$  elements in a normal distribution has a normal distribution. So for  $X_i$  drawn randomly from a normal distribution:

$$\sum_{i=1}^{n} X_i \sim N(n\mu, n\sigma^2)$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$$
  $E(\sum_{i=1}^{n} X_i) = n\mu$ 

Multiplying by constant  $\frac{1}{n}$ :

$$E(\bar{X}) = \frac{1}{n}n\mu = \mu$$

Variance:

$$Var(\bar{X}) = Var(\frac{1}{n}\sum_{i=1}^{n}X_i) = \frac{1}{n^2}Var(\sum_{i=1}^{n}X_i) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

so:

$$\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$$

What if  $X_i$  is not normally distributed?

**Central Limit Theorem:** The average (or sum divided by n) of a large number of variables always has a normal distribution. How Large? n = 30 is sufficient

so we have 3 means:

- 1. sample
- 2. probability distribution (expected value)
- 3. Population

Note sample mean is  $\bar{X}$  because  $\bar{x}$  is a specific value.

Sample mean based on a large n has a smaller variance so closer to  $\mu$ 

#### 2.13 Work Unit 2: Confidence Intervals

**Point Estimate:** No information regarding the uncertainty of the estimate

vs **Interval:** range of values, communicate how much precision or uncertainty is present in the estimate

$$Pr(\bar{X}-z^{(\frac{\alpha}{2})}\frac{\sigma}{\sqrt{n}}<\mu<\bar{X}+z^{(\frac{\alpha}{2})}\frac{\sigma}{\sqrt{n}})=1-\alpha=2L$$

we are  $100(1-\alpha)\%$  confident that the value of  $\mu$  lie in this interval choose  $z^{(\frac{\alpha}{2})}$  from table such that  $P(0 < z < z^{(\frac{\alpha}{2})}) = L$ , given  $\sigma^2$  from population

$100(1-\alpha)\%$	$\mathbf{z}^{(\frac{\alpha}{2})}$
95%	1.96
90%	1.645
98%	2.33
99%	2.58

$$L = z^{(\frac{\alpha}{2})} \frac{\sigma}{\sqrt{n}}$$

$$n = \left(\frac{z^{(\frac{\alpha}{2})}\sigma}{L}\right)^2$$

width of confidence interval = 2L, if  $\alpha = 0.05$  then confidence interval = 95%.

increase n, confidence interval narrows. Increase z confidence interval widens

# 2.14 Work Unit 3: Testing whether the mean is a specific value

6 Step Hypothesis test

1. Define null hypothesis  $(H_0)$ 

$$H_0: \mu = 0$$

2. Define alternative hypothesis  $(H_1)$ 

$$H_1: \mu < a$$

Or > (one-sided) or  $\neq$  (2 sided)

3. Set significance level

$$\alpha = 0.05$$

You will erroneously reject  $H_0$   $\alpha\%$  of the time

- 4. Set up rejection region Find  $z^{\alpha/2}$  (2-sided) or  $z^{\alpha}$  (one-sided)
- 5. Calculate the test statistic (Assume  $H_0$  is true)

$$X \sim N(\mu_0, \frac{\sigma^2}{n})$$

$$z = \frac{\bar{X} - \mu_0}{\sqrt{\frac{\sigma^2}{n}}}$$

6. Draw a conclusion:

If test statistic fall in rejection region reject  $H_0$  in favour of  $H_1$  else do not reject  $H_0$ 

The default is a two-sided test unless you have good reason to suspect a departure in a specific direction.

Remember to split  $\alpha$  in a two sided test

### 2.14.1 Errors

- 1. Type 1
  - Reject  $H_0$  erroneously
  - $\bullet$  controlled by  $\alpha$
  - $\alpha$  small reduces probability of this error
  - $P(T_1E) = \alpha$
- 2. Type 2
  - Accept  $H_0$  erroneously
  - $\alpha$  small increases probability of this error
  - $P(T_2E)$  varies dependent on how close  $H_0$  is to the true situation, so difficult to control

## 2.15 Work Unit 4: Comparing 2 Sample means

EG Test if 2 dies come from same or different populations

To compare look at difference:

$$\bar{X}_1 - \bar{X}_2 \sim N(\mu_1 - \mu_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2})$$

Hypothesis Test

- 1.  $H_0: \mu_1 = \mu_2 \text{ or } H_0: \mu_1 \mu_2 = 0$
- 2.  $H_1: \mu_1 \neq \mu_2$
- 3. set  $\alpha$
- 4. Find rejection region (in this case 2 sided test but can be one-sided)
- 5. Calculate test statistic:

$$z = \frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

## 2.15.1 The Modified Approach

Don't specify a significance level, instead observe significance level based on the test statistic.

