

Inference

- Let y_i denote observations from some process, + let $f(x_i; \vec{\theta})$, where x_i is independent variable + $\vec{\theta}$ parameters.
- We assume f approximates the true process. Then

$$y_i = f(x_i; \vec{\theta}) + \varepsilon_i, \quad i=1, \dots, N$$

where ε_i accounts for difference between y_i + f .

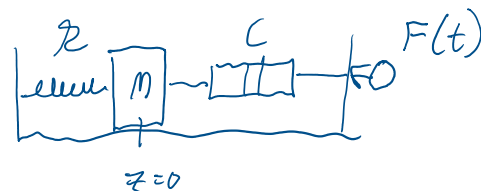
- y_i : random observations (real y_i)
- ε_i : random measurement errors (real ε_i)
- $\vec{\theta}$: Parameters
 - Frequentist: fixed, unknown
 - Bayesian: random variables

Observation Models

- Let us consider $f(x_i; \vec{\theta})$ where x_i represents ind. variable.

- y_i denotes observations.

Ex) Spring eq. $F=ma$



z : Displacement of mass

$$\Rightarrow m \frac{d^2 z}{dt^2} + c \frac{dz}{dt} + k z = F(t) = f_0 \cos(\omega_F t) \quad (I)$$

$$z(0) = z_0, \quad \left. \frac{dz}{dt} \right|_{t=0} = v_0, \quad t \in [0, T]$$

• Recall that we can rewrite (I) as

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$$\frac{d\vec{u}}{dt} = \underset{\sim}{A} \vec{u} + \vec{F}, \quad \vec{u} = \begin{bmatrix} z \\ v \end{bmatrix}, \quad \vec{u}_0 = \begin{bmatrix} z_0 \\ v_0 \end{bmatrix}$$

$$\underset{\sim}{A} = \begin{bmatrix} 0 & 1 \\ -\frac{g}{m} & -\frac{c}{m} \end{bmatrix}, \quad \vec{F} = \begin{bmatrix} 0 \\ \frac{F}{m} \cos(\omega_F t) \end{bmatrix}$$

• Note $\vec{\theta} = [m, c, g]$ are parameters if $\vec{u}_0 + F(t)$ are known.

– Suppose we have a sensor that provides measurements at discrete times t_j . Then:

$$f(t_j; \vec{\theta}) = \vec{c} \cdot \vec{u}(t_j; \vec{\theta})$$

• we call \vec{c} the observation operator

□ $\vec{c} = [0, 1] \Rightarrow$ velocity measurements

□ $\vec{c} = [1, 0] \Rightarrow$ displacement measurements

□ $\vec{c} = [1, 1] \Rightarrow$ Full System

$$\left. \begin{array}{l} \square \vec{c} = [0, 1] \Rightarrow \text{velocity measurements} \\ \square \vec{c} = [1, 0] \Rightarrow \text{displacement measurements} \\ \square \vec{c} = [1, 1] \Rightarrow \text{Full System} \end{array} \right\} C: \mathbb{R}^{n_u} \rightarrow \mathbb{R}^{n_y}$$

Statistical Observation Models

– Here, we relate $f(y_i; \vec{\theta}) = C(u(y_i; \vec{\theta}))$, to y_i + its possible measurement error.

Additive Noise Models

– For additive measurement error, we have:

$$y_i = f(y_i; \vec{\theta}) + \varepsilon_i, \quad i=1, \dots, n_y$$

$$y_i = f(x_i; \theta) + \varepsilon_i, \quad i=1, \dots, n_y$$

• we typically assume $\varepsilon_i \sim N(\mu_\varepsilon, \sigma^2)$, $\mu_\varepsilon \equiv \text{bias}$

• we also like to assume $\mu_\varepsilon = 0$

- If ε_i are additive with $\varepsilon_i \sim N(0, \sigma^2)$, we call these errors "independent + identically distributed", $\varepsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$

- under these assumptions, the sampling distribution is given by a Gaussian pdf

$$P(y|\theta) = \frac{1}{(2\pi\sigma^2)^{n_y/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{n_y} (y_i - f(x_i; \theta))^2\right)$$

Inference with noise models

- once we determine \mathcal{L} + ε distribution, we can begin inferring (estimating/learning) the model parameters, θ .

- under the assumption of $\varepsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$, the ordinary least squares (OLS) estimator + estimate for θ are

$$\text{(estimator)} \quad \tilde{\theta}_{OLS} = \arg \min_{\theta \in \mathbb{R}^p} \sum_{i=1}^{n_y} [y_i - f(x_i; \theta)]^2$$

$$\text{(estimate)} \quad \hat{\theta}_{OLS} = \arg \min_{\theta \in \mathbb{R}^p} \underbrace{\sum_{i=1}^{n_y} [y_i - f(x_i; \theta)]^2}_{\text{Cost function / Loss}}$$

• Note that $\hat{\theta}_{OLS}$ can be obtained without ε_i information.

• However, without ε_i or observation error assumptions, we cannot analyze $\tilde{\theta}$ (i.e., sampling distribution)

Maximum Likelihood Estimation (MLE)

- We can also consider the likelihood function for a random sample $\vec{y} = [y_1, \dots, y_{n_y}]$ ($\vec{\theta} \in \Gamma$)

$$L(\vec{\theta} | \vec{y}) = f_y(\vec{y} | \theta), \quad L: \Gamma \rightarrow [0, \infty)$$

where $f_y(\vec{y} | \theta)$ is the parameter-dependent joint pdf for \vec{y} .

- Note that if y_i are iid, then

$$L(\vec{\theta} | \vec{y}) = f_y(\vec{y} | \vec{\theta}) = \prod_{i=1}^{n_y} f_{y_i}(y_i | \vec{\theta})$$

- Note: Since $L: \Gamma \rightarrow [0, \infty)$, we often look at the log-likelihood,

$$\ln(L(\vec{\theta} | \vec{y})) = \text{LL}(\vec{\theta} | \vec{y})$$