MECH7710 - HW1

Random Variables and Probability

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Problem 1

Part A - 6 dice numbered 1,2,3,4,5,6

```
pdf_matrix = 1/6 * ones(6,6);
pdf_combined = nfoldconv(pdf_matrix);
[mean, variance] = statistics(pdf_combined);
true_norm = normpdf([1:length(pdf_combined)], mean, sqrt(variance));
if plots,
    createfigure(true_norm, pdf_combined, 'Problem 1 - Part A')
end
```

```
mean =
   21.0000
variance =
   17.5000
```

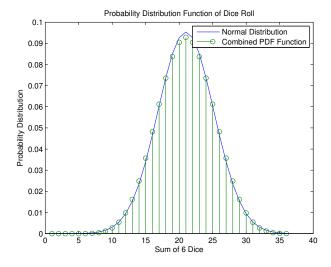


Figure 1: PDF of a Dice Roll - Problem 1A

Part B - 6 dice numbered 4,5,6,7,8,9

```
pdf_matrix = 1/6 * [zeros(6,3),ones(6,6)];
pdf_combined = nfoldconv(pdf_matrix);
[mean, variance] = statistics(pdf_combined);
true_norm = normpdf([1:length(pdf_combined)], mean, sqrt(variance));
if plots,
    createfigure(true_norm, pdf_combined, 'Problem 1 - Part B')
end
```

```
mean =
    39.0000
variance =
    17.5000
```

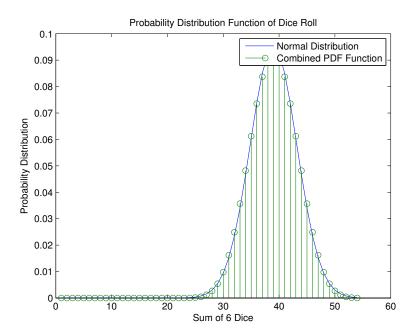


Figure 2: PDF of a Dice Roll - Problem 1B

Part C - 6 dice numbered 1,1,3,3,3,5

```
mean =
    16
variance =
    11.3333
```

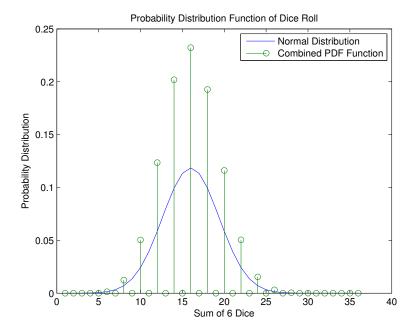


Figure 3: PDF of a Dice Roll - Problem 1C

Part D - 3 dice numbered 1,2,3,4,5,6 and 3 numbered 1,1,3,3,3,5

mean = 18.5000 variance = 14.4167

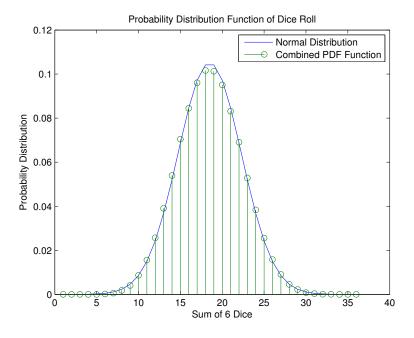


Figure 4: PDF of a Dice Roll - Problem 1D

Part A - Mean, Central Moment, Mean Squared, Variance and Covariance

```
pdf_die_1 = 1/6 * ones(1,6);
pdf_die_2 = 1/6 * ones(1,6);

die = 1:6;
joint_pdf = pdf_die_1' * pdf_die_2;

[mean_1, variance_1, c_moment_1, mean_sq_1] = statistics(pdf_die_1);
```

```
mean_v1 =
    3.5000

variance_v1 =
    2.9167

c_moment_v1 =
    0

mean_sq_v1 =
    15.1667
```

Part B - Covariance Matrix

```
% Covariance of 2 independant variables is 0.
covariance = sum(sum((die-mean_1)'*(die-mean_1) * joint_pdf));
P = [variance_1, covariance; covariance, variance_1]
```

```
covariance = 0
P = 2.9167 0
0 2.9167
```