Test Statistic:

- if  $H_1$  is  $\neq$  (2-sided):  $p-value = P(z > |teststat|) <math>\times 2$
- if  $H_1$  is > (1-sided): p value = P(z > TestStat)
- if  $H_1$  is < (1-sided): p-value = P(z < TestStat)

If  $H_0$  is true we would observe a difference of at least the size  $X_1-X_2$  p-value% of the time.

Small p-value means  $H_0$  unlikely

reject  $H_0$  if p-value < 0.05 (remember p-value is prob of type 1 error)

## 2.16 Work Unit 5: Tests about the mean when we don't know the variance

Estimate  $\sigma^2$  from  $s^2$ 

- Now two random variables  $\bar{X}$  and  $s^2$
- Test statistic now t-test
- still bell shaped and symmetric but flatter-fatter tails
- increase n looks more normal, smaller n heavier tails
- t distribution

$$t = \frac{X - \mu}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$$

has n-1 degrees of freedom

In table: degrees of freedom along left column, % > along top

Confidence interval

$$\bar{X} \pm t_{n-1}^{\alpha/2} \frac{s}{\sqrt{n}}$$

Hypothesis test the same, just use t-table

Alternative Hypothesis test, for p-value look along n-1 row fro largest value that the test statistics exceeds

Why n-1 Degrees of freedom?

If given  $\bar{x}$  and  $x_1, ..., x_{n-1}$  can determine  $x_n$ , this is why sample variance is multiplied by 1/n-1

General Rule: For each parameter we need to estimate  $(s^2)$  prior to evaluating the current parameter of interest  $(\bar{X})$ , we lose 1 degree of freedom

for 
$$n > 30$$
:  $s^2 \approx \sigma^2$ 

# 2.17 Work Unit 6: Comparing Means of 2 Independent Samples

1. Define null Hypothesis

$$H_0: \mu_1 = \mu_2$$

$$H_0: \mu_1 - \mu_2 = 0$$

2. Define Alternative Hypothesis

$$H_1: \mu_1 \neq \mu_2$$

$$H_0: \mu_1 - \mu_2 \neq 0$$

- 3. Define significance level  $\alpha$
- 4. Rejection region
- 5. Test Statistic (t-test)

$$t = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

But this is wrong because it does not have a t-dist

6. so use pooled variance:

$$s_p = \sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{(n_1 + n_2 - 2)}}$$

Assuming population variances are equal i.e.  $s_1^2$  and  $s_2^2$  are viewed as estimates of the same true variance

7. t becomes:

$$t = t = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1 + n_2 - 2}$$

8. conclusion:

$$|test\ statistic| > t_{n_1+n_2-2}^{lpha/2}$$

Use closest degrees of freedom if the one you are looking for isn't in the table

Also can use modified approach

# 2.18 Work Unit 7: Comparing Means of 2 Dependent Samples

What matters is difference (change) so Before - After Data is paired

Hypothesis Test

1. Define null Hypothesis

$$H_0: \mu_B = \mu_A$$

$$H_0: \mu_B - \mu_A = 0$$

2. Define Alternative Hypothesis

$$H_1: \mu_B \neq \mu_A$$

$$H_0: \mu_B - \mu_A \neq 0$$

can be 2-sided

- 3. Define significance level  $\alpha$
- 4. Rejection region
- 5. Test Statistic

$$d = X_B - X_A$$

degrees of freedom is n-1, where n is number of pairs

$$t = \frac{\bar{d} - \mu}{\frac{s_d}{\sqrt{n}}}$$

6. conclusion

if two-sided double p-value

Critical Point: paired data are not independent of each other - repeated measures (i.e. same people)

confidence interval

$$\bar{d} \pm t_{n-1}^{\alpha/2} \frac{s_d}{\sqrt{n}}$$

# 2.19 Work Unit 8: Testing whether data fits a specific distribution

Goodness of fit test: check what we observe in a sample with what we expect under a specific hypothesis