Part C, D, E, and F

Find the PDF matrix for the variables $v_1 = x_1$ and $v_2 = x_1 + x_2$. Use [1:6, 1:12], with the first column zeros.

```
v_1 = [1:6]; v_2 = [1:12];
   joint_pdf = zeros(6,12); joint_v = zeros(6,12);
  for ii=1:6,
       joint_pdf(ii, ii+[1:6]) = 1/36;
       for jj=1:6,
           joint_v(ii,ii+jj) = ii+jj;
       end
  \quad \text{end} \quad
   [mean_v1, variance_v1, c_moment_v1, mean_sq_v1] = statistics(pdf_die_1)
  mean_v2 = sum(sum(joint_v .* joint_pdf));
mean_sq_v2 = sum(sum(joint_v.^2 .* joint_pdf));
  c_moment_v2 = sum(sum((joint_v - mean_v2) .* joint_pdf));
  variance_v2 = sum(sum((joint_v - mean_v2).^2 .* joint_pdf));
  covariance_12 = sum(sum((v_1 - mean_v1))'*(v_2 - mean_v2).*joint_pdf));
20
  P = [variance_v1, covariance_12;
       covariance_12, variance_v2]
   rho_{12} = covariance_{12}/(sqrt(P(1,1)) * sqrt(P(2,2)))
```

```
mean_v2 =
   7.0000
mean_sq_v2 =
   54.8333
c_moment_v2 =
  1.9151e-15
variance_v2 =
    5.8333
covariance_12 =
    2.9167
    2.9167
            2.9167
   2.9167
            5.8333
rho_12 =
    0.7071
```

Two random vectors are called uncorrelated if P=0

Part A

Show that independent random vectors are uncorrelated.

$$P(x_1, x_2) = E\{(x_1 - \bar{x_1})(x_2 - \bar{x_2})^T\} = 0$$
(1)

$$= E[x_1x_2 - \bar{x_1}x_2 - x_1\bar{x_2} + \bar{x_1}\bar{x_2}$$
 (2)

$$= E[x_1 x_2] - E[\bar{x_1} x_2] - E[x_1 \bar{x_2}] + E[\bar{x_1} \bar{x_2}]$$
(3)

$$= E[x_1 x_2] - \bar{x_1} \bar{x_2} - \bar{x_1} \bar{x_2} + \bar{x_1} \bar{x_2} \tag{4}$$

$$= E[x_1 x_2] - \bar{x_1} \bar{x_2} \tag{5}$$

$$= \iint x_1 x_2 p(x_1, x_2) dx_1 dx_2 - \bar{x_1} \bar{x_2}$$
 (6)

$$= \int x_1 p(x_1) dx_1 \int x_2 p(x_2) dx_2 - \bar{x_1} \bar{x_2}$$
 (7)

$$= \bar{x_1}\bar{x_2} - \bar{x_1}\bar{x_2} = 0 \tag{8}$$

Part B

Show that uncorrelated Gaussian random vectors are independent.

$$P(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(\frac{-1}{2(1-\rho^2)} \left[\frac{(x_1 - \bar{x_1})^2}{\sigma_1^2} + \frac{(x_2 - \bar{x_2})^2}{\sigma_2^2} - \frac{2\rho(x_1 - \bar{x_1})(x_2 - \bar{x_2})}{\sigma_1\sigma_2} \right] \right)$$
(9)

Uncorrelated implies that $\rho = 0$, reducing this to:

$$P(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2} \exp\left(\frac{-1}{2} \left[\frac{(x_1 - \bar{x_1})^2}{\sigma_1^2} + \frac{(x_2 - \bar{x_2})^2}{\sigma_2^2} \right] \right)$$
(10)

$$= \frac{1}{\sqrt{2\pi}\sigma_1} \exp\left(\frac{-(x_1 - \bar{x_1})^2}{2\sigma_1^2}\right) \frac{1}{\sqrt{2\pi}\sigma_2} \exp\left(\frac{-(x_2 - \bar{x_2})^2}{2\sigma_2^2}\right)$$
(11)

$$P(x_1, x_2) = P(x_1)P(x_2)$$
(12)

The joint probability distribution function reduces to two single-variable probability distribution functions, which are independent from each other.

Consider a sequence created by throwing a pair of dice and summing the numbers which are $\{-2.5, -1.5, -0.5, 0.5, 1.5, 2.5\}$.

```
v_0 = [-2.5, -1.5, -0.5, 0.5, 1.5, 2.5];
pdf_0 = 1/6 * ones(1,6);
pdf_new = conv(pdf_0,pdf_0);
v_new = [-5:5];
mean_v0 = sum(pdf_new .* v_new)
variance_v0 = sum((v_new - mean_v0).^2 .* pdf_new)
```

```
mean_v0 =
    0
variance_v0 =
    5.8333
```

If we generate a new random sequence as: $V_N(k+1) = (1-r)V_N(k) + rV_o(k)$ The mean in steady state can be determined by writing the V_N formula as an infinite sum.

$$\mu_{Vn} = \sum_{-\infty}^{\infty} r(1-r)^n V_o[n] * PDF(V_o[n]) = E[V_o] * \sum_{0}^{\infty} r(1-r)^n = 0$$

Likewise, the variance can be calculated.

$$\sigma_n^2 = \sum_{-\infty}^{\infty} r^2 (1 - r)^n V_o[n]^2 * PDF(V_o[n]) = \sigma_o^2 * \sum_{n=0}^{\infty} r^2 (1 - r)^n = r\sigma_o^2$$

Also, from the properties of the autocorrelation function, we know that $E[V_n]^2 = R_X(0)$.

$$\frac{V_N(s)}{V_o(s)} = \frac{r}{s + (r - 1)} = r \exp[(r - 1) * |\tau|]$$
$$E[V_n]^2 = r\sigma_o^2$$

The covariance function follows from the autocorrelation function:

$$R_v(k) = r \exp\left[(r-1) * |L|\right]$$

The practical constraints on r is that it may not exceed 2.0 and it should not go under 0.0. Going outside these two bounds will cause the random signal to exponentially grow, which will sacrifice BIBO stability. There is no other reasonable analysis that we could do on the system if it was determined to be unstable.

Given a random variable \mathbf{x} , determine the mean and variance.

The mean can be determined with the definition. The integral will only be evaluated between 0 and 2, where the PDF is not equal to zero.

$$PDF(x) = \begin{cases} 0 & x < 0 \\ x/2 & 0 \le x < 2 \\ 0 & x \ge 2 \end{cases}$$

$$\bar{x} = \int_{-\infty}^{\infty} x \cdot PDF(x) dx \tag{13}$$

$$\bar{x} = \int_0^2 x \cdot x/2dx \tag{14}$$

$$\bar{x} = \frac{x^3}{6} \Big|_0^2 \tag{15}$$

$$\bar{x} = 4/3 \tag{16}$$

Similarly, the variance may be found using the definition.

$$\sigma_x^2 = \int_{-\infty}^{\infty} (x - \bar{x})^2 \cdot PDF(x) dx \tag{17}$$

$$\sigma_x^2 = \int_0^2 (x - 4/3)^2 \cdot x/2dx \tag{18}$$

$$\sigma_x^2 = \int_0^2 (1/2x^3 - 4/3x^2 + 8/9x)dx \tag{19}$$

$$\sigma_x^2 = 2 - 32/9 + 16/9 \tag{20}$$

$$\sigma_x^2 = 2/9 \tag{21}$$

Consider a normally distributed two-dimensional vector x, with mean value zero and

$$P_x = \begin{bmatrix} 2 & 1 \\ 1 & 4 \end{bmatrix}$$

```
P_x = [2, 1; 1, 4];
Mu = [0;0];
[V,D] = eig(P_x);
likelihood(Mu, P_x, [0.25, 1.0, 1.5]);
```

The eigenvalues of this matrix are found using the eig command in MATLAB.

The principle axes in the case of this matrix would be along the eigenvectors. This can also be found using the eig command in MATLAB.

$$V = -0.9239$$
 0.3827 0.9239

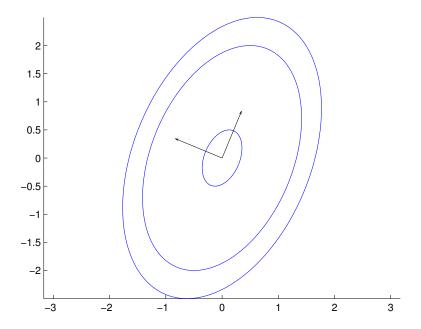


Figure 5: Likelihood Ellipses for c = 0.25, 1.0, 1.5

\mathbf{C}	0.25	1.0	1.5
Table Value	0.5987	0.8413	0.9332
Likelihood	19.74%	68.26%	86.64%

If $x \sim N(0, \sigma_x^2)$ and $y = 2x^2$, then:

$$p_x[y] = p_x[f^{-1}(y)] \left| \frac{\delta f^{-1}(y)}{\delta y} \right|$$
(22)

$$x = (y/2)^{1/2} (23)$$

$$\left| \frac{\delta f^{-1}(y)}{\delta y} \right| = \frac{1}{2\sqrt{2y}} \tag{24}$$

$$p_x[y] = \frac{1}{2\sqrt{2y}} \left(f_x \left(\frac{y}{2} \right)^{1/2} - f_x \left(\frac{y}{2} \right)^{1/2} \right)$$
 (25)

$$p_x[y] = \frac{1}{\sigma_x \sqrt{4\pi y}} \exp\left[\frac{-y}{4\sigma_x^2}\right] \qquad y \ge 0$$
 (26)

$$=0 y < 0 (27)$$

The new probability distribution function is zero up to the origin, then infinite at the origin, and declines as y approaches positive infinity. This is due to the idea that this new PDF is similar to distribution of power, with the most probable power being 0. The variable y is no longer a normal random variable.

```
X = linspace(0,5,1000);
sigma = 2;
fx_pdf = normpdf(linspace(-5,5,1000),0,sigma);
fy_pdf = 1./(sigma * sqrt(4*pi*X)) .* exp(-X./(4*sigma^2));
plot(X,fy_pdf)
hold
plot(linspace(-5,5,1000),fx_pdf)
```

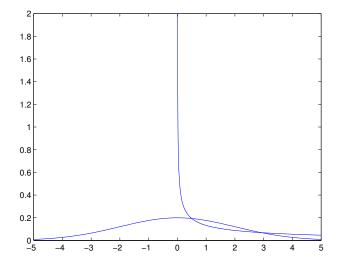


Figure 6: Normal PDF vs. PDF of y

Appendix A - Calculate Statistics of a Vector

```
function [ mean, var, c_moment, mean_sq ] = statistics( pdf )
s = length(pdf);
mean = 0;
mean_sq = 0;
var = 0;
c_moment = 0;

for i=1:length(pdf),
    mean = mean + i*pdf(i);
    mean_sq = mean_sq + (i^2) * pdf(i);
end
for i=1:length(pdf),
    c_moment = c_moment + (i - mean) * pdf(i);
    var = var + (i - mean)^2 * pdf(i);
end
end
```

Appendix B - Perform N-fold Convolution

```
function [ pdf_total ] = nfoldconv( pdf_matrix )
   %NFOLDCONV Perform a convolution sum of PDF functions.
      pdf_total = NFOLDCONV(pdf_matrix) iterates over a matrix
      of PDFs and convolves each into a single PDF
      representing the sum of the random variables.
  s = size(pdf_matrix, 1);
  pdf_total = conv(pdf_matrix(1,:),pdf_matrix(2,:));
10
  for i=3:s,
      pdf_total = conv(pdf_total,pdf_matrix(i,:));
  end
  % Correct MATLAB's convolution function chopping leading zeros off.
  pdf_total = [zeros(1,s-1),pdf_total];
  assert(sum(pdf_total) - 1.0 <= eps, ...
      'Combined PDF has a sum greater than 1, recheck PDF matrix');
20
  end
```

Appendix C - Draw Likelihood Ellipses