6 Step Approach

- 1.  $H_0$ : X has some pdf/ pmt
- 2.  $H_1$ : X has some other distribution (always 1 sided test, don't split  $\alpha$ )
- 3.  $\alpha = 0.05$
- 4. Chi-squared distribution
  - has degrees of freedom
  - skewed to the right
  - always positive
  - df = number of categories-num parameters we estimate-1
- 5. Test Statistic

$$D^{2} = \sum_{i=1}^{k} \frac{(O_{i} - E_{i})^{2}}{E_{i}} = \sum_{i=1}^{k} \frac{(O_{i})^{2}}{E_{i}} - n$$

k = number of comparisons, O = observed, E = expected value

6. conclusion

Modified approach: calculate p-value from table: find largest value which is smaller than the test statistic

Lose a degree of freedom for each parameter you estimate

Need an expected value of at least 5 in each category, otherwise collapse categories. e.g. Poisson dis either choose  $\lambda$  and test or get  $\lambda$  from data, in which case df = n-2

For normal  $\mu$  and  $\sigma^2$  are estimated so df = n-3

There is a mathematical relationship between the normal and chisquared distributions

# 2.20 Work Unit 9: Testing for an association between 2 categoric variables

Table, there is an association between rows (e.g. gender) and columns (e.g. job level). Counts in cells, assume data is random and representative

### 6 Step Approach

- 1.  $H_0$ : there is no association between rows and columns
- 2.  $H_1$  there is an association
- 3.  $\alpha = 0.05$
- 4.  $D^2 > \chi_{df}^{2\alpha}$  assume  $H_0$  is true. one sided, compare observed and expected values DF = ([no rows]-1)([no cols]-1)
- 5. Test Statistic want  $E_i$  is the variables are independent. Remember

$$P(A \cap B) = P(A)P(B)$$

if A and B independent. so:

$$E_{ij} = \frac{Row_i Total \times Col_j Total}{Grand Total}$$

now use:

$$D^2 = \sum \left(\frac{O_i^2}{E_i}\right) - n$$

6. Conclusion

Modified approach: calculate p-value from table or excel

# 2.21 Work Unit 10: Testing for a predictive relationship between 2 Numeric Variables

Linear relationship between 2 quantitative random variables:

$$y = a + bx$$

y = dependent variable

x = independent variable

a + b = regression coefficients

Use correlation coefficient - true value unknown so estimate from sample:

 $\rho = \text{population correlation (parameter)}$ 

r = Sample Correlation (statistic)

$$-1 \le r \le 1$$

r=-1 perfect negative (x inc, y dec), r=1 perfect positive (x inc, y inc), r=0 variables independent

$$r = \frac{\sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})}{(n-1)s_x s_y}$$

Coefficient of determination:

$$R^2 = r \times r \quad 0 \le R^2 \le 1$$

measure the property of variation in y that x is able to explain.  $1-R^2$  is the proportion of variation in Y that is explained by factors other than X

$$y = \alpha + \beta x$$

 $\alpha = y$  intercept

 $\beta =$  slope. estimate from sample

Closer —r— or  $\mathbb{R}^2$  to 1 the better the regression model fits the data, closer to 0, the worse the fit

Hypothesis Testing:

- 1.  $H_0: \beta = 0$  no linear relationship
- 2.  $H_1: \beta \neq 0$  or  $H_1: \beta < 0$  or  $H_1: \beta > 0$
- 3.  $\alpha = 0.01$
- 4. Rejection Region: test stat  $\sim t_{n-2}$  where n = number of pairs of x and y (check other notes for tests)
- 5. Test statistic

## 3 Excel

- 1. = Rand()
- 2. = If(cond, value, or)
- 3. =countif(start:end, equals